SPSS FOR PSYCHOLOGISTS

HALLER

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SEVENTH EDITION

SPSS for Psychologists

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Seventh Edition

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To Emily Rae – my little statistical anomaly! <3 Virginia

To Penelope, Joseph, Francesca, Frankie and Wilb, and to my wonderful colleagues Nicky and Rosie who have worked with me for so many years on previous editions of this book. *Richard*

To my lovely family, my siblings Sue, Meg and Tim, and to my nieces, nephew, great-nieces, great-nephews and great-great-nieces – they've been so supportive and a continuing delight through all the editions. *Rosemary*

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Preface

How to use this book

This book is designed to help you analyse psychological data using the software package IBM® SPSS® statistics software (SPSS). SPSS Inc. was acquired by IBM in October 2009. With the exception of the first few sections of Chapter 1, we recommend that you read this book while sitting at a computer that is running SPSS.

Chapter 1 serves as a reminder of some issues related to research design, while Chapter 2 shows you how to enter data into SPSS. In Chapters 3 and 4, we show you how to explore, clean and manage your data. In Chapters 5–12, you will learn how to undertake a variety of statistical procedures using SPSS. The order of these chapters is designed to reflect the way many university psychology departments teach statistics to undergraduates, but the chapters are largely free-standing, so you do not need to read them in sequence. We hope that, once you are familiar with the process of data entry, you will use this book as a reference to assist you to undertake the analysis you need. To help you select the most appropriate statistical procedure, we have included a brief description of each test in Chapter 1 alongside a helpful decision tree on page 18 and 19. The procedures covered in Chapters 5–8 are likely to be taught in most undergraduate psychology research methods courses, whereas those covered in Chapters 9–13 are slightly more advanced. We hope that the statistical procedures described in these later chapters will be of particular help to students undertaking their final-year research project, postgraduate students and researchers.

Within each of the analysis chapters, we briefly describe the procedure, show you how to undertake the analysis, explain some of the options available and teach you how to obtain suitable graphs and descriptive statistics. Critically, each section ends with an explanation of how to interpret the output produced by SPSS, and write it up in APA format.

Chapter 14 describes the use of syntax to control SPSS and also provides some useful information on topics such as printing and importing and exporting files. Finally, we have included an extensive Glossary of the statistical terms used in this book.

The content of each chapter is now explained in a little more detail.

Chapter 1

Chapter 1 provides a brief overview of the basic concepts and terminology used in psychological research and statistical analysis, and introduces SPSS. We describe some basic methods of data collection and the types of data that are collected in quantitative research. We then consider data analysis and provide you with an introduction to the windows and buttons you will use when analysing your data with SPSS. Finally, we show you how to start and exit SPSS.

Chapters 2–4

Chapter 2 shows you how to set up SPSS to receive your data, and how to create and save a data file. In Chapter 3, you learn how to explore the data you have entered to calculate some simple descriptive statistics, how to check and clean your data files and how to use SPSS to produce useful graphical representations of your data. Chapter 4 introduces you to some useful functions you can use to manage your data set. We show you how to do things such as split the data file or group participants together, and how to recode existing variables or compute new ones.

Chapters 5–9

Chapter 5 covers some simple inferential statistical tests that can be used to determine: whether a sample mean differs from a known population mean; whether there is a difference between the scores from two groups of participants; or whether there is a difference between the scores from participants when tested under two different conditions. Chapters 6 and 7 focus on statistical tests of association, while Chapters 8 and 9 describe a family of statistical procedures used to analyse data from more complex experimental designs involving several conditions or variables.

Chapters 10–13

Chapters 10–13 look at tests that are appropriate for experiments involving more complex designs or for data obtained using non-experimental methods such as surveys or questionnaires.

In general, Chapters 5–13 cover a family of related tests, each of which is described in its own section. We introduce each statistical test with a brief description. This description is not intended to replace that which you will find in a statistics textbook; rather, it is intended to act as a reminder of when and how to use the test. We also include an example of a typical or real piece of research that might be analysed using the test, to give you a concrete psychological example. We then give detailed, stepby-step instructions on how to perform the analysis using SPSS. In the earlier chapters, we include screenshots (pictures) and a full description of every step required to perform the test. In later chapters, we assume that you will have become a little more familiar with the workings of SPSS, and therefore summarise some of the simpler operations. Each chapter includes an annotated example of the output produced by SPSS to help you understand the results of your analyses. Finally, we include a note on how you should report the results of your analyses. The data we use to demonstrate the statistical tests covered in Chapters 5–7 can be found in the Appendix, and all the data can be downloaded from the companion website (macmillanihe.com/ harrison-spss-7e).

Chapter 14

Chapter 14 introduces you to syntax and shows how using this text language to control SPSS can increase your efficiency, especially when working on larger, more complex data files. In addition, we describe some useful option settings, how to print files, how to import and export spreadsheet files, and how to incorporate SPSS output into other documents. This chapter also describes how to access the extensive SPSS help libraries.

Differences between versions of SPSS

This book was written using SPSS Version 26, which is quite similar to earlier versions by IBM, and indeed to even earlier versions of SPSS (or PASW as it was briefly known). However, the changes that have been introduced do not relate much to the tools we introduce in this book. You may see small differences between the dialogue boxes and the output we show and those produced by your version of SPSS. In most cases, these are likely to be purely cosmetic, affecting only the appearance of the program (indeed, even the same version of SPSS can look quite different on different computers, depending on the operating system and general display options). Furthermore, differences in appearance may result from changes made to the option settings in SPSS, or as a result of different 'Add-Ons' being installed. We show you how to use the option settings to control SPSS in Chapter 14. Regardless, any appearance differences that do occur are not likely to affect the advice and steps we present in this book.

We hope that you find our book useful and that it helps you to enjoy psychological research.

Acknowledgements

Welcome to the seventh edition of *SPSS for Psychologists*. This new edition of the book brings with it a new author, Virginia Harrison, who is very grateful to the original authors for letting her come on board! In turn, the original authors are very glad that Virginia agreed to join the team!

As with the previous editions, we would all like to take the opportunity to thank the many people who have contributed to this book. In particular, we would like to thank our colleagues, past and present, for their invaluable advice, and the various reviewers who have provided feedback on earlier editions. These contributions have again helped to shape the new edition. We thank you for your time and hope that you can see where we have made amendments or additions in light of your comments. We would also like to thank the various colleagues, supervisors, students and friends who have contributed to our knowledge of statistics and data analysis either through formal teaching, informal discussion or by presenting us with challenging data analysis problems. Some of these people were also kind enough to allow us to utilise their data to illustrate the use of particular statistical techniques. Richard and Gini would particularly like to thank William Hoang for his help in preparing some of the screenshots in this edition.

We first thought of writing a book about SPSS some 10 years before we got around to doing it. With each new version of SPSS and each new group of students, we revised our teaching notes and talked about what our book would look like. So, when our first edition was published, we were delighted by the positive response it received. Our objective was to enable students to actively engage in the discipline by undertaking their own research. Reading about psychology can be interesting, but the real fun is in doing it, and that inevitably involves the collection and analysis of data and the use of software such as SPSS. We are therefore particularly pleased to hear from students who tell us that our book has enabled them to complete their own research. It is this more than anything else that keeps us motivated to update and expand our text for each new edition. The fact that we are now into the seventh edition has amazed and delighted us (although this demonstrates how even older we now are). Thank you for taking the time to contact us.

Finally, we would like to thank the team at Red Globe Press for their support. Over the seven editions of this book, we have worked with many different members of their team. All have contributed greatly to this book, but we would particularly like to thank Luke Block and Verity Rimmer for their help and support in producing this edition.

> Virginia Harrison Richard Kemp Nicola Brace Rosemary Snelgar

Introduction

In this chapter

- Psychological research and SPSS
- Guide to the statistical tests covered
- Working with SPSS
- Starting SPSS
- How to exit from SPSS

SPSS for Psychologists online

Visit macmillanihe.com/harrison-spss-7e for data sets, online tutorials and exercises.

Section 1: PSYCHOLOGICAL RESEARCH AND SPSS

- IBM SPSS statistics software (SPSS) is a widely used computer program designed to aid the statistical analysis of data, particularly data collected in the course of research.
- SPSS® has been around in various different forms for many years and has become the 'industry standard' software for data analysis.
- SPSS is the program most widely used by university researchers, especially those working in psychology and the social sciences. SPSS is also widely used in private and government research organisations and many large private companies and nongovernmental organisations.
- Being able to describe yourself as a competent user of SPSS will enhance your employment prospects considerably.
- Although at first sight SPSS might appear rather complex, it is not difficult to learn how to undertake a wide range of statistical analyses, and once you have mastered these basics, they will enable you to undertake far more sophisticated research than would be possible without the aid of such software.

But I am studying psychology, not statistics – why do I need to learn to use SPSS?

This is a common question, and the answer lies in the nature of psychology and the type of research that many, if not most, psychologists undertake. It's important to remember that our current understanding of human behaviour and experience is based on the body of psychological research that has gone before us, so knowing how this research was carried out will allow us to better understand and evaluate the knowledge that we have; and equip us with tools to develop new knowledge.

Much of the research that has been carried out in psychology has used the scientific method, borrowing systematic research frameworks from the sciences. The problem is that people are not the easiest of things to study because they vary both between individuals and over time. That is, different people can react differently to a particular situation; and how a person reacts in a situation today might be quite different from how they react tomorrow. This means that the data collected by psychologists are much more 'noisy' than that collected in some other sciences. Consider, for example, a chemist investigating the properties of magnesium. The chemist knows that, under constant conditions, every bit of magnesium will react in the same way as every other bit of magnesium, and that how the magnesium reacts today will be the same as how it reacts tomorrow. Thus, the chemist might only need to do an experiment once using one piece of magnesium to draw firm conclusions about the nature of magnesium in general, and is unlikely to need to use statistics to help explain the results of the observations. The situation facing the psychologist is quite different, and in order to be able to determine how, in general, people react in a given situation, the psychologist will probably need to test a range of different individuals and then make use of statistical techniques to determine what trends are present in the data. So psychologists are particularly likely to need to use statistics in their research.

In recent years, the complexity of the statistical techniques routinely used in psychological research and taught to undergraduate students has increased considerably. This routine reliance on more complex statistical analysis is made possible by the widespread availability of sophisticated statistical analysis software such as SPSS. Thus, in order to be able to undertake psychological research, either as a student or a professional, you need to be able to use software such as SPSS. This book is designed to introduce you to SPSS so that you can use the program to undertake the statistical analyses you need for your course or research.

In the remainder of this section, we will provide a brief overview of some of the concepts relating to research methods, data and statistical analysis that are important to bear in mind when using SPSS to analyse your data. We refer to many of these in later chapters. (You are probably already familiar with many of them from your lectures on research methods.) Additionally, in each chapter we include an overview of some statistical issues, but this book is not intended to be a statistics reference, and you should also refer to a statistics text for full guidance.

Asking questions and collecting data

The aim of psychology research is to try to better understand human experience and behaviour. This is not an easy task when you think about all of the complexities involved in what it is to be a human being. Just think about all of the different feelings, thoughts and behaviours involved in everyday things like managing your workload, engaging with social or mainstream media, maintaining friendships and relationships... and that barely scratches the surface. There are a multitude of different questions that psychologists might want to ask about human behaviour and experience... What makes someone more likely to experience depression than another? What treatments work best for anxiety? Do people act differently on their own and in public? How does social media affect the way we think or act? Does multitasking reduce productivity? What factors predict criminality? Do we make good eyewitnesses? Why do we like some people and not others?

The different types of questions that we can ask about human behaviour are so wide and varied that psychologists have devised a huge array of different methods that we can use to help us better understand this vast topic. Which method we need to use to answer a particular question depends on the type of question we are asking and the type of data we need to answer it.

Some questions we might ask lend themselves to more qualitative methods of research. For example, if we wanted to explore how people subjectively experience or give meaning to certain events, we might want to collect rich qualitative data through interviews or focus groups to answer our question. Alternatively, other questions may be better explored using quantitative methods. Quantitative methods allow you to measure, score or count the things that you are interested in. This allows you to collect numerical data, identify patterns and make inferences about the relationships between different variables. For example, investigating how the amount of exercise people do relates to a measure of their mental health.

Methods commonly used in psychological research include questionnaire studies, interviews, observation and experiments. Each of these methods can result in the collection of quantitative data suitable for analysis using SPSS. What we do with that data will depend on what questions we are trying to answer, or what hypotheses we are trying to test. This book will focus on some of the quantitative methods commonly used in psychological research.

Levels of measurement

SPSS is designed to aid the analysis of quantitative data, which usually consists of a series of measurements of one or more variables. A variable is simply some quantity that varies and can be measured; so, height, time, weight, sex (male or female) and IQ test score are all examples of variables. A typical data set in psychological research will consist of several variables each measured for a set of different participants or cases.

We use a scale to make the measurements of a variable, and the characteristics of the scale determine the characteristics of the data we collect and ultimately what descriptive and inferential statistics we can legitimately undertake on this data. Many statistics texts written for psychology students describe four different types of scales, or levels of measurement: nominal, ordinal, interval and ratio. When entering your data into SPSS you will need to tell SPSS what types of variable you have. SPSS uses the following terms to classify variable types: nominal, ordinal or scale ('scale' refers to either interval or ratio variables). We describe them briefly below, along with an illustration of the icons SPSS uses to denote these different levels of measurement.

1. Nominal level 💞

Nominal (also referred to as 'categorical') variables are categorised rather than measured in the strictest sense. As SPSS needs us to enter numbers to represent our data, we have to assign a numerical value to the different groups or categories within our variable. For example, we might decide to record the sex of our participants and to enter these data into SPSS by adopting the coding scheme 1 = Male, 2 = Female. Assigning 1 or 2 allows us to differentiate between these two groups of participants, but the numbers themselves are meaningless in this context; these values should not be taken to imply any more than a label (or a name, hence the term 'nominal'). With nominal data just about the only thing we can do is count or measure frequency. We can report how many men and women we have, but there is little else we can do with these data and it would make no sense to calculate statistics such as the mean sex of our participants.



SPSS does not automatically know about the level of measurement used to collect your data. It is up to you to make sure that you do not ask SPSS to undertake any inappropriate analysis, such as calculating the mean of nominal data.

2. Ordinal level

Ordinal variables are measured using a scale that allows us to imply order or rank. An ordinal variable tells you more than a nominal variable; for example, it may be legitimate to say that the student who was assigned a rank of 1 performed better than the student with rank 2. However, what we can do with these data is still limited because we do not know how much better the rank 1 student performed than the rank 2 student, and because it is unlikely that the difference in ability indicated by ranks 1 and 2 will be the same as that between ranks 2 and 3, or ranks 107 and 108. Thus an ordinal scale is still a rather limited measure of performance.

3. Interval and ratio (scale) levels

Measuring at interval or ratio levels provides us with numbers that are more number-like. If we have an interval or ratio scale, then we can be certain that 3 is more than 2 and 4 is more than 3. Furthermore, we know that the interval between all points on the scale is the same; for example, the difference between 23 and 24 is the same as the difference between 102 and 103.

An interval scale is a scale that has an arbitrary zero so that a value of zero doesn't indicate that you have none of the quantity being measured. In a ratio scale, zero means that there is none of the quantity being measured. In practice, the difference between these two types of data is not critical and SPSS does not distinguish between them, using the term *scale* to describe both interval and ratio variables. For your research methods modules, however, you will probably need to know the difference between interval and ratio levels of measurement.

Hypotheses

As mentioned above, psychology research follows the scientific method – formulating ideas (or hypotheses) about the way the world works, and then testing these ideas against reality by observing, measuring and analysing data. We are all able to hypothesise or speculate about the underlying causes of people's behaviour – we do it every day. Whether it's assuming the stumbling person on the street has had too much to drink, or assuming that genetics are responsible for your statistics ability (we'll let you decide whether that's a positive or a negative assumption...), we are making inferences about the causes of behaviour. In psychological research we essentially do the same thing, but we use statistics to test these speculations (hypotheses) in a formal and rigorous way.

In a research context, a hypothesis is a prediction about the outcome of the research you plan to carry out. The hypothesis, often known as a 'research hypothesis', 'experimental hypothesis' or 'alternative hypothesis', makes a very specific prediction about something that can be defined, measured and tested. Depending on your study design, it predicts that there will be an association between variables, or that there will be a difference between the conditions you are investigating (which allows you to say whether there is an effect of one (or more) variable(s) on another). The *null hypothesis*, by contrast, expresses the possibility that there is no effect or association.

When analysing data using SPSS you need to be clear about whether your research hypothesis is one-tailed or two-tailed. A one-tailed hypothesis makes predictions about both the presence of an effect (e.g. there will be a difference in the performance of young and old participants on a memory test) and also about the direction of this difference or association (e.g. young participants will obtain a higher score on a memory test than elderly participants). In contrast, a two-tailed hypothesis predicts only the presence of an effect, not its direction. With certain statistical procedures, SPSS will ask you to specify whether your hypothesis is one- or two-tailed.

Operationalisation

To examine how a hypothesis is formulated, let's imagine you are a psychologist interested in investigating how human memory works. This is an extremely broad research question which could be interpreted in a number of different ways, depending on how it is defined. For example, are you interested in recognition or recall; do you want to know if we remember some things better than others, and if so, why; perhaps you're interested in how short-term memory changes according to personal characteristics, such as age; or are you more interested in the content of memories themselves? It is such a broad topic, it is difficult to know where you would start to carry out research in this area. And this is where hypotheses come in. Before a researcher can explore their topic of interest (e.g. how people remember things), they first need to turn the concepts they are interested in from something vague and abstract into something more concrete. This is a process known as operationalisation, and essentially involves identifying, narrowing down and defining specific factors or psychological constructs in a way that allows them to be directly accessed or measured.

In this example, you would need to find a way to measure the aspects of memory (and any other variables) that you are interested in. For example, if you were interested in how age is related to memory, you might want to look at how many words people are able to remember from a long word list and look at how that changes as a function of their age (measured in years).

Once your specific variables and measures have been identified and properly defined, you can produce your specific hypothesis (or prediction). While a research question is often exploratory, a hypothesis makes a specific prediction about the pattern of results that are expected. This prediction should always be theoretically grounded in previous research. For this example, previous research suggests that short-term memory declines with age. As such, a suitable hypothesis for this research might be something like: *Participants will be able to recall fewer words as their age increases*.

You now have a specific statement that makes a prediction about people's behaviour. This statement lends itself to being investigated in a single study and clearly identifies the variables to be tested. Once a testable hypothesis has been identified, you then need to think carefully about how you can design an appropriate and scientifically rigorous study that can address it. The aim is to find evidence to either support or refute your prediction.

Types of study design

Broadly speaking, there are two types of study design that can be used to test hypotheses: correlational designs and experimental designs. Which one you choose will depend on your specific hypothesis and how you have operationalised your variables.

Correlational designs

Correlational designs are used to investigate hypotheses that predict associations or relationships between variables. That is, whether changes that happen to one variable are related to changes that happen in another (like memory scores decreasing as age increases). To take another example, imagine that you predict a relationship between time spent doing statistics and general happiness scores. You could measure people's happiness using a scale so that each individual receives a happiness score, and then you could also measure the amount of time they spend doing statistics each week. You might find that, as the amount of time spent doing statistics increases, people's happiness levels decrease; in other words, as one variable changes so does the other. This is an example of a correlational study. (Disclaimer: This is an entirely fictional example, and is in no way suggesting that doing statistics is bad for your health. Obviously, everybody loves statistics!)

While this type of design can tell you whether or not scores on two variables change in a synchronised manner, they cannot tell you whether one variable *causes* changes in the other. One reason for this is that it is not always possible to establish the direction of the relationship between the variables. Common sense might tell you that having to do lots of statistics will cause people to become unhappy; however, you cannot rule out an alternative explanation. It could be that people who are more unhappy seek comfort in the concrete logic and reason of numbers – unlikely, but possible!

It is also possible that any correlation you find might actually be the result of a third factor. In other words, the link between statistics and happiness might arise because both variables are related to something else. For example, people who spend more time doing statistics might spend less time outdoors, meaning they are exposed to less sunlight (a factor known to affect mood). Again, this prevents you from being able to draw causal conclusions about the relationship between your variables.

Experimental designs

Experimental designs are different from correlational designs, as they allow you to establish cause and effect. As such, they are used to investigate hypotheses that predict one variable will have an influence (or an effect) on another. This is done by manipulating (or introducing a change in) one variable (the **independent variable**) and then monitoring its effect on another (the **dependent variable**). Analysing the data typically involves looking for differences in dependent variable (DV) scores between the different experimental groups or conditions.

For example, imagine you hypothesise that doing statistics will cause people to become unhappy (or will have an effect on happiness scores). The trick is to design an experiment where everything except the independent variable (IV) is kept constant. You can manipulate the IV by changing the conditions under which people take part in the experiment. In one condition, participants could be asked to score their happiness according to a scale after doing complicated statistics for half an hour; while in another condition participants could be asked to score their happiness after 30 minutes of doing nothing. This allows you to say that any differences you observe in your DV (happiness score) is the result of your manipulation of the IV.

While not always the case, experiments are typically carried out in laboratories because this allows experimenters to manipulate their variables of interest in a controlled environment, thus minimising the possible effects of confounding variables (i.e. variables that are not controlled by the researcher but that can affect the results of the study). This allows the experimenter to focus on the specific relationship between the independent and dependent variable.

Quasi-experimental designs

While experiments can be extremely useful in helping us to establish an influence of one variable on another (i.e. establishing cause and effect), there are situations where experimental research is simply not possible. For a study design to be strictly experimental, the researcher has to directly manipulate a variable (the IV) to investigate its impact on another variable (the DV). The implication being that if the researcher does not manipulate the IV, then the study cannot be an experiment.

But what happens when the experimenter is unable to directly manipulate the IV? Does this mean that we are unable to explore its effects on (or relationship with) the DV? Of course not! For example, you may want to look for differences in behaviour or performance between naturally occurring grouping variables, such as gender or age bands. Or you may be interested in exploring a topic where the direct manipulation of an IV would be unethical. For example, conducting an experiment to establish a causal relationship between lung cancer in humans and smoking would require you to randomly allocate people to two groups and make one of the groups smoke cigarettes (which would involve knowingly causing harm to those people). Or if you wanted to investigate neural activity in murderers compared with a control group... you could not split a group of people into two groups and then ask one of these groups to commit murder! The same kinds of problems are experienced when investigating things like the impact of brain damage or mood disorders on different abilities.

In situations where a researcher wants to investigate the effect of an existing grouping characteristic (such as gender, age, height, IQ, mood, smoking behaviour or a personality trait) on another variable, psychologists can use what is known as a quasi-experimental design. In these cases data is treated and analysed in the same way as data from actual experiments, although causal conclusions need to be more tentatively drawn as strictly controlled manipulations have not been carried out on the IV.

Once you have come up with a suitable design to test your hypotheses, you need to collect your data and then interpret it using the appropriate inferential statistics. Section 2 contains two handy decision trees to help you choose which test is likely to be most appropriate for the data you have, for situations where you are exploring differences (i.e. in experimental and quasi-experimental designs) or relationships (i.e. in correlational designs).

Related and unrelated designs in psychological research

One important characteristic of any research method is whether it involves a *related* or *unrelated* design. To keep you on your toes, these designs are referred to by many different names. Related designs are also commonly known as 'within-participants', 'repeated measures', 'correlated' and 'paired' designs; while unrelated designs are often referred to as 'between-participants', 'independent groups' and 'uncorrelated'. In a related research design, two or more measures on a variable are linked in some way – typically because they are measured from the same participant. In contrast, if we take only one measure on each variable from each participant, this would be an unrelated design. This distinction is important because data collected using related and unrelated data must be coded differently in SPSS data files (see Chapter 2).

Populations and samples

When we plan a piece of research, we will have in mind a population we want to explore. For statistical purposes, a population is the total set of all possible scores for a particular variable. For some research the population might be quite small. For example, if we are researching occupational stress among professional taxidermists, then our population is all professional taxidermists. Given that there are relatively few professional taxidermists, we might be able to survey them all and so measure the entire population of scores. However, for most research, it would never be possible to collect the entire population of all possible scores for a variable, either because the population is too large or because of practical limitations. In these situations we rely on testing a sample. That is, we collect a smaller subgroup of scores that we hope will be representative of the whole population. A bigger sample has a better chance of being representative of the population, so sample size is an important consideration (see Statistical power later in this section). You will need to seek guidance on this from your statistics text as SPSS will analyse your data even when your sample size is too small for the analysis you want to undertake.

Parameters and statistics

If we measure the entire population of scores for a variable, then we can calculate a *parameter* of the population, such as the mean. However, usually we only sample the population and so can only calculate the mean of the sample. A measure, such as the mean or standard deviation, based on a sample is known as a *statistic*. The important distinction is that we can calculate a statistic, but usually can only estimate a parameter because we cannot measure the entire population of scores. Assuming that our sampling procedure was sound and our sample was not too small, we would expect the mean of the sample to provide a fairly accurate estimate of the mean of the population. The bigger the sample, the more accurate our estimate is likely to be; but unless we test the entire population, we will never know the true population parameter for certain.

Descriptive statistics

Descriptive statistics summarise large volumes of data. The role of descriptive statistics is to provide the reader with an understanding of what the data look like by using a few indicative or typical values. Depending on the type of data, this can involve a measure of central tendency (e.g. mean) and a measure of spread (e.g. standard deviation). In Chapter 3 we cover how to obtain descriptive statistics. Additionally, in each of the chapters covering a statistical procedure we show you how to obtain descriptive statistics; however, you will need to decide which is most appropriate for your data.

Confidence intervals and point estimates

When estimating a parameter such as a population mean, there are two approaches we could adopt. In psychological research the most common approach is to cite a single value, or point estimate, which represents the best estimate of the parameter. For example, we might calculate that the mean height of a sample is 1.73 m and we might use this as our estimate of the mean of the population. Clearly, it is unlikely that the true value for the population will be exactly 1.73 m, but this point estimate represents

our best guess at the value. Although they are widely employed in psychology, point estimates are limited because they do not tell us anything about the likely accuracy of our estimate. The alternative approach uses a device called a 'confidence interval', which consists of two values, an upper and lower limit, and these two values define the range within which we expect the true population value to fall. These values give an indication of how uncertain we are about our estimate of the population parameter; the more widely spread the values, the less certain we are about our estimate. For example, we might calculate that the upper and lower limits of the 95% confidence interval are 1.6 m and 1.8 m. These confidence intervals define the range of values, which is likely to include the unknown population parameter. More precisely, these values tell us that, if we took repeated samples from the population and calculated the 95% confidence intervals for each, then in 95% of these samples the calculated interval would include the true population parameter. Thus, we estimate that there is a 95% probability that the true value lies between these two values.

The confidence interval approach has several advantages over the point estimate approach. First, it serves to remind us that the population estimate is just that – an estimate which is associated with a certain level of error. Second, the confidence interval approach conveys information about the likely magnitude of the real value. If we only provide a point estimate value for the population mean, then we have no way of knowing what the likely limits are. Another valuable use of confidence intervals is in the graphing of results. By marking the confidence intervals when plotting different group means, we provide the reader with a clear indication of the extent to which the estimated means for different conditions overlap. This use of confidence intervals in graphing is illustrated in Chapter 5.

Bootstrapping

Recent versions of SPSS have included a powerful procedure called 'bootstrapping'. The name 'bootstrapping' refers to the impossible idea of picking yourself up off the floor by pulling on your shoelaces, and is used in various contexts to describe a process that achieves something without external assistance; for example, we talk about 'booting' a computer when we start it up. In statistics, the term is used to describe a particular approach to estimating parameters. Normally, when we estimate a parameter such as variance, we make assumptions about the nature of the distribution. Bootstrapping provides an alternative method of estimating these same parameters without making such assumptions.

Imagine that we are interested in the height of a population of people and have measured a sample of 100 individuals. We can now use the bootstrapping method to estimate the mean height of the population. We do this by collecting lots of new samples, each of 100 observations, and then calculate the mean for each of these samples. But where do we get these new samples from? SPSS doesn't collect more data for you, instead it samples from your sample using a procedure called 'random resampling with replacement'. To understand this, imagine that each of your 100 observations are written on a scrap of paper and placed into a tub. Now imagine reaching into the tub, picking out one piece of paper and recording the value on the paper before returning it to the tub. Now repeat this 99 more times to give you a new sample of 100 observations. This is called our 'bootstrap sample'. Note that some of the original values will not be included in the bootstrap sample, while others might appear several times, and as result the mean of this new sample will probably not be the same as the mean of the original sample. If we repeat this whole process lots of times (typically at least 1,000 times), we will end up with a set of means of our bootstrap samples and can use these to estimate the mean of the population from which the original sample was drawn.

Bootstrapping is particularly useful when we don't know about the underlying population distribution or don't want to assume that the distribution is normal. SPSS now includes a bootstrapping option in many statistical procedures to estimate mean, standard deviation, variance skewness, kurtosis and various other statistics. We will not cover the use of the bootstrapping option here, but once you have familiarised yourself with SPSS, you might like to explore the bootstrapping option and compare the output with that from the conventional method we describe in detail in each chapter.

Inferential statistics and probability

Inferential statistics allow us to go beyond simply describing the data. Through the use of a variety of different inferential statistical tests, we can answer questions such as: 'Can caffeine improve memory ability?', 'Is there a difference between the reading ability of the male and the female participants?' or 'Is there a relationship between participants' age and memory?' What these statistical tests have in common is that they use mathematical procedures to attempt to estimate the probability of obtaining the data if the null hypothesis were true.

Inferential statistical tests allow us to decide which of two possible explanations is the most likely: are the results of our study 'real' or simply down to a poor sample selection. This is best illustrated using an example. Imagine you wanted to explore the effects of caffeine on memory, so you test people's memory for a list of words after drinking no caffeine or after drinking 50 mg. If you could study every person in the population in every experimental condition, you could be fairly certain that any differences found were really the result of the experimental manipulation (that is, they were really due to differences in the amount of caffeine participants drank). However, if you can only include a sample of the population in the study, it is quite possible that any differences found between the two groups could be attributed to the specific samples chosen rather than the experimental manipulation itself.

For example, imagine you drew a sample of 20 people to represent a population. It could be that, by chance, you picked 20 people who just so happened to have an extraordinary memory. This sample would therefore not be representative of the population as a whole. Taking this one stage further, say these 20 people made up one experimental group in your study (your caffeine drinking group) and you then drew another 20 people from the population to make up your second group (in this case the control group, who drank no caffeine). This time, your sample contained people who just so happened to be particularly bad at remembering things. The problem is, because you have drawn two different samples from the general population, how can you then tell whether any differences you found between the two groups' scores was due to your experimental manipulation (caffeine) or simply due to the accidental differences between the samples that you chose (sampling error)?

To make sampling useful, you need a way of estimating how well your samples represent the population. In addition, you need some means of determining whether any difference found between the two samples is a real difference (i.e. one that would occur if you could test the entire population in both conditions) or if it is just the result of the specific samples selected. This is where inferential statistics come in!

Inferential tests do this by calculating the probability that the apparent difference was just down to luck in our sampling; that the pattern of results observed in our data could have arisen by chance alone. For this reason, all inferential statistical tests will result in the computation of a probability value or a p value, which ranges between 0 and 1.

The probability value calculated for a statistical test tells us how likely it is that any patterns we can see in our data are just down to chance fluctuations and not indicative of a 'real' effect. When using SPSS the output from the inferential tests will include this probability value and this is often labelled the 'Sig. Value'.

But how should you interpret the value? There are two approaches that can be taken here: report the probability level and leave the reader to decide how to interpret this; or set a criterion below which the null hypothesis can be rejected and the research hypothesis accepted. This second approach is the one traditionally adopted in psychological research and is the basis of the concept of statistical significance. Psychologists usually set the criterion (also known as the Type I error rate) at p = .05. If the probability that the pattern we have observed in our data would arise just by chance (if the null hypothesis was true) *is less than .05* (or 5%), then we reject the null hypothesis and declare that we have a *statistically significant* result. In each chapter covering a statistical test, we will show you how to report the probability value provided by SPSS.

Adjusting *p* values for one- and two-tailed hypotheses

By default, SPSS usually calculates the p value for two-tailed hypotheses, as this is the safest, most conservative option to adopt. However, if before undertaking the research, you established a one-tailed hypothesis and did so on the basis of sound reasoning or previous research, then it is possible to either request the calculation of a p value appropriate to a one-tailed hypothesis (as you can do for the tests of correlation described in Chapter 6) or to calculate the one-tailed p value by halving the two-tailed value.

Exact and asymptotic significance

SPSS provides the option of either exact or asymptotic significance for certain inferential statistical tests. Exact significance values are calculated using the true probability distribution of a test statistic. However, as this type of calculation can be computationally intensive and time consuming, it is often easier and quicker to use an approximation of the true distribution for the calculation. As such, SPSS often uses asymptotic significance, which is tested against an asymptotic distribution in which the probability of occurrence of extremely high or low values never reaches zero (i.e. in a graph of the distribution the tails never touch the horizontal axis). The normal distribution is an example of an asymptotic distribution. If we have a large sample, then it is safe to use the asymptotic distribution to calculate probability values (as the difference between the asymptotic and exact p values is minimal). However, with small sample sizes this approach can be problematic, as the asymptotic p-values can be quite different from the exact p values, leading to different possible interpretations of the results. To overcome this problem versions of several inferential statistical tests have been developed, which allow you to produce the 'exact' significance. One common example is Fisher's exact test, used in chi-square designs that involve small numbers of participants (see Chapter 7, Section 4). If you choose to make use of an exact test, then you should make this clear when reporting the results of your analysis.

The fact that the exact tests allow you to test hypotheses with small numbers of participants does not mean that it is acceptable to test few participants. The statistical power of your design (see below) will be greatly reduced by a small sample size.

Confidence intervals and statistical inference

The confidence interval approach described earlier can also be employed to make statistical inference. SPSS can calculate the *mean difference* (e.g. between two groups or two conditions) – this is an estimate of the real difference between the two population means – and we can state this mean difference using either a point estimate or a confidence interval. For example, if the upper and lower limits of the 95% confidence interval for the mean difference were 1.3 and 2.2, this would tell us that, according to our estimates, there is a probability of .95 (or 95% probability) that the real mean difference will lie somewhere between these two values.

Expressing the significant difference using confidence intervals has the advantage that it also provides an indication of the magnitude of the likely difference between two groups or two conditions, and the variability of the difference. Using the confidence interval approach, a difference is taken to be significant if the upper and lower limits of the confidence interval are either both positive or both negative – that is, if the range of values does not include zero. If the range of values does include zero, this tells us that it is quite likely (more than 5% in the case of 95% confidence intervals) that the true difference between the population means might be zero, and hence any difference we see in our data is just down to sampling error. This is equivalent to saying that the difference is non-significant.

At present, SPSS provides confidence intervals for the mean difference as part of the default output for the *t*-test (see Chapter 5), and in this case, these confidence intervals can be used to infer statistical significance. Confidence intervals for the mean are also available as optional output for some tests; for example, ANOVA, ANCOVA and MANOVA. However, the use of confidence intervals to infer statistical significance in these routines is not well supported in SPSS and is beyond the scope of this book. Interested readers are advised to consult Bird (2004), which provides comprehensive coverage of this complex issue.

Effect size

Most psychological journals require that authors quote *effect size* alongside a statement of whether the result is significant. If we are undertaking research into the effectiveness of a new form of psychological treatment for depression, then we will not only want to know whether the new form of treatment is more effective than the old, but also *how much* more effective the new treatment is. This is the role of a measure of effect size; it tells us how big an effect we can expect from an intervention, treatment, manipulation of an independent variable or difference between groups.

Effect sizes are measured and quoted using a variety of scales. The most basic, and sometimes most useful, measure of effect size is a raw difference of means expressed in the same units used to measure the dependent variable. For example, when reporting the results of an intervention for children with reading difficulties, we could report that the mean difference between the treated and untreated group was 8.2 points on a standardised test of reading ability. Researchers familiar with this test of reading ability will be able to determine whether this is a large, medium or small effect size and thus whether the intervention is worthwhile.

Sometimes, it is not very useful to express effect size in the units used to measure the dependent variable. For example, suppose a researcher measures how long it takes a group of elderly and a group of young participants to memorise a piece of text. The dependent variable is time measured in minutes, but the absolute time taken will depend on the passage selected and is therefore rather arbitrary and difficult to interpret. The most widely adopted solution in cases like this is to standardise the effect size measure by expressing it in standard deviation units, a procedure advocated by Cohen (1988). One such measure, called Cohen's d, tells us how big the difference in the means is relative to the standard deviation of the scores. Cohen (1988) suggested a classification for effect sizes using this measure. He suggested that an effect size of 0.2 should be regarded as 'small', 0.5 should be regarded as a 'medium' sized effect, and an effect size of 0.8 should be regarded as 'large'. It is important to realise that there is nothing special about these values and the labels attached to them – this is merely the naming convention suggested by Cohen.

Effect size measures, like confidence intervals, are a way of providing the reader with more meaningful information. There are several different measures of effect sizes in addition to Cohen's *d*. We show you how to obtain a measure of effect size in each chapter covering an inferential statistical test.

Statistical power

The power of an inferential statistical procedure is the probability that it will yield statistically significant results. Power is the ability of a procedure to accurately discriminate between situations where the null hypothesis is true and situations where the null hypothesis is false. Statistical power is influenced by several factors, including the effect size (see above) and the sample size (usually the number of participants tested). If we are hoping to find statistically significant evidence for a small effect, then we will need to test a large number of participants. It is possible to calculate, for a given effect size, how many participants we will need in order to have a good chance of achieving statistical significance. Cohen (1988) suggests that, in order to avoid wasting their own time and that of their participants, experimenters should aim to

achieve a power of 0.8. That is, we should design our experiments so that we have at least an 80% chance of obtaining a significant result should one exist. For this reason, it is good practice to estimate the power that a particular design will achieve, and how many participants will be needed, to have a particular probability of revealing a significant effect, should there be one. Many research funding bodies will only consider applications that include this type of power calculation.

There is another reason to calculate the power of a given design. If we fail to obtain a significant result, there are two possible explanations. It could be that the null hypothesis is true – that is, that there is no difference or no association. Alternatively, it could be that there is a difference or an association but that we failed to detect it because we had insufficient statistical power. Unless you know about the power of your design, there is no way you can distinguish between these two alternatives. However, if you know that your design had a power of about 0.8 to detect a small effect size (d = 0.2) and you have not found a significant effect, then you know with a probability of 0.8 that *either* there is no effect *or*, if there is an effect, it is very small (*d* is less than 0.2). Thus, there is only a relatively small probability that there is an effect greater than 0.2, which you did not find.

Statistical power calculations can appear daunting because they involve the interrelationship between four variables: the statistical power of the design, the sample size, the criterion p value and the effect size. However, there are several specialist software packages available to undertake these calculations for you. We particularly recommend a piece of software called G*Power, which can be downloaded free from www.gpower.hhu.de.

Practical equivalence of two samples

In advance of undertaking our research, we might decide that we are only interested in an effect that exceeds a certain effect size. For example, if comparing two treatments for depression, we might decide that, given the additional 'costs' associated with our new treatment, we are only interested in an effect size of at least d = 0.15. In this situation, we have decided, in advance, that an effect size of less than 0.15 is 'trivial' and that two treatments that differ by less than this amount are of *practical equivalence*. Given this, if our power was 0.8 to detect an effect of size 0.15, then we can confidently conclude, if we fail to find an effect size of at least 0.15, that there is no practical difference between the two treatments. If, on the other hand, the statistical power of our design to detect such a small effect was low, say 0.5, then we should conclude that the issue of whether this new treatment is worthwhile is still undecided (and we should chastise ourselves for having conducted such an important piece of research with such low statistical power!).

Statistical power and SPSS

As with confidence intervals, recent editions of SPSS have increased the availability of measures of power. For example, measures of 'Observed power' are now available from the 'Options' dialogue box for the ANOVA, ANCOVA and MANOVA routines. Where available, users may like to incorporate this additional information into their research reports.

Section 2 provides a guide to the tests covered in this book, and we then move on to describing SPSS.

Section 2: GUIDE TO THE STATISTICAL TESTS COVERED

Choosing the correct statistical procedures

SPSS will not tell you which descriptive statistic/s you should use to describe your data, or which inferential statistic to use to test your hypothesis. Broadly speaking, you need to consider the nature of the hypothesis being tested, the method and/or experimental design employed, the number of variables manipulated and/or measured, and the type of data collected.

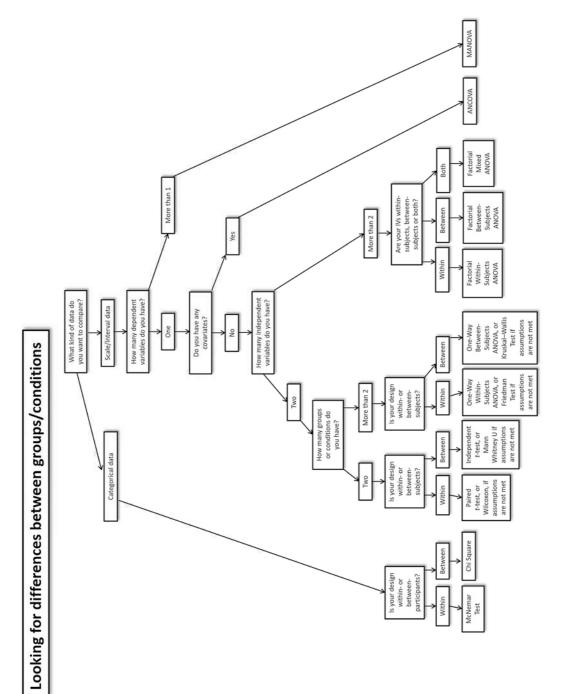
Most of the inferential statistical tests covered in this book are *parametric tests*, as these are inferential tests that have greater statistical power. However, their use is limited to situations where the data meet certain requirements, including the following:

- 1. The data are measured using an interval or ratio scale.
- 2. The data are drawn from a population that is normally distributed.
- 3. The samples being compared are drawn from populations that have the same variance.

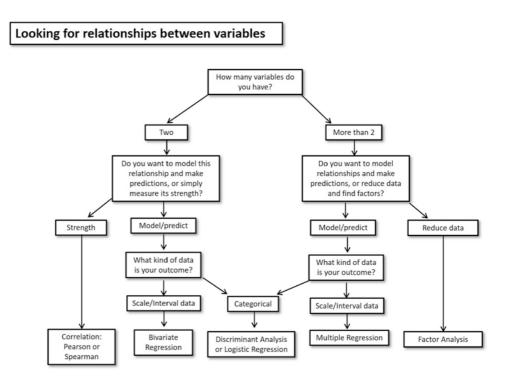
As we shall see, in some cases, SPSS includes information that you can use to assess whether the data violate any of these assumptions. Also, certain tests make additional or different assumptions from those listed above, and we alert you to these cases in the relevant chapters.

There are occasions in psychology when we collect data that do not satisfy all these requirements. *Nonparametric* tests are inferential tests that make very few assumptions about the data and in particular its distribution. However, they are less powerful than their parametric equivalents.

On pages 17–18 you will find two tools to help you choose which statistical test is most appropriate for your data, as well as a guide to tell you where you can find these tests in this book.



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Once you have established which test is most appropriate for your data, you can find out more about it in one of the below chapters.

Chapter 5 Covers tests that look for a difference between two groups or two conditions (or levels) of the independent variable (IV) or factor	 One-sample <i>t</i>-test: determines whether a mean is different from a set value Independent <i>t</i>-test: suitable for independent groups (or between-subjects) designs Paired <i>t</i>-test: suitable for repeated measures (or within-subjects) and matched-subjects designs Mann–Whitney test: nonparametric equivalent to the independent <i>t</i>-test Wilcoxon test: nonparametric equivalent to the paired <i>t</i>-test 				
Chapter 6 Covers tests that look for a relationship between two variables	Pearson's <i>r</i> : used to test for correlation between two variables; assumes linear relationship Spearman's r_s : nonparametric equivalent to Pearson's <i>r</i> ; used to test for correlation between two variables using ranked data Bivariate regression : a simple, parametric linear regression that allows you to investigate the relationship between two variables and predict the values of one variable from another.				
Chapter 7 Covers tests suitable for nominal data	Multidimensional chi-square: can be used as a test of association or as a test of differences between independent groups McNamar test: used to test for differences on a dichotomous variable for related designs				

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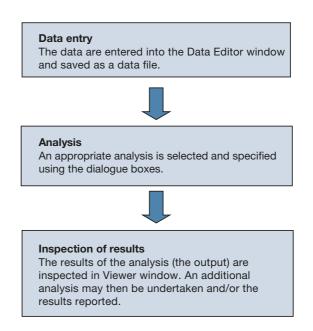
Chapter 8 Covers tests of difference for experimental designs involving one IV with more than two levels	One-way ANOVA: analyses data from both within- and between-subjects experimental designs, involving one IV with more than two conditions Kruskal–Wallis: nonparametric test suitable for independent groups designs involving one independent variable with three or more groups or conditions Friedman: nonparametric test suitable for repeated measures designs involving one independent variable with three or more conditions
Chapter 9 Covers tests of difference for experimental designs involving more than two or more IVs	Factorial ANOVA : various versions (two-way, three-way etc.) Used to analyse data from complex experimental designs involving two or more IVs with within-subjects, between-subjects or mixed designs
Chapter 10 Covers a test looking at the relationship between several IVs and one dependent variable	Multiple regression : allows us to predict a score on one dependent variable (the criterion variable) from a number of independent (or predictor) variables
Chapter 11 Covers two tests related to ANOVA	ANCOVA: similar to ANOVA but controls for the effect of one or more covariatesMANOVA: similar to ANOVA but used when the design includes more than one dependent variable
Chapter 12 Covers two tests looking at the relationship between several IVs and category membership	Discriminant analysis : similar to multiple regression but used when the dependent variable is a categorical (nominal) variable
Chapter 13 Covers a test looking at the interrelationships among a large number of variables	Factor analysis: allows you to identify whether a factor structure underlies correlations between a number of variables and to extract the factors

Section 3: WORKING WITH SPSS

SPSS (originally Statistical Package for the Social Sciences) is an enormously powerful program for statistical analysis. SPSS has been around in various forms for over 50 years, and, for a brief period, was known as PASW, before being taken over by IBM in 2009. The package is now known as IBM SPSS Statistics[®]. Knowing how to use SPSS will allow you to perform a wide range of statistical analyses, which would otherwise require a great deal of complex calculation.

Data analysis using SPSS

There are three basic steps involved in data analysis using SPSS: First, enter the raw data and save as a data file; Second, select and specify the analysis you require; Third, examine the output produced by SPSS. These steps are illustrated below. The special windows used by SPSS to undertake these steps are described next.



The different types of window used in SPSS

SPSS utilises several different window types. However, new users of SPSS only need to be familiar with two of these windows: the Data Editor window and the Viewer window. We will be using these two windows in this and the next chapter. The other window types are explained briefly here and will be covered in more detail elsewhere in the book.

https://cochrana.ir/SPSS The Data Editor window

The Data Editor window (or data window) is the window you see when you start up SPSS. This window looks rather like a spreadsheet and is used to enter all the data that is going to be analysed. You can think of this window as containing a table of all your raw data. We will examine this window in more detail when we start up SPSS.

The Viewer window

The Viewer window is used to display the results or output of your data analysis. For this reason we will sometimes refer to it as the 'Output window'. We will examine this window in more detail when we perform our first simple analysis in Chapter 2.

Some other windows used in SPSS

- 1. The Syntax Editor window is used to edit special program files called 'syntax files'. The use of this window is explained in Chapter 14 and will be of particular interest to more advanced users.
- 2. The Chart Editor window is used to edit standard (not interactive) charts or graphs. The use of this window is explained in Chapter 3.
- 3. The Pivot Table Editor window is used to edit the tables that SPSS uses to present the results of your analysis. The use of this window is explained in Chapter 14.

Section 4: STARTING SPSS

Start SPSS by double-clicking on the icon on your desktop or by selecting **IBM SPSS Statistics 26** (or the latest version you have) from the applications on your machine.

If you do not have an SPSS icon on your desktop then:

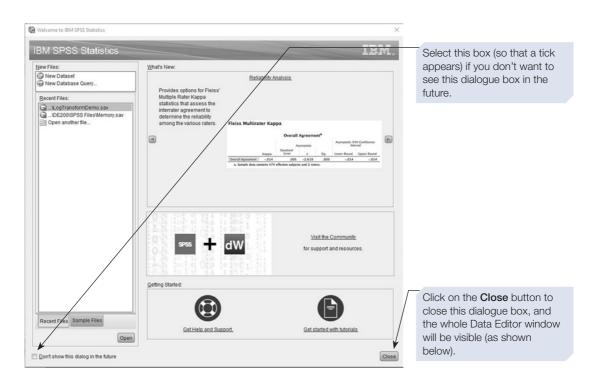
For PC, click on the Start button at the bottom left-hand corner of the screen, then select **All Programs**. If the SPSS application is not visible in the list, it may be inside a folder. You can search for it by pressing Windows key on your keyboard and typing 'SPSS' in the search box.

For Mac, select Finder, then Applications.

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The image above is the opening screen for SPSS Statistics 26. Unless a previous user has switched it off, a box will appear in the centre of the opening screen, as shown above. This box is an example of a dialogue box. SPSS makes extensive use of dialogue boxes to allow you to control the program. This dialogue box is designed to help get new users started, but we aren't going to use it. Click on the Cancel button to close the dialogue box (see below) so that we can see the Data Editor window behind it.

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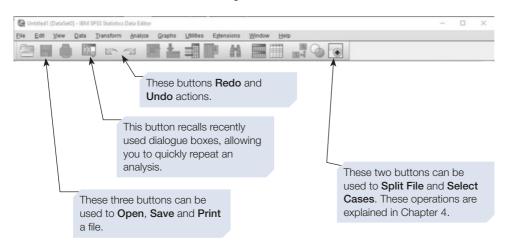


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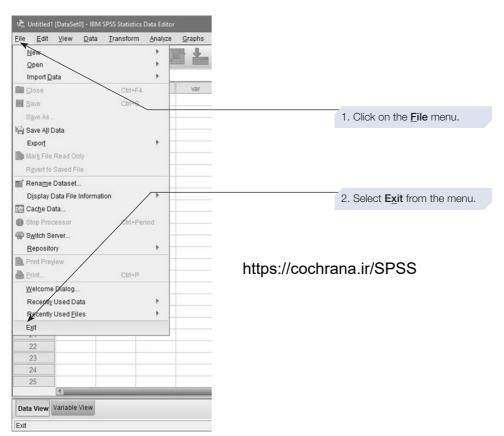
The menu and toolbars of the Data Editor window

The menu and toolbars from the Data Editor window of SPSS Statistics 26 are shown below. The toolbar buttons duplicate functions that are also available from the menus. Some of the more useful buttons are explained below.

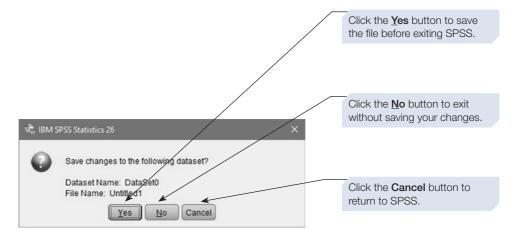


Section 5: HOW TO EXIT FROM SPSS

- 1. Click on the **File** menu item.
- 2. Select Exit from the File menu.



If you have made any changes since you last saved the file, SPSS will ask you if you want to save the file before you exit (see below).



SUMMARY

- This chapter provided an overview of some of the statistical concepts you need to understand when using SPSS and introduced you to SPSS.
- Section 2 gave an overview of the statistical procedures covered in detail in Chapters 5–13. Section 3 outlined the SPSS windows, and Sections 4 and 5 showed you how to start and exit from SPSS.
- Chapter 2 will cover how to enter data into SPSS.
- Chapter 3 will show how to explore your data.
- Chapter 4 will demonstrate different ways of manipulating your data, for example to create new variables.

Data entry in SPSS

In this chapter

- The Data Editor window
- Defining a variable in SPSS
- Entering data
- Saving a data file
- Opening a data file
- Data entry exercises
- Answers to data entry exercises
- Checking and cleaning data files

SPSS for Psychologists online

Visit macmillanihe.com/harrison-spss-7e for data sets, online tutorials and exercises.

Section 1: THE DATA EDITOR WINDOW

What is the Data Editor window?

- The Data Editor is the active window when you start SPSS. It is used to record all the data we want to analyse.
- It has two views: the Variable View and the Data View.
- The Variable View allows us to name each column in the Data table and specify what sort of values the column will contain.
- The Data View contains a table with a large number of cells in rows and columns. The table can be very large with only a small part of it visible, in which case use the scroll bars on the edges of the window to move round the table.
- In psychology, we almost always enter data in the same way. Normally, each row represents an individual participant and each column represents a variable.

The arrangement of the data in the Data Editor window

The precise way that the data are entered in the Data Editor window is critical and will depend, in part, on the details of your study. If you are entering data from an experiment, then you need to consider the design employed. In an independent groups design, each participant will provide one measure of performance. In addition, you will need to indicate which of your experimental groups each participant was assigned to. Thus, the most basic independent groups design will require that you use one column of your data table to record which group your participant was in and a second column to record that participant's score. By comparison, in a repeated measures design, each participant's performance will be assessed several times. Thus, for each participant you will need several variables to record the performance of that participant under each condition.



In SPSS, the word 'variable' means a column in the data table; it does not have the same meaning as it does in experimental design. For example, in a repeated measures design where there are two levels of the independent variable, we use two columns in the data table to record the values for the single dependent variable.

Section 2: DEFINING A VARIABLE IN SPSS

Our first job is to set up the data file with important information about each of our variables. This process of defining the variables is described here.

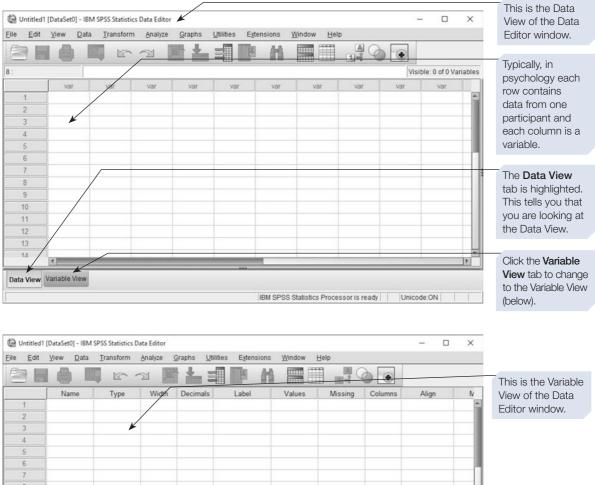
The Data View and Variable View

If you look at the bottom left-hand corner of the Data Editor window, you will notice two 'tabs'. One tab is labelled 'Data View' and the other 'Variable View'. You can think of these as the index tabs for two different pages of information. When you first enter the Data Editor window, you will probably find the Variable View tab selected. If you click on the Data View tab, you will be presented with the empty data table. If you click on the Variable View tab, a different screen of information will be displayed. These two views are illustrated below.

The Data View is the screen you will use when entering your data into SPSS. At present, this view shows an empty data table in which each of the variables (columns) is labelled 'var'. Before you can enter your data into this data table, you must set up each variable ready to receive the data. The first thing SPSS needs to know is the name of each of your variables; these names will be inserted at the top of the columns of the data table. In addition, you need to give SPSS other important information about each variable. This process of defining the variables is undertaken in the Variable View. If you click on the Variable View tab, you will notice that in this view the columns are headed **Name**, **Type**, **Width**, **Decimals** etc. In the Variable View of the data table, the variables are in rows arranged down the side of the table and each column gives

information about a variable. For example, in the column headed **Name** we will provide the name of each variable, in the **Type** column we are going to tell SPSS what type of variable this is, and so on.

In the Data View of the Data Editor window, each row of the data table represents data from one case (usually a participant) and each column contains data from one variable. However, in the Variable View, the columns and rows are used differently. In this view, each row gives information about one variable. Don't let this confuse you; remember, once you have set up all your variables and are ready to enter your data, you will work in the Data View, where a row is a case and a column is a variable.



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8								
9								
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1 2 3 4 5								 Note that the Variable View tab is highlighted to indicate you are in the Variable View.
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	variable view							
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Data



An alternative way to switch from the Data View to the Variable View is to double-click on the grey header (which will probably be labelled 'var') at the top of the column you wish to define. This will take you to the appropriate row of the Variable View.

Henceforth, when we refer to 'the Data Editor window' without specifying which view, we will be referring to the Data View.

Setting up your variables

If you are not already in the Variable View of the Data Editor, click on the Variable View tab now. We will now use this view to set up each of the variables we need.



It will take a little while to explain what each of the variable settings do, but in practice you can accept the default setting in most cases. If you are in a hurry, just set the **Name**, **Label**, **Values**, **Missing** and **Measure** options for now.

Variable name

The first thing we need to do is to give the variable a meaningful name. Type the name of your first variable into the first row of the Name column. You should choose a variable name that is unique and makes sense to you and which you are not likely to forget. Students often use the variable name 'score'. This is not a good choice as it tells us almost nothing about the variable. Examples of more useful variable names might include 'MemScore' (for participants' scores in a memory experiment), 'Introver' (a participant's introversion score), 'Sex' or 'FamFaces' (the number of famous faces named by a participant). Your variable name can be up to 64 characters long, but keep it much shorter than that so it is easy to read at the top of each column, and in the output. You can't use spaces in a variable name but you can use mixed case (capitals and lowercase characters), and this can be used to improve the readability of the variable name as in the example 'MemScore'. It is also OK to use the underscore character (as in 'Mem_Score'). Variable names must start with a letter of the alphabet (i.e. not a number), and they cannot contain some special characters such as colons, semicolons, hyphens or commas (full stops, the @, #, \$ and _ characters are allowed provided they are not the first character). If you enter an invalid variable name, SPSS will warn you when you try to move from the Name column.



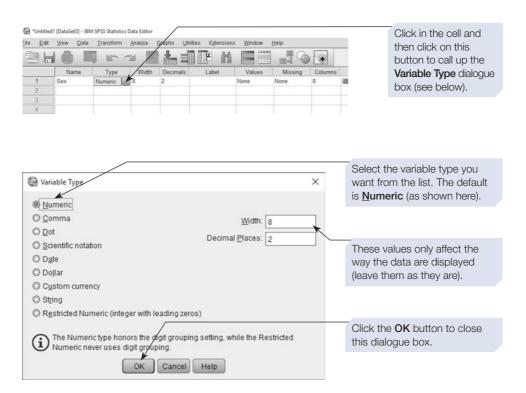
Your variable name should be short and meaningful. It cannot start with a number and cannot contain spaces or some special characters. Upperand lower-case characters and the underscore character (_) are OK and make the name easier to read. For example, the name 'MemScoreTrial1' might be used to code the number of items recalled on trial 1, and 'Q1_1' could be used for the scores from Question 1 Part 1. There are a few other restrictions, but SPSS will warn you if you try to use an illegal name.

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Once you have entered the variable name, use either the mouse (point and click) or the tab key to move to the next column of the table. As you move the cursor, several of the other columns of the table will be filled with either words or numbers. These are the default settings for the variable 'Sex'. You can leave these settings as they are, or you can change some or all of them before moving on to define your next variable. Below we explain each of the settings and how to adjust them.

Variable type

The second column in the Variable View table is headed **Type**. SPSS can handle variables of several different types. For example, variables can be numeric (containing numbers) or string (containing letters) or even dates. The **Type** column is used to indicate what type each variable is. The **Type** will now be set to **Numeric** (unless the default settings have been changed on your copy of SPSS). If you want to change the variable type, click on the word **Numeric** and then click on the button that appears in the cell. This will call up the **Variable Type** dialogue box (see below).



We strongly recommend that, until you are an experienced user, you only use numeric variables. It is easy to use numbers to represent categories and this will save you trouble later (e.g. you can use the numbers 1 and 2 rather than 'm' and 'f' to record the sex of your participants). Until you are a much more experienced SPSS user, you are unlikely to need to use any of the other variable types.



As far as possible, avoid using string variables in SPSS – if you ignore this advice you may regret it later!

Variable width and decimal places

As we saw above, the Variable Type dialogue box allows you to set the <u>W</u>idth and Decimal <u>Places</u> of the variable (see above). Alternatively, these settings can be changed in the third and fourth columns of the Variable View (see below).

These settings adjust the width of the column and the number of decimal places used to display the variable in the Data Editor. These settings do not affect the way the value is stored or the number of decimal places used in statistical calculations. Changing decimal places, however, does affect the number of decimal places shown in SPSS output. With numeric data, the default settings are for a total <u>Width</u> of 8 characters with 2 Decimal <u>Places</u> (e.g. 12345.78). If you attempt to input a data value that will not fit into the width, then SPSS will round it in order to display the value. However, the actual value you entered is stored by SPSS and used in all calculations. One effect of this is that, unless you set Decimal <u>Places</u> to zero, all values, even integers (whole numbers) will be displayed with 2 decimal places. Thus, if you enter a value of '2' in the Data Editor window, SPSS will display '2.00'. This might look a little untidy but is of little consequence and it is only worth changing these settings if you like things to look neat.

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You can change the variable Width and number of Decimals places by changing these values. However, this only alters the look of the table, not the way the values are stored or used in calculations.



Unless you want to change the appearance of the data table, you can probably leave the **Type**, **Width** and **Decimals** settings at their default values.

Variable label

The fifth column in the Variable View table is headed **Label**. This column is used to enter a variable label.

A variable label is simply a phrase associated with the variable name and which helps you to remember what data this variable contains. If you have called a variable something like 'Sex', then you probably don't need to be reminded about what it is describing. If, however, you have a large number of variables, then variable labels can be useful. For example, if you are entering the data from a questionnaire, you might have a variable named 'q3relbef'. In this case a variable label might be invaluable, as it could remind you that this variable coded the responses to question 3 on your questionnaire, which asked about religious belief. You can type in any phrase using any characters that you like, but it is best to keep it fairly short. SPSS will not try to interpret this label; it will simply insert it into the output next to the appropriate variable name when you perform any analysis. It is also noteworthy that when you have to select variables for inclusion in an analysis, SPSS will list the variables by these variable labels, not the variable names. This is another reason to keep the labels short and meaningful. To add a variable label, type it in to the column **Label**.

Variable labels are included in the output produced by SPSS. Although they are not essential, they act as a reminder about the variables and can be helpful when you are interpreting the output. We recommend you take the time to use them whenever appropriate.

Value labels

A value label is a label assigned to a particular value of a variable. You are most likely to use value labels for nominal or categorical variables. For example, we might want to use value labels to remind ourselves that, when entering values for the religion of our respondents, we used the codes 0 = Atheist, 1 = Buddhist, 2 = Christian, 3 = Hindu, 4 = Muslim and 5 = Other. Value labels can also be used in a similar way for ordinal variables, for example when using codes such as 1 = strongly disagree, 2 = disagree, 3 = neither agree nor disagree, 4 = agree, 5 = strongly agree.

Another common use for value labels is with a grouping or independent variable. For example, you might want to compare the reaction time of participants who were tested under one of several different doses of alcohol. You could use a value label to remind yourself that group 1 received no alcohol, group 2 received 1 unit of alcohol and group 3 received 2 units. The value labels you use will be inserted into the SPSS output to remind you what these values mean.

Value labels are entered using the Values column of the Variable View table. At present, this column will probably contain the word **None**. Click on this cell, or use the tab key to move to this cell. As you do so, a button will appear at the right-hand side of the cell. Click on this button to call up the Value Labels dialogue box (see below).

Value labels can be a great help when interpreting SPSS output. Although they are not essential, we recommend that you use them when appropriate. It would not be appropriate to add value labels to some variables. For example, you would not want to add a label to every possible value of a continuous variable such as reaction time. A good rule of thumb is that you should add value labels to all nominal variables and should consider adding them to ordinal variables. They are unlikely to be needed for interval or ratio variables.

32 SPSS for Psychologists

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Missing values



The Missing Values option is an important option to consider. It's good practice to specify at least one missing value even if you don't expect any gaps in your data set.

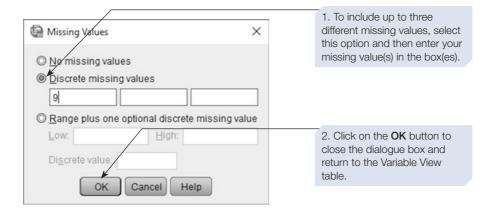
Sometimes you will not have a complete set of data. For example, some participants might decline to tell you their religion or their age, or you might lose or be unable to collect data from some participants, perhaps because of equipment failure. These gaps in the data table are known as 'missing values'. It is vital that SPSS can differentiate between real data and missing values.

When we have a missing value, we need to be able to tell SPSS that we do not have valid data for this variable for this participant. We do this by choosing a value that cannot normally occur for this variable. In the religion example above, we might choose to code religion as 9 when the participant does not state their religion. Thus, 9 is the missing value for the variable religion. The missing value can be different for each variable but does not have to be. The important thing is that this value cannot normally occur for this variable. For age, you could use 99 (unless you are testing very old people). Alternatively, you can use a negative number (e.g. -9), assuming that negative values cannot occur for the variable you have measured.

Before you specify any missing values, the cell in the **Missing** column of the Variable View table will contain the word 'None'. To specify a missing value, click in the **Missing** column of the Variable View table. A button will appear at the right-hand end of the cell. Click on this button to call up the **Missing Values** dialogue box (see below).

Note that we can see the value labels we added in the previous step. *Untitled1 [DataSet0] - IBM SPSS Statistics Data Editor File Edit View Data Transform Analyze Graphs Utilities Extensions Window Help μ âΆ R Label Values Missing Name Туре Width Decimals Colu 1 Sex Numeric 8 2 Sex of Participant {1.00, Male}. None 8 2

> Click on this button in the **Missing** cell to call up the **Missing Values** dialogue box (see below).



SPSS allows you to specify the missing values in several ways:

- 1. <u>No missing values</u>: This is the default setting for this dialogue box. If this option is selected, SPSS will treat all values for this variable as valid.
- 2. <u>Discrete missing values</u>: This option allows you to enter up to three discrete values. For example, 7, 9 and 11 could all be set as missing values by selecting this option and entering the values in the three boxes. If you choose to use only one missing value to code your missing data, enter it into the first of the three boxes (as we've done above).
- 3. <u>Range plus one optional discrete missing value</u>: This option allows you to indicate that a range of values is being used as missing values. For example, selecting this option and entering the values 7 and 11 in the Low and <u>High</u> value boxes would instruct SPSS to treat the values 7, 8, 9, 10 and 11 as missing values. If, in addition to this range of values, the value 0 was typed into the **Discrete value** box, then SPSS would treat the values 7, 8, 9, 10, 11 and 0 as missing.



The **Missing Values** dialogue box does not allow you to label the missing values. Once you have entered them, however, you can label them in the **Value Labels** dialogue box if you wish. For example, you could add labels to show that 9 = unanswered and 10 = illegible.

You should always assign at least one missing value. In most cases this is all you will need, but occasionally we might want to differentiate between several different types of missing data. For example, by assigning different missing values we could distinguish between an unanswered question, an indecipherable answer and an item that was not applicable to this participant.

Select either the <u>D</u>iscrete missing values or the <u>R</u>ange plus one optional discrete missing value option and enter the value(s) you have chosen to represent missing values. If you have only one missing value, enter it into the first of the <u>D</u>iscrete missing values boxes. Now click on the OK button to return to the Variable View table.



Many of the problems encountered by novice SPSS users stem from a failure to think about missing values and to specify at least one missing value for each variable.

Column format

The next column of the Variable View table is labelled **Columns**. This entry in the table is used to specify the width of the column that the variable occupies in the Data View table of the Data Editor window. You can leave this value at its default setting unless you want to change the appearance of the Data View table. You may, for example, want to fit more columns onto the screen in order to see more variables without having to scroll. In this case you could reduce the width of each column. To adjust the

settings, click on the cell and then use the up and down buttons that will appear at the right-hand end of the cell to adjust the value. You can look at the effect of the change you have made by switching to the Data View.

You can more easily change the width of a column by dragging it with the mouse. Switch to the Data View and then move the mouse to the top row of the table and hover over the border between two columns. The mouse pointer will change to a double-headed arrow. You can now click and drag to adjust the width of the column. The changes you make here will be reflected in the Variable View settings.

Column alignment

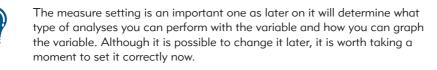
The column of the Variable View labelled **Align** allows you to specify the alignment of the text within the cells of the Data View table. This setting has no effect on the operation of SPSS and only changes the appearance of the Data View table. The default setting is right alignment, in which the decimal points of the values in the column are lined up. In left alignment, the values are flush to the left-hand end of the cell. In centre alignment, the values are centred in the cell (and thus the decimal points will not necessarily line up).

If you wish to change the column alignment, click in the **Align** cell and then click on the menu button that will appear in the cell and select the required alignment from the drop-down list (see below).

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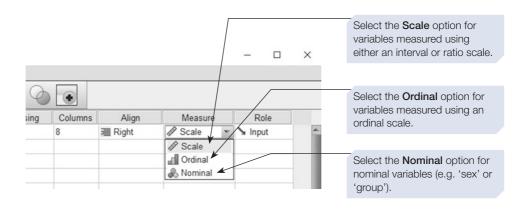
Measure

This is used to specify the level of measurement for the variable. SPSS offers three options: Nominal, Ordinal and Scale.



Psychologists usually distinguish four levels of measurement: nominal, ordinal, interval and ratio (see Chapter 1, Section 1). SPSS does not distinguish between interval and ratio data and uses the term **Scale** to describe a variable measured using either of these levels of measurement.

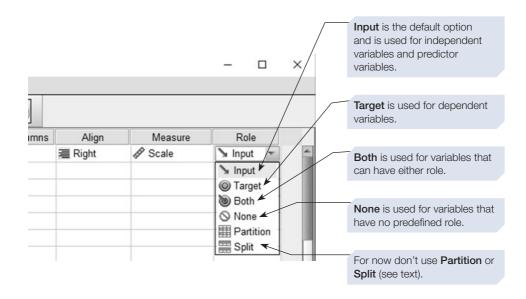
To set the measurement option, click in the **Measure** cell of the Variable View table and then click on the button that appears in the cell and select from the dropdown list (see below). The relevant icons will appear in the SPSS dialogue boxes as a reminder of the level of measurement of this variable.



If you open a data file created using an earlier version of SPSS or some other program, the **Measure** will be set for you – variables with value labels will be set as **Nominal**, while variables with only a small number of values will be set as **Ordinal**. All other variables will be set as **Scale**.

Role

The final column of the Variable View table is called **Role**. This is a fairly recent addition to SPSS and is intended to help users select appropriate analyses. The idea is that you can identify a group of variables as having a particular role in your analysis. For example, you might have several variables that are going to be used as dependent variables and others that will be independent variables. If you indicate this through the **Role** setting, these variables will be automatically assigned when you come to undertake some kinds of analyses. The six **Role** options are **Input** (for independent variables), **Target** (for dependent variables), **Both** (for variables that may take on either role), **None** (for variables with no role set), **Partition** (used to divide a data set so different parts can be used for different purposes) and **Split** (used by other analysis packages). We recommend novice users leave **Role** at its default setting of Input, as we won't be using this function in our analyses.



Remember, for most variables, you can accept most of the default settings. In the first instance just set the variable name, variable label missing values and measure – and add value labels if appropriate. It takes much longer to explain this process than it does to do it!

Check your settings

Once you have completed the definition of your first variable, switch to the Data View by clicking on the Data View tab at the bottom left-hand corner of the table. You will now see the name of your new variable appear at the top of the appropriate column of the Data Editor window (see below). If you changed the column width and/or alignment, you will see the effect of these changes.

Now switch back to the Variable View and repeat this process for each of the variables required for your data file.

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Eile Edit	View Data	Iransform	Analyze	<u>G</u> raphs	Utilities	In the Data View the new variable name appears at the top of the column. This column is now ready to accept data.
	A Sex	var	var	var	Va	
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5 6 7					-	

Copying variable settings

It is easy to copy the settings from one variable and 'paste' these onto one or more new variables.

Suppose, for example, that you have administered a questionnaire that contains 20 items. Each item consists of a statement to which the participant is asked to respond by choosing from one of several options, such as 'Strongly Disagree', 'Disagree', 'Neither Agree or Disagree', 'Agree' and 'Strongly Agree'. In our SPSS data table, each question will be represented by a variable, which we might name Q1, Q2 etc. For each of these variables, it would be useful to enter the value labels 1 = 'Strongly Disagree', 2 = 'Disagree' and so on. This would be rather time-consuming. However, if we enter these value labels for the first variable, we can then move the cursor to the Values cell of the Variable View table and select Copy from the Edit menu (or right click and select Copy). If we now click in the cell (or select the range of cells) we want to copy these labels to and select Paste from the Edit menu (or right click and select Paste), the value labels will be copied to all the selected cells.

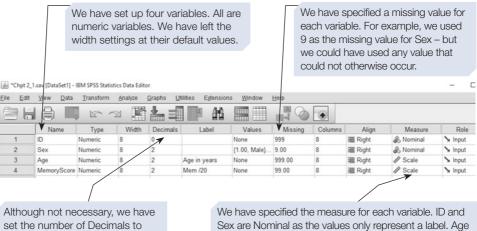
Section 3: ENTERING DATA

zero for the variable ID.

A first data entry exercise

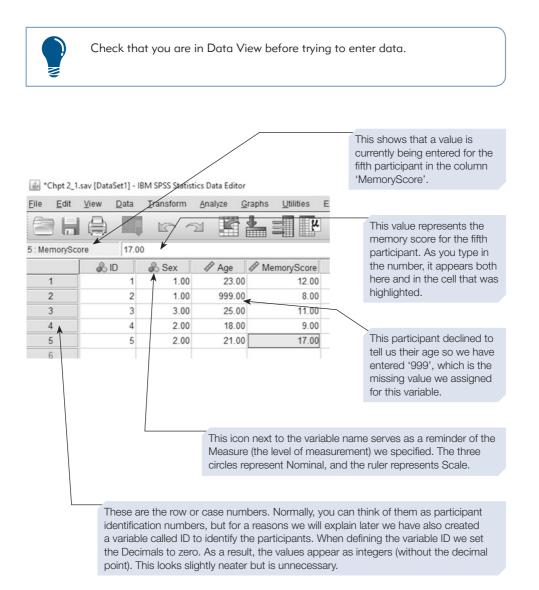
As a data entry exercise, we will enter the data from a simple study in which we have recorded each participant's Sex (coded as 1 = Male, 2 = Female and 3 = Other), Age (in years) and the MemoryScore (number of words recalled from a list of 20) of each participant. In addition we have assigned each participant an ID number.

Before we can enter these data, we need to define the four variables to be used (see the previous section for details of how to define a variable). As sex is a nominal variable, we will use value labels to remind ourselves what the values 1, 2 and 3 represent. We also need to set the Measure for each variable. In this case ID and Sex are nominal while Age and MemoryScore are scale variables. See below.



Sex are Nominal as the values only represent a label. Ac and MemoryScore are set to Scale. Note the different icons associated with these settings. These icons will appear next to the variable name in the Data View. Once the variables have been defined, we can switch to the Data View (click on the **Data View** tab at the bottom left-hand corner of the data table) and begin entering the data. You can copy the data for the first five participants from the screenshot shown below.

Click on the top left-hand cell of the table (ensure that you are at the top left-hand corner of the window by checking the scroll bars). This cell will become highlighted (it will have a bold border). Any number you now type will appear in the bar above the variable names at the top of the window. If you press the Enter key or the Tab key, or use the mouse or cursor keys (up, down, left and right arrows) to move to another cell, this number will be inserted into the cell.



Moving around the Data Editor window

You can use either the mouse or the cursor keys to move around the data table. Alternatively, you can press the Enter key to move down to the next participant for the current variable, or the Tab key to move across to the next variable for the current participant.

It is best to enter the data one participant at a time, working across a row of the data table. For example, you should enter the sex, age and memory score for the first participant in row one, then for the second participant in row two and so on. If you enter the data a column at a time working down the columns (e.g. the sex of all the participants first, then their ages etc.), you may make a mistake, and such an error is likely to result in the data from one participant being assigned to another participant.

Once you have entered all your data into the data table, you should carefully check there are no errors. Cross-checking the data file against the original record of the data is an important stage in the process of analysis. Either cross-check the original records against the data on the screen or against a printout of the data (see Chapter 14, Section 5 for details of how to print a copy of your data).

You may accidentally enter an empty row of data, which will appear as a row of cells filled with dots. If this has happened, it is worth taking the time to remove the blank line(s) as SPSS will interpret each blank line as a participant for whom you have no data. Thus, SPSS will tell you it has more cases than you expect. To delete the blank case, click on the case number associated with the extra row; the case will become highlighted. Now either press the delete key or right click and select **Clear**.

Sometimes, novice users of SPSS panic that they have 'lost' their data because they cannot see it on the screen. This is often because the data have scrolled out of view. Check that the scroll bars are set to the top left-hand corner of the window.

The value labels button

If you have assigned value labels to one or more of your variables, you can choose whether you want SPSS to display the values you enter or the labels associated with the values. For example, in this file, we have assigned the value labels 'Male', 'Female' and 'Other' to the values 1, 2 and 3 of the variable 'Sex'. SPSS can either display the values (i.e. the numerals '1', '2' or '3') or the labels. Clicking on the Value Labels button on the toolbar of the Data Editor window will toggle between these two display states (see below). This option affects only the way the data are displayed in the Data Editor window and not the way they are entered or analysed.

Furthermore, if you select to display the variable labels rather than the data values, then SPSS will help with the data entry process by providing a drop-down list of possible values. See below to see how this works.

<u>F</u> ile	<u>E</u> dit	View	<u>D</u> ata	Transform	<u>A</u> nalyze	<u>G</u> raphs <u>U</u> tilities	Extensions	Window	<u>H</u> elp	
			00,		1	¥ = 4	<u>Ana</u>		A 14	
5 : Me	morySc	ore	17.00)					1	
		8	D	& Sex	Age Age	A MemoryScore	var	var	var	var
1	1		1	1.00	23.00	12.00				
	2		2	1.00	99.00	8.00			/	
	3		3	3.00	25.00	11.00		/	/	
	4		4	2.00	18.00	9.00				
	5		5	2.00	21.00	17.00				
	c									

Click on the **Value Labels** button to toggle between displaying the values entered (as shown here) and the associated labels (see below).

*Chpt 2_1.sav [DataSet1] - IBM SPSS Statistics Data Editor

<u>F</u> ile	<u>E</u> dit	View	Data	Transform	Analyze G	raphs <u>U</u> tilities	Extensions	Window	Help		
6						*	81 8		A 14		
5:Sex			2.00	D					1		
		&1	D	& Sex	Age 🖉	A MemoryScore	var	var	var	var]
1			1	Male	23.00	12.00					
2	!		2	Male	99.00	8.00					
3			3	Other	25.00	11.00					
4			4	Female	18.00	9.00					
5			5	Female 🔽	21.00	17.00					
6				Male							
7				Female							
8				Other							
0											
						In this mode, if click in a cell, S you a drop-dov value labels ass this variable.	PSS will o vn list of t	offer he	L	is depresse	Yalue Labels button ed, the values in the e are replaced by ated labels.

When you have finished entering your data and have carefully checked it, you should save a copy of the data file. We describe how to save the data file in Section 4.



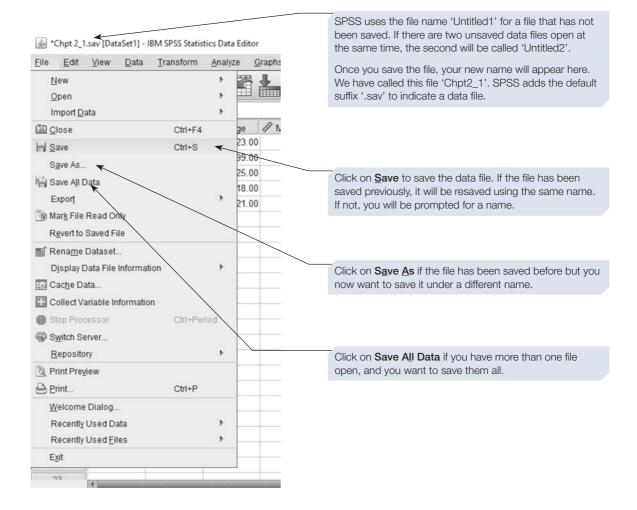
It is possible to have several data files open at a time; however, this can become confusing, so we recommend that you work with just one data file until you have become familiar with SPSS.

Section 4: SAVING A DATA FILE

You will have spent a lot of time entering your data, so remember to save the data file. There is currently no autosave feature in SPSS, so if you are entering a large amount of data, it is a good idea to save the file every few minutes.

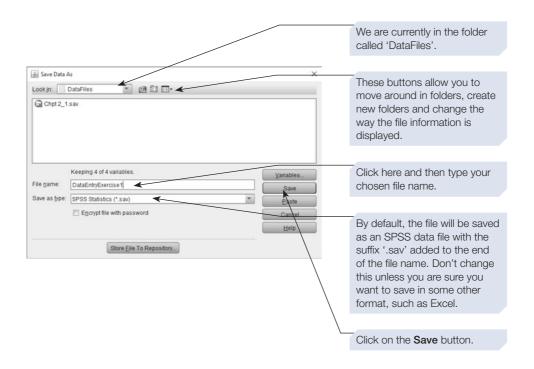
To save the data to a file

Click on the menu item File at the top of the screen. Now click on either Save or Save As.



Select <u>Save</u> to resave the file using the existing name. The resaved file will replace the old version. If the file has not been saved previously, or if you click on <u>Save As</u>, you will be presented with the <u>Save Data As</u> dialogue box (see below).

Type the name for the file into the **File name:** box. The file name you choose should be reminiscent of the study from which the data originated (for example, 'DataEntryPractice'). You should not use a full stop in the file name and should not attach a suffix to the file name. By default, SPSS will attach the suffix '.sav' to the name you enter. Do not change this suffix, or SPSS might not recognise the file as a data file. Before you click the <u>Save</u> button, check which drive and which directory the file is going to be saved to. You may want to save the file to a different drive, or to a USB stick or cloud drive.



Section 5: OPENING A DATA FILE

To open a data file, follow the instructions below:

1. Ensure that the Data Editor window is the active window. If this is not the case, select the SPSS icon from the taskbar at the bottom of the screen.

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		V Output (Document) - IBM S	DataEntryExercise1.sav (D
©≊ € ∎	o 🖻 🗐 🏷 5 🖻		2. 5
	Select the SPSS icon from the taskbar.		If you have more than one SPSS window open, select the Data Editor window.

- 2. To open a different data file, click on the **File** menu.
- 3. Select Open.
- 4. Select Data. The Open File dialogue box will now appear (see overleaf).

DataEntryExercise1.sav [DataSet]		C Ctatic	tics Data Editor			2. Click on the File menu.
	-	Analyz		Utilities	Exter	3. Select <u>O</u>pen . 4. Choose D<u>a</u>ta .
 Rename Dataset Display Data File Information Cache Data Collect Variable Information 		•				

Cook jn: DataFiles Chpt 2_1.sav DataEntryExercise 1.sav	5. Move to the correct folder either using this drop-down list or by selecting the folder from the list below.
File name: Open Files of type: SPSS Statistics (*.sav, *.zsav) Encoding: Cancel Help	6. Click on the name of the file and then click the Open button. Alternatively, double-click the name of the file to open.
Betrieve File From Repository	You can choose which type of file to show in the box above. By default, only SPSS data files will be shown.

SPSS can read and write files of various formats, including Excel spreadsheet files. In Chapter 14 we describe how to import Excel spreadsheets into SPSS and how to save SPSS data files as Excel spreadsheets.

If the file you are looking for has a suffix other than '.sav', SPSS will not recognise it as a data file and will not display it in the list of data files. If you can't find the file you are looking for and think that it may have been saved with some other file name suffix, click on the button at the right-hand end of the **Files of type** box and select 'All files (*.*)' from the list of file types offered. All the files in the current directory, regardless of type or suffix name, will now be displayed in the dialogue box. If you find that your data file was saved with some other suffix, either resave the file with the correct suffix or make a copy and change the suffix to '.sav'.

Section 6: DATA ENTRY EXERCISES

In this section, we are going to practise entering data from studies with two different types of experimental design. Later in this chapter and in subsequent chapters we will use these data files to demonstrate other procedures. Take the time to complete these exercises, as they will help you to appreciate the way that the design employed in a study influences the shape of the data file. When you have completed these two data files, compare them to the files we made, which are shown in Section 7.

Data from an unrelated (independent groups) design

As explained in Chapter 1, the independent groups design is an unrelated design in which we compare the performance of two or more groups of different participants. In the example below, we have used this design to investigate the effect of a mnemonic instruction given to a group of participants before they were asked to learn a total of 20 words. The dependent variable was the number of words correctly recalled.

RODENTS IN SPACE: A SIMPLE MEMORY EXPERIMENT

Twenty-one first-year undergraduates participated in a simple memory experiment designed to investigate the effect of a mnemonic strategy on memory for paired words. The participants were randomly divided into two groups. All participants were given two minutes to memorise a list of 20 words presented in pairs. All the participants were told to memorise the words, but those in the mnemonic instruction group were advised to try to form a mental image to link the two words in a pair (e.g. for the word pair Rocket – Hamster, a participant might imagine a small furry rodent being launched into space). The participants in the other group, the non-mnemonic group, were not given this instruction. After learning the words for two minutes, the participants were then required to complete some simple mental arithmetic problems for two minutes. Finally, all participants tried to recall the words in any order. The number of words correctly recalled was recorded. The data are summarised below.

Memory scores (out of 20) for participants in the mnemonic instruction group:

20, 18, 14, 18, 17, 11, 20, 18, 20, 19, 20

Memory scores (out of 20) for participants in the non-mnemonic group:

10, 20, 12, 9, 14, 15, 16, 14, 19, 12

Using these data, attempt to do the following:

- 1. Set up an SPSS data file to record these data. Give appropriate names to the variables you are using.
- 2. Apply value and variable labels where appropriate and specify the level of measurement and the missing values for each variable.
- 3. Enter and check the data, then save the file using an appropriate file name.
- 4. Ensure that you can reopen the file.
- 5. Compare the data file you have constructed to the one illustrated in Section 7.

Data from a related (repeated measures) design

In a repeated measures design, every participant is exposed to each condition and thus contributes a data point for each level of the independent variable. This will be reflected in the structure of the data file, which will have a column for each level of the independent variable. In the example below, we have used this experimental design to investigate mental representation.

COMPARING MENTAL IMAGES

If you ask someone the question 'How many windows are there in the front of your home?', most people will report that they attempt to answer the question by 'inspecting' a mental image of the building. Does this mean that we store information in the form of mental images? Some psychologists think not, arguing that information is actually stored in a more abstract form and that our experience of inspecting mental images is illusory. However, several lines of evidence support the idea that we are able to manipulate information utilising a form of representation that shares many qualities with mental images. This experiment is modelled on one such line of evidence.

Imagine you were asked to decide whether or not a lion was bigger than a wolf. You could make your decision by recalling abstract information about the average size of each animal. Alternatively, you could form a mental image of these two animals standing side by side and decide which was the taller. If you adopted the mental imagery approach, then you might expect the decision to take longer when the two animals were of similar sizes than when they were very different sizes. If the decision were based on a more abstract form of representation, then you would expect the relative size of the animals to have no impact on the speed of the decision. Thus, if it takes longer to compare the size of two similar-sized animals than two dissimilar-sized animals (e.g. lion and mouse), this will offer some support for the idea that these decisions are based on the manipulation of image-like forms of mental representation.

In our experiment, each of the 16 participants undertook 20 trials. In each trial, the participant was presented with a pair of animal names and had to decide as quickly as possible which animal was the largest. The dependent variable was the time to make this decision (in milliseconds – ms). For half of the trials the difference in size between the two animals was large (e.g. Mosquito – Elephant), and for the other half of the trials the difference in size was small (e.g. Horse – Zebra). In the data table below, we have recorded each participant's mean decision time (in ms) for the large- and small-difference trials.

DATA

Participant	Large diff.	Small diff.
1	936	878
2	923	1005
3	896	1010
4	1241	1365
5	1278	1422
6	871	1198
7	1360	1576
8	733	896
9	941	1573
10	1077	1261
11	1438	2237
12	1099	1325
13	1253	1591
14	1930	2742
15	1260	1357
16	1271	1963

- 2. Apply value and variable labels where appropriate and specify the level of measurement and missing values.
- 3. Enter and check the data, then save the file to disk using an appropriate file name.
- 4. Ensure that you can reopen the file.
- 5. Compare the data file you have constructed to the one illustrated in Section 7.

Section 7: ANSWERS TO DATA ENTRY EXERCISES

A data file for an unrelated (independent groups) design

Below is a screenshot of the data file we constructed for this simple memory experiment. Your data table might not look identical but should have the same basic characteristics. Note that there are two critical variables in this design. The first is a nominal variable (or grouping variable) that we have used to record whether the participant was in the Mnemonic or Non-mnemonic group. Thus, it indicates the level of the independent variable (or factor) the participants belong to. The other critical variable is a ratio variable and has been used to record the dependent variable, the number of

<u>F</u> ile <u>E</u> dit	<u>V</u> iew <u>D</u> ata	Transform A	nalyze <u>G</u> raphs	Util
		5 2		
4 :				
	& ParticipantID	& Condition	A MemoryScore	
1	1	1	20.00	
2	2	1	18.00	
3	3	1	14.00	
4	4	1	18.00	
5	5	1	17.00	
6	6	1	11.00	
7	7	1	20.00	
8	8	1	18.00	
9	9	1	20.00	
10	10	1	19.00	
11	11	1	20.00	
12	12	2	10.00	
13	13	2	20.00	
14	14	2	12.00	
15	15	2	9.00	
16	16	2	14.00	-
17	17	2	15.00	
18	18	2	16.00	
19	19	2	14.00	
20	20	2	19.00	
21	21	2	12.00	
22	1			

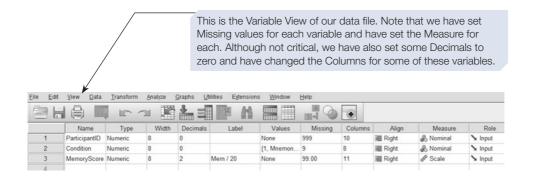
🐏 *DataEntryExIndGrps.sav [DataSet2] - IBM SPSS Statistics Data Editor

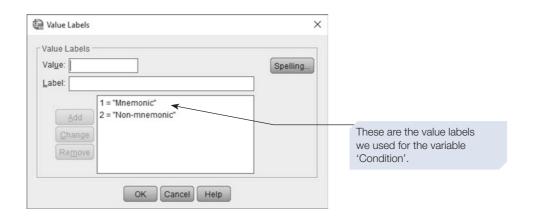
We have called this variable 'MemoryScore'. Being lazy, we have left the variable width as 8 with 2 decimal places, but we have increased the Columns setting to 10 so that the variable name fits. This isn't essential.

We have used the name 'Condition' for this variable. Because it is a nominal variable, we have used value labels to show that the value 1 indicates the participant was in the mnemonic condition, and the value 2 that they were in the non-mnemonic condition (see screenshot of the Value Labels dialogue box below). We have set Decimals to 0 (to display no decimal places), but this isn't essential.

For this screenshot, we did not have the **Value Labels** button depressed; if it is depressed on your system, then your screen will show the value labels rather than the values.

We have included a variable called 'ParticipantID', which gives each participant a number. This is not essential but is good practice; we will explain why later. We have also increased the column width so the variable name fits in. words each participant recalled. In addition to these two variables, we have also included a variable called 'Participant', which assigns a number to each participant. This is good practice. If you have the **Value Labels** button (on the toolbar) depressed, then the 'Condition' column will display value labels rather than values (i.e. Mnemonic or Non-mnemonic rather than 1 or 2).





Remember, the data file constructed for an experiment that employed an independent groups design will always require a nominal variable that is used to indicate the condition under which each participant was tested.

The data file for a related (repeated measures) design

Below is a screenshot of the data file we constructed to record the data from our mental imagery experiment. Your data table might not look identical but should have the same basic characteristics. As with the independent groups design, there are a total of three variables, but in this case two are used to record the performance of the participants

when tested under each experimental condition. As this is a repeated measures design, participants were all tested under both conditions, so we have two data points for each participant. In this design, there is no need for a nominal variable.

Compare this data file to the one for the independent groups design. Make sure that you understand why these two files have a different structure.



Remember, in a data file constructed for a repeated measures design, there must be a variable (and hence a column in the data file) for each condition.

<u>F</u> ile <u>E</u> dit	<u>V</u> iew <u>D</u> ata	<u>T</u> ransform <u>A</u> r	nalyze <u>G</u> raph	
8 H		5 3		We have called this variable 'SmallDif' and are using it to
16 : SmallDif	1963.0	10		record the participants' mean decision times on the small- difference trials. We have left the variables Width and
	& ParticipantID		SmallDif	Columns at their default values.
1	1.00	936.00	878.00	
2	2.00	923.00	1005.00	
3	3.00	896.00	1010.00	We have called this variable 'LargeDif' and are using it to
4	4.00	1241.00	1365.00	record the participants' mean decision times on the large-
5	5,00	1278.00	1422.00	difference trials. All other variable settings are left at their
6	6.00	871.00	1198.00	default values.
7	7.00	1360.00	1576.00	
8	8.00	733.00	896.00	
9	9.00	941.00	1573.00	
10	10.00	1077.00	1261.00	We have used the variable name 'ParticipantID' for the
11	11.00	1438.00	2237.00	participant number variable. This is not essential but is good practice; we will explain why later.
12	12.00	1099.00	1325.00	good practice, we will explain why later.
13	13.00	1253.00	1591.00	
14	14.00	1930.00	2742.00	Participant 16, for example, had a mean decision time of
15	15.00	1260.00	1357.00	1271 ms for the large-difference trials and 1963 ms for the small-difference trials.
16	16.00	1271.00	1963.00	
17				

Section 8: CHECKING AND CLEANING DATA FILES

In this chapter we have described how to enter and save data in SPSS. Having done that, your next task should always be to carefully check the file and 'clean' it to remove the errors it will inevitably contain.

Never start analysing your data before you have checked and cleaned it. Ignoring this advice could have one of two outcomes. If you don't check the data and later find an error, you will have to repeat all the analyses you have undertaken so far. This can cost you hours, days or even weeks of work. Alternatively, you may never notice the errors. The consequences of this could include a failed course, a misleading publication or even an incorrect policy decision. Carpenters have a saying: 'measure twice and cut once'. The equivalent rule in statistics should be 'check your data twice and analyse once'.

Careful data checking should lead to the detection of all the errors in your data. These errors can then be corrected in a process known as 'data cleaning'. Checking and cleaning a data file may be a quick job, or it may take a great deal of time. Checking and cleaning the data file for a simple experiment may only take a few minutes. On the other hand, it may take a researcher weeks of work to check and clean the data files for a big project, which involves combining data from different sources across thousands of participants and hundreds of variables. In some projects, the data checking and cleaning takes much longer than the analysis of the final data set.

The best way to check your data for errors is to use some of the simple descriptive statistics procedures we describe in Chapter 3. These simple analyses will allow you to detect unexpected, implausible and invalid values in your data file and correct them. Because we need to use descriptive statistics commands to check your data, we are going to introduce these commands first before explaining in Chapter 3, Section 6 how to use the commands to check the data. Remember that, in practice, you should always do your data checking and cleaning first before attempting any final statistical analysis of the data.

To avoid wasting hours of work, always carefully check and clean your data before undertaking any analysis. We describe how to check and clean a data file in Chapter 3, Section 6.

Summary

- This chapter showed you how to create a data file in SPSS.
- Sections 1 and 2 explained the different parts of the Data Editor window and showed how to define a variable.
- Section 3 walked you through the process of setting up a data file, and Sections 4 and 5 showed how to save and open a data file.
- Sections 6 and 7 provided two data entry exercises to highlight differences between the data files used to code the data from unrelated and related designs. In Chapter 3, we will use these data files to learn how to explore data using descriptive statistics.
- Section 8 described the importance of taking time to check and clean your data file. The procedures for doing this will be described in Chapter 3, after we have introduced you to some of the descriptive statistics commands you will need to use to do this checking.
- For guidance on how to incorporate your data file in a report, or print it, see Chapter 14.

 \prec

Exploring, cleaning and graphing data in SPSS

In this chapter

- Descriptive statistics
- The Descriptives command
- The Viewer window
- The Frequencies command
- The Explore command
- Using descriptive statistics to check your data
- Introducing graphing in SPSS
- Chart Builder
- Graphboard Template Chooser



SPSS for Psychologists online

Visit macmillanihe.com/harrison-spss-7e for data sets, online tutorials and exercises.

Section 1: DESCRIPTIVE STATISTICS

- Descriptive statistics allow us to summarise data by using numbers or graphs and charts. Descriptive statistics can help us understand important aspects of a data set, and critically can help identify errors in our data – as described in Section 6.
- Summary descriptives can accurately describe a large volume of data with just a few values. Common summary descriptives include measures of central tendency (e.g. mean, median and mode), confidence intervals, and measures of dispersion (e.g. range, minimum and maximum, interquartile range, standard deviation and variance). In this chapter we show some of the ways to calculate these values using SPSS.

- Data can also be described through the use of graphs or charts. In this chapter we show how to use SPSS to produce some types of graphs, and briefly introduce others that are covered in more detail in subsequent chapters.
- Most inferential commands in SPSS include options to generate some relevant descriptive statistics, but SPSS also includes several commands specifically designed to produce descriptive statistics. In this chapter we describe the use of three of these commands, <u>Descriptives</u>, <u>Frequencies</u> and <u>Explore</u>.
- The <u>Descriptives</u> and <u>Frequencies</u> commands produce a variety of useful descriptive statistics, but these are calculated across all participants and cannot easily be broken down by one or more grouping variables.
- The <u>Explore</u> command can provide descriptive statistics broken down by one or more grouping variables.
- We use the first of these commands, the **Descriptives** command, to introduce you to the Viewer window, which displays all the output in SPSS.
- A research report should always include appropriate descriptive statistics to assist the reader to understand the data and any effects reported.

Section 2: THE DESCRIPTIVES COMMAND

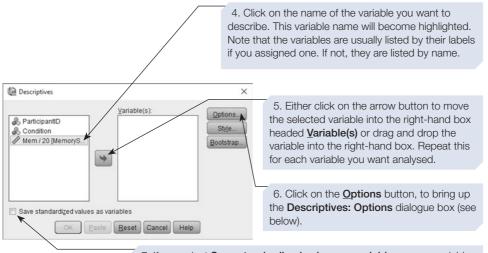
In this section we will use the data from the independent groups study of memory, which you entered and saved in Chapter 2, Section 6. Employing this data set we will learn how to use the **Descriptives** command to produce some basic descriptive statistics, to illustrate some principles of working in SPSS, and how the Viewer window is used to study output.

To obtain output from the Descriptives command

Once you have entered, checked and saved the data from our simple independent groups memory experiment (see Chapter 2, Section 6) follow the steps on the next page:

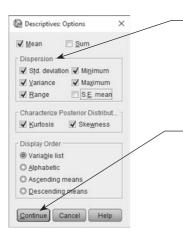
1.	Select Analyze		2. Select Descriptive Statistics	3.	Select Descriptiv	es 🗍
			PSS Statistics Data Editor			
ile <u>E</u> dit	<u>V</u> iew <u>D</u> ata	Transform	Analyze Graphs Utilities	Extensions	Window Help	_/
			Reports Descriptive Statistics	4	Erequencies	P
:			Bayesian Statistics	•	Descriptives	
	& ParticipantID	& Condition	Ta <u>b</u> les		-Q Explore	
1	1		Compare Means			
2	2		General Linear Model	*	Crosstabs	
3	3		Generalized Linear Models		TURF Analysis	
4	4		Mixed Models		Ratio	
5	5		Correlate		P-P Plots	
6	6		Regression		Q-Q Plots	
7	7					
8	8		L <u>og</u> linear			
9	9		Neural Networks			
10	10		Classify			
11	11		Dimension Reduction			
12	12		Sc <u>a</u> le	•		
13	13		Nonparametric Tests			
14	14		Forecasting			
15	15		<u>S</u> urvival	•		
16	16		Multiple Response			
17	17		Missing Value Analysis			
18	18		Multiple Imputation			
19	19		Complex Samples	*		
20	20		₽ Simulation			
21	21		Quality Control			
22			Spatial and Temporal Modelin	ng 🕨		
22	1		Direct Marketing	*		
Data View	Variable View		IBM SPSS Amos			

SPSS will now present you with the **Descriptives** dialogue box shown below. This dialogue box follows a basic structure common to most commands in SPSS. Within the dialogue, there are two boxes: the left-hand box lists all the variables in the data file, listed either by variable name or variable label (you can choose whether to view variable names or labels by right-clicking on this list); while the right-hand box, which is headed $\underline{V}ariable(s)$, lists all the variables which will be analysed by this command.



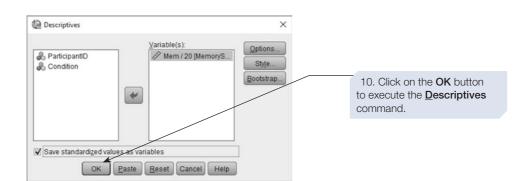
7. If you select **Save standardized values as variables**, a new variable will be created in the data file, which will contain the *z*-score for each value of the variable/s that you select. See below for an example.

You can select more than one variable by holding down either the <shift> key or the <Ctrl> key while clicking on the names of the variables. If you <shift> click the first and last variables in a list, all the variables in between will be selected. If you hold down the <Ctrl> key while clicking on the name of a variable, that individual variable will be selected or deselected.



8. Click in the boxes next to the names of the descriptive statistics you require. A tick in a box indicates that a descriptive has been requested. In this example, we have requested the Mean, Standard deviation, Variance, Range, Minimum, Maximum, Kurtosis and Skewness.

9. Once you have chosen your descriptive statistics, click on the <u>Continue</u> button to return to the **Descriptives** dialogue box.



Finally, click on the **OK** button to execute the **Descriptives** command.

We have only requested descriptive statistics for 'MemoryScore', which was the dependent variable in our study. You can request statistics for any of your variables – just drag them into the right-hand box. However, just because you can request descriptive statistics doesn't mean you should. If a variable is nominal or ordinal, then most of the statistics given by the **Descriptives** command are unsuitable – for example, mean Sex or mean Condition is not meaningful. In these cases the **Frequencies** command (see Section 4) provides more appropriate descriptive statistics.

Most of the results of the **Descriptives** analysis will be presented in the Viewer window, which we will explore in Section 3. However, if you selected the **Save stand-ardized values as variables** option (see step 7 above), a change will also be made to your data file, as shown below.

le <u>E</u> di	t <u>V</u> iew <u>D</u> ata	Transform A	nalyze <u>G</u> raphs	Utilities Extensions
ZMem	oryScore Print	5 3		∎ Pr HA
	& ParticipantID	& Condition	A MemoryScore	ZMemoryScore
1	1	1	20.00	1.09682
2	2	1	18.00	.54841
3	3	1	14.00	54841
4	4	1	18.00	.54841
5	5	1	17.00	.27420
6	6	1	11.00	-1.37102
7	7	1	20.00	1.09682
8	8	1	18.00	.54841
9	9	1	20.00	1.09682
10	10	1	19.00	.82261
11	11	1	20.00	1.09682
40	40	2	40.00	4 64600

If you selected the option Save standardized values as variables (see step 7 above), then this new variable will be added to our data file.

CHAPTER 3

A *z*-score is a useful way of describing performance on a test. The *z*-score indicates by how many standard deviations a score differs from the mean score. A *z*-score of 0 indicates that the participant's test score is equal to the mean for the test, while a *z*-score of +1 indicates that the test score is one standard deviation above the mean, and a *z*-score of -2 indicates performance is two standard deviations below the mean. In the example above, we can see that participant 6 recalled 11 words. This is hard to interpret, but the *z*-score of -1.37 tells us that this performance is 1.37 standard deviations below the mean performance on this test. *Z*-scores are useful for three reasons: First, they allow us to easily identify participants who are scoring above or below a threshold value; second, they allow us to track changes in performance over time.

In Section 3, we will explain the workings of Viewer window before examining the output of the **Descriptives** command.

Section 3: THE VIEWER WINDOW

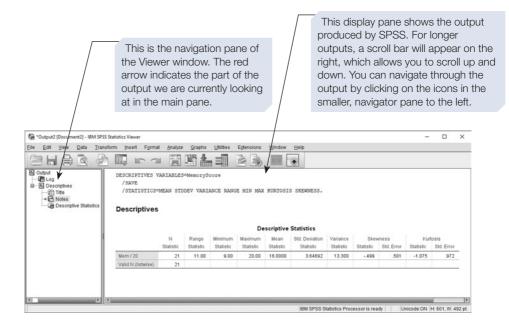
Once you click on the own button to execute a command, SPSS will open a new window called the Viewer window. The Viewer window is the SPSS window in which the results of all analyses are presented. If a Viewer window is already open, SPSS will append the output to the bottom of the window. If there isn't a Viewer window open, SPSS will open one for you.

The Viewer window comprises two distinct parts or 'panes'. The left-hand pane acts as a 'navigator' or 'outline'. This is a bit like a table of contents that lists all the components of the output that are shown in the larger pane. Clicking on an icon in the navigator pane moves you to that part of the output in the main or 'display' pane. This is useful when you have used a command that produces large amounts of output, or when you have executed several different commands.

Before examining the output in detail, we need to highlight several characteristics of the Viewer window.



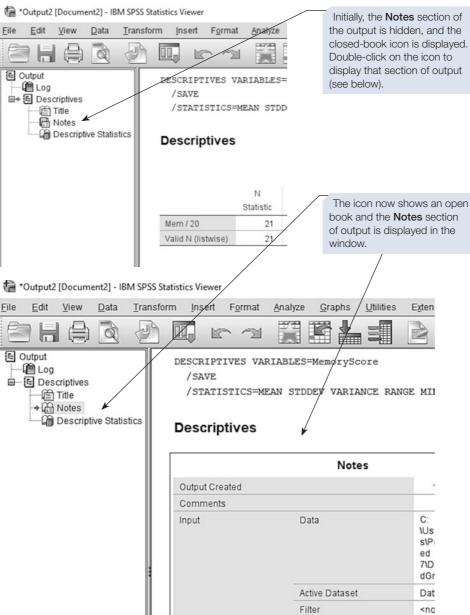
If you find the icons or text in the output window too small to see clearly, you can enlarge them by selecting **Outline Size** from the <u>View</u> menu. Now select <u>Small</u>, <u>Medium</u> or <u>Large</u>. You can also change the relative width of the two panes by clicking and dragging the line that separates the two panes.





Sections of the output can be expanded or contracted by clicking on the small blue squares in the navigation pane. A square with a '+' symbol will expand the output, and a '-'symbol will contract the output. This can be useful if you have lots of complex output in the same window. Also, double-clicking on the close-book icons such as the one next to 'Notes' will expand that section of output.





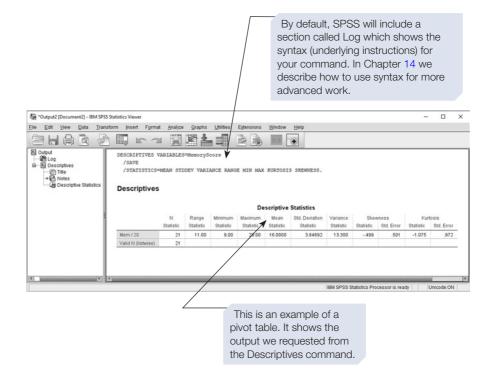
Most of the output produced by the SPSS commands is in the form of tables known as 'pivot tables'. The Descriptives Statistics table in the output above is an example of a pivot table. Pivot tables can be edited in various ways - see Chapter 14, Section 4 for full details. Other parts of the output are in the form of plain text. The title Descriptives is an example of text output.

You can select output either by dragging the mouse over the section in the main pane or by clicking on the appropriate icons in the navigator pane. In the navigator pane, click on the name of a command to select all the output from that command. You can select all output in the Viewer window by clicking on the top icon in the navigation pane (labelled **Output**). Lower-level icons (e.g. **Descriptive Statistics**) select only that particular part of the output.

Selected output can be cut, copied and pasted using the relevant options on the <u>E</u>dit menu or by right clicking and selecting the desired operation. Output can be printed using the <u>Print</u> command available under the <u>File</u> menu (see Chapter 14, Section 5 for details of printing output).



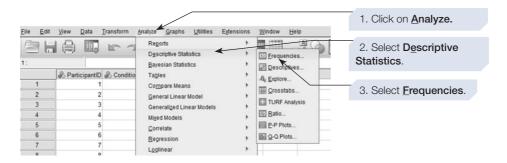
Sometimes, you will want to delete all the output in the Viewer window – for example, when you realise you have made an error and want to start your analysis again. The easiest way to do this is to click on the **Output** icon in the navigator pane and then press the Delete key on your keyboard. You can now start a new analysis with a blank output window.



The descriptive statistics shown in the output of the **Descriptives** command are based on all cases. For example, the mean value of 16.00 shown is calculated across all 21 participants in the data file. In reality, you may also want some of these statistics calculated separately for each condition. This is more easily done using the **Explore** command, which is introduced in Section 5 of this chapter.

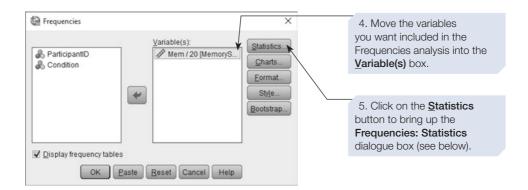
Section 4: THE FREQUENCIES COMMAND

The **Frequencies** command generates frequency distribution tables showing the number of cases (participants) with a particular score on each variable. For example, a frequency distribution table of the variable age would tell you how many of your participants were 20-year-olds, how many 21, and so on, for each of the ages represented in the group of participants. In addition, the **Frequencies** command will also produce a range of summary descriptives, including measures of central tendency and measures of dispersion, and some charts. One limitation is that the **Frequencies** command calculates descriptive statistics for all cases and will not give descriptives for subgroups broken down by some other variable (unless we use it in combination with commands such as **Select Cases or Split File**; see Chapter 4, Section 4).

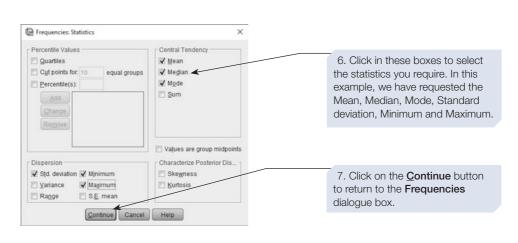


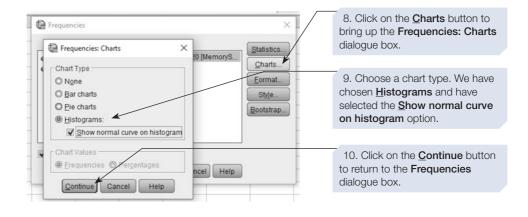
To obtain a Frequencies output

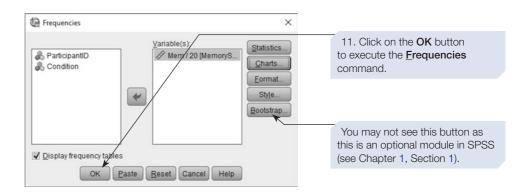
Now select the variable(s) you want included in the frequency analysis and move them into the Variable(s) box either by clicking the arrow button or by dragging and dropping them.



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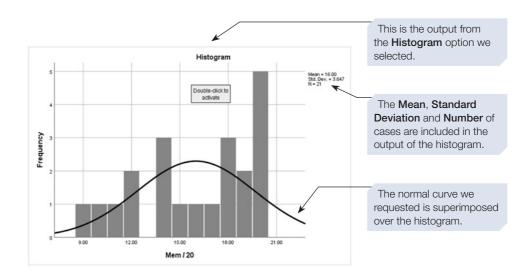


Clicking on the **Format** button in the **Frequencies** dialogue box will allow you to adjust the appearance of the output. Experiment with these settings.

command This is the title for the output produced by the **Frequencies** command. Frequencies [DataSet2] C:\Users\Richard\Documents\Publishing\Books\SPSS ed 7\DataFiles\DataEntryExIndGrps.sav Note SPSS tells us which data Statistics file is being analysed. This is Mem / 20 important as it is possible to have N Valid 21 several data files open at once. Missing 0 Always check this. Mean 16.0000 17.0000 Median Mode 20.00 This table lists the descriptive Std. Deviation 3.64692 statistics we requested calculated Minimum 9.00 across all participants. We can see Maximum 20.00 that the mean number of words recalled was 16 and that the maximum number and Mode were both 20. This is the Frequencies table for the dependent variable, 'MemScore'. Note SPSS is using our variable label ('Mem/20') as the heading. The Frequency column tells us how many participants Mem / 20 remembered a certain number of Cumulative words - for example, 3 people Frequency Percent Valid Percent Percent remembered 14 words. Valid 9.00 1 48 48 48 10.00 1 4.8 4.8 9.5 The Percent column tells us 11.00 1 4.8 4.8 14.3 what percentage of participants 12.00 2 9.5 9.5 23.8 remembered a certain number 14.00 14.3 14.3 38.1 3 of words, and the Valid Percent 15.00 1 4.8 4.8 42.9 column adjusts this percentage 16.00 4.8 4.8 47.6 1 for any missing cases. There are 52.4 17.00 1 4.8 4.8 no missing cases here so the two 18.00 14.3 66.7 3 14.3 columns are the same. 76.2 19.00 9.5 9.5 2 20.00 5 23.8 23.8 100.0 Total 21 100.0 100.0 From the Cumulative Percent column, we can see that (for example) 52.4% of our participants recalled 17 words or less.

The output produced by the Frequencies

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So far in this chapter we have looked at the **Descriptives** command and the **Frequencies** command. These commands provide useful output, but are limited, in that they cannot easily produce descriptive statistics separately for different groups of participants. Next we will introduce the **Explore** command, which is designed for use in these situations.

Section 5: THE EXPLORE COMMAND

Unlike the **Descriptives** or the **Frequencies** command, using the **Explore** command makes it easy to obtain descriptive statistics for separate groups of participants. For example, in the case of our memory experiment, we could use the **Explore** command to calculate summary statistics, such as the mean memory score, separately for the participants in the Mnemonic and Non-mnemonic groups. The variable that defines which group participants belong to (the variable Condition in this case) is referred to as the Factor.

Using the Explore command to analyse data from an independent groups design

1. Select <u>A</u> naly		lect D <u>e</u> scriptive stics.		3. Select	Explore.	
*DataEntryExIn	ndGrps.sav [DataSet2] - IBM	SPSS Statistics Data Edit	or			
<u>File Edit Vie</u>	w <u>D</u> ata <u>T</u> ransform	Analyze Graphs	⊈tilities	Extensions	Window	Help
19: 19: 1 2 3 4 5 6 7 °	ParticipantID & Condition	Regorts Descriptive Statist Bayesian Statisti Tables Compare Means General Linear M Generalized Line Mixed Models Correlate Regression Loglinear	cs Iodel	4 4 4 4 4 4 4 4 4 5 4 5	Image: Explore Image: Explore <t< td=""><td>ives bs nalysis</td></t<>	ives bs nalysis
Explore	Dependent List	MoryS Statistics Plots Options	w ei th or bo	I. Select the ant descrip ther drag ar the Depende r select ther row, and the row button ependent I	tive statistic and drop the ent List boom from the a click on the to move the	cs for and em into x (as here) left-hand
	Eactor List:	Bootstrap				
Display @ Both	Label <u>Cases by</u>	Help	m bo de br as	5. Select the factor variable and move it across to the Factor List box. This is the variable which defines how we want the statistics broken down. In this case we are asking for the statistics broken dow by Condition.		
	6. Click on the OK k	outton to		Note the me the variable		

execute the command.

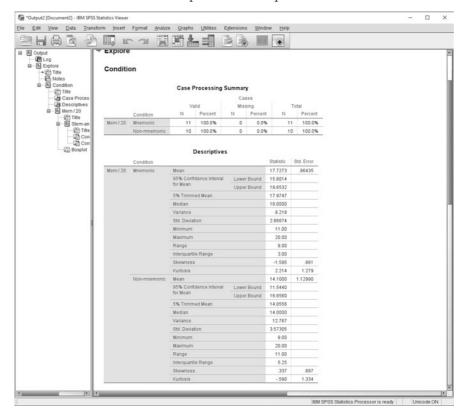
Note the measurement icon next to the variable name helps you to select a suitable variable for this role. Suitable factor variables are likely to be nominal or ordinal.

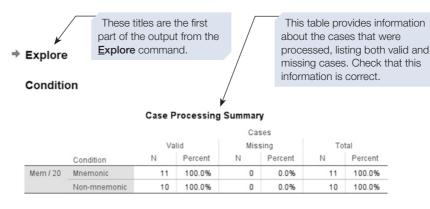
You could also click on the **Plots** button to change the type of graph requested. We will explain some of the plot options when analysing the data from our repeated measures design later in this section.

The Viewer window will now become the active window; if not, select it from the taskbar menu at the foot of the screen. The output of the **Explore** analysis using the data from the memory experiment is presented next.

The output produced by the Explore command for an independent groups design

The output of the **Explore** command will appear in the Viewer window, as shown below. We will now examine each part of this output in detail.





This is the label we gave to the dependent variable 'MemoryScore'.		participants in ea the mean, the 95	s the descriptive s ach condition. The 5% confidence inte rd deviation (SD) a	statistics (rval, the n	given incluc nedian,
		This is the factor grouping variable			
	Condition	Descripti	ves	Statistic	Std. Error
Mem / 20	Mnemonic	Mean		17.7273	.86435
Ment/20	1	95% Confidence Inter	val Lower Bound	15.8014	.00433
		for Mean	Upper Bound	19.6532	
This first part of the	/	5% Trimmed Mean	oppor bound	17.9747	
output gives the statistics		Median		18.0000	
for the participants in the		Variance		8.218	
Mnemonic group.		Std. Deviation			
		Minimum	11.00		
		Maximum		20.00	
		Range		9.00	
		Interquartile Range		3.00	
		Skewness		-1.595	.661
		Kurtosis		2.214	1.279
	Non-mnemonic	Mean		14.1000	1.12990
	7	95% Confidence Inter	val Lower Bound	11.5440	
		for Mean	Upper Bound	16.6560	
And in the second part of		5% Trimmed Mean		14.0556	
the table these statistics are given for those in the Non-		Median		14.0000	
mnemonic condition.		Variance		12.767	
		Std. Deviation		3.57305	
		Minimum		9.00	
		Maximum		20.00	
		Range		11.00	
		Interquartile Range		5.25	
		Skewness		.337	.687
		Kurtosis		590	1.334
We can report that the mean (and SD) for the Mnemonic an Non-mnemonic conditions we 17.73 (2.87) and 14.10 (3.57),		ku sa ar	xplore calculates Irtosis for each cor Imple sizes, as her e not very useful. F ou should consider	ndition. Fo e, these v For larger s	r small alues samples,

respectively.

Before writing up your results, consider how many decimal places you should report. In the example above, we have reported values rounded to two decimal places in line with the *Publication Manual of the American Psychological Association* (APA, 2009).

data.

indicate about the distribution of the

Mem / 20

Stem-and-Leaf Plots

Mem / 20 Stem-and-Leaf Plot for Condition= Mnemonic

Frequency	Stem &	Leaf
1.00 Ext:	remes	(=<11)
1.00	1.	4
5.00	1.	78889
4.00	2.	0000
Stem width:	10.	00

Each leaf: l case(s)

Mem / 20 Stem-and-Leaf Plot for Condition= Non-mnemonic

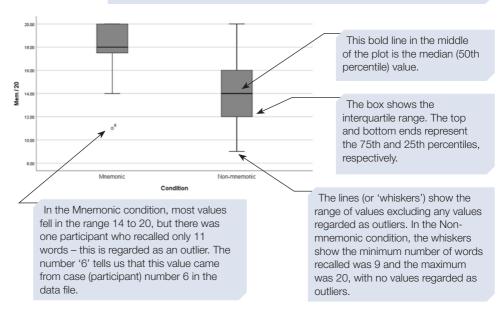
Frequency	Stem &	Leaf
1.00	ο.	9
5.00	1.	02244
3.00	1.	569
1.00	2.	0
Stem width: Each leaf:	10.0)0 case(s)

The default **Explore** output includes Stem-and-Leaf plots, as shown here. Each plot can be thought of as a histogram on its side. The advantage over a histogram is that all the individual data values can be seen. These plots can help you understand your data, but boxplots or histograms are more commonly used in psychology reports (see below).

This is the Stem-and-Leaf plot for the Non-mnemonic condition. The data are represented in two parts, the leaf, which contains the last digit of the value, and the stem, which contains the other digits. In this case, the stem has a value (or stem width) of 10. The number 9 is represented as a stem of 0 plus a leaf of 9, while 12 is a stem of 1 plus a leaf of 2, and 20 is a stem of 2 and a leaf of 0

The plot allows us to see the number of words recalled by the participants in the Non-mnemonic condition. We can see that our data include 1 participant who recalled 9 words, 5 participants who recalled between 10 and 14 words (10, 12, 12, 14 and 14), 3 participants who recalled between 15 and 19 words (15, 16 and 19) and 1 participant who recalled 20 words.

This is the boxplot of our data. Boxplots are a useful way to represent a data set. In a single plot, we can show many different characteristics of the data, including the minimum and maximum values, the interquartile range (25th and 75th percentile values), the median and any outliers.



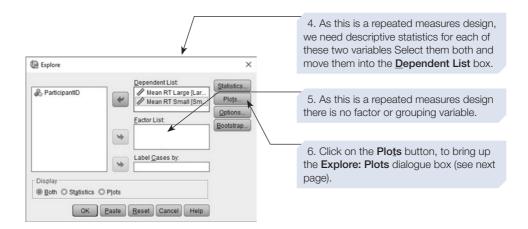
CHAPTER 3

The boxplot provides a powerful way of summarising our data and showing a range of descriptive statistics all in one plot. In this case, the boxplot suggests there is a ceiling effect for the Mnemonic condition, with many participants recalling either the maximum possible 20 words or close to this value.

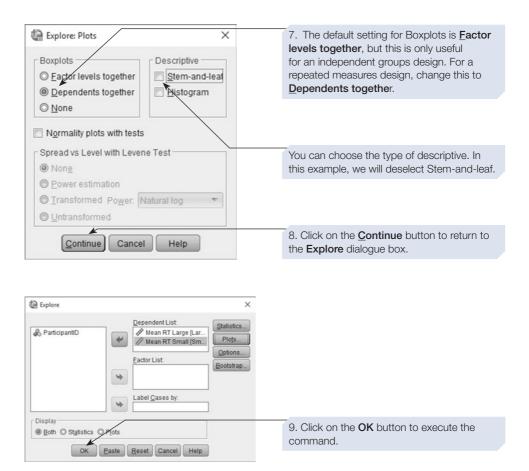
Using the Explore command to analyse data from a repeated measures design

To illustrate how to use the **Explore** command to analyse repeated measures data, we will use the data file we made and saved in Chapter 2, Section 6. Open that data file and follow the steps below.

1	. Selec	ct <u>A</u> naly	/ze.		2. Sele Statist	ct D<u>e</u>scr i cs .	iptive	\square	3. 5	Select <u>E</u> x	plore.	7
()	DataEntry	yExRptMea	s.sav [D	ataSet1] - IBM	SPSS Statis	tics Data Edi	itor					
<u>F</u> ile	Edit	View	Data	Transform	Analyze	<u>G</u> raphs	<u>U</u> tilities	Extens	ions	Window	Help	
E					Rep D <u>e</u> s	orts criptive Stati	stics	+	4	Ereque		
					Baye	sian Statist	tics		*	Descri	iptives	
		& Parti	icipantl	D 🖉 LargeD	Ta <u>b</u> l	es			*	-Q Explor	· •	Va
2	1]	1.0	936.0	Com	pare Mean	S					
	2]	2.0	923.0	Gen	eral Linear I	Model			Crosstabs		
	3]	3.0	0 896.0	Gen	eralized Lin	ear Models			TURF		
	4]	4.0	0 1241.0		d Models				Ratio	23	
	5		5.0	0 1278.0	-	-				P-P PI	ots	
	6]	6.0	0 871.0	-	ression				Q-Q P	lots	
	7]	7.0	0 1360.0		inear						
	8	1	8.0	0 733.0	Logi	ineai			1			

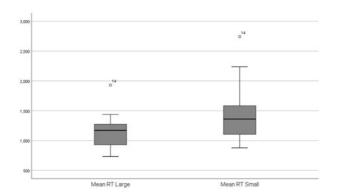


The Explore: Plots dialogue box allows you to alter how boxplots are produced. We need to change the setting to Dependents together. This change is needed because, in the data file for a simple repeated measures design, there is no SPSS variable that denotes the factor (IV). Instead, the levels of the independent variable are represented by separate columns of data (i.e. separate SPSS variables). SPSS calls these variables 'dependents', hence we need to select the Dependents together option (see below).

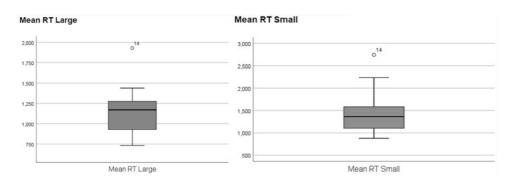


The boxplots produced by the Explore command for a repeated measures design

The boxplot produced by the Explore command is shown below. This plot allows us to visually compare the performance of our participants when they were completing trials under each of our two conditions. Looking at this figure, it is apparent that the participants tended to respond faster (i.e. shorter mean response time) when completing large-difference trials compared with small-difference trials. This is reflected in the relative median values, marked by the bold horizontal lines in the figure. We can also see that participant 14 was an outlier, having particularly slow mean response times in both conditions.



It is worth noting how this figure would have looked if we hadn't changed the default setting of Factor levels together in the Explore: Plots dialogue box. The default setting would produce two separate boxplots, as shown below (for this demonstration we have positioned the graphs alongside each other and resized them).



Presented this way, the performance of the participants in the two conditions appears to be quite similar; however, this is an artefact of the way SPSS plots graphs like these. If you look carefully, you will see that the scales of the *y*-axes of the two charts are different. Although it is possible to standardise the scale (we will explain how later), it is better to get SPSS to plot the two charts on one axis – as it does when we choose the **Dependents together** option.

Section 6: USING DESCRIPTIVE STATISTICS TO CHECK YOUR DATA

As we explained at the end of Chapter 2, before undertaking any analysis, it is critical to carefully check your data and correct any errors. This data checking and cleaning is most easily performed using the descriptive statistics commands, <u>Descriptives</u>, <u>Frequencies</u> and <u>Explore</u>, outlined above in Sections 2, 4 and 5.

To illustrate this process, we will use a file that we will be exploring in more detail in Chapter 4. For the purposes of this exercise, we have deliberately included errors in this file. If you would like to follow along as we check and clean this file, you can download your own copy (see below). Alternatively, use the procedures we describe here to check your own data files.

You may need to refer back to the instructions in Chapter 2 and in the earlier sections of this chapter to help you follow these steps

Checking variables in Variable View

First check the format of your data file:

- Open your data file and go to Variable View. Check that you set Value labels for any nominal or ordinal variables, and remind yourself what values you used. For example, in the case of Sex, what values did you use to represent Male and Female? This will remind you what values to expect for these variables.
- Next check the Missing Values. Did you set at least one missing value for each variable? And, if so, what value(s) did you use?
- Switch to Data View and check there aren't any rows of cells with dots. This is especially likely to occur below your last participant. If you find any such rows, delete them (see below for an example of this).

Using descriptive statistics to check the data

Next, we will use the **<u>F</u>requencies** command to check for errors in your data file. The same process could be completed using the output of the **<u>D</u>escriptives** and **<u>Explore</u>** commands, but the frequency tables of the **<u>F</u>requencies** command are particularly useful for checking nominal or ordinal variables.

Following the instructions in Section 4, undertake a Frequencies analysis for each of the variables in your data file. Select options to give you Mean, Median and Mode values, together with Minimum and Maximum values and Range. Some of the output from this analysis of our data file is reproduced below.

This is the first part of the output from the **Erequencies** command. This table reports each of the statistics we requested for each of the variables.

Frequencies

[DataSet2] C:\Users\Richard\Documents\Publi	shing\Books\SPSS ed 7\DataFiles\adoption
---	--

			Statistics								
		ID	sex	ethnicity	religion	adopted	q1	q2	q3		
Ν	Valid 🖌	19	20	21	20	19	19	20	18		
	Missing	2	1	0	1	2	2	1	3		
Mean	1	10.58	2.7500	7.2857	3.6500	1.5789	2.6316	2.4500	9.8333		
Media	an	11.00	2.0000	3.0000	3.0000	1.0000	2.0000	2.0000	3.0000		
Mode	•	1 ^a	2.00	2.00 ^a	3.00	.00	1.00	1.00	1.00 ^a		
Rang	le	19	23.00	98.00	6.00	4.00	4.00	4.00	129.00		
Minin	num	1	1.00	1.00	1.00	.00	1.00	1.00	1.00		
Maxir	num	20	24.00	99.00	7.00	4.00	5.00	5.00	130.00		

The first thing to look at is the **Number** of cases (N). Two values are given: the number of **Valid** cases (excluding missing cases) and the number of **Missing** cases. How many cases (participants) do you think are in your data set? Is this the same as the number reported here? We collected data from 20 participants, so we should expect a total of 20 cases. However, our first variable (ID) shows 19 valid cases and 2 missing cases, a total of 21 cases. Furthermore, given that this variable is the participant identifier number we allocated, we shouldn't expect any missing values for this variable. We must have made one or more errors here.

Now look at the data file and see if you can spot the errors for this variable.

di auc	puon_s	urrey n	in cirois	eu nsav (Dati	aSet3] - IBM SH	55 Statistic	5 Data L	ancon
<u>F</u> ile I	<u>E</u> dit	View	<u>D</u> ata	Transform	<u>A</u> nalyze	Graphs	<u>U</u> tilities	E;
			00,		M .			Ľ
22:ado	pted							
		æ	ID	& sex	& ethnicity	& relig	ion d	Bac
1			1	2.00	2.00)	3.00	
2			2	1.00	2.00)	7.00	
3			3	23.00	4.00)	2.00	
4			4	2.00	3.00		2.00	
5			5	2.00	2.00	/	6.00	
6			6	2.00	99.00)	5.00	
7			7	1.00	4.00)	3.00	
8			8	1.00	1.00)	6.00	
9			9	2.00	2.00)	3.00	
10			10	2.00	3.00)	3.00	
11			11	2.00	4.00)	5.00	
12	8		12	1.00	4.00)	5.00	
13			13	2.00	4.00)	3.00	
14			14	1.00	1.00)	6.00	_
15			15	1.00	3.00)	3.00	
16			16	2.00	4.00		1.00	
17			17	2.00	5.00		2.00	
18			18	1.00	1.00	5	3.00	
19			19	2.00	1.00)	2.00	
20			20	2.00	2.00)	3.00	
21				K	2.00)		

adoption_survey with errors ed 7.sav [DataSet3] - IBM SPSS Statistics Data Editor

The first error is quite easy to spot. We have an additional row of data for a phantom participant (row 21 of the data table). This occurred because we accidently entered an extra value for the variable 'ethnicity'. SPSS has then set all the other variables for this row to system missing – shown by dots in the cells. Importantly, this single error will have resulted in a missing case being recorded for every variable. It is also important to think about how this error occurred and whether it might indicate other problems. For example, does it suggest that the ethnicity data are out of sequence with the other variables? And, if so, is it possible that we have assigned the wrong ethnicity to some of our participants? You may need to go back to your original data records to check how this error occurred. For the purposes of this exercise, we are going to assume that it was just a typing error and doesn't indicate anything more worrying. The solution is simple – click on the row number 21 to highlight the whole row and press the delete key to remove it from the file. Then resave the data file. The second error for the variable 'ID' is harder to spot. When we set up the data file, we assigned the missing value of 9 to all the variables, including this one. However, for this variable, the value 9 is legitimate as there was a 9th participant. As a result, SPSS has treated our 9th participant as having a missing value for the variable 'ID'. The solution is to go to the Variable View of the data file and change the missing value for this variable. In this case, we could set it to 99 as we only have 20 participants. An alternative would be to set it to 0 as we would never use this value as a participant ID number. When you have made the change, save the data file again.

Having fixed these two errors, save the corrected data file and rerun the **<u>Frequencies</u>** command. You will note that when you select Frequencies all the options will be set as before, so all you need to do is click on the OK button. Now check the output to make sure you have fixed the problem with the ID variable.

The new output is shown below. We now have the correct number of participants, with no missing values for the variables 'ID', 'sex', 'ethnicity' and 'religion'. The variables 'adopted' and 'q1' each show 1 missing value, and 'q3' shows 2 missing values, but checking our records reveals that this is correct as some participants didn't respond to these questions.

							Statis	tics	
		ID	sex	ethnicity	religion	adopted	q1	q2	q3
N	Valid	20	20	20	20	19	19	20	18
	Missing	0	0	0	0	1	1	0	2
Mean		10.50	2.7500	7.5500	3.6500	1.5789	2.6316	2.4500	9.8333
Media	an	10.50	2.0000	3.0000	3.0000	1.0000	2.0000	2.0000	3.0000
Mode		1 ^a	2.00	4.00	3.00	.00	1.00	1.00	1.00 ^a
Minim	num	1	1.00	1.00	1.00	.00	1.00	1.00	1.00
Maxin	num	20	24.00	99.00	7.00	4.00	5.00	5.00	130.00

Frequencies

a. Multiple modes exist. The smallest value is shown

Next look at the **Minimum** and **Maximum** values for these variables. Can you see any errors here? The variable 'sex' has a maximum value of 23. This doesn't look right as sex was coded using 1 = Male, 2 = Female and 3 = Other. To investigate this further, scroll through the Frequencies output to the Frequencies table for the variable 'sex' (shown below).

			sex		
		Frequency	Percent	Valid Percent	Cumulative Percent
Valid	Male	7	33.3	35.0	35.0
	Female	12	57.1	60.0	95.0
	23.00	1	4.8	5.0	100.0
	Total	20	95.2	100.0	
Missing	System	1	4.8		
Total		21	100.0		

From this table we can see we have 7 male and 12 female participants in our data file, and 1 person who has been coded with a value of 23. Checking the data file shows that participant 3 is the one with the value 23, and checking our records shows that we actually had 13 female respondents, not 12, so it seems we accidently typed '23' instead of '2' for this participant. Make the change and resave the data file.

Now look at frequencies tables for the next two variables, 'ethnicity' and 'religion' (below). In the 'ethnicity' variable, we appear to have one erroneous value of 99 and in the 'religion' variable one erroneous value of 7. If we can identify what the correct values should have been, we can correct the data file. Alternatively, if we were not certain what the true values should have been, we could change these values to 9, the missing value for these two variables.

ethnicity										
		Frequency	Percent	Valid Percent	Cumulative Percent					
Valid	Asian	4	20.0	20.0	20.0					
	African	5	25.0	25.0	45.0					
	Chinese	3	15.0	15.0	60.0					
	European	6	30.0	30.0	90.0					
	Other	1	5.0	5.0	95.0					
	99.00	1	5.0	5.0	100.0					
	Total	20	100.0	100.0						

	religion										
		Frequency	Percent	Valid Percent	Cumulative Percent						
Valid	Atheist	1	5.0	5.0	5.0						
	Buddhism	4	20.0	20.0	25.0						
	Christianity	8	40.0	40.0	65.0						
	Muslim	3	15.0	15.0	80.0						
	Other	3	15.0	15.0	95.0						
	7.00	1	5.0	5.0	100.0						
	Total	20	100.0	100.0							

After locating and correcting each error, resave your data file. We recommend using a slightly different name each time you save the file, so that the previous version isn't overwritten and is available in case you make a mistake. For example, we could use the names 'adoption_survey v2', adoption_survey v3' and so on.

Following this process of carefully examining the output of the descriptive statistics commands such as Frequencies and Explore, we can check each variable in our data file looking for implausible or impossible values.

Checking scale variables

So far we have checked nominal or ordinal variables. When checking scale variables, the **<u>Frequencies</u>** command may be less useful. For example, imagine a file containing data from 100 participants where one of the variables was reaction time measured in milliseconds. In this case, it is likely each participant will have a unique value for the 'reaction time' variable, so the **<u>Frequencies</u>** command will produce output, including a long and relatively uninformative frequency table. In these cases, you may find it simpler to use the **<u>Descriptives</u>** command, and carefully check the mean, median, mode, minimum and maximum values of the variable to ensure that all are plausible. A useful alternative is to check the boxplot produced by the **<u>Explore</u>** command looking for unexpected values or outliers.

Finish cleaning the file

Now work through the remaining variables in your file and check each one. Once you are satisfied that you have identified and corrected all the errors, resave the data file one last time using a new file name to indicate that this is the cleaned version of the data file, for example 'adoption_survey_cleaned'.



The two most common sources of errors in data files are 'finger errors', where a simple typing mistake results in the wrong value being recorded, and errors relating to the use of missing values. Always think carefully about what missing values to set for a variable and make sure you use these values correctly when entering your data.

In the remaining sections we will introduce the topic of how to produce some simple graphs using SPSS.

Section 7: INTRODUCING GRAPHING IN SPSS

Producing graphs in SPSS

There are several ways to create graphs using SPSS. Many of the statistical commands accessed from the <u>Analyze</u> menu include graphing options in their output. We have already seen examples of graphs produced this way in Sections 4 and 5 covering the <u>Explore</u> and <u>Frequencies</u> commands. Many other statistical commands also include graphing options, and these will be covered in later chapters. However, SPSS also includes separate graphing commands. In the remaining sections of this chapter we will introduce you to these graphing commands, but first we will briefly describe some of the graphs types available in SPSS.

Graph types

Boxplots

Boxplots are useful when initially exploring your data, as described in Section 5. Boxplots provide a visual representation of the median and the dispersion (interquartile range and outliers) of the data in each condition of a study.

Histogram

Histograms are used to plot the frequency distribution of data measured on an ordinal, interval or ratio level of measurement. For an independent groups design, data should be plotted separately for each condition. We saw an example of a histogram produced by the **<u>Frequencies</u>** command in Section 4.

Bar charts

Bar charts are used to summarise data collected for different groups and are suitable for nominal and ordinal data. For example, we could use a bar chart to plot responses to a survey question about religious belief. Each bar would show the number of respondents (the frequency) who were followers of each religion.

Bar charts can also be used to summarise data across several levels of an independent variable or factor. In this case, the dependent variable should be of ordinal, interval or ratio level of measurement and each bar could represent the mean (or some other summary statistic) of the dependent variable. The different bars would represent the different levels of the independent variable. Error bars can be added to bar charts to indicate the variability in the data (e.g. standard error or confidence intervals; see Chapter 8, Sections 2 and 3), which can serve a similar function to an Error bar chart.

Error bar chart

Error bar charts can be thought of as a simpler version of a boxplot. The central point marks the mean (or another measure of central tendency), and the error bars show a suitable measure of dispersion. For example, you could display the mean and standard error, or the mean and the 95% confidence intervals, or the median and the interquartile range. See Chapter 9, Sections 2 and 3 for an example of this type of graph.

Line charts

Line charts are used to plot a series of data points that are joined together with a line. This is most appropriate when plotting the value on one variable (on the *y*-axis) against a second variable which is measured on an interval or ratio scale (on the *x*-axis). However, in psychology, it is quite common to use line charts to plot the data from more complex designs where we want to illustrate an interaction between factors. In this situation, the line chart has the advantage of clearly linking together related data points. We refer to these as 'interaction graphs' and give examples in Chapter 9, Sections 2 and 3.

Scatterplot

Scatterplots are used to illustrate the relationship between two variables. Each variable should be of ordinal, interval or ratio level of measurement.

The table below shows where you can find examples of each of these types of graph.

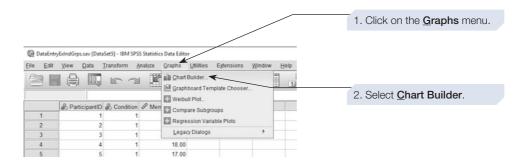
Graph type	Chapter and Section
Boxplots	This chapter, Sections 5 and 7
Histogram	This chapter, Sections 4 and 7
Bar charts	Chapter 7, Sections 4 and 5 Chapter 8, Section 2 and 3
Error bar graph	Chapter 5, Section 3 Chapter 9, Sections 2 and 3
Line charts/interaction graphs	Chapter 9, Sections 2 and 3
Scatterplot	Chapter 6, Section 2

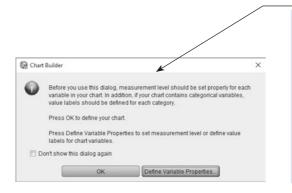
Section 8: CHART BUILDER

To use Chart Builder

We will demonstrate how to use <u>Chart Builder</u> by producing a boxplot for the memory study described in Chapter 2, Section 6. The result will be very similar to the boxplot we produced as an option when using the <u>Explore</u> command in Section 5 of this chapter. As we shall see, <u>Chart Builder</u> is slightly more flexible, allowing us to alter the look of the chart. In addition, we will show how you can edit a chart after it has been produced.

Open the appropriate data file and then follow these steps.

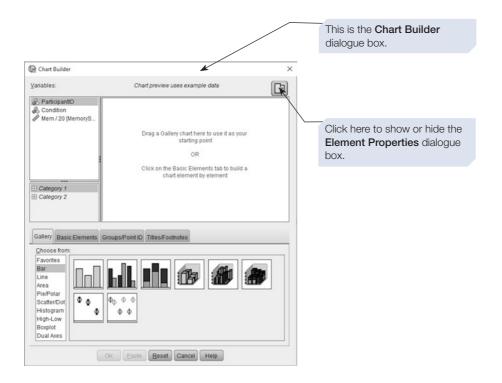




3. Before allowing you to proceed, SPSS gives you this reminder. To be able to use **Chart Builder** correctly, two properties of the variables must be set in the Variables view of the data window. You must set the level of measurement for each variable, and for nominal or ordinal variables you should give value labels. If you haven't already done this, do it now by closing this dialogue box and going to the Variable View of the data window. If the variables are set up correctly, then click on the OK button.

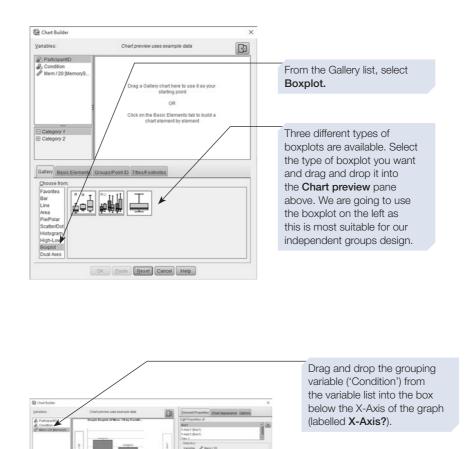
If you need to set the level of measurement or the value labels for a variable, it is easiest to close this dialogue and go back to the Variable View of the data window to set these characteristics. However, it is also possible to do this using the **Define Variable Properties** button on the dialogue box, but novice users will probably find this method confusing.

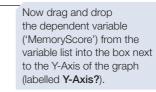
SPSS will now present you with two linked dialogue boxes: the Chart Builder and the Element Properties (see below). We will examine the **Chart Builder** dialogue box first.



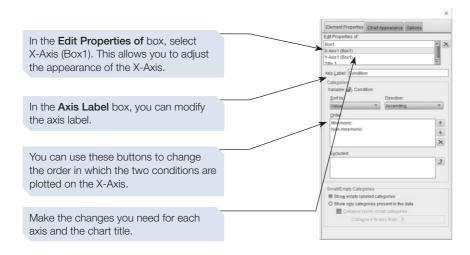
80 SPSS for Psychologists

OK Basta Beest Cancel Help





We can now adjust some of the settings for our graph in the **Element Properties** dialogue box (if necessary open this by clicking on the button on the top right-hand corner of the Chart Builder dialogue box).



Finally, click on the **OK** button on the **Chart Builder** dialogue box to produce your chart. The output is shown below.

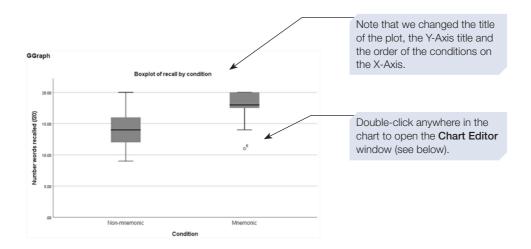
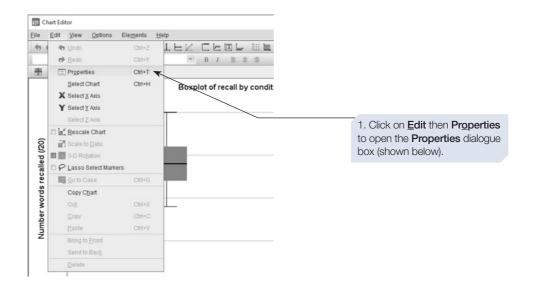


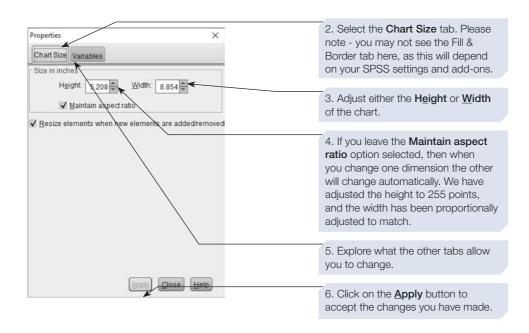
Chart Editor × Elle Edit Yew its Help * B / E \$ 1 A - M. 🖶 La Litt Boxplot of recall by condition This is the Chart Editor window. recalled (/20) vords r Number Non-mnemonic Mnemonic Condition

Editing charts in the Chart Editor window

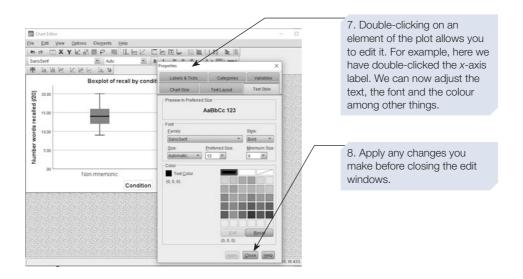
The menus and toolbars in the Chart Editor window can be used to improve the initial appearance of the chart. Items such as chart titles, subtitles, legends, axis titles, labels, line characteristics and markers can all be altered. Changes can be made through the menus, or by double-clicking directly on the item, or by clicking on the toolbar icons or buttons. If you hover your cursor over an icon or button, a pop-up message indicates its function.

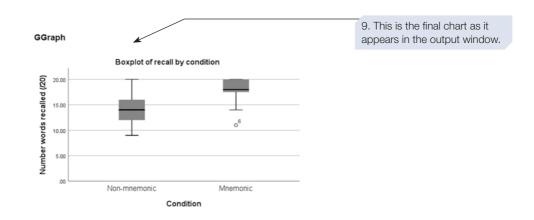
For example, we will reduce the size of our chart. It is best to do this in the Chart Editor commands to ensure that text and symbol size remain legible. Start by clicking on <u>Edit</u> then Properties, as shown below.





It is also possible to edit the features of the chart by double-clicking on the feature. For example, if we double-click on the *x*-axis title, then this dialogue box appears.







The Chart Editor window is used to edit charts produced from Chart Builder, and also to edit charts produced from commands on the Analyze menu. For charts produced using Graphboard Template Chooser, there is a separate editor, shown in Section 9.

Section 9: GRAPHBOARD TEMPLATE CHOOSER

The <u>Graphboard Template Chooser</u> is an alternative way to produce high-quality charts and includes innovative features such as an animation option, which could be useful in presentations. However, it does not currently include all the graph types we cover in this book.

We will demonstrate how to use <u>Graphboard Template Chooser</u> by producing histograms for the memory study described in Chapter 2, Section 6. Earlier, in Section 4, we produced a histogram using the <u>Frequencies</u> command. At that time we noted that we couldn't easily produce a separate histogram plot for participants in the Mnemonic and Non-mnemonic conditions. We will now show that it is easy to achieve this using the <u>Graphboard Template Chooser</u>.

To use Graphboard Template Chooser

- 1. Click on the Graphs menu item.
- 2. Click on <u>G</u>raphboard Template Chooser. The dialogue box appears, as shown on the next page.

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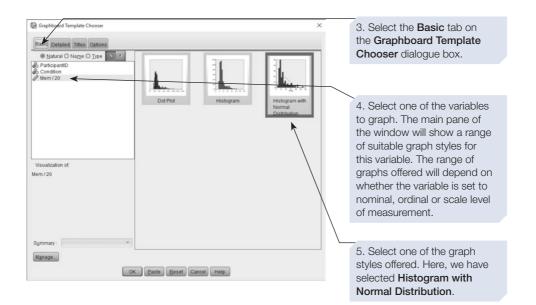
🖗 DataEn	tryExIndGrps.sav [Data	Set5] - IBM SPSS Stati	stics Data Editor			1. Clic item.
<u>Eile Edi</u>	t <u>V</u> iew <u>D</u> ata	Transform Analyze	e <u>G</u> raphs <u>U</u> tilitie	es Extensions	Window He	
16:	ParticipantID	& Condition & N	Chart Builder	Template Choose	H	2. Sel Temp bring
2	2	1	Regression Variable Plots			Temp
3	3 1		Legacy Dialo	Legacy Dialogs		box (s
4	4	1	18.00			
5	5	1	17.00			
6	6 1		11.00			
7	7 1		20.00			

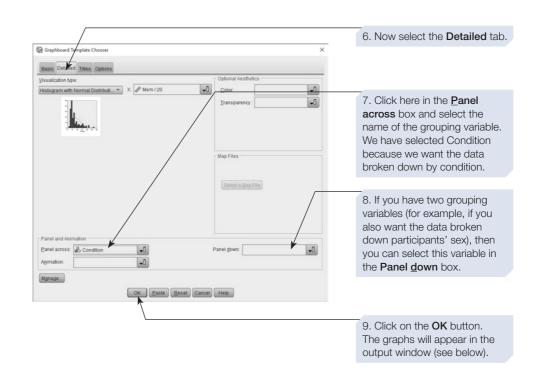
1. Click on the **<u>G</u>raphs** menu item.

2. Select <u>Graphboard</u> Template Chooser. This will bring up the Graphboard Template Chooser dialogue box (see below).

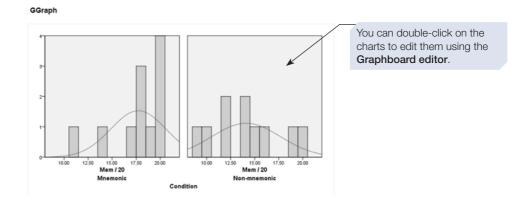


As with **<u>Chart Builder</u>**, **<u>Graphboard Template Chooser</u>** requires the level of measurement and value labels to be set for each of the variables you are going to use. However, **Graphboard Template Chooser** does not remind you of this.



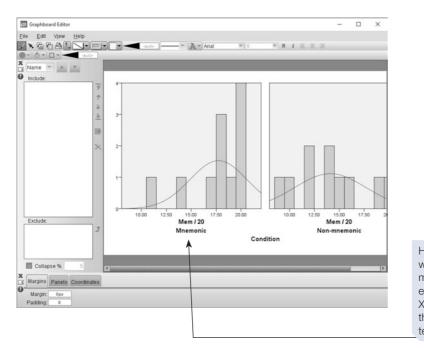


The two histograms produced by Graphboard Template Chooser are shown below.



The Graphboard Editor window

The **Graphboard Editor** window allows you to change many aspects of the graphs you have produced. Explore the menus and buttons to discover what you can do.





Summary

- This chapter introduced you to several statistical commands that can be used to produce a range of descriptive statistics in SPSS and demonstrated various ways of producing graphs.
- Section 3 introduced the Viewer window, which is the SPSS window in which the results of all analyses are displayed.
- Descriptive statistics can be obtained in several different ways. Sections 2, 4 and 5 demonstrated the <u>Descriptives</u>, <u>Frequencies</u> and <u>Explore</u> commands.
- It is vital to check and clean data files. Section 6 illustrated how to use the <u>Descriptives</u>, <u>Frequencies</u> and <u>Explore</u> commands to check and clean your data file.
- Graphs can be created in SPSS as options in many of the statistical analysis commands found under the <u>Analyze</u> menu. In addition, graphs can be produced using the <u>Graphs</u> menu.
- The <u>Graphs</u> menu includes Chart Builder and Graphboard Template Chooser. These methods of producing graphs were demonstrated in Sections 7–9.

Data handling

In this chapter

- An introduction to data handling
- Sorting a file
- Splitting a file
- Selecting cases
- Recoding values
- Computing new variables
- Counting values
- Ranking cases
- Data transformation
- Data file for scales or questionnaires



SPSS for Psychologists online

Visit macmillanihe.com/harrison-spss-7e data sets, online tutorials and exercises.

Section 1: AN INTRODUCTION TO DATA HANDLING

- SPSS includes a series of commands that can be used to modify, manipulate or transform the data. We refer to these as the Data Handling commands.
- These commands are very useful, especially when working with large data files containing many variables for each participant. Files such as these often arise from survey or questionnaire research.
- Questionnaires often contain items (questions) that can be grouped into a number of subscores. We can use certain Data Handling commands to calculate the subscores.

CHAPTER 4

Data Handling commands can also be used to transform variables, for example a log transformation converts a variable from a raw score to a log score. This type of transformation can reduce distortions, such as skewness, which might otherwise invalidate some analyses.

Another common use of Data Handling commands is to limit our analysis to a select group of participants. For example, we might decide to analyse the responses from male and female participants separately, or to exclude participants who had not met a set of criteria we have set for inclusion.

An example data file

To illustrate the use of these commands, we have created a small data file, which is printed in the Appendix and available from macmillanihe.com/harrison-spss-7e. The data file contains the results of a fictitious survey of people's attitudes to issues around the topic of adoption. The data file contains participant number, demographic data, such as the participant's age, sex, ethnic origin, religious belief and experience of adoption, together with their responses to 10 statements concerning aspects of adoption. These responses were made using a five-point scale ranging from 'Strongly Agree' (1) to 'Strongly Disagree' (5). The response to each of these items has been recorded in variables q1 to q10.

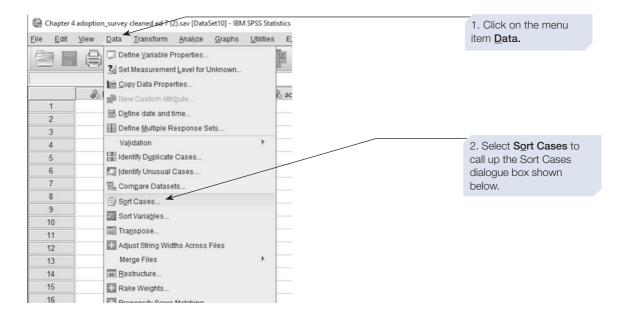
Section 2: SORTING A FILE

Novice users of SPSS sometimes worry about the order in which participants' data are entered into the SPSS data file. For example, if you have two independent groups do you have to enter all the data from participants in one condition before entering the data from the other condition? Normally, the order of the cases does not matter, but there are occasions when you might want to sort a data file so that the cases are in some meaningful order, perhaps to make it easier to check the accuracy of the data file.

If you are likely to change the order of the cases in a file by using either the **Sort Cases** or **Split File** commands, then you will not be able to rely on the SPSS case numbers (the numbers on the extreme left of the data window) to identify participants. For this reason, it is always best to create your own participant identification variable and record this number both in the data file and on your original records so that you can check the data later if necessary.

The Sort Cases command

In this example we will sort the data by two variables, first by participants' sex, and then, within sex, we will sort by ethnicity.



Cases X	3. Select the name of the sorting variable.
Sort by:	4. Either click on the arrow button or drag and drop the variable into the Sort by box.
G q1 g2 g3 g4 g4 Sort Order ▲ @ Ascending @ Descending	5. You can choose to sort in Ascending or Descending order (see below). To sort by a second variable, repeat steps 3 to 5.
Save Solied Data Save Solied Data File ☐ <u>Oreate an index</u> OK_ Paste Reset Cancel Help	6. You can choose to save the sorted file now (leave this option unselected to save later as normal).
Laste Leser Calicel Help	7. Finally, click on the OK button –

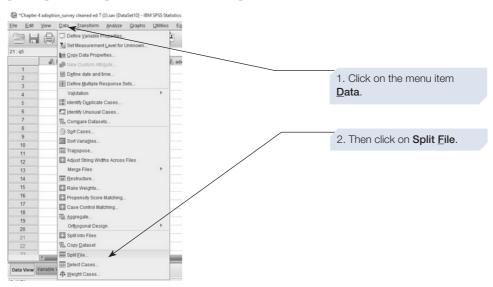
SPSS will now re-sort the data table.

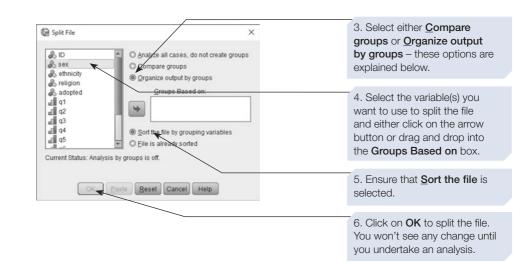
You can sort in either ascending or descending order. **Ascending** order puts participants with a low value on the sort variable before participants with a higher value (e.g. Male before Female if we used the code Male = 1 and Female = 2). **Descending** would sort in the reverse order. You will probably want **Ascending** order.

Chapter -	4 adoption_survey	cleaned ed 7	(2).sav [DataSet]	101 - IBM SPSS 5	tatistics	
Eile Edit	View Data	Iransform	Analyze Gr	raphs Utiliti	s E	We sorted the data file by two variables – Sex and Ethnicity.
4:q4	2.00			/		Note that, after sorting the
	& ID	& sex	& ethnicity	Breligion	ad ad	, 0
1	2	1.00	2.00	9.00		file, the participant number
2	7	1.00	4.00	3.00		('ID') no longer matches the
3	8	1.00	1.00	6.00		case number. This is why we
4	12	1.00	4.00	5.00		recommend including a
5	14	1.00	1.00	6.00		variable for participant
6	15	1.00	3.00	3.00		number.
7	18	1.00	1.00	3.00		namber.
8	1	2.00	2.00	3.00		
9	3	2.00	4.00	2.00		
10	4	2.00	3.00	2.00		
11	5	2.00	2.00	6.00		

Section 3: SPLITTING A FILE

The **Split <u>File</u>** function is a particularly useful feature of SPSS. **Split** <u>File</u> semipermanently splits a data file into groups, and in subsequent analysis the output can be organised according to these groups. For example, you can request SPSS to organise all subsequent output so that statistics are presented separately for male and female participants. To split a file, follow the steps shown below.



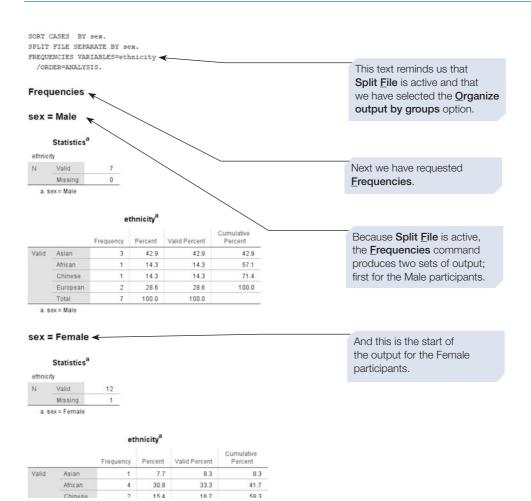


Options

The difference between the options <u>Compare groups</u> and <u>Organize output by</u> groups is worth exploring. The former contrasts the two groups within one section of output, whereas the latter produces two different sections of output. We usually find <u>Organize output by groups</u> more useful, but try each so that you understand the difference.

Output

Initially, the only consequence of the **Split** <u>File</u> command is that two lines of text are added to the output window. This text serves as a reminder that the **Split** <u>File</u> is active. It tells you that the file has been sorted (this is a necessary first step that SPSS undertakes for you) and split using the separate output option. Subsequently, any command will be performed separately for each of the two groups. For example, if we **Split** <u>File</u> by sex, and then use the <u>Frequencies</u> command (see Chapter 3, Section 2) to obtain some information about the variable 'ethinicity', the output will be shown first for the Male participants and then for the Female participants.



Unsplitting a file

Remember that **Split File** is a semi-permanent change. Once you have executed the **Split File** command, the output from all other commands will be broken down by the selected variable. This will remain in force until you reverse the process by unsplitting the file. You can tell whether Split File is on or not by looking at the bottom right-hand corner of the Data View window. In this case, when the data has been split SPSS displays "Split by Sex" in the pane of the SPSS window. To undo this, first repeat steps 1 and 2 above and then select the option <u>Analyze all cases</u>, do not create groups. Then click the own button. From this point onwards, all analyses will be performed for all cases and the output will return to the normal format.





Another option under the **Data** menu item is **Split into Files**. As the name implies, this allows you to divide the data file on the basis of one variable (e.g. sex) and save the different parts of the data set in separate data files. Although this might sound useful, in practice it is usually more efficient to keep the data together in a single file and use **Split File** or **Select Cases** (described in the next section) to achieve the analysis you want.

Section 4: SELECTING CASES

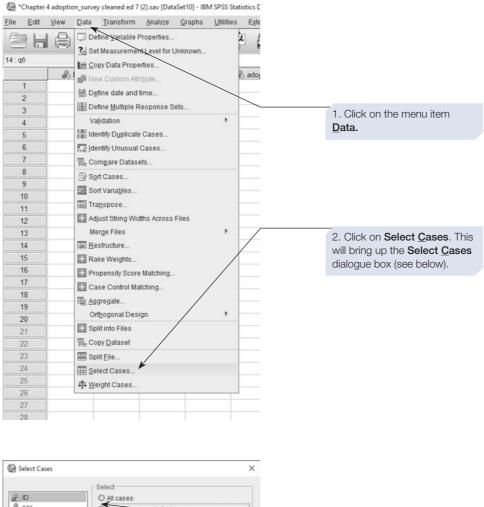
An alternative to splitting a file is to select certain cases and use only these in subsequent analyses. For example, we might be particularly interested in the survey responses made by our respondents who had themselves been adopted. Select Cases will allow us to analyse just these participants' responses. All other data will be temporarily suppressed.

Comparing the Select Cases and Split File commands

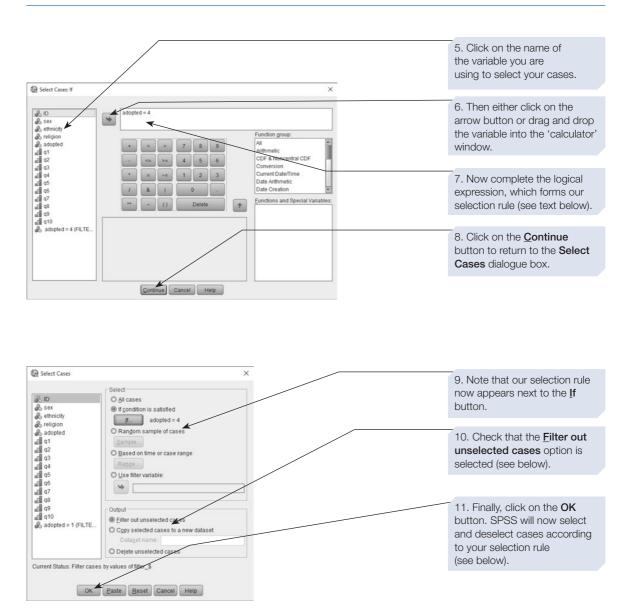
Select Cases is different from Split File. Select Cases suppresses analysis of nonselected cases, whereas **Split <u>F</u>ile** analyses all cases but arranges output by the sorting variable. Use Select Cases when you want to consider only some of your data. Use Split File when you want to contrast two or more groups of participants.

The Select Cases command

To Select Cases, perform the following steps.



O All cases & sex & ethnicity IIX & religion 3. Select the If condition & adopted C Random sample of cases 4 4 4 4 4 4 4 4 4 4 4 4 5 4 4 4 4 5 4 7 4 8 4 9 4 9 4 10 7 is satisfied option. O Based on time or case range The other options are described in the text below. O Use filter variable: \$ Output Eilter out unselected cases 4. Click on the **If** button. This Copy selected cases to a new dataset will bring up the Select Cases: If dialogue box (see below). O Dejete unselected cases Current Status: Do not filter cases OK Paste Reset Cancel Help



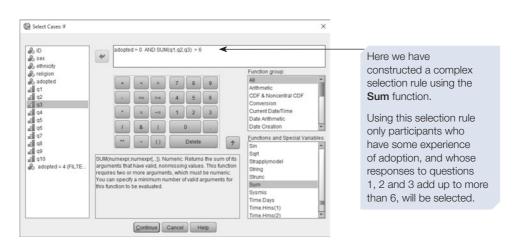
In step 10 above, there are three options. The default is **Eilter out unselected cases**. The second alternative, **Copy selected cases to a new dataset**, allows us to open a new data file that we can use to save just the selected cases. This is useful if we want to draw out a small subset of cases for separate analysis, leaving the original file unchanged. The third option, **Delete unselected cases**, will permanently delete the unselected cases from your data file. If you use this option, make certain that you have a backup copy of your data file.

ile <u>E</u> dit	⊻iew <u>D</u> ata	Transform	<u>A</u> nalyze <u>G</u>	raphs <u>U</u> tilitie	s Extensions	Wi			SPSS has put a line
	A	Er -			# #				through the case number
									of the deselected cases.
5:q10		-	-			-	-	- X	Note that only cases 7,
13	8 ID	a sex	& ethnicity	& religion	& adopted	al.	q10	& filter_\$	8, 16 and 19 are selected
	1	2.00	2.00	3.00	1.00		2.00	0	because they are the only
2	2	1.00	2.00	9.00	.00		5.00	0	respondents who
-3	3	2.00	4.00	2.00	1.00		1.00	0	reported they were
4	4	2.00	3.00	2.00	.00		1.00	0	
-5-	5	2.00	2.00	6.00	.00		3.00	0	adopted.
-6	6	2.00	9.00	5.00	.00		1.00	0	
7	7	1.00	4.00	3.00	4.00		5.00	1	
8	8	1.00	1.00	6.00	4.00		1.00	1	
-0	9	2.00	2.00	3.00	2.00		1.00	0	
10	10	2.00	3.00	3.00	3.00		2.00	0	
11	11	2.00	4.00	5.00	.00		5.00	0	
12	12	1.00	4.00	5.00	1.00		3.00	0	SPSS has created a
13	13	2.00	4.00	3.00	1.00		4.00	0	
14	14	1.00	1.00	6.00	2.00		5.00	0	new variable called
15-	15	1.00	3.00	3.00	1.00		2.00	0	'filter_\$', which it
16	16	2.00	4.00	1.00	4.00		2.00	1	uses to choose which
17	17	2.00	5.00	2.00	.00		5.00	0	cases are selected.
18	18	1.00	1.00	3.00	2.00		4.00	0	
19	19	2.00	1.00	2.00	4.00		1.00	1	
20	20	2.00	2.00	3.00	9.00		3.00		

Selection rules

You can construct very complex selection rules by using the logical operators AND, OR and NOT. The selection rules can either be typed in using the keyboard or can be built up using the calculator keypad that appears in the dialogue box. For example, if we wanted to select only Chinese Christians who had some experience of adoption, we could construct the following expression: religion = 3 and ethnicity = 3 and adopted > 0.

The **Select Cases:** If dialogue box also contains useful functions you can use in constructing your selection rule. The functions are organised into groups; for example, there is a statistical group containing various functions to calculate the mean, standard deviation and related values. Click on a Function group to see the available functions. If you click on a function name, a description will appear in the box next to the functions.



Selection methods

The **Select Cases** dialogue box offers a total of four methods of selecting cases (see step 3 above). These are:

- 1. The default, If <u>Condition is satisfied</u> method, is the one you will probably use most often.
- 2. The **Random sample of cases** method allows you to sample cases at random from your data file, and SPSS allows you to specify either the number or percentage of cases to be selected. This can be useful in some more advanced statistical operations.
- 3. The **Based on time or case range** method allows you to select cases either by case number or on the basis of a time or date range.
- 4. In the <u>Use filter variable</u> method, a case is selected if the value of the chosen variable is not zero (and is not missing) this option can be useful if you have a yes/no variable coded as 1/0. Using this method, you could easily select only the 'yes' responses.

It is useful to note that a line of text at the bottom of the Select Cases dialogue box indicates the current selection rule. Finally, remember to reselect <u>All cases</u> after you have completed your analysis of the selected cases.

Reselecting all cases

The **Select Cases** function can be very useful, but it is important to remember that it is semi-permanent. **Select Cases** will stay in force until you either make some other selection or choose the <u>All cases</u> option in the **Select Cases** dialogue box.

Select Cases	Select	
ethnicity eligion adopted add q1 q1 q2	Based on time or case range Range Use filter variable:	Once you have finished working with your subset of cases, remember to select <u>All</u> cases so that subsequent analyses are based on your entire data set.
교 q8 교 q9 교 q10 중 adopted > 0 AND S	Output	
Current Status: Filter cases	by values of filter_S Paste Reset Cancel Help	

CHAPTER 4

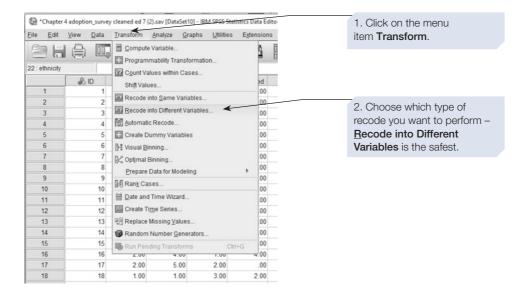
Section 5: RECODING VALUES

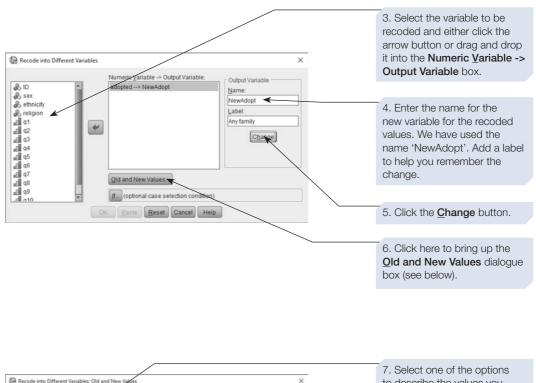
There are many occasions when you need to recode some of your data. This might be because you made an error when entering the data, but it is more likely that you will want to recode your data in light of results from a preliminary analysis, or to allow you to undertake an additional analysis.

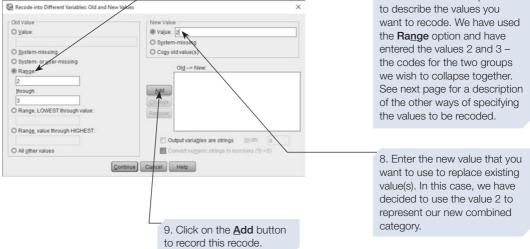
For example, early analysis of our adoption survey might reveal that we have very few participants who report experience of adoption through either 'immediate family' or 'other family'. In light of this, we might decide to collapse these two categories together into one new category. We could do this manually, but for a large data set that would be time-consuming. SPSS provides the **Recode** command for this purpose.

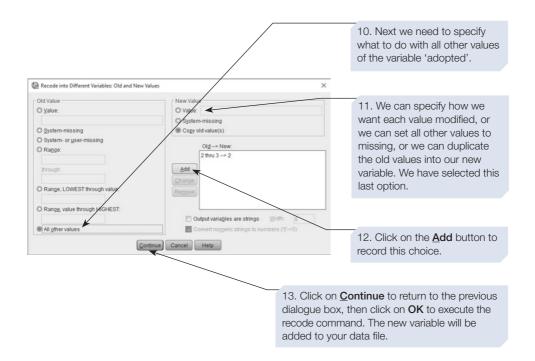
SPSS offers two **Recode** options. We can either change the values in the existing variable, or we can create a new variable for our recoded data. These two options are called **Recode into Same Variables** and **Recode into Different Variables**, respectively. It is usually safer to recode into a different (new) variable rather than overwriting the original data – that way, if you make a mistake, you will be able to go back to the original values and try again. To recode a variable, follow the steps outlined below.

Recode into Different Variables









Specifying the values to be recoded

In the **Recode into Different Variables: Old and New Values** dialogue box (see step 7 above), you are offered seven different methods of specifying the values you want to recode, and you can use a combination of these methods if required. The **Range**, **LOWEST through value:** and the **Range**, **value through HIGHEST:** are often very useful – for example, if you want to collapse together all categories 6 and above, you could use the **Range**, **value through HIGHEST:** option entering the value 6 in the box. When specifying the values to be recoded, you should give careful consideration to your missing values. If, for example, you used 9 as the missing value, then recoding using the **Range**, **value through HIGHEST:** would result in the missing observation being included in the new category, which you may not want.

The <u>Value</u>: option allows you to specify a single value that you want to recode. The All <u>other values</u> option translates as 'and for everything I haven't yet specified' and allows you to tell SPSS how to recode all the values not covered by one of the previous recode instructions.

The <u>System-missing</u> and the <u>System- or user-missing</u> options are very powerful. System-missing values are rather like user-missing values (which, in Chapter 2, we simply called missing values). Both are used to indicate that there is no valid value for a variable. However, a system-missing value indicates that SPSS rather than you (the 'user') has declared a value non-valid – perhaps, for example, because, for this participant, it is not possible to calculate a valid value for the variable. These two options allow you to recode these two types of missing values but, before you use them, think carefully about the implications of your actions. Note if you enter a value into the **New Value** box, which has previously been specified as a missing value (see step 8 above), you can effectively remove a range of values from an analysis by recoding valid responses into missing values. Similarly, by clicking on the **System-missing** option in the **New Value** box, you can instruct SPSS to regard any value or range of values as system-missing from this point onwards.



When using the **Recode** command, think carefully about what will happen to any missing values.

Don't forget to tell SPSS how to deal with all the other values from the original variable. The **Copy old value(s)** option is very useful, as it allows you to tell SPSS that any values which you haven't otherwise specified should just be copied over into the new variable. If you forget to do this, all these unspecified values will be set as system-missing in your new variable.



Remember: the big advantage of using **<u>Recode into Different Variables</u>** (rather than **<u>Recode into Same Variables</u>** described below) is that you do not lose anything. If you make an error, the original data are still available in the old variable and you can simply try again.

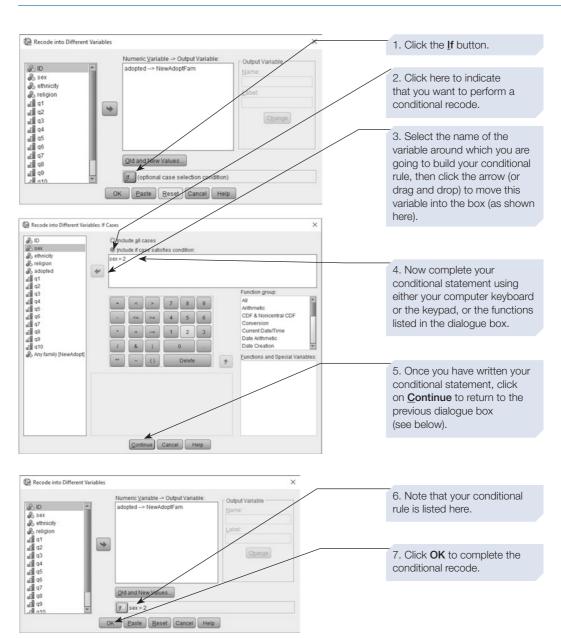
Recode into Same Variables

If you are certain that you know what you are doing, and you have a backup of your data file, you might decide that you can overwrite the existing data rather than create a new variable. To do this, select **Recode into Same Variables** at step 2 above. From this point onwards, the procedure is very similar except that you omit steps 4 and 5 as there is no new variable to name. The results of this recode will overwrite the original data in the data file.

Conditional recode

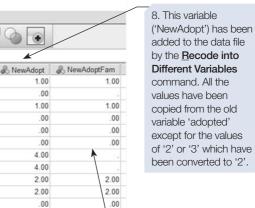
On some occasions, you might want to recode a variable only if a particular condition is satisfied for that participant. For example, you might want to perform the recode described above, but only for the female participants. This can be achieved by using the <u>If</u> button, which appears in both the **Recode into Different Variables** and the **Recode into Same Variables** dialogue boxes. Follow the procedure described above up to step 13. Now follow the steps described below.





ile <u>E</u> d	it <u>V</u> iew	Data	Transform	Analyze G	raphs <u>U</u> tilitie	es Extensions
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7:sex						
	æ	ID	& sex	& ethnicity	& religion	& adopted
1		1	2.00	2.00	3.00	1.00
2		2	1.00	2.00	9.00	.00
3		3	2.00	4.00	2.00	1.00
4		4	2.00	3.00	2.00	.00
5		5	2.00	2.00	6.00	.00
6		6	2.00	9.00	5.00	.00
7		7	1.00	4.00	3.00	4.00
8		8	1.00	1.00	6.00	4.00
9		9	2.00	2.00	3.00	2.00
10		10	2.00	3.00	3.00	3.00
11		11	2.00	4.00	5.00	.00
12		12	1.00	4.00	5.00	1.00
13		13	2.00	4.00	3.00	1.00
14		14	1.00	1.00	6.00	2.00
15		15	1.00	3.00	3.00	1.00
16		16	2.00	4.00	1.00	4.00
17		17	2.00	5.00	2.00	.00
18		18	1.00	1.00	3.00	2.00
19		19	2.00	1.00	2.00	4.00
20		20	2.00	2.00	3.00	9.00

*Chapter 4 adoption_survey cleaned ed 7 (2).sav [DataSet10] - IBM SPSS Statistics Data Edito



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4 00

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8. And this variable ('NewAdoptFam') was the one we created using the conditional recode. In this case the recode has only been applied to female respondents. For male respondents the values are set as system-missing (represented by a dot in the cell).

The rules for constructing the conditional statement (or logical expression) are the same as in the Select If command described earlier. You can construct quite complex logical expressions by using a combination of the functions provided and the operators (add, subtract etc.) available on the calculator-style buttons. Some of the less obvious buttons are listed below:

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- ** Raise to the power (for example, $'3^{**}2'$ is equivalent to $3^2 = 9$)
- Less than or equal to <=
- Greater than or equal to >=
- Not equal to ~=
- & AND
- OR
- NOT ~

Section 6: COMPUTING NEW VARIABLES

Sometimes we need to calculate a new variable based on the values of existing variables. For example, you may have entered the response given by each participant to each question in a questionnaire. You could now use SPSS to calculate the overall score for the questionnaire or several separate scores for subscales within the questionnaire. In our fictitious survey of attitudes to adoption, we administered a 10-item questionnaire, which was made up of two subscales. We therefore need to sum the responses to the items that contribute to each of the subscales. SPSS can do this for us using the **Compute Variable** command.

e Edit	⊻iew	Data	Transform Analyze	Graphs	Utilities	Extension	s		
-	1.0		Compute Variable				F		
			Programmability Tra	notorm No.			E		
			Count Values within						1. Click on menu item
	3	q4		Cases		q 8	3		Transform.
1	4.00	5	Shift Values				5.0		
2	5.00	5	Recode into Same V				3.1		
3	2.00	2	Recode into Differen	t Variables.			1.0		
4	9.00	3	Automatic Recode				1.0		2. Select <u>Compute Variable</u> .
5	3.00	4	Create Dummy Varia	ables			5.0		
6	9.00	1	Visual Binning				1.0		
7	1.00	2	🖧 Optimal Binning				3.0		
8	2.00	1	Prepare Data for Mo	deling			2.0		
9	2.00	1	Rank Cases				1.0		
10	1.00	2	Date and Time Wiza	rd			3.0		
11	5.00	5	Create Time Series.				5.0 3.0		
12	3.00	3	Replace Missing Val				2.0		
13	5.00	4	Random Number Ge				1.0		
14	2.00	1					2.0		
16	1.00		Run Pending Transf	2.00	Ctrl+0	3	2.0		
17	4.00		00 3.00	4.00	3.00		4.0		
			,				We	e have ca	ne name of the new variable. led it 'TransNatAdopt', as it is a attitudes to transnational adoptio
							We me The	e have cal easure of e Type &	lled it 'TransNatAdopt', as it is a attitudes to transnational adoptio
Compute Vi	ariable						We me The	e have cal easure of e Type &	led it 'TransNatAdopt', as it is a attitudes to transnational adoption Label button can be used to ado the type for the new variable.
arget Variabl ransNatAdo Type & Label & ID & sex & ethnicity & religion & adopted dl q1 q2 dl q2 dl q4 dl q5 dl q6 dl q6			mgric Epression: 1 + q3 + q4 + q8 + q10)5 + ≤ > 7 = 8 - ≈ > ≈ 4 = 5 + = − 1 = 2 / & 1 = 0 = − () Detet		Conversi Current D Date Arith Date Cre	c oncentral CDF on Date/Time Imetic	We me lab	e have cal easure of e Type &	led it 'TransNatAdopt', as it is a attitudes to transnational adoption Label button can be used to add
arget Variabi ransNatAdo ()ppe & Label 6, ID 6, Sex 6, ethnicity 6, sex 6, ethnicity 6, religion 6, adopted 11 q1 11 q2 11 q2 11 q2 11 q3 11 q5 11 q5	te: ot		* < > 7 8 - cs >s 4 5 - cs >s 4 5 - z - s 1 2 / & 1 0 * - () Dete		All Arithmetii CDF & Ni Conversi Current D Date Arith Date Cre	c oncentral CDF on Date/Time timetic ation	We me lab	e have cal easure of e Type &	 Ided it 'TransNatAdopt', as it is a attitudes to transnational adoptio Label button can be used to address the type for the new variable. 4. Enter the formula for the computation of the new variable here. You can type variable names, use the arrow button, or drag and drop them from the list on the left. Use the keyboard or the screen buttons to add the other symbols you need. You
arget Variabi ransNatAdo Type & Label & D & D & Sex & ethinicity & religion & adopted d 1 d 2 d 2 d 2 d 4 d 4 d 5 d 65 d 75 d 75			* < > 7 8 - cs >s 4 5 - cs >s 4 5 - z - s 1 2 / & 1 0 * - () Dete		All Arithmetii CDF & Ni Conversi Current D Date Arith Date Cre	c oncentral CDF on Date/Time timetic ation	We me lab	e have cal easure of e Type &	 Ided it 'TransNatAdopt', as it is a attitudes to transnational adoption Label button can be used to ado the the type for the new variable. 4. Enter the formula for the computation of the new variable here. You can type variable names, use the arrow button, or drag and drop them from the list on the left. Use the keyboard or the screen buttons to add the other symbols you need. You can also use the functions. 5. It is possible to perform a



When entering the name of the new variable (see step 3 above), it is possible to enter a variable label to act as a reminder of what the new variable means. Do this by clicking on the **Type & Label** button. You can then either type in a text label, or, by selecting the **Use expression as label** option, you can ask SPSS to use your numeric expression as the variable label. In this case, the label would be 'COMPUTE TransNatAdopt=(q1+q3+q4+q8+q10) / 5'.

taSet10] - IBM SPSS Statistics Data Editor

Graphs Utilities Extensions Window

7. T		× ×		- 1			r. I	17
'Tra	va	TransNatAdopt	q10	q3	g2	q1	adopted	
aver		4.00	2.00	4.00	4.00	4.00	1.00	0
item		4.40	5.00	5.00	1.00	4.00	.00	0
		1.60	1.00	2.00	3.00	2.00	1.00	0
			1.00	9.00	1.00	2.00	.00	D
		3.80	3.00	3.00	5.00	4.00	.00	0
			1.00	9.00	1.00	2.00	.00	0
		2.40	5.00	1.00	1.00	1.00	4.00	0
 8. N		1.40	1.00	2.00	2.00	1.00	4.00	0
dot i		1.20	1.00	2.00	1.00	1.00	2.00	0
syste		2,00	2.00	1.00	2.00	2.00	3.00	0
canr		5.00	5.00	5.00	4.00	5.00	.00	0
(see		3.00	3.00	3.00	4.00	4.00	1.00	D
		3.00	4.00	3.00	3.00	3.00	1.00	0
		3.40	5.00	5.00	3.00	2.00	2.00	0
		¥ .	2.00	2.00	1.00	9.00	1.00	0
		1.60	2.00	1.00	2.00	1.00	4.00	0
— 		4.40	5.00	4.00	4.00	5.00	.00	0
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ansv ente		1.20	1.00	1.00	1.00	1.00	4.00	0
(9). /		2.20	3.00	3.00	2.00	1.00	9.00	0

Help

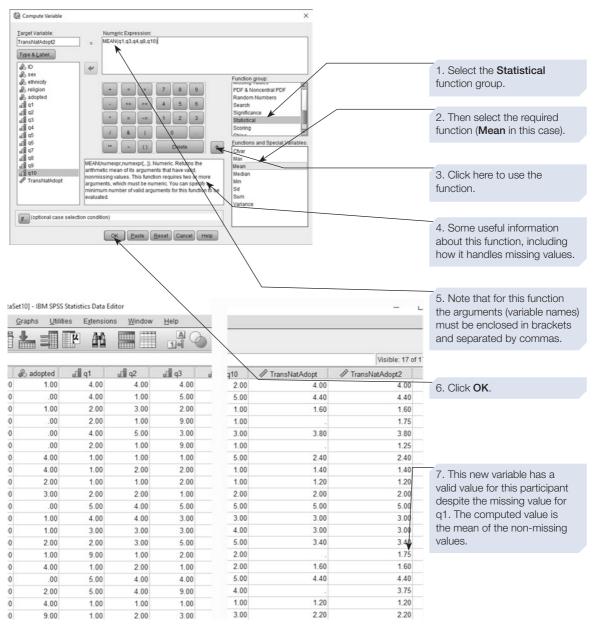
7. This is our new variable TransNatAdopt' – the average of the scores on tems 1, 3, 4, 8 and 10.

3. Note that SPSS has put a dot in this cell to represent a system-missing value. SPSS cannot calculate a valid value see step 9).

 9. This participant did not answer question 1, so we entered a missing value
 (9). As a result, SPSS can't calculate a value for our new variable for this participant.

Compute and Missing Values

When using **Compute Variable**, it is important to think carefully about missing values. SPSS will not be able to compute the value of the new variable if any of the values for the variables involved in the compute statement are missing. In the example above, participant 15 had not answered question 1 (q1) and we had entered a missing value (9) in this cell of the data table. As a result, SPSS is unable to compute a value for the new variable 'TransNatAdopt' for this participant. With complex compute statements involving lots of variables, this can be a major problem. One way round this is to make use of some of the functions built into SPSS, such as **Mean**, which make automatic allowances for missing values. This is illustrated below.



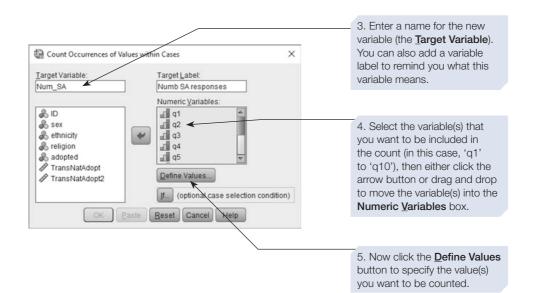


It is important to consider whether the new values computed using functions in this way are legitimate. For example, if you use the **Sum** function, then in cases where one or more values are missing, the new value will be calculated based only on the non-missing values. This may or may not be what you want. Before undertaking a **Compute Variable** command, think carefully about what will happen in cases with missing values.

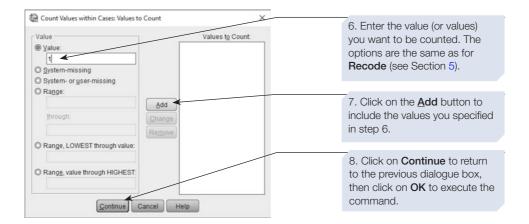
Section 7: COUNTING VALUES

Sometimes it is useful to be able to count for each participant how many times a particular value occurs over a range of variables. If, as in our example data set, you have a series of variables that represent the responses to questionnaire items, you might want to find out how many times each participant has answered 'Strongly Agree'. You could do this by asking SPSS to count the number of times the value 1 (the value used to code the response 'Strongly Agree') has occurred in variables 'q1' to 'q10'. Using **Count Values within Cases**, SPSS will create a new variable that will contain a value representing the number of times the value 1 occurs in variables 'q1' to 'q10'.

Chapter	<u>View</u> Data	y cleaned ed 7 (2) sav [DataSet16] - IBM SPSS Statistics Data Edit Iransform Analyze Graphs Utilities Extensions	1. Click on menu item <u>Transform</u> .
26 : TransNat	Adopt2	Programmability Transformation Count Values within Cases	2. Select Count Values within Cases.
1	2.00		
2	1.00	Recode into Same Variables	
3	2.00	Recode into Different Variables	
4	2.00	Automatic Recode	
5	2.00	Create Dummy Variables	
6	2.00	Visual Binning	
7	1.00	C Optimal Binning	
8	1.00	Prepare Data for Modeling	
9	2.00	Rank Cases	
10	2.00		
11	2.00	Date and Time Wizard	
12	1.00	Create Time Series	
13	2.00	Replace Missing Values	
14	1.00	Random Number Generators	
15	1.00	Run Pending Transforms Ctrl+G	
16	2.00	4.00 1.00 4.00 1.00	
47	2.00	003 00 000	



When selecting more than one variable – as in step 4 above – you can select them all in one go by clicking on the first and then holding down the <shift> key while you click on the last of the variables. You can then use the mouse to drag and drop the variables as a block.



V									
NatAdopt	A TransNatAdopt2	& Num_SA							
4.00	4.00	.00	-						
4.40	4.40	3.00							
1.60	1.60	5.00							
	1.75	4.00							
3.80	3.80	.00							
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2.40	2.40	6.00							
1.40	1.40	5.00							
1.20	1.20	8.00							
2.00	2.00	1.00							
5.00	5.00	.00							
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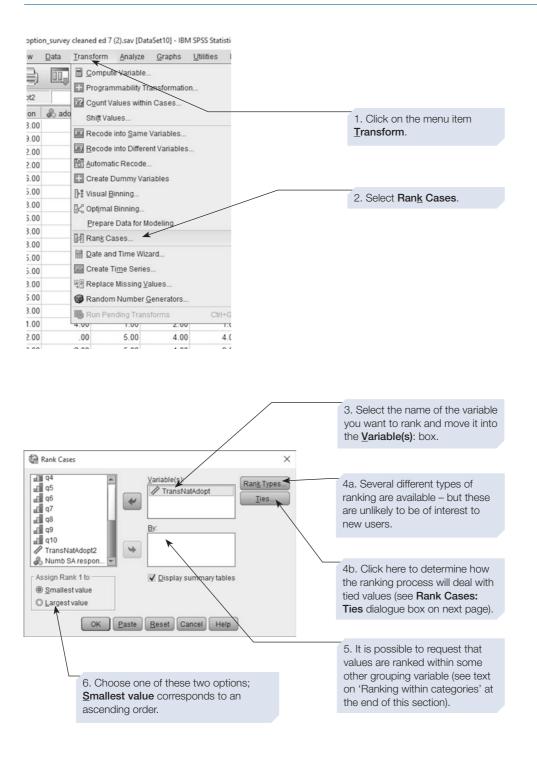
9. This is the new variable. We can see that participant 1 never responded 'Strongly Agree', while participant 2 used this response a total of 3 times across variables 'q1' to 'q10'.

Conditional Count

It is possible to perform a conditional count – where we only count the occurrences of a value(s) for participants who satisfy some particular criterion. This is done by clicking on the <u>If</u> button after step 5 above. This will bring up a dialogue box almost identical to the one we used for the conditional recode described in Section 5. You can now specify your conditional rule and then click on the **Continue** button.

Section 8: RANKING CASES

Sometimes it is useful to convert interval or ratio scores (scale data) into ordinal scores. We might, for example, want to convert the variable 'TransNatAdopt', which we calculated using <u>Compute</u> in Section 6, into a rank score. That is, we might want to rank all our participants on the basis of their score on this variable. The participant who had the highest overall 'TransNatAdopt' score would be given a rank of 1, the next highest a rank of 2 and so on. The **Rank Cases** command calculates the ranks for us and generates a new variable to contain the ranks. We can rank in either ascending or descending order, and can even rank on the basis of more than one variable.



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Rank Cases: Ties X	7. You will probably want tied values represented by the Mean rank (see text below for a discussion of the alternatives).
© Sequential ranks to unique values	8. Click on the <u>Continue</u> button to return to the previous dialogue box. Then click OK to perform the ranking.

File Edit	<u>V</u> iew <u>D</u> ata]	'n	
27 : q8			_
	and ID	TransNatAdopt	
1	1	0 4.0	00
2	2	0 4.4	10
3	3	0 1.6	50
4	4	0	
5	5	0 3.8	30
6	6	0	
7	7	0 2.4	10
8	8	0 1.4	0
9	9	0 1.2	20
10	10	0 2.0	00
11	11	0 5.0	00
12	12	0 3.0	00
13	13	0 3.0	00
14	14	0 3.4	10
15	15	0	
16	16	0 1.6	50
17	17	0 4.4	0
18	18	0	
19	19	0 1.2	20
20	20	0 2.2	
04			

A TransNatAdopt	

latAdopt		RTransNa	
4.00)	13.000	
4.40)	14.500	
1.60)	4.500	
)		
3.80)	12.000	
)		
2.40)	8.000	
1.40)	3.000	
1.20)	1.500	
2.00)	6.000	
5.00)	16.000	
3.00)	9.500	Y
3.00)	9.500	
3.40)	11.000	
)		/
1.60)	4.500/	/
4.40)	14.500	
)	۴.	
1.20)	1.500	
2.20)	7.000	

9. This is the new variable containing each participant's rank on the variable 'TransNatAdopt'. SPSS has created a new variable name by prefixing the first few characters of the old name with the letter 'R' (for rank). You should amend this or add a variable label to remind you what it means.

10. Participants 12 and 13 both had a score of 3.0 on 'TransNatAdopt' so have been given the mean rank of 9.5.

11. Note that, if the value of the original variable is missing, then SPSS will assign a systemmissing value to the rank.

Ranking tied values

SPSS provides four alternative methods of dealing with tied values:

- The default, <u>Mean</u> method, gives the tied values the mean of the available ranks. You can see this option in operation above where participants 12 and 13 both scored 3.0 on 'TransNatAdopt' and were given a rank of 9.5 – the mean of ranks 9 and 10. This is the ranking method described in most introductory statistics books.
- 2. The **Low** option assigns the tied participants the lowest of the available ranks so in this case both would have been ranked 9.
- 3. The <u>**High**</u> option would award both participants the highest of the available ranks -10 in this case.
- 4. The <u>Sequential ranks to unique values</u> option would assign a rank of 10 to both participant 12 and 13, but would then assign a rank of 11 to the next highest-scoring participant, thus ensuring that all the sequential ranks are awarded this means that the highest rank will *not* be equal to the number of valid cases (as it would for the other three methods).

Ranking within categories

By specifying a second variable in the $\underline{B}y$: box (see step 5 above), it is possible to request SPSS to rank the scores on the first variable within categories formed by the second variable. For example, if we specified the variable 'sex' in this box, then SPSS would first rank all the male participants and then rank all the female participants. Thus, in this case, we would have two participants (one male and one female) with each rank.

Section 9: DATA TRANSFORMATION

Data transformation involves applying a mathematical function to every value in a data set to systematically change the values. At its simplest, data transformation could involve adding a constant, or converting from a raw score into a percentage score. For example, we might give students a simple statistics test scored out of 11. We could then transform these raw scores into percentage values by applying the formula PercentScore = $(RawScore/11)^*100$.

One of the most common data transformations used in psychology is the log transformation. Log transformations (or log transforms) involve replacing every value in a data set with the log of the value. You will remember from your high school mathematics classes that the logarithm of a value is the exponent to which the base must be raised to produce that number. Common logs (written as log_{10}) use a base of 10, while natural logs (log_{e} , sometimes written as ln) use the mathematical constant *e* (approximately 2.71828) as their base. For example, because $10^2 = 100$, the common log of 100 is 2. This is written as $log_{10} (100) = 2$. The natural log of 100 is approximately 4.605. That is, loge (100) = 4.605, and $e^{4.605} = 100$.

The table below shows the relationship between raw values (*x*), and the \log_{10} and \log_e of these values.

x	1	10	100	1000	10000	100000	1000000
Log ₁₀ (x)	0	1	2	3	4	5	6
Log _e (x)	0	2.30	4.61	6.91	9.21	11.51	13.82

Looking at this table, we can see one useful characteristic of logs. With each order of magnitude increase in the value *x*, the log of this value increases by a much smaller constant amount (in the case of common logs increasing by just 1). We can use this characteristic of logs to our advantage to rescale certain types of data. Often data are not normally distributed, showing either a positive or a negative skew. For example, a data set in which most values are close to the mean value, with a few much higher values, is said to have a 'positive skew' and when plotted in a histogram will appear asymmetric around the mean, with a longer tail to the right than the left. Skewed data such as these can invalidate certain inferential statistical tests. A possible solution is to undertake a log transform of the data by calculating the log of each raw value. This is often enough to reduce the skew in the data, allowing us to proceed with our analysis using the transformed data. We demonstrate this process below.

Log transformation of decision latency data

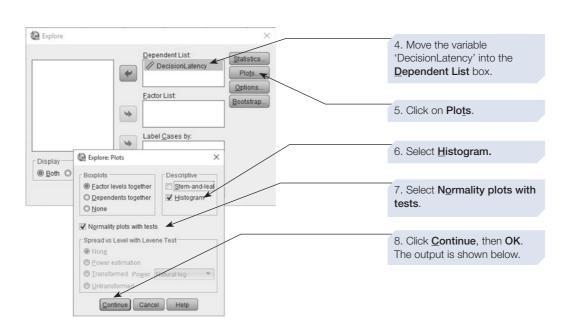
To illustrate the use of log transformations, we will work with a file containing decision latency data from 129 trials in which participants were required to decide whether two photographs were of the same or different people. The time taken to reach this decision was recorded. Time-based data such as these are often positively skewed because there is no upper limit to how long participants can take to respond. Participants need a finite minimum amount of time to respond to a stimulus, but they could lose concentration and take seconds, minutes or even hours to respond. As a result, decision time data typically have a strong positive skew.

To follow this example, download the data file called 'LogTransformDemo.sav' from macmillanihe.com/harrison-spss-7e.

First, we will use the **Explore** command to check whether the raw data are normally distributed.



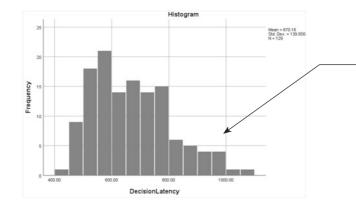




DecisionLatency						
Decision stancy					Statistic	Std. Error
DecisionLatency	Mean				670.1628	12.31360
	95% Confidence I	Interval	Lower	Bound	645.7982	
	for Mean		Upper	Bound	694.5274	
	5% Trimmed Mea	in			663.7127	
	Median				653.0000	
	Variance				19559.590	
	Std. Deviation				139.85561	/
	Minimum				437.00	
	Maximum				1057.00	
	Range				620.00	
	Interquartile Rang	le			193.00	
	Skewness				.618	.213
	Kurtosis				228	.423

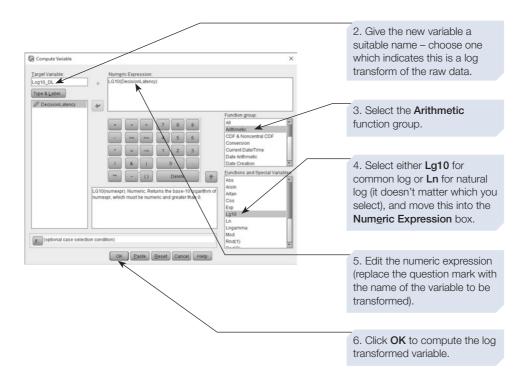
This table reports the Skewness statistic. The value of 0.618 tells us that the data are positively skewed and there is a moderate degree of skew (a value greater than 1 indicates a high degree of skew, a value of zero indicates no skew).

This table reports two tests of the 'normality' of the distribution (the Kolmogorov– Smirnov and Shapiro–Wilk tests). Both are significant (Sig value less than .05), indicating that this distribution is significantly different from a normal distribution.



The problem is evident from this histogram. The positive skew is seen as a longer tail to the right than to the left of the distribution. We need to transform this data. This process is described next.

ile	<u>E</u> dit	View	Data	Transform	Analyze	Graphs	Utilities	Extensi	on	
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:				Count V	alues within	n Cases			-	
		& De	cisionLat	Shift Val	lues			ra	1	-
	1		5	Recode	into Same	Variables				1. Select Transform , th
	2		5	Recode						<u>C</u> ompute Variable.
	3		~				5			
	4			Automa						
	5		8	Create I	Dummy Var	riables				
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		103	
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			Ln_DL
1	510.00	2.71	6.23
2	586.00	2.77	6.37
3	509.00	2.71	6.23
4	718.00	2.86	6.58
5	853.00	2.93	6.75
6	996.00	3.00	6.90
7	620.00	2.79	6.43
8	590.00	2.77	6.38
9	568.00	2.75	6.34
10	533.00	2.73	6.28
11	1057.00	3.02	6.96
12	469.00	2.67	6.15
13	534.00	2.73	6.28
14	520.00	2.72	6.25
15	549.00	2.74	6.31
16	629.00	2.80	6.44
17	437.00	2.64	6.08
18	468.00	2.67	6.15
19	608.00	2.78	6.41
20	534.00	2.73	6.28
21	612.00	2.79	6.42
22	585.00	2.77	6.37

*LogTransformDemo.sav [DataSet11] - IBM SPSS Statistics Data Editor Eile Edit View Data Transform Analyze Graphs Utilities

The new variable is added to the data file. To illustrate this process, we have computed two new variables: the first is a log_{10} transform, the other a log_e transform. Note that these two transformations have different values, but as we will see shortly, they produce similarshaped distributions. Now repeat the **Explore** analysis described above. This time include both the original variable and the transformed variable(s).

Explore	×	
Dependent List Dependent List DecisionLatency Log10_DL Eactor List Eactor List Label Gases by: Display @ Both © Statistics © Plots OK Paste Reset Cancel Help	Statietros Piots Options Bootstrap	Repeat the Explore analysis with the new transformed variable(s) included in the Dependent List .

Sections of the output of the Explore analysis are shown below.

			Statistic	Std. Error	_
DecisionLatency	Mean		670.1628	12.31360	
	95% Confidence Interval	Lower Bound	645.7982		-
	for Mean	Upper Bound	694.5274		
	5% Trimmed Mean		663.7127		
	Median		653.0000		
	Variance		19559.590		
	Std. Deviation		139.85561		
	Minimum		437.00		
	Maximum		1057.00		
	Range		620.00		
	Interquartile Range		193.00		
	Skewness		.618	.213	The Skewness value of the
	Kurtosis		228	.423	original variable was 0.618,
	Mean		2.8171	.00779	indicating a moderate degree
	95% Confidence Interval	Lower Bound	2.8017		of positive skew. Compare th
	for Mean	Upper Bound	2.8325		to
	5% Trimmed Mean		2.8155		
	Median		2.8149		
	Variance		.008		the Skewness value of the
	Std. Deviation		.08851	/	transformed variable. The
	Minimum		2.64		value is now 0.225, indicating
	Maximum		3.02		a greatly reduced degree of
	Range		.38		skew.
	Interquartile Range		.13		
	Skewness		.225	.213	
	Kurtosis		690	.423	

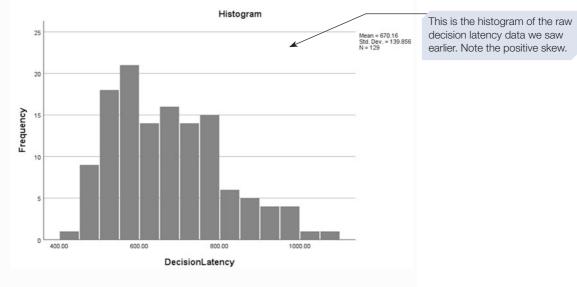
Descriptives

Tests of Normality

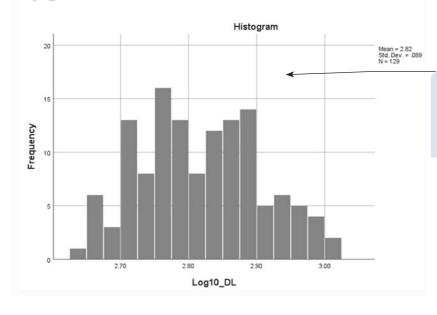
DecisionLatency	.084	df 129	Sig.	Statistic	df	Sig.	
DecisionLatency	.084	129	0.05				
			.025	.859	129	.001	The Kolmogorov–Smirnov
Log10_DL	.064	129	.200	.981	129	.070	and Shapiro-Wilk tests
Ln_DL	.064	129	.200	.981	129	.070	show no significant devia from normality for the two

a. Lilliefors Significance Correction

DecisionLatency







This is the histogram of the data after a log₁₀ transformation. It is apparent that the distribution is now much more symmetric around the mean.

Section 10: DATA FILE FOR SCALES OR QUESTIONNAIRES

In this section, we demonstrate how SPSS can be used to help you handle data obtained using scales or questionnaires. We describe a simple data check, and how to recode responses from reversed items. Checking the reliability and dimensionality of a scale is covered in Chapter 13, Sections 4 and 5.

We have entered data obtained with Larsen's (1995) Attitudes Toward Recycling (ATR) scale, used for a research methods exercise with first-year psychology students at the University of Westminster. There are 20 items, used in the order they are printed in Larsen (1995, Table 1). Two changes were made, as Larsen developed his scale in USA: 'styrofoam' was replaced with 'polystyrene' and 'sorting garbage' was replaced with 'sorting rubbish into different containers'. We used a Likert-type scale with responses from 1 (strongly agree) to 5 (strongly disagree). The data file 'ScaleV1.sav' (available from the Appendix or macmillanihe.com/harrison-spss-7e) includes data from 50 cases, which we will use to demonstrate some issues around the use of scales in psychology. Normally one would need many more cases.

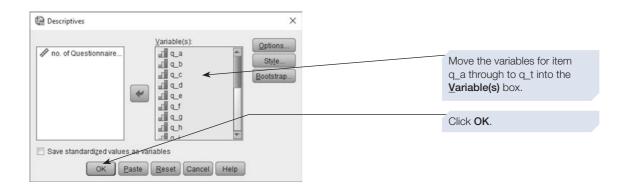
Open the data file and note that:

- 'Qnum' (Questionnaire Number) is the identification number we wrote on the paper copies of the questionnaires as we entered the data. We have included it so we can check our data file against our paper records, even if the order of cases in the data file is changed by sorting the file (see Section 2).
- 2. The responses to each item have been entered separately: 'q_a' to 'q_t'. Students often prefer to calculate the total or mean score by hand and then enter these values only. However, it is better practice to enter the response to each item. This allows us to undertake various checks on the data before using <u>Compute</u> to calculate the total or mean response for each participant.
- 3. In this data set we have used 9 as the missing value for each of the scale items.

A simple check on data entry

Using the process described in Chapter 3, Section 6, we will check the data using the **Descriptives** command. Follow the steps below.

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ا م_4 ا	Ta <u>b</u> les			*	- a Explor	re		d .	
2	Compare			*	T Cross	tabs		_	
2	General			*	TURF	Analysis			
2		Generalized Linear Models				Ratio			
2	Mixed Mo			*	P-P Plots				
2	Correlate			1	Q-Q P				
2	Regress					4			
1	Loglinea				2	5	2		
2	Neural N	etworks		1	3	5	3		
1	Classify				2	5	1		



Descriptive Statistics

	Ν	Minimum	Maximum	Mean	Std. Deviation	
q_a	50	2	5	3.96	.832	This is a section of the output.
q_b	50	1	5	2.08	.913	Note that the maximum value
q_c	50	2	6 🖊	3.94	1.058	for the variable q_c is 6 but all
q_d	50	2	5	4.30	.707	these variables were coded 1 to 5 so this must be an error.
q_e	50	2	5	4.48	.646	5 SO THIS MUST DE AIT EITOI.
q_f	49	1	5	4.02	1.051	
q_g	50	1	4	1.64	.749	
q_h	50	1	4	2.22	.737	
q_i	50	1	5	2.44	.993	
a i	50	1	5	1 32	262	

We now need to find and correct this error in the variable q_c.

🕼 scaleV1 Ed7.sav [DataSet14]	- IBM SPSS Statistics Data Editor
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		🖉 Qnum	ြ ရ_a	q_b	 c	q_d	al
	1	1	4	2	3	4	
	2	2	2	2	6	3	
	3	3	4	2	4	4	
	4	4	4	2	4	5	
	5	5	4	2	5	4	
	6	6	4	2	3	3	
	7	7	5	1	5	4	
	8	8	5	2	5	5	
	9	9	4	1	5	5	
-	10	10	5	2	5	5	
-	11	11	4	2	3	4	
-	12	12	4	1	4	4	
-	13	13	4	1	3	4	
	14	14	4	1	3	4	
		i		-	-		

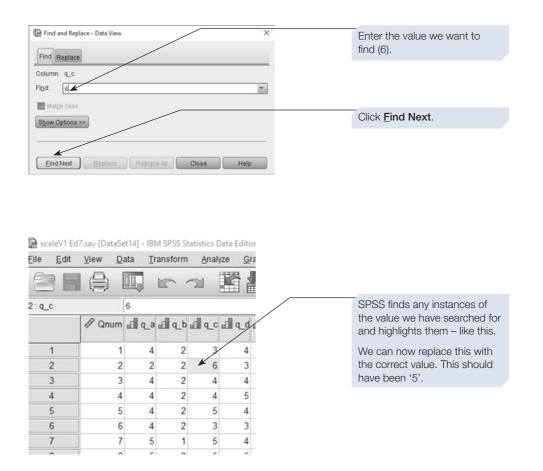
Click here on the header for the variable q_c to select the column.

a scaleV1 Ed7.sav [DataSet14] - IBM SPSS Statistics Data Edite



Click <u>Edit</u> \Rightarrow <u>Find</u>.





Now save the corrected data file. Call it 'ScaleV2.sav', so that you can use it in the next exercise.

Reversals

Good scales or questionnaires often include items that are reverse coded to avoid a participant response bias. We can use SPSS to reverse the coding of these items. First consider what you want a high score to indicate. For example, consider these two items from a library satisfaction questionnaire:

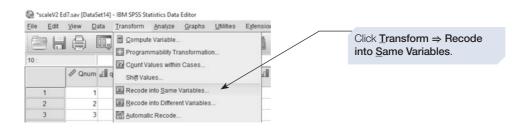
- 1. The university library is an excellent place to make notes for coursework.
- 2. I find it very difficult to study in the university library.

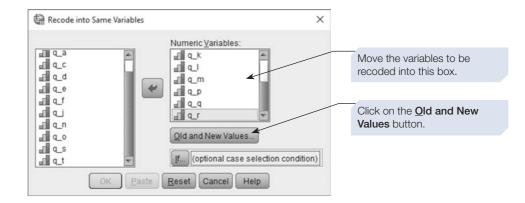
With responses on a scale of 1 (strongly agree) to 5 (strongly disagree), an individual with a strong positive view of the library should respond in opposite directions to those two items. Do you want your final score to represent overall *satisfaction* with the library or overall *dissatisfaction*? This will determine which items you reverse.



This type of data can be scored in either direction. Use the variable label to keep a note of the direction of the scoring.

In the case of the 'ScaleV2.sav' data file, we want a high score to indicate that the participant has positive attitudes towards recycling and related environmental issues, so items b, g, h, i, k, l, m, p, q and r need to be reversed. We can do this using the **Recode** command. Normally it is safer to use **Recode** into Different Variables because the original variables are preserved in case we make a mistake. However, when we need to recode lots of variables in the same way, as here, it is much quicker to use **Recode** into <u>Same</u> Variables, so that is what we will do here. Remember to save a backup copy of the data file before making these changes in case of errors.





Recode into Same Variables: Old and New Valu Old Value Zalue:	New Value Value: System-missing	× Enter the Old and New values. Note that in addition to reversing the scale we have also recoded System-or <u>u</u>ser-
System-missing System- or yser-missing Range: prough	Old→ New: 1→5 2→4 3→3 4→2	missing values into System missing values.
Range, LOWEST through value: Range, value through HIGHEST:	Remove 5->1	Click here to add the last recode.
O All other values		
 [20	tinue Cancel Help	Click <u>Continue</u> then OK.

The relevant variables should now be reverse coded. The first few lines of the data file should look like this.

	A Qnum	يا م	يتا وه	يا وه	يطا وره	يا وه	يا ال	يا م	بدأأ ورا	نه ال	نه الت	يطأ وراد	يا ب	dl ^۹ m	يطا وره	يا و ه	du		dle	des	يا عرد
1	1	4	4	3	4	4	4	4	4	4	4	4	4	3	5	4	3	1	3	3 3	5
2	2	2	4	5	3	4	4	5	4	3	4	4	4	2	- 4	5	3	3	3	3 3	2
3	3	4	4	4	4	4	5	4	4	4	5	3	4	3	2	5	4	3	3	3 3	4

Resave the data file as 'ScaleV3.sav'. The file is now ready to use in other exercises on assessing the reliability and dimensionality of the scale (Chapter 13, Sections 4 and 5).

Summary

- This chapter demonstrated a series of useful commands that allow us to manipulate the contents of the data file.
- Using these commands, it is possible to change the coding of variables, to compute new variables, or to select only certain cases for subsequent analysis.
- These commands are particularly valuable when managing large data sets such as those produced in survey research.
- The compute command can also be used to transform data so these meet the assumptions of inferential tests. One of the most common transformations, the log transform, was demonstrated, and we saw how this normalised a set of previously skewed data.
- Section 10 illustrated the use of some of these commands in a survey study.

Tests of difference for one- and two-sample designs

In this chapter

- An introduction to the *t*-tests
- The one-sample t-test
- The independent *t*-test
- The paired *t*-test
- An introduction to nonparametric tests of difference
- The Mann–Whitney test
- The Wilcoxon test



SPSS for Psychologists online

Visit macmillanihe.com/harrison-spss-7e for data sets, online tutorials and exercises.

Section 1: AN INTRODUCTION TO THE t-TEST

- The *t*-test is used to determine whether two means are significantly different from one another.
- There are three types of *t*-test:
 - The one-sample *t*-test, which is the simplest, determines whether the observed mean is different from a part particular value (such as a population mean).
 - The independent *t*-test is used when comparing means from two independent groups of individuals.
 - The paired *t*-test is used when comparing the means of two sets of observations from the same individuals (e.g. repeated measures design) or from pairs of individuals (e.g. when using a matched-subjects design).
- All forms of the *t*-test are parametric tests and make certain assumptions about the data: that they are measured at interval or ratio level, meet the assumption of homogeneity of variance and are drawn from a population that has a normal distribution (see Chapter 1, Section 2).

CHAPTER 5

- When reporting descriptive statistics to accompany the results of the *t*-test, you should give the mean and standard deviation as the measures of central tendency and dispersion.
- In some textbooks you might find this test referred to as the Student's *t*-test. This is because William Gossett, who devised the test, worked for the Guinness Brewing Company, which did not permit him to publish under his own name, so he wrote under the pseudonym of 'Student'.

Section 2: THE ONE-SAMPLE t-TEST

We mentioned in Section 1 that the t-test is used to determine whether two means are significantly different from one another. The one-sample t-test allows us to compare the sample mean against a specific score. For example, we might want to see whether the mean from a group of participants differs significantly from chance performance, or from some other predefined value, such as the population mean IQ score, or an expected value based on previous research. This test should be used when the data meet the assumptions for a parametric test.

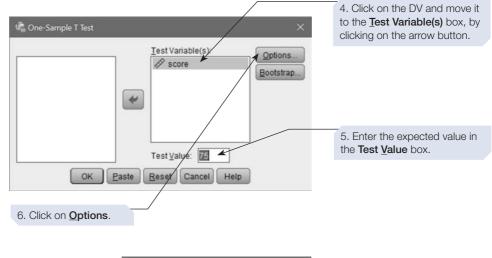
Example study: assessing memory

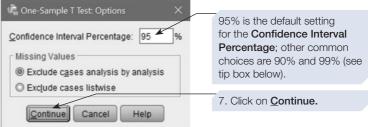
For the purposes of demonstrating how to perform the one-sample *t*-test, we will use fictitious data from a made-up study. An enthusiastic teacher, who also happens to be a psychology graduate, noticed that some of the children in the final year of primary school class consistently forgot to bring in their sports kit, their lunch box or their homework, and frequently forgot the instructions given to them in class. She wondered if this was because they were generally more forgetful due to poorer memory skills. She therefore asked permission to administer a standardised test of long-term memory, which had a published norm of 75% for children of this age group. She performed the one-sample *t*-test on the data she collected to test the hypothesis that the mean of her sample differed from the expected value of 75%. The results showed that her sample mean was significantly lower than the expected value.

If you use this fictitious data set and follow the instructions given next, you will be able to compare the output you produce with the annotated output we give at the end of this section. (These data are available in the Appendix or from macmillanihe. com/harrison-spss-7e.)

To perform a one-sample t-test

ile <u>E</u> dit	View Data Transform		Extensions F		2. Click on Compare Means .
		Descriptive Statistics	,		2. Choix off Compare Meane.
		Bayesian Statistics		Visible: 1 of 1 Variables	
	& score var	Tables		var var var	
1	68.00	Compare Means		Means	
2	62.00	General Linear Model	Jeneral Linear Model One-Sample T Test.		
3	58.00	Generalized Linear Models			
4	67.00	Mixed Models			
5	65.00	Correlate Regression			
6	69.00			Qne-Way ANOVA	
7	72.00				3. Click on One-Sample T Tes
8	76.00	Neural Networks			
9	62.00	Classity Dimension Reduction Scale Nonparametric Tests			
10	64.00				
11	69.00				
12	70.00				
13	71.00				
14	66.00	Forecasting			
	4	Survival		14	
Data View	/ariable View	Multiple Response	,		
	T Test_	Multiple Imputation		ocessor is ready Unicode ON	





8. Finally, click in the **One-Sample T Test** dialogue box. The output for the *t*-test will appear in the Output window, and we show this with annotations below. We also show you how you would describe the results of the test, were you to write a report on this study.

Confidence intervals describe the limits within which the population mean is likely to fall (see Chapter 1, Section 1). The default, 95%, indicates that there is a 95% probability of the population mean falling between the limits shown in the output. SPSS allows us to make these confidence intervals narrower or wider. For example, if we are conducting a study with important clinical implications, where we want to estimate the population mean with a higher degree of precision, we might choose 99% confidence intervals.

SPSS output for one-sample t-test

Obtained using menu items: Compare Means > One-Sample T Test

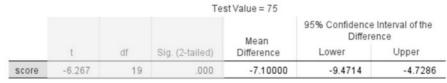
test.

T-Test

One-Sample Statistics

	Ν	Mean	Std. Deviation	Std. Error Mean	
score	20	67.9000	5.06692	1.13300	

One-Sample Test



So, t = 6.27. Note that we ignore the minus sign.

The degrees of freedom (df) = 19.

Sig. (2-tailed) is the *p*-value for a two-tailed test. This needs to be smaller than .05 for the result to be significant. While this value is displayed as .000, a *p*-value can never equal zero. SPSS rounds *p*-values to three decimal places, so here *p* must be less than .0005 (see tip box below on reporting *p*-values).

Useful descriptive statistics showing the number of children and their mean score (*M*) and standard deviation (*SD*) on the memory



Reporting the results

In a report you might write:

Performance on a memory test by a group of forgetful children was on average lower (M = 67.9%, SD = 5.07) than the expected value of 75%. A one-sample *t*-test showed that this difference was significant (t = 6.27, df = 19, p < .001).



Guidance provided by the American Psychological Association recommends that we report statistics such as *t* to two decimal places, and *p*-values less than .001 as *p* < .001. If *p*-values are greater than .001, the exact value should be reported up to two or three decimal places. Note: A leading zero is not used when reporting values that can never be greater than 1, such as *p*-values.

Section 3: THE INDEPENDENT t-TEST

The independent *t*-test compares the performance of the participants in one group with the performance of the participants in a different group. This test should be used when the data meet the assumptions for a parametric test and are obtained using an independent groups design. These two groups could constitute a male and a female group if you wanted to examine sex differences, or they could constitute two groups of participants who undergo different drug conditions: one a low-dose drug condition and one a high-dose drug condition. This type of *t*-test is often also called an *unrelated t*-test.

Example study: the memory experiment

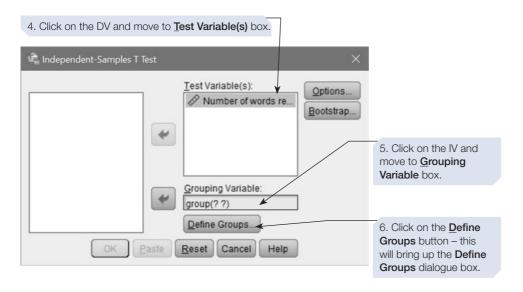
In the example shown next, we use the data from the memory experiment used in the data entry exercise in Chapter 2, Section 6. It was hypothesised that the group receiving mnemonic instructions would remember more words than the group who did not receive any specific mnemonic instructions. As this hypothesis is directional, it is a one-tailed hypothesis. If you use these data and follow the instructions given next, you will be able to compare the output you produce with the annotated output we give at the end of this section.

	View Data	Transform	Analyze Graphs Utilities B	Extensions Window Help Analyze.
38	B group	e score	Reports Descriptive Statistics Bayesian Statistics Tables	2. Click on Compare
1	go group	20.00	Compare Means	Means.
2	1	18.00	General Linear Model	Means
3	1	14.00	Generalized Linear Models	One-Sample T Test
4	1	18.00	Mixed Models	Independent-Samples T Test
5	1	17.00	Correlate	💹 Paired-Samples T Test
6	1	11.00	Regression	One-Wax ANOVA
7	1	20.00	Loglinear	
8	1	18.00	Neural Networks	3. Click on Independent-
9	1	20.00	Classify	
10	1	19.00	Dimension Reduction	Samples T Test.
11	1	20.00	Scale	
12	2	10.00		
13	2	20.00	Nonparametric Tests	
14	2	12.00	Forecasting	
15	2	9.00	<u>S</u> urvival	
16	2	14.00	Multiple Response	
17	2	15.00	Missing Value Analysis	
18	2	16.00	Multiple Imputation	>
ata View	Variable View		Complex Samples	•

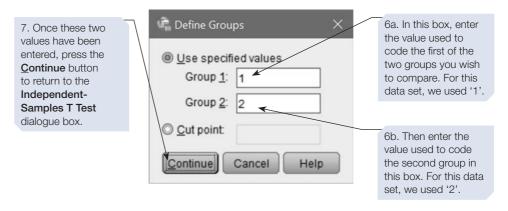
To perform an independent *t*-test

- 4. You will now be presented with the **Independent-Samples T Test** dialogue box (see below). As is typical in SPSS, the box on the left lists all the variables in your data file. Click on the name of the dependent variable (DV) in your analysis and then click on the arrow button to move this variable name into the box marked **Test Variable(s)**.
- 5. Now click on the name of the independent variable (IV) and then click on the arrow button to move this into the box marked <u>Grouping Variable</u>.

Once you have entered the dependent and independent variables into their appropriate boxes, the dialogue box will look like this.



6. Click on the Define Groups button to bring up the Define Groups dialogue box (see below). This dialogue box is used to specify which two groups you are comparing. For example, if your independent variable is 'sex', which you have coded as 1 = Male and 2 = Female, then you need to enter the values 1 and 2 into the boxes marked Group 1 and Group 2, respectively. This might seem rather pointless, but you might not always be comparing groups that you had coded as 1 and 2. For example, you might want to compare two groups who were defined on the basis of their religious belief (Atheists and Christians, who could be coded as 0 and 2, respectively – see Chapter 2, Section 2, on value labels). In this case, we would enter the values 0 and 2 into the two boxes in this dialogue box. (We will not be describing the use of the <u>Cut point</u> option here.)



- 7. Clicking on the <u>Continue</u> button in the <u>Define Groups</u> dialogue box will return you to the <u>Independent-Samples T Test</u> dialogue box. You will see that your two values have been entered into the brackets following the name of your independent variable (you may have noticed that previously there were question marks inside these brackets).
- 8. Finally, click on **or** in the **Independent-Samples T Test** dialogue box. The output of the *t*-test will appear in the Output window.

The output from this independent *t*-test is shown, with annotations, below. We also show you how you would describe the results of the test, were you to write a report on this experiment.

SPSS output for independent groups t-test

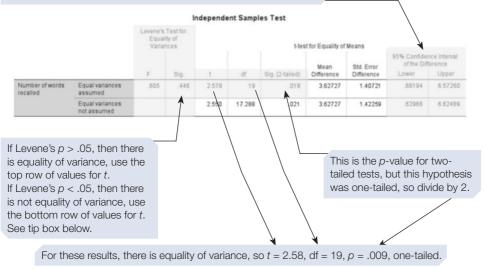
Obtained using menu items: Compare Means > Independent-Samples T Test

T-Test

Useful descriptive statistics showing that those in the Mnemonic condition remembered the most words. You can calculate the effect size from these descriptives (see next page).

	Group S	tatistics			
	Condition	Ν	Mean	Std. Deviation	Std. Error Mean
Number of words	Mnemonic condition	11	17.7273	2.86674	.86435
recalled	No mnemonic condition	10	14.1000	3.57305	1.12990

These lower and upper confidence limits enclose the confidence interval. There is a 95% probability that the difference between the means for the population will fall between 0.68 and 6.57. Note, 0 is not within those limits: remember that indicates a significant difference (Chapter 1, Section 1).



Equality (or at least similarity) of variance is one of the requirements for using parametric statistical tests. SPSS, however, carries out two versions of the independent groups t-test: the top row for when there is equality of variance and the bottom row for when the variances are unequal. If you use the latter in a report, you must note that fact.

Measure of effect size

The output for the independent groups *t*-test does not include an estimate of the size of the effect, and this is not an option on SPSS that you can select. However, from the output you can calculate this, and we show you how to do so next.

Cohen's d (see Chapter 1, Section 1) is a measure of effect size that is frequently reported in journal articles when reporting t-tests (Fritz, Morris and Richler, 2012). Essentially, this involves working out the difference between the two means from each condition and dividing this by the two standard deviations (*SD*) combined. You use the following formula to calculate d when the number of participants in each condition (*N*) is identical or similar. Consult a statistics text on an alternative way of combining standard deviations if there is a large discrepancy in N in your data.

$$d = \frac{\left(x_1 - x_2\right)}{\text{mean } SD}$$

This formula requires you to carry out the following steps:

1. Look at the SPSS output and identify the mean and standard deviation for each condition, provided in a table called **Group Statistics**. For the data set used in this section:

The mean for the mnemonic condition, $x_{1,} = 17.73$ and the SD = 2.87. The mean for the nomnemonic condition, $x_{2,} = 14.10$ and the SD = 3.57.

2. Take the mean of one condition from the mean of the other condition (it is not important which mean to take from which, so you can ignore the sign). For the data set in this section:

$$17.73 - 14.10 = 3.63.$$

3. Find the mean *SD* by adding the *SD* for each condition together and dividing by two: (*SD* of condition 1 + SD of condition 2) / 2). For the data set in this section:

$$(2.87 + 3.57) / 2 = 3.22.$$

4. Use the formula to calculate *d*. For the data set in this section:

$$d = (17.73 - 14.10) / 3.22 = 3.63 / 3.22 = 1.13.$$

This would be considered to be a large effect size. Cohen (1988) provided the following guidance on how to interpret *d*:

- Small effect size: 0.2 (or more)
- Moderate effect size: 0.5 (or more)
- Large effect size: 0.8 (or more).

Reporting the results

In a report you might write:

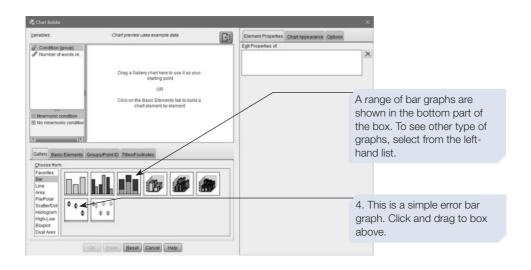
More words were recalled in the Mnemonic condition (M = 17.73 words) than in the Non-mnemonic condition (M = 14.10 words). An independent *t*-test showed that the difference between conditions was significant, and the size of this effect was large (t = 2.58, df = 19, p = .009, one-tailed, d = 1.13).

It would be helpful in your report to include more than just the descriptive statistics that are automatically generated as part of the independent *t*-test output. In Chapter 3, Section 5 we showed you how to use the **Explore** command, and the output using this command with the data from the memory experiment is included at the end of that section. This output provides you with the 95% confidence interval for the mean of each condition. Our estimate of the population mean for the Mnemonic condition (based on our sample) is 17.73, and the confidence intervals tell us that there is a 95% probability that the true value will fall between 15.80 and 19.65. Our estimate of the population mean for the Non-mnemonic condition is 14.10, and the confidence intervals tell us that there is a 95% probability that the true value will fall between 11.54 and 16.66. A useful way of presenting this information in your report is to show it graphically using an error bar graph. Next, we show you how to create such a graph using SPSS.

Creating an error bar graph: independent groups design

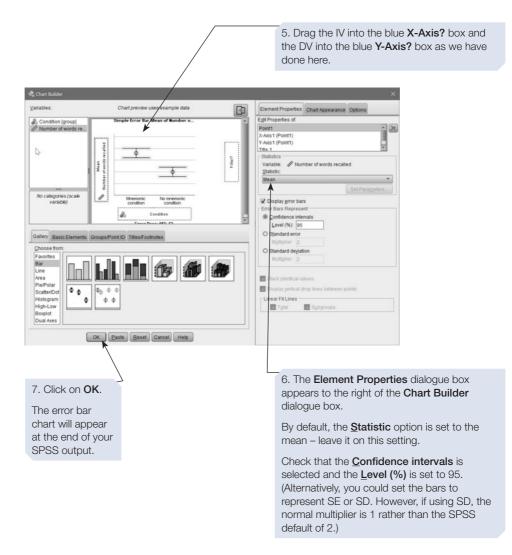
To obtain an error bar graph using Chart Builder

- 1. On the menu bar, click on Graphs.
- 2. Click on Chart Builder.
- 3. Close the top **Chart Builder** dialogue box, which reminds you that measurement level should be set properly for each variable in your chart (see Chapter 3, Section 8). The **Chart Builder** dialogue box below will be visible.

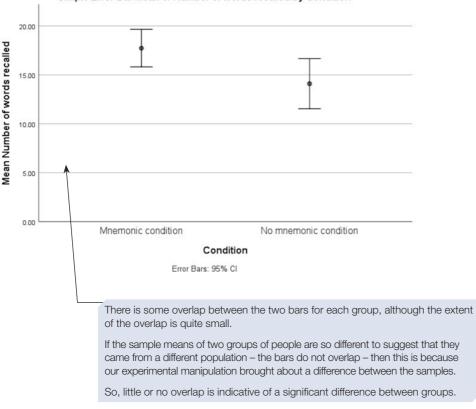




We are showing you how to create an error bar graph where the error bars are selected to represent the 95% confidence intervals. The confidence interval tells us the limits within which the population mean (the true value of the mean) is likely to fall. SPSS allows alternative options where the bars represent a measure of dispersion such as the standard deviation (see below).



SPSS output for bar chart



Simple Error Bar Mean of Number of words recalled by Condition

Section 4: THE PAIRED t-TEST

In the repeated measures design, data are collected from each participant in all conditions (or levels) of the independent variable. For example, we might compare participant 1's memory performance under noisy conditions with participant 1's memory performance under quiet conditions. In this situation, it is likely that the data from participants will be correlated; for example, if participant A has a good memory, then their scores on a memory test will be high regardless of condition. It is for this reason that a repeated measures *t*-test is sometimes called a *correlated t*-test. With a repeated measures design, it is essential that the data are kept in the correct order, so that participant 1's data on variable A are indeed compared with participant 1's data on variable B. The test itself considers pairs of data together, and for this reason this test is also known as a *paired t*-test.

Example study: the mental imagery experiment

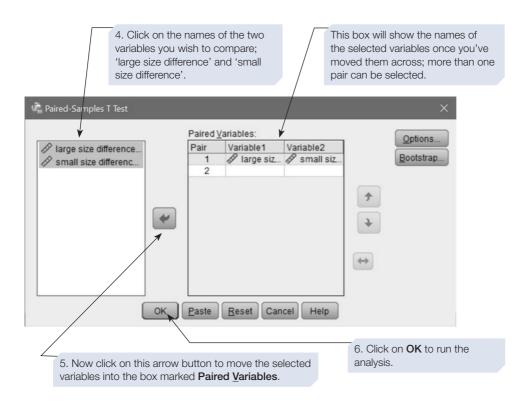
To demonstrate the use of the paired *t*-test, we are going to analyse the data from the mental imagery experiment, the second data entry exercise in Chapter 2, Section 6. It was hypothesised that, as participants would compare their mental images of the two animals to determine which was the larger, their decision times for the small-size-difference trials would be longer than for the large-size-difference trials. A paired *t*-test is conducted to test this hypothesis.

1. Click on the word Analyze

To perform a paired t-test

				T. Click of the word Analyze .
💼 paired t-	test.sav [DataSet2]	- IBM SPSS Sta	itistics Data Editor	- 🗆 ×
Eile Edit	⊻iew <u>D</u> ata	Transform	Analyze Graphs Utilities Extensions	Window Help
8			Reports Descriptive Statistics	
2 :			Bayesian Statistics	2. Click on Compare Means .
	/ large	& small	Tables +	ar var
1	936.00	878.00	Compare Means	Means
2	923.00	1005.00	General Linear Model	One-Sample T Test
3	896.00	1010.00	Generalized Linear Models	
4	1241.00	1365.00	Mixed Models	Independent-Samples T Test
5	1278.00	1422.00	Correlate	Paired-Samples T Test
6	871.00	1198.00	Regression >	Qrie-Way ANOVA
7	1360.00	1576.00	Loglinear	
8	733.00	896.00	Neural Networks	
9	941.00	1573.00	Classify	
10	1077.00	1261.00		
11	1438.00	2237.00	Dimension Reduction	3. Click on the words Paired-
12	1099.00	1325.00	Scale M	Samples T Test.
13	1253.00	1591.00	Nonparametric Tests	
14	1930.00	2742.00	Forecasting	
15	1260.00	1357.00	Survival •	
16	1271.00	1963.00	Multiple Response	
17			Missing Value Analysis	
18	4		Multiple Imputation	18
	in the second se		Complex Samples	[1]
Data View	Variable View		聞 Simulation	
Paired-Sam	ples T Test		Quality Control	tistics Processor is ready Unicode:ON

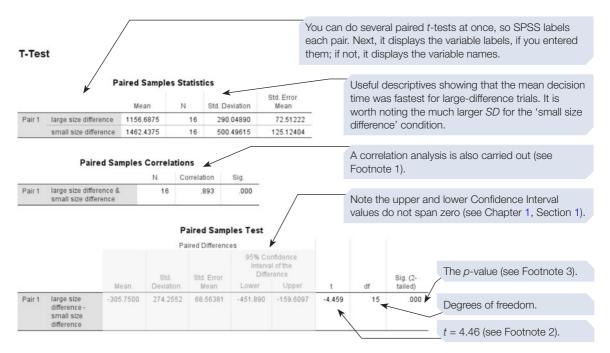
4. You will now see the **Paired-Samples T Test** dialogue box (see below). You need to choose the names of the two variables you want to compare. As before, all the variables in your data file are listed in the left-hand box. Click on each of the two variables you want to compare. These variable names will now be highlighted.



SPSS will perform the paired *t*-test. The annotated output is shown below.

SPSS output for paired (or related) t-test

Obtained using menu items: > Compare Means > Paired-Samples T Test



Footnotes

- SPSS performs a Pearson's correlation (see Chapter 6, Section 3), but you can ignore it if you only want a *t*-test. A significant positive correlation tells you that participants who were fast on large-size-difference trials were also fast on small-sizedifference trials. It does not mean that the scores are significantly different.
- 2. The minus sign tells you that the mean value for the first variable name in the **Paired Variables** box is lower on average than the mean value for the second variable name.
- 3. A *p*-value can never equal zero. SPSS rounds *p*-values to three decimal places, so *p* must be less than .001 or it would appear as .001. In a report, put p < .001 if the hypothesis was two-tailed. Here, the hypothesis was one-tailed, so divide by 2, which gives p < .0005. However, as we report *p*-values less than .001 as p < .001, we would write: p < .001, one-tailed.

Measure of effect size

As with the independent *t*-test, the output for the paired *t*-test does not include an estimate of the size of the effect and this is not an option on SPSS that you can select. However, from the output, you can calculate Cohen's *d*, which we described in Section 3, following the steps below and using the formula: $d = (x_1 - x_2) / \text{mean } SD$.

1. Look at the SPSS output and identify the mean and standard deviation (*SD*) for each condition, which are provided in a table called **Paired Sample Statistics**. For the data set used in this section:

The mean for the large size difference condition, $x_1 = 1156.69$ and the SD = 290.05.

The mean for the small size difference condition, $x_2 = 1462.44$ and the SD = 500.50.

2. Take the mean of one condition from the mean of the other condition (it is not important which mean to take from which, so you can ignore the sign). For the data set used in this section:

$$1156.69 - 1462.44 = 305.75.$$

3. Find the mean SD by adding the SD for each condition together and dividing by two: (SD of condition 1 + SD of condition 2) / 2. For the data set used in this section:

$$(290.05 + 500.50) / 2 = 395.28.$$

4. Use the formula to calculate *d*. For the data set used in this section:

$$d = (1156.69 - 1462.44) / 30.95.28 = 305.75 / 395.28 = 0.77.$$

This is a moderate to large effect size. Cohen (1988) provided the following guidance on how to interpret d:

- Small effect size: 0.2 (or more)
- Moderate effect size: 0.5 (or more)
- Large effect size: 0.8 (or more).

Reporting the results

In a report you might write:

The average time to decide which of the pair of animals was larger was greater for small-size-difference trials than for large-size-difference trials (1462.44 ms and 1156.69 ms, respectively). A paired *t*-test showed that the difference between conditions was significant, and the size of this effect was moderate to large (t = 4.46, df = 15, p < .001, one-tailed, d = 0.77).



We will not show you how to create an error bar graph with SPSS for the data analysed with the paired *t*-test. Unfortunately, SPSS does not allow Chart Builder to create error bar graphs for data from a repeated measures design. With a repeated measures design the same participants complete both conditions, so this must be taken into account. Brysbaert (2011, 233–4) explains how the confidence intervals from experiments using such a design can be 'corrected' to allow them to be interpreted appropriately, in line with an independent groups design.

Section 5: AN INTRODUCTION TO NONPARAMETRIC TESTS OF DIFFERENCE

- The Mann–Whitney test and the Wilcoxon matched-pairs signed-ranks test are nonparametric tests of difference and are used to explore whether two data samples are different.
- The Wilcoxon test is the nonparametric equivalent of the paired *t*-test, and is used for data gathered in experiments involving repeated measures and matched-pairs designs.
- The Mann–Whitney test is the nonparametric equivalent of the independent *t*-test, and is used to compare data collected in an experiment involving an independent groups design.

- These nonparametric tests should be used in preference to the equivalent *t*-tests when data are only of ordinal level of measurement or do not meet the other assumptions required for parametric tests.
- Both tests involve ranking the data, and the calculations are carried out on the ranks. In the annotated output pages for these tests, there is a brief explanation of how each test is performed.
- When reporting descriptive statistics to accompany the results of a nonparametric test of difference, such as the Mann–Whitney or Wilcoxon test, you should normally give the median and range (not the mean and standard deviation) as the measures of central tendency and dispersion. The median and range are more appropriate descriptives for nonparametric tests because these are distribution-free tests and do not assume normal distribution.

Section 6: THE MANN–WHITNEY TEST

Example study: sex differences and emphasis on physical attractiveness

To demonstrate how to perform the Mann–Whitney test, we shall use the data from an experiment based on research conducted by one of our past students, which was designed to determine whether males and females differ in the emphasis they place on the importance of the physical attractiveness of their partner. Previous research has reported that men are more concerned than women about the physical attractiveness of their heterosexual partner. However, current advertising trends and societal pressure may have altered the emphasis placed on physical attractiveness and, more specifically, the importance they attach to 'body' or physique compared with other characteristics of their ideal partner.

The hypothesis tested is two-tailed: that men and women will differ in the importance they attach to physique. The design employed was an independent groups design. The independent variable was whether the participant was male or female, operationalised by asking equal numbers of males and females to take part in the experiment (only one partner from a relationship participated). The dependent variable was the importance attached to body shape, operationalised by asking participants to rank order 10 characteristics of an ideal partner, one of these being body shape. (These data are available in the Appendix or from macmillanihe.com/ harrison-spss-7e.)

How to do it

🖬 Mann-W	hitney.sav [DataSet3] - IBM SPSS S	Statistics Deta Editor		
ile <u>E</u> dit	<u>V</u> iew <u>D</u> ata	Iransform	Analyze Graphs Utilities Extension	ns <u>W</u> indow <u>H</u> elp	
8			Reports Descriptive Statistics		۲
			Bayesian Statistics	2	Click on Nonparametric
	a sex	a rating	Ta <u>b</u> les 9	ar var	
1	1	4	Compare Means		sts.
2	1	6	General Linear Model		
3	1	5	Generalized Linear Models		
4	1	8	Mixed Models		Click on Legacy Dialogs.
5	1	5	<u>C</u> orrelate	0.	Sher on <u>E</u> egacy Dialogs.
6	1	2	Regression		
7	1	4	Loglinear		
8	1	4	Neural Networks	•	Chi-square
9	1	5	Classify	•	Dinomial
10	1	7	Dimension Reduction	· · · · · · · · · · · · · · · · · · ·	Runs
11	1	5	Scale ,		1-Sample K-S
12	1	4	Nonparametric Tests	One Sample	2 Independent Samples
13	1	3	Forecasting	A Independent Samples	Kindependent Samples
15	1	5	Survival	Related Samples.	2 Related Samples
15	1	3	Multiple Response	Legacy Dialogs	K Related Samples
17	1	3	Missing Value Analysis	Legacy Dialogs	EAST
18	1	8	Multiple Imputation	7	
	4		Complex Samples		•
Data View	Variable View		B Simulation		
Independe	nt Samples		Quality Control	IBM SPSS Statistics Pro	cessor is ready Unicode:ON

The data have been entered with the variable names 'sex' and 'rating'. Follow steps 5 to 11, shown below, then click on or . The SPSS output, which will appear after a short delay, is shown on the following page with explanatory comments.

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5. Highlight the relevant variables by clicking on them.	7		nis arrow to move highlighted variable vant box.
🖷 Two-Independent-Samy	iles Tests	×	
Sex of subject [Sex]	Test Variable List	Exact Options	7. Test Variable is the dependent variable, so move the 'Rating of' into this box.
	Grouping Variable:		8. Grouping Variable is the independent variable, so move the 'Sex of subject [sex]' into this box.
Test Type	Kolmogorov-Smirnov Z		
	ons 🔄 Wald-Wolfowitz runs		O After step 9 this how will
	Paste Reset Cancel Help		9. After step 8, this box will become bold: click on it, and the dialogue box below will appear.
🚔 Two Group	Independent Samples: ×		
Group	2: 2	SI 1 (fe	D. Enter the codes used in the PSS data file; in this example, (male) for Group 1 and 2 emale) for Group 2. Then click n <u>Continue</u> to return to the lain dialogue box.
11. Make sure the Mann–Whitney the small box. Ye to require the ot Click on OK .	U has a tick in ou are unlikely		
Chick off OR.			

SPSS output for Mann-Whitney U test

Obtained using menu items: Nonparametric Tests > Legacy Dialogs > 2 Independent Samples

NPar Tests

Mann-Whitney Test

	Rank	s		
	Sex of subject	Ν	Mean Rank	Sum of Ranks
Rating of the importance	Male	20	17.88	357.50
of body as characteristic in a partner	Female	20	23.13	462.50
	Total	40	1	

Results section, but it gives you some information about the calculations for the Mann–Whitney U test: first, all the data from both groups combined are assigned ranks from the lowest to the highest; then, the ranks given to one group are compared with the ranks given to the other group; the mean ranks shown here indicate whether there are more high ranks in one group than in the other.

Test Statisti	Rating of the
	importance of body as characteristic in a partner
Mann-Whitney U	147.500
Wilcoxon W Z	357.500
Asymp. Sig. (2-tailed)	.150
Exact Sig. [2*(1-tailed Sig.)]	.157 ^b
a. Grouping Variable: Se b. Not corrected for ties.	



Reporting the results

In a report you might write:

There was no significant difference between men and women in the importance they attached to body shape in a partner (U = 147.50, $N_1 = 20$, $N_2 = 20$, p = .157, two-tailed).



If you are reporting descriptive statistics, you should avoid reporting the mean rank or sum of ranks given in this SPSS output and that for the Wilcoxon test described next. Instead, obtain the median and range for the two conditions by using the **Explore** command (Chapter 3, Section 5).

Section 7: THE WILCOXON TEST

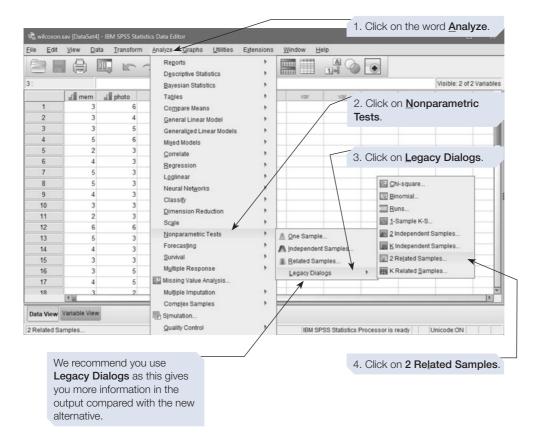
Example study: quality of E-FIT images

The police frequently use a computerised facial composite system to help eyewitnesses recall the face of a perpetrator. One such system is E-FIT (Electronic Facial Identification Technique). In a study by Newlands (1997), participants were shown a short video clip of a mock crime scenario depicting an instance of petty theft. Participants were then asked to generate an E-FIT composite of the perpetrator. On completion, they were asked to rate the likeness of their E-FIT image to the person they remember seeing in the video. They were then shown a photo of the perpetrator and again asked to rate the likeness of their E-FIT to that person.

The hypothesis tested was one-tailed: that the likeness ratings of the E-FIT to the perpetrator would be more favourable when recalling the perpetrator from memory than when seeing a photograph of the perpetrator. The design employed was a repeated measures design. The independent variable was the presence or absence of a photograph of the perpetrator, operationalised by asking participants to rate the likeness of their E-FIT, first to their recall of the perpetrator and then to a photo of the perpetrator. The dependent variable was measured on an ordinal scale and was the likeness rating, operationalised by the response on a seven-point scale, where point 1 was 'very good likeness' and point 7 'no likeness'.

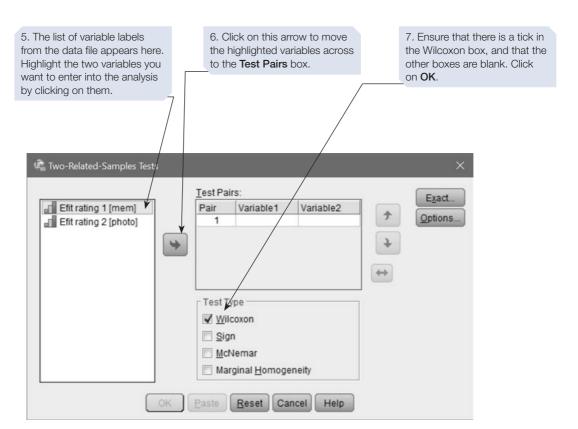
For the purposes of this book, we have created a data file that will reproduce some of the findings of this study. (These data are available in the Appendix or from macmillanihe.com/harrison-spss-7e.)

How to do it



The dialogue box shown below will appear. The variable labels, and the variable names ('mem' and 'photo'), used in the data file appear in the left-hand box. Follow steps 5 to 7, shown below, then click on **or**. The SPSS output, which will appear after a short delay, is shown below with explanatory comments.

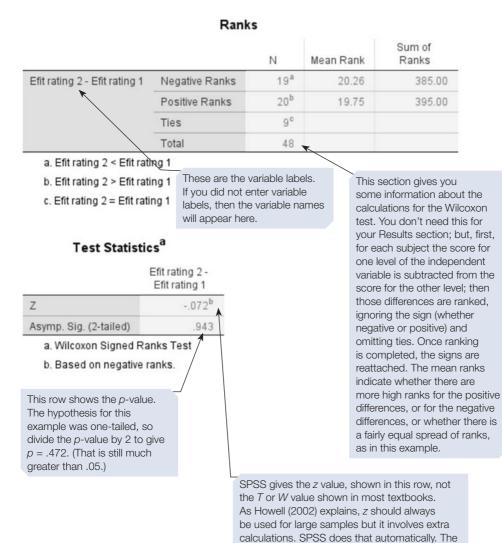
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SPSS output for Wilcoxon matched-pairs signed-ranks test

Obtained using menu items: Nonparametric Tests > Legacy Dialogs > 2 Related Samples

Wilcoxon Signed Ranks Test



negative sign can be ignored (as for the *t*-test).



Reporting the results

In a report you might write:

There was no significant difference between the likeness ratings of the E-FITs that were made with a photo of the perpetrator visible and those that were made from memory (z = 0.07, N - Ties = 39, p = .472, one-tailed).

Summary

- This chapter introduced you to statistical tests that will tell you if there is a significant difference between two means.
- This could involve comparing the performance of a single group of participants with a specific score, or comparing the performance of two groups of participants with one another (such as in an independent groups design), or comparing the performance of one group of participants who perform in two conditions (as in a repeated measures design).
- Your choice of which test to use will depend on the design of your experiment, and whether the data are parametric.
- Remember our advice on the appropriate descriptive statistics to accompany the results of these tests. See Chapter 3 for guidance on obtaining these.
- For guidance on incorporating SPSS output into a report, or on printing the output, see Chapter 14.
- Chapter 8 introduces you to Analysis of Variance (ANOVA), a test of difference that is appropriate for designs that involve more than two groups or conditions, or more than one independent variable.

6 Tests of correlation and bivariate regression

In this chapter

- An introduction to tests of correlation
- Producing a scatterplot
- Pearson's r. parametric test of correlation
- Spearman's *r*_s: nonparametric test of correlation
- Partial correlations
- Comparing the strength of correlation coefficients
- Brief introduction to regression
- Bivariate regression



SPSS for Psychologists online

Visit macmillanihe.com/harrison-spss-7e for data sets, online tutorials and exercises.

Section 1: AN INTRODUCTION TO TESTS OF CORRELATION

- Researchers often wish to measure the strength of the relationship between two variables. For example, there is likely to be a relationship between age and reading ability in children. A test of correlation will provide you with a measure of the strength and direction of such a relationship.
- In a correlation, there is no independent variable: you simply measure two variables. So, if someone wished to investigate the relationship between smoking and respiratory function, they could measure how many cigarettes people smoke and their respiratory function, and then test for a correlation between these two variables.
- Correlation does not imply causation. In any correlation, there could be a third variable which explains the association between the two variables you measured. For example, there may be a correlation between the number of ice creams sold and the number of people who drown. Here, temperature is the third variable, which could explain the relationship between the measured variables. Even when there seems to be a clear cause-and-effect relationship, a correlation alone is not sufficient evidence to prove this causal relationship.

- Francis Galton carried out early work on correlation and one of his colleagues, Pearson, developed a method of calculating correlation coefficients for parametric data: Pearson's product moment correlation coefficient (Pearson's r). If the data are not parametric, or if the relationship is not linear, a nonparametric test of correlation, such as Spearman's r_s should be used.
- Note that, for a correlation coefficient to be reliable, one should normally make at least 100 observations; otherwise a small number of extreme scores could skew the data and either mask a real relationship or give the appearance of a correlation that does not really exist. The scatterplot is a useful tool for checking such eventualities and for checking that the relationship is linear.

Section 2: PRODUCING A SCATTERPLOT

A scatterplot (or scattergram) will give a good indication of whether two items are related in a linear fashion. Figure 6.1 shows a hypothetical example. Each point on the scatterplot represents the age and the reading ability of one child. The line running through the data points is called a 'regression line'. It represents the 'best fit' of a straight line to the data points. The line in Figure 6.1 slopes upwards from left to right: as one variable increases in value, the other variable also increases in value and this is called a *positive correlation*. The closer the points are to being on the line itself, the stronger the correlation. If all the points fall along the straight line, then it is said to be a *perfect correlation*. The scatterplot will also show you any outliers.

It is often the case that as one variable increases in value, the other variable decreases in value: this is called a *negative correlation*. In this case, the pattern of points on the scatterplot will slope downwards from top left to bottom right.

As well as telling you the direction of the relationship between two variables (i.e. positive or negative), a correlation can also tell you how strong the relationship

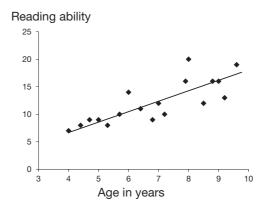


Figure 6.1 Scatterplot illustrating a positive correlation: hypothetical data for the relationship between age and reading ability in children

between your variables is. In this way, it's very similar to the concept of effect size which you came across in Chapter 1 (and, in fact, correlation coefficients are sometimes used as a measure of effect size). In this case:

- a value of 1 shows a perfect positive correlation
- a value of -1 a perfect negative correlation
- a value of zero shows that there is no relationship between the two variables.

In reality, correlation coefficients are rarely exactly 1, -1 or 0, but instead fall somewhere in between. As with effect size, you can use a general rule of thumb to help interpret the strength of these correlations. We suggest using the following as a guide:

- 0.7 to 1 shows a strong relationship between the variables
- 0.3 to 0.6 suggests a moderate relationship
- 0 to 0.2 indicates any relationship is weak

The size (or magnitude) of the correlation coefficient is directly related to how linear (or line-like) your data points appear on a scatterplot. If the points form a perfectly straight line, then the correlation coefficient will be either 1 or -1, depending on which direction it is sloping in. However, if the points are completely random, so that they appear distributed almost equally across the scatterplot, then the correlation coefficient will be nearer 0. This is illustrated in Figure 6.2.

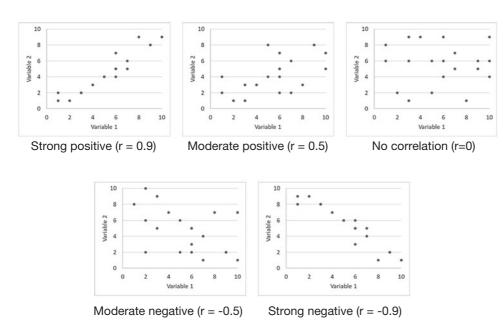


Figure 6.2 Scatterplots illustrating different directions and strengths of correlations between two variables

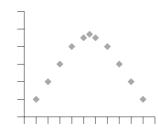


Figure 6.3 Scatterplot showing two variables with an inverted U-shape relationship

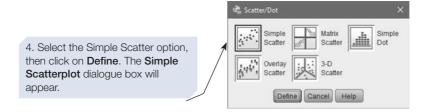
In this book we will only show procedures for dealing with linear relationships, like that shown in Figure 6.1, and also the relationship in the example study below. Sometimes relationships are nonlinear, for example, an inverted U-shape relationship, illustrated in Figure 6.3, which might be found between two variables such as stress and exam performance. For procedures to deal with nonlinear relationships, read the appropriate textbooks.

Example study: relationship between age and CFF

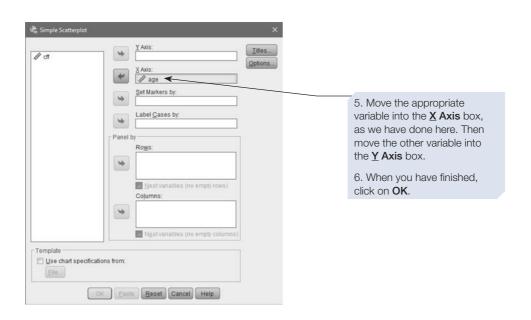
A paper by Mason et al. (1982) described an investigation of (among other things) whether the negative correlation between age and critical flicker frequency (CFF) is different for people with multiple sclerosis than for control participants. For this example, we have created a data file that will reproduce some of the findings for the control participants. CFF can be described briefly and somewhat simplistically as follows: If a light is flickering on and off at a low frequency, most people can detect the flicker; if the frequency of flicker is increased, eventually it looks like a steady light. The frequency at which someone can no longer perceive the flicker is called the *critical flicker frequency* (CFF). (These data are available in the Appendix or from macmillanihe.com/harrison-spss-7e.)

How to obtain a scatterplot using Legacy Dialogs

- 1. On the menu bar, click on Graphs.
- 2. Click on Legacy Dialogs.
- 3. Click on Scatter/Dot, and the Scatter/Dot dialogue box will appear.





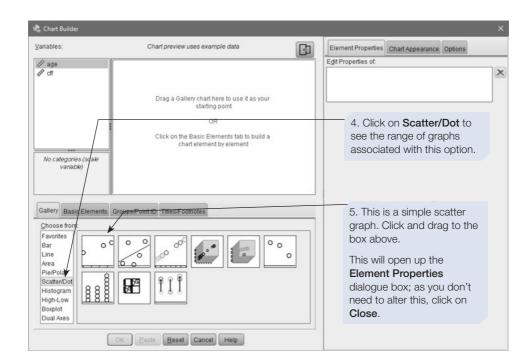


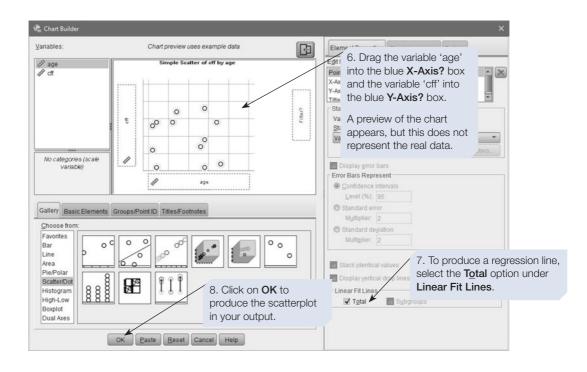
The scatterplot will appear in the SPSS output window. You can edit it, in the way we describe below. However, using a different menu option to produce your graph will allow you to directly add a regression line to your scatterplot, and give you a number of other options in the dialogue box. We will describe this in the next section.

Producing a scatterplot with a regression line using Chart Builder

- 1. On the menu bar, click on Graphs.
- 2. Click on Chart Builder.
- 3. A **Chart Builder** dialogue box will remind you that measurement level should be set properly for each variable in your chart (see Chapter 3, Section 8). If you are sure that they are set correctly, just click on OK. (If you want to add a regression line to the scatterplot, then both variables must be set at Scale in the Measure column of Variable View.)

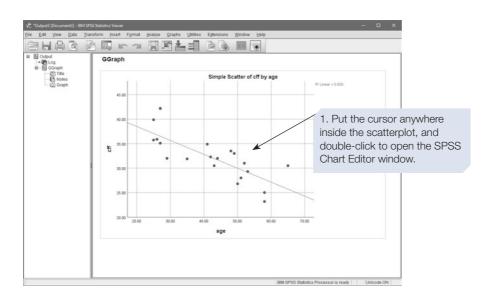
The Chart Builder dialogue box shown below will now be visible.

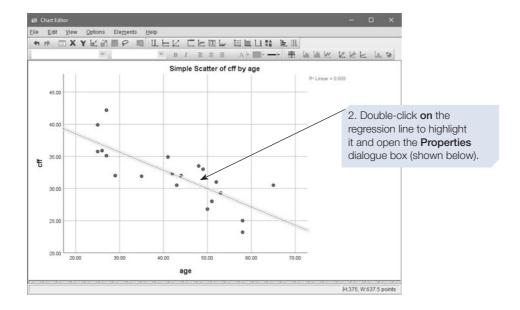




How to add a regression equation to the scatterplot

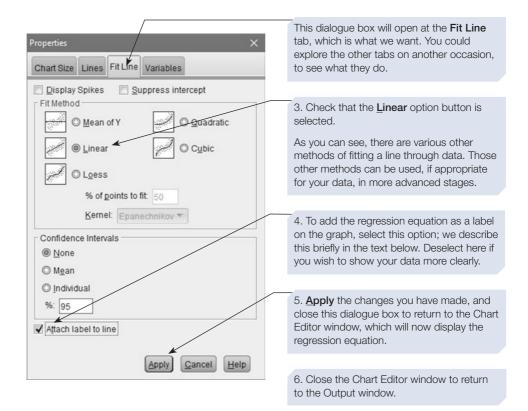
To add the regression equation to your scatterplot, you have to edit the graph SPSS produces. You can also use this method to add a regression line to a scatterplot produced using **Legacy Dialogs**. Start by double-clicking in the scatterplot, and the SPSS Chart Editor window, shown below, will appear.







If you produced your scatterplot using the <u>Legacy Dialogs</u>, you can add a regression line here by selecting **Elements** and then <u>Fit Line at Total</u>. If this option is not available, check that in your data file both variables are set to **Scale** in the **Measure** column, then start again.



You can copy the scatterplot and paste it into a report, adding a suitable figure legend. For example, see Figure 6.4.

Figure legends should be suitable for the work into which you are incorporating the figure. The legend to Figure 6.4 might be suitable for a report about the study into age and CFF. The legends to Figures 6.1, 6.2 and 6.3, however, are intended to help you follow the explanation in this book, and would not be suitable for a report.

In addition to adding the regression line, you can edit other elements of the chart to improve appearance. For example, SPSS charts are usually rather large. If you leave them large, the report will be spread over more pages than necessary, which can hinder the ease with which the reader follows your argument. You can shrink charts easily in a program such as MS Word, but it is much better to change the size in Chart Editor, because then the font and symbol size will automatically be adjusted for legibility. We show you how to do this in Chapter 3, Section 8. Editing would also be useful when a number of cases all fall at the same point. The data we use to illustrate the use of Spearman's r_s (Section 3) demonstrate that situation. To clearly illustrate the data you can edit the data symbols in Chart Editor, so that they vary in size according to the number of cases at each point. Guidelines on the appearance of figures are given in the Publication Manual of American Psychological Association (APA, 2019).

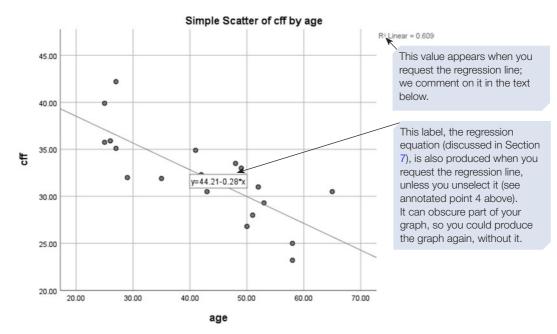


Figure 6.4 Critical flicker frequency (in Hz) plotted against participant's age (in years)

A scatterplot is a descriptive graph that illustrates the data, and can be used to check whether the data are suitable for analysis using a test of correlation. For example, if there are a few cases in one corner of the scatterplot, and most of the other cases are clustered together at the opposite end of the regression line, those outliers may produce a significant correlation even though there is no real relationship. A scatterplot would also indicate if there is a relationship but it is nonlinear; for example, if the relationship is U-shaped. If there does appear to be a linear relationship (and Pearson's r makes the assumption that any relationship will be linear), we can find out whether or not it is significant with an inferential statistical test of correlation (as outlined above).

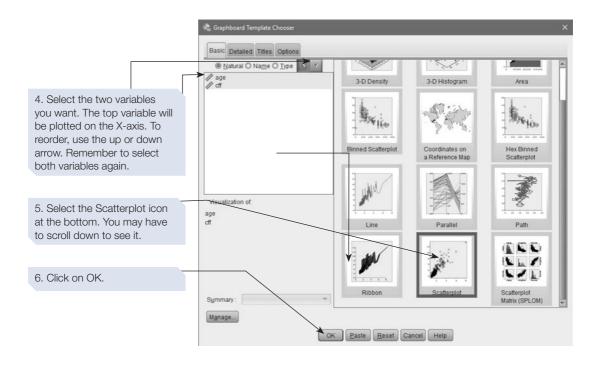
Note the R^2 Linear value that appears next to the scatterplot above. This is not the correlation coefficient itself; it is the square of Pearson's *r* and is itself a useful statistic (described in Section 3). You can remove the R^2 legend if you wish: in the Chart Editor window double-click on the legend, so that it is selected, then press delete key.

Next, we show you how to obtain a scatterplot using <u>Graphboard Template</u> Chooser, which has been available since Version 17. It is not yet possible to add a regression line using this relatively new SPSS graphing option or to edit the size of the graph.

How to obtain a scatterplot using Graphboard Template Chooser

Graphboard Template Chooser was described in Chapter 3, Section 9.

- 1. On the menu bar, click on Graphs.
- 2. Click on Graphboard Template Chooser.
- 3. In the dialogue box, ensure that the variables you want to plot are indicated as ordinal or scale. If they are not, go to your data file and set them.



Here we have shown you how to produce the scatterplot using the **Basic** tab. You could instead use the **Detailed** tab, described in Chapter 3, Section 9. You would select **Scatterplot** from the **Choose** list, and then set the X and Y variables.

The graph will appear in Viewer window. Unlike other scatterplot commands, Graphboard does not add R^2 Linear. Double-click in the scatterplot if you wish to edit it, and the Graphboard Editor window will appear. However, you cannot add a regression line using this method.

CHAPTER

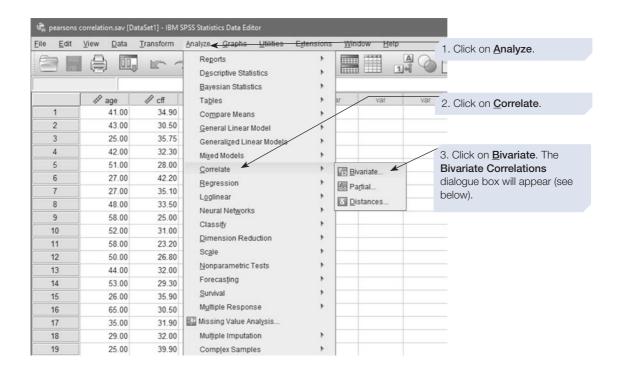
Section 3: PEARSON'S r : PARAMETRIC TEST OF CORRELATION

Example study: critical flicker frequency and age

To illustrate how to carry out this parametric test of correlation, we will continue using the CFF and age data. Note that the data do not meet the guidelines for correlation of a sample size of around 100. The hypothesis tested was that there would be a negative correlation between CFF and age.

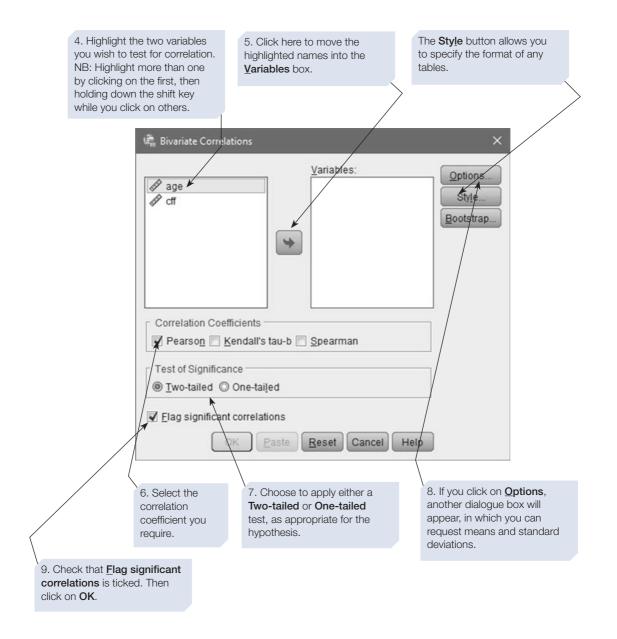
The study employed a correlational design. Two variables were measured. The first was age, operationalised by recruiting volunteer participants who ranged in age from 25 to 66 years. The second variable was CFF, operationalised by using a flicker generator to measure CFF for each participant; six measures were made, and the mean taken to give a single CFF score for each participant.

How to perform a Pearson's r





SPSS will correlate each variable you include with every other variable you include. Thus, if you included three variables A, B and C, it will calculate the correlation coefficient for A * B, A * C and B * C. In the Pearson's *r* example we have just two variables, but in the Spearman's *r*_s example (in Section 4), we include three variables so that you can see what a larger correlation matrix looks like.



CHAPTER 6

In the **Bivariate Correlations** dialogue box, you have the option of choosing either a one- or two-tailed test, and SPSS will then print the appropriate value of *p*. Here, we selected two-tailed. In the statistical tests we have covered previously, SPSS will only print the two-tailed *p*-value, and if you have a one-tailed hypothesis, you halve that value to give the one-tailed *p*-value.

The annotated output for Pearson's r is shown below.

SPSS output for Pearson's r

Obtained using menu item: Correlate > Bivariate

CORRELATIONS

matrix.

Descri	ptive	Statistics	
--------	-------	------------	--

	Mean	Std. Deviation	Ν
age	42.4000	12.55891	20
cff	32.1375	4.58249	20

Useful descriptives obtained by using the **Options** button in the **Bivariate Correlations** dialogue box.

The Pearson's correlation cff age coefficient or Pearson's r. -.780 age Pearson Correlation 1 000 Sig. (2-tailed) The *p*-value. (See Chapter 5, 20 Section 4, Footnote 3.) N 20 -.780 ** 🖌 cff Pearson Correlation 1 .000 Sig. (2-tailed) N, the number of cases. 20 N 20 **. Correlation is significant at the 0.01 level (2-A complete matrix is printed. tailed). Two of the cells are for each variable with itself (for these cells, p is not calculated). The other two cells contain the same information about the correlation between the two In addition to the *p*-values in the matrix, SPSS prints this message. variables. Significant correlations are flagged by asterisks; this is particularly useful if you have entered several variables and so have a large correlation

Correlations

What you might write in a report is given below, after we tell you about effect sizes in correlations.



For correlations, the sign of the coefficient indicates whether the correlation is positive or negative, so you must report it (unlike the sign in a *t*-test analysis).

Effect sizes in correlation

The value of r indicates the strength of the correlation, and it is a measure of effect size (see Chapter 1, Section 1). As a rule of thumb, r values of 0 to .2 are generally considered weak, .3 to .6 moderate and .7 to 1 strong. The strength of the correlation alone is not necessarily an indication of whether it is an important correlation: normally, the significance value should also be considered. With small sample sizes this is crucial, as strong correlations may easily occur by chance. With large to very large sample sizes, however, even a small correlation can be highly statistically significant. To illustrate this, look at a table of the critical values of r (in the back of most statistics textbooks). For example, if you carry out a correlational study with a sample of 100 and obtain r of .20, it is significant at the .05 level, two-tailed. Yet .2 is only a weak correlation. Thus, we recommend you report the effect size, the statistical significance and the proportion of variation, which we explain next.

The concept of 'proportion of variation explained' is described in Chapter 8, Section 1. Briefly, a correlation coefficient allows us to estimate the proportion of variation within our data that is explained by the relationship between the two variables. (The remaining variation is down to extraneous variables, the situation and participants.) The proportion of variation explained is given by r^2 . Thus, for the age and CFF example, in which r = .78, $r^2 = (.78 \times .78) = 0.6084$. Multiplying r^2 by 100 allows us to turn this into a percentage, and we can say that 60.84% of the variation in the CFF data can be attributed to age. Note that, logically, we can just as easily say that 60.84% of the variation in the age data can be attributed to CFF. The latter statement should make it clear that we are not implying a causal relationship: we cannot do so with correlation. The important practical point is that the two variables have quite a lot of variation in common, and one could use a person's age to predict what their CFF might be. If their measured CFF is outside the lower confidence limit for their age, we could investigate further.

Note that the proportion of variation explained does not have to be large to be important. How important it is may depend on the purpose of the study (see Howell, 2013, 312–13). We will come back to this concept in Section 5 of this chapter, and in Chapter 10, which covers multiple regression.



In a report you might write:

There was a significant negative correlation between age and CFF (r = -.78, N = 20, p < .001, one-tailed). It is a fairly strong correlation: 60.84% of the variation is explained. The scatterplot (Figure 6.3) shows

that the data points are reasonably well distributed along the regression line, in a linear relationship with no outliers.

Section 4: SPEARMAN'S r_s : NONPARAMETRIC TEST OF CORRELATION

When the data for one or both of the variables are not parametric, for example they are measured at ordinal level, or if the scatterplot suggests that the relationship between the two variables is not linear, then we use a nonparametric measure of correlation. (See Chapter 1 for more about choosing the correct statistical procedure, and information about levels of measurement and parametric tests).

Here, we describe two such tests, Spearman's r_s and Kendall's tau-b. The $_s$ on Spearman's r_s is to distinguish it from Pearson's r. This test was originally called Spearman's ρ (the Greek letter rho), and SPSS still calls the output Spearman's rho.

Example study: the relationships between attractiveness, believability and confidence

Previous research using mock juries has shown that attractive defendants are less likely to be found guilty than unattractive defendants, and that attractive individuals are frequently rated more highly on other desirable traits, such as intelligence. In a study undertaken by one of our students, participants saw the testimony of a woman in a case of alleged assault. They were asked to rate her, on a scale of 1–7, in terms of how much confidence they placed in her testimony, how believable she was and how attractive she was. (These data are available in the Appendix or from macmillanihe. com/harrison-spss-7e.)

The design employed was correlational, with three variables each measured on a seven-point scale. Although it is often accepted that such data could be considered interval in nature (see Chapter 1, Section 1), for the purpose of this section we will consider it as ordinal data. The hypotheses tested were that:

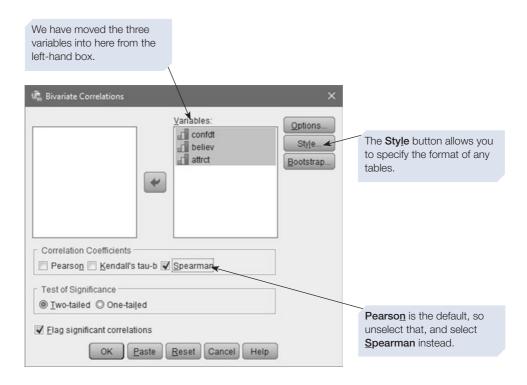
- 1. There would be a positive relationship between attractiveness and confidence placed in testimony.
- 2. There would be a positive relationship between attractiveness and believability.
- 3. There would be a positive relationship between confidence placed in testimony and believability.



We are using this study to illustrate the use of Spearman's r_s and some other aspects of correlation. However, multiple regression (Chapter 10) would usually be more appropriate for three or more variables in a correlational design.

How to perform Spearman's r_s

Carry out steps 1 to 5 as for Pearson's r (previous section). At step 6, select **Spearman** instead of **Pearson** (see **Bivariate Correlations** dialogue box below). This example also illustrates the fact that you can carry out more than one correlation at once. There are three variables, and we want to investigate the relationship between each variable with each of the other two. To do this, you simply highlight all three variable names and move them all into the <u>Variables</u> box.



The SPSS output for Spearman's r_s is shown on the next page.

SPSS output for Spearman's *r_s*

Obtained using menu item: Correlate > Bivariate

NONPARAMETRIC CORRELATIONS

				confdt	believ	attrct	
Spearman's rho	confdt	Correlati	on Coefficient	1.000	.372	.157	
		Sig. (2-ta	ailed)		.000	.143	
		Ν		89	89	89	
	believ	Correlati	on Coefficient	.372	1.000	.359	
		Sig. (2-ta	ailed)	.000		.001	
		Ν		89	89	89	
	attrct	Correlati	en Coefficient	.157	.359	1.000	
		Sig. (2-ta	ailed)	.143	.001		
		Ν		89	89	89	
**. Correlation	is signific	ant at the O).01 level (2-taile	ed).	,	1	
This cell contains the values for the correlation between variables 'confdt' and 'believ': .372 is r _s .000 is p 89 is the number of cases.			output for e correlation i	output for each bivariatecomplete matcorrelation is given twice, oncethis matrix is labelow the diagonal and oncethree variables		earson's output, a natrix is printed, but is larger because oles were entered.	

Correlations

To find the correlation between two variables (for example, confdt and believ), you move down the **column** that represents your first variable (confdt), and move across the **row** that represents your second variable (believ). The cell where they **intersect** contains the relevant correlation statistics.

Reporting the results

When reporting the outcome for each correlation, at the appropriate points, you would write:

There was a significant positive correlation between confidence in testimony and believability ($r_s = .37$, N = 89, p < .001, two-tailed). There was no significant correlation between confidence in testimony and attractiveness ($r_s = .16$, N = 89, p = .143, two-tailed). There was a significant positive correlation between attractiveness and believability ($r_s = .36$, N = 89, p = .001, two-tailed).

You could illustrate each pair of variables in a scatterplot (see Section 1). Note that the R² Linear value, given in the scatterplot when you add a regression line, is the square of Pearson's r (r^2) and not the square of Spearman's r_s . As described in Section 2, r^2 indicates the proportion of variation explained, but this may not be appropriate for ordinal data.

How to perform Kendall's tau-b

Some researchers prefer to use Kendall's tau instead of Spearman's r_s . To undertake a Kendall's tau-b, follow the same steps as for Pearson's r, but at step 6 select <u>Kendall's</u> tau-b. The output takes the same form as that for Spearman's r_s . Kendall's tau-b takes ties into account. Kendall's tau-c, which ignores ties, is available in Crosstabs (see Chapter 7, Section 4).

Section 5: PARTIAL CORRELATIONS

There may be times, when carrying out research, that you want to examine the unique relationship between two variables, excluding any potential influence of a third variable.

Example study: academic achievement

Imagine you are a researcher interested in exploring the relationship between how much revision your students do, and the score they get on their final exam. Obviously, a correlation would be able to tell you the strength and direction of the relationship between participants' exam scores and how hard they revised (e.g. time spent revising in hours). However, there are other variables that might be influencing this relationship. For example, how much your students enjoy their subject is likely to be related to both how well they do in that subject, and how much they engage with revision.

So, if you wanted to explore the direct relationship between revision time and exam outcome, you might want to partial out the influence of subject enjoyment. To do this, you can use a partial correlation. This allows you to look at the strength and direction of the relationship between two variables while controlling for the influence of a third variable. This process involves looking at the proportion of variation explained by the different variables (see Figure 6.5) and effectively removing the influence of the third variable, producing a correlation between your two variables of interest, while the effects of the third variable are held constant.

As with Pearson's correlation, partial correlations require that all variables are measured at the interval level. In this example, both exam score and time spent revising naturally meet this criteria. However, for a partial correlation to be valid, you would also need to find a measure of subject enjoyment that is appropriate. As simply asking students how much they enjoy their subject on a scale of 1–5 would only produce ordinal data, you would instead need to find a standardised and validated measure that produces data that can be treated as an interval scale. In this case, the (fictional) Subject Enjoyment Questionnaire, which contains a number of different items and produces a score on a scale from 0–50 (with larger scores indicating greater subject enjoyment).



The question of whether standardised scales can be assumed to be interval (rather than ordinal) data is still hotly debated in the literature. However, much psychology research assumes that this is a legitimate claim, and uses parametric tests to analyse the data produced by them.

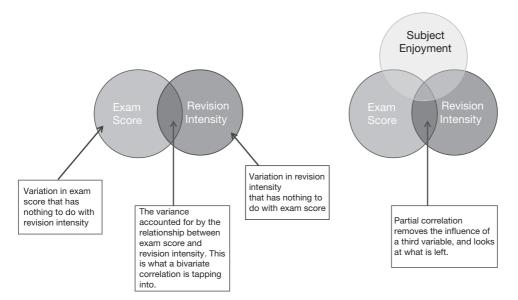
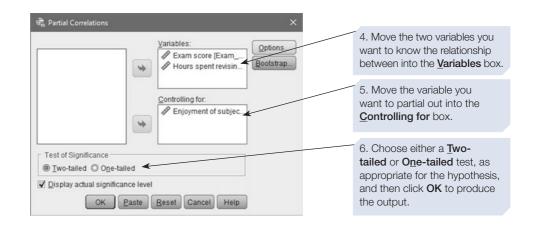


Figure 6.5 Venn diagrams illustrating the role of shared variance in partial correlations

How to perform a partial correlation

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SPSS output Obtained using: Correlate > Partial

PARTIAL CORRELATIONS

Control Variables			Exam score	Hours spent revising
Enjoyment of subject	Exam score	Correlation	1.000	.461
		Significance (2-tailed)		.012
		df	0	27
	Hours spent revising	Correlation	.461	1.000
		Significance (2-tailed)	.012	
		df	27	0

Correlations

Here you can see the values for the partial correlation between Exam score and Hours spent revising, taking into account Enjoyment of subject:

.461	is r _s
.012	is p
27	is the degrees of freedom

Reporting the results

When reporting the outcome for a partial correlation, at the appropriate points, you would write:

There was a significant moderate positive correlation between revision time and exam score whilst controlling for subject enjoyment (r = .46, N = 30, p = .012, two-tailed).

If you wanted to see how much influence subject enjoyment had on the relationship between these two variables, you could also carry out a standard Pearson correlation between revision time and exam score and compare the two correlations and the proportion of variance they explain. If we did this following the steps in Section 3, we would find the following:

There was a significant moderate positive correlation between revision score and exam score (r = .57, N = 30, p = .001, two-tailed).

So, controlling for subject enjoyment reduced the correlation coefficient (although not by a huge amount, as both correlations are still moderate and significant).

Looking at the variance explained can help to interpret this further. While the bivariate (or zero order) correlation accounts for 32% (.57 x .57) of the variance in the data, the partial correlation only accounts for 21% (.46 x .46) of the variation. This tells us that, while revision intensity alone does explain some of the variance in exam scores (and vice versa), there is some influence of subject enjoyment on this relationship.

Partial correlation is very closely related to multiple regression (which you will cover in Chapter 10). As multiple regression allows you to do more with your data, and explore the relationship between your variables in more depth, it is often the preferred analysis.

Section 6: COMPARING THE STRENGTH OF CORRELATION COEFFICIENTS

Sometimes you may have data on the same variables for participants from each of two populations. We may have reason to hypothesise that the correlation between two variables will differ between the two populations. To illustrate this, we will use a study we carried out with student participants as part of a module. It is in the area of environmental psychology, and for this purpose we will use three of the variables we recorded. These were: 'threat', the response to a question about the perceived threat from environmental problems to the participant's own health and wellbeing; 'recycling score', the mean of self-reported recycling rates for a range of materials; and 'gender'. Below, we show the SPSS output for the correlation between the two variables for all participants, for the female group and for the male group.

	Correlation	Correlations			
		Threat	Recycle_score		
Threat	Pearson Correlation	1	.401		
	Sig. (2-tailed)		.000		
	N	219	219		
Recycle_score	Pearson Correlation	.401	1		
	Sig. (2-tailed)	.000			
	N	219	219		

For all participants.

Note: the male N and female N do not sum to the total N as two participants did not record their gender on the questionnaire.

**. Correlation is significant at the 0.01 level (2-tailed).

	Correlation	ns		For famala participanta
		Threat	Recycle_score	For female participants.
Threat	Pearson Correlation	1	.279**	
	Sig. (2-tailed)		.004	
	Ν	105	105	
Recycle_score	Pearson Correlation	.279	1	
	Sig. (2-tailed)	.004		
	Ν	105	105	

*. Correlation is significant at the 0.01 level (2-tailed).

	Correlatio	ns		
		Threat	Recycle_score	For male participants.
Threat	Pearson Correlation	1	.452	
	Sig. (2-tailed)		.000	
	N	112	112	
Recycle_score	Pearson Correlation	.452	1	
	Sig. (2-tailed)	.000		
	Ν	112	112	

**. Correlation is significant at the 0.01 level (2-tailed).

You should have a rationale for any comparison of correlation coefficients you make. For this example, differences between the genders are often reported in the area of environmental psychology. The differences may, however, be due to differences in other variables such as age, social class, type of residence area, level of education and so on. Thus, if you do find a difference between correlation coefficients for samples that are not matched on other variables, you should check whether your two samples differ on other variables which might explain your finding.

A test that Fisher devised (e.g. see Howell, 2013, 284–5) allows use of the z tables to assess whether the difference between two r values is significant. There are three stages to applying this test:

- 1. Transform each r value to a value called r' (pronounced 'r prime'). This is required because the distribution of the difference between two r values is used to assess whether a particular difference is significant; that distribution can be skewed, and then use of z would not be valid.
- 2. Use the r' values to calculate the z value.
- 3. Use *z* to determine whether or not there is a significant difference between the two *r* values.

You can do this yourself using the formulae shown here, or by a SPSS syntax procedure, described in Chapter 14, Section 2.

Using equations

Equation 1
$$r' = (0.5) \log_e \left| \frac{1+r}{1-r} \right|$$

Equation 2
$$z = \frac{r_1' - r_2'}{\sqrt{\frac{1}{N_1 - 3} + \frac{1}{N_2 - 3}}}$$

First, calculate r' from each r:

$$r_1$$
 (for women) = .279

 $r_2 \text{ (for men)} = .452$

$$r_{1}' = (0.5) \log_{e} \left| \frac{1 + .279}{1 - .279} \right|$$

$$r_{2}' = (0.5) \log_{e} \left| \frac{1 + .452}{1 - .452} \right|$$

$$r_{1}' = (0.5) \log_{e} \left| \frac{1 - .279}{0.721} \right|$$

$$r_{2}' = (0.5) \log_{e} \left| \frac{1 - .452}{0.548} \right|$$

$$r_{1}' = (0.5) \log_{e} 1.774$$

$$r_{2}' = (0.5) \log_{e} 2.649$$

$$r_{1}' = (0.5) 0.573$$

$$r_{2}' = (0.5) 0.974$$

$$r_{2}' = 0.487$$

Second, calculate z:

$$z = \frac{r_1' - r_2'}{\sqrt{\frac{1}{N_1 - 3} + \frac{1}{N_2 - 3}}}$$
$$z = \frac{.287 - .487}{\sqrt{\frac{1}{105 - 3} + \frac{1}{112 - 3}}}$$
$$z = \frac{-.200}{\sqrt{0.010 + 0.009}}$$
$$z = \frac{-.200}{0.138}$$
$$z = -1.451$$

Note that the sign of the z simply indicates whether r'_1 is larger or smaller than r'_2 .

Third, compare our observed z with the critical z of 1.96. We use the absolute value (that is, ignore the sign). If our observed z is less than 1.96, the difference between the two r values is not significant (p > .05); whereas if the observed z is greater than 1.96, the difference is significant (p < .05).

In our example, the absolute value of the observed *z* is 1.451, and therefore p > .05.

Reporting the results

When reporting the outcome of the comparison, if you have only used the critical value of *z*, you would write:

There was no significant difference between the correlation coefficients of .287 for women and .487 for men (z = 1.45, p > .05). If you used *z* tables to find *p* more specifically, you would report: There was no significant difference between the correlation coefficients of .287 for women and .487 for men (z = 1.45, $p \simeq .147$).

Section 7: BRIEF INTRODUCTION TO REGRESSION

Like correlational analysis, regression is also concerned with the relationship between variables. But while correlation is just used to describe the relationship between two variables (i.e. description), regression can be used for prediction. Regression is a statistical technique that allows us to predict someone's score on one variable from their scores on one or more other variables. This is particularly helpful when we want to make predictions that extend beyond our current data range (extrapolation).

- Unlike the situation with correlation, in regression we attempt to specify the variables in terms of being dependent or independent variables.
- Regression involves one dependent variable which we want to predict, known as the 'outcome' or 'criterion' variable, and one or more independent variables, which we refer to as the 'predictor variables'.
- When we only have one predictor variable, we have a bivariate regression; in contrast, multiple regression (see Chapter 10) involves two or more predictor variables.
- The predictor variable can be measured using a range of scales, although ideally at interval or ratio level, but the criterion variable must be measured using a ratio or interval scale.
- Human behaviour is inherently noisy, and therefore it is not possible to produce totally accurate predictions; however, regression allows us to assess how useful a predictor variable is in estimating a participant's likely score on the criterion variable.
- As with bivariate correlation, bivariate regression does not imply causal relationships unless a variable has been manipulated.

Regression as a model

As we discuss in Chapter 8, Section 1, human behaviour is variable and therefore difficult to predict. A model is an attempt to explain and simplify data we have measured in a way that allows prediction of future cases. Say, for example, that we have measured how confident each student feels about how to use SPSS after they have completed a module. The simplest model is the mean; if the mean confidence of students is 33.58 (on a standardised confidence scale from 1 to 50), then we can predict that other students who complete the module the following year will have a confidence score of 33.58. However, there will be much error! In other words, the difference between each observed value and the predicted value (the mean) will be large for many students. Of course, students vary in other ways, and other variables can affect their confidence in using SPSS; for example, how much time they have spent practising with SPSS. If we measure that as well as confidence, we can use the technique of regression to model the relationship between those two variables. We can then predict how confident a student will be from how much time they have spent practising. There will still be error, but it will be less than in the situation when we only used the mean student confidence. Figure 6.6 illustrates how well (or not!) using the mean as a model works for these data, compared with using the regression line (or the line of best fit).

The amount of error in the model is indicated by the residuals (the difference between the observed value and predicted value for each case). We describe residuals in more detail below.

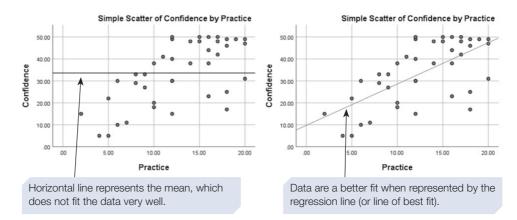


Figure 6.6 Comparing the mean and regression line as models

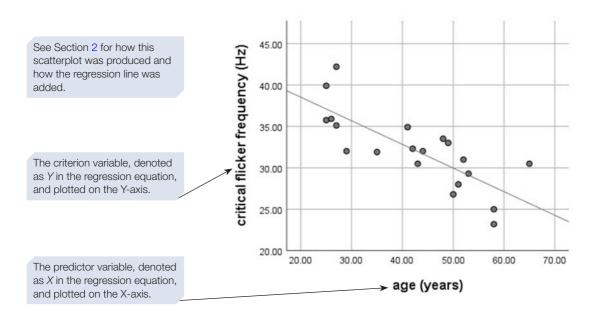
Section 8: BIVARIATE REGRESSION

From bivariate correlation to bivariate regression

Earlier sections covered correlation between two variables. Here, we use the term 'bivariate', which indicates that only two variables are involved, to distinguish from 'multiple', which indicates that there are more than two variables. To illustrate this section, we will use the example previously used when obtaining a scatterplot with a regression line, and also for Pearson's correlation coefficient. Age and CFF (*critical flicker frequency:* the frequency at which someone can no longer perceive a light flickering) were measured; the study is described in Section 2.

In bivariate correlation, we consider the strength of association between two variables, and do not consider whether one might be independent and the other a dependent variable. The technique of regression, however, allows prediction of one variable from another; thus we need to distinguish between the two variables. Some authorities use the terms 'independent' and 'dependent' variable; however, prediction does not necessarily mean direct causation, and other authorities use a different labelling system, which we prefer. In this system, the terms are the 'predictor variable' and the 'criterion variable'. The criterion variable is said to be predicted by the predictor variable. Another labelling system, used in equations and graphs, is *X* and *Y*. The predictor variable is denoted *X*, and the criterion variable is denoted *Y*. A scatterplot with regression line for the example data is shown again below. Age is classified as the predictor, or *X*, variable, and CFF as the criterion, or *Y*, variable. It is that way round because we assume that something to do with the ageing process affects CFF, rather than that CFF affects age or ageing. There can often be a logical reason such as that for classifying the variables. Mathematically, however, regression works

equally well in the opposite direction; if we used CFF as the predictor and age as the criterion, we can predict someone's age from their CFF with the same degree of accuracy.



The bivariate regression equation

In Section 2 we introduced the regression line that can be added to the scatterplot (using the Fit Line at Total option, see Section 2), and we will now explain the equation underlying the regression line. The relationship between any two variables that have a linear relationship is given by the equation for a straight line:

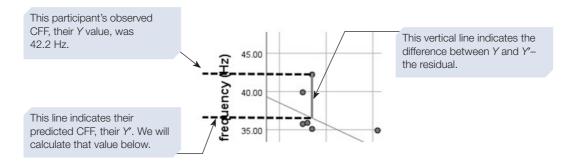
$$Y = a + bX$$

- *Y* is the criterion variable. When we use the equation to predict values of *Y* from observed values of X, we can use the symbol Y' (pronounced Y prime).
- X is the predictor variable.
- a is the intercept; it is the value of Y' when X = 0; when the regression line is added to the scatterplot, a is the value of Y at the point where the line intercepts the Y-axis (if the X-axis starts at 0). Thus, in our example, a is the CFF in Hz for someone who is aged 0 years. For purposes of prediction, you should not try to extrapolate much beyond the range of values you measured. Nonetheless, for purposes of the equation, *a* is the value of Y' when X = 0. Remember that, as default, SPSS applies axis scales around the values occurring in the data as in the example above, so scatterplot axes may not include 0. In such cases, a cannot be read from the graph; however, SPSS will provide the value of a, and also of b, as we show below.

The regression process involves finding a solution to the equation (that is, identifying values of a and b) for which the residuals are at a minimum, as described next.

Residuals

A residual is the difference between the observed value of Y for a participant and the value predicted for them by the regression equation (Y'). This section of the scatter-plot illustrates that.



If the two variables are perfectly correlated, then all the points will fall exactly along the straight line and the residuals will all equal zero. It is unlikely, however, that two variables measured in psychological research will be perfectly correlated; normally, there will be a difference between most Y values and their Y'. The difference, Y-Y', is considered error, and known as the 'residual' for each case. Note that residuals can be negative or positive, so they are usually squared when used (see next paragraph). There is error for a number of reasons. There is always error in measurement; this may be due to error in the scale (for example, questionnaires are unlikely ever to be perfect measures of a construct), but it is also due to extraneous individual or situational variables. In addition, the criterion variable is likely to be affected not just by the predictor variable but also by variables we have not measured.

The best solution to the regression equation will involve values of *a* and *b* that minimise residuals – that is, the predicted values are as close as possible to the observed values, on average. The least squares criterion is most commonly used to find the best solution; in this criterion $\Sigma(Y-Y')^2$ (the sum of the squared differences between observed and predicted scores) is at a minimum.

Proportion of variance explained

In regression, we wish to explain the degree of variation or dispersion in our data, usually referred to as the variance; we can ask: 'How much of the variance in the criterion variable is explained, or accounted for, by the predictor variable?' For bivariate relationships, r^2 gives the proportion of variance explained as described in Section 3. We will return to the concept of variance explained when describing ANOVA in Chapter 8, Section 1, and for multiple regression in Chapter 10.

How to perform bivariate regression in SPSS

For this purpose, we will continue to use the age and CFF data. Click on <u>Analyze</u>, <u>Regression</u>, <u>Curve Estimation</u>. You will be presented with the Curve Estimation dialogue box (shown below). You now select *Y*, the criterion (dependent) variable, and *X*, the predictor (independent) variable. SPSS allows you to add more than one *Y* variable, but we currently wish to predict a single *Y* ('cff') from the *X* ('age').

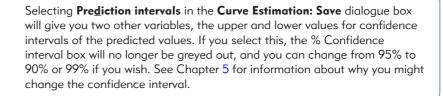
📽 Curve Estimation	Dependent(s):	X.	Select the criterion (or dependent) variable and click here to move it into the Dependent box.
	Independent Yariable: Image: Time Image: Time		Select the predictor (or independent) variable and move it into <u>V</u>ariable in the Independent area.

You can enter two or more Y variables into the **Dependent** box, and then SPSS will carry out the bivariate regression procedure separately for each Y with the X. If you do have more than one Y, to obtain all the output for all the Ys, you must select **Display ANOVA table** in the dialogue box.

We leave the other settings in the **Curve Estimation** dialogue box as they are. In the Models area of the dialogue box, the default is **Linear**, which applies a bivariate linear regression model and is appropriate for the example we are using. In the future, if you have data with nonlinear relationships, you could explore the other types of curve.

Next click on the S<u>a</u>ve button (top right) to obtain the Curve Estimation: Save dialogue box (shown below). SPSS calculates certain values for each case, and the S<u>a</u>ve command allows you to save these to your data file. We have selected <u>Predicted</u> values and <u>Residuals</u>.

💼 Curve Estimation: Save	>	<	
Save Variables	Predict Cases		
Residuals Prediction intervals 95 % Confidence interval	Predict through: Observation:	Select Predicted values and Residuals to add these to your data file.	•
The Estimation Period is: All cases			
	ontinue Cancel Help		



Click on **Continue** to return to the **Curve Estimation** dialogue box, then click on **OK**. You will be reminded that you have asked to add variables to your data file – if you are sure all is well, then click **OK**. (If you inadvertently added unwanted variables to your data file, you can delete them or just not save the amended file; SPSS does not automatically save the data file.)

The Output window will contain information and a graph (described below). The new variables will be in the Data file (as shown below), and described in the annotations.

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1		100			
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	/ age	Ø cff	P FIT 1	@ ERR 1	9 19
1	41.00	34.90	32.53616	2.36384	
2	43.00	30.50	31.96665	-1.46665	\backslash
3	25.00	35.75	37.09222	-1.34222)
4	42.00	32.30	32.25140	.04860	
5	51.00	28.00	29.68862	-1.68862	
6	27.00	42.20	36.52271	5.67729	
7	27.00	35.10	36.52271	-1.42271	
8	48.00	33.50	30.54288	2.95712	
9	58.00	25.00	27.69534	-2.69534	
10	52.00	31.00	29.40386	1.59614	
11	58.00	23.20	27.69534	-4.49534	

The new variable 'FIT_1' holds the predicted values, Y'. Thus, for participant 1, who is 41 years old, CFF is predicted from the regression equation to be 32.5 Hz.

The second new variable 'ERR_1' holds the residual (Y, the observed value, minus Y', the predicted value). See text below for explanation.



If you selected **Prediction intervals** in the **Curve Estimation: Save** dialogue box, the upper and lower values for confidence intervals of the predicted values will also be added to the data file. We show these on the last page of this chapter.

The residual is equal to Y (the observed value) minus Y' (the predicted value). It is sometimes called the 'error' (indicated by the SPSS variable name 'ERR_1') as it can be considered the amount of error in the prediction. The residual values can be interesting, as we will explain now, but they are not necessarily useful for our purposes. In this example, the residual 'ERR_1' = 'cff' - 'FIT_1'. For participant 1, the predicted value is 2.4 Hz less than their observed value of 34.9. If you scan the 'ERR_1' column, you will see that the absolute value (that is, ignoring whether they are negative or positive) of the residuals varies from effectively 0 to nearly 6. The mean of the absolute residuals is indicative of the strength of the relationship between the two variables, but residuals are unstandardised. That is, they are on the same scale as the original data and not standardised (*z*-scores are an example of a standardised measure, as described in Chapter 3). Thus, residuals can be difficult to interpret because you must consider the scale on which the variable is measured.

As mentioned above, you can enter more than one *Y* in the **Curve Estimation** dialogue box, in order to carry out separate bivariate regressions of *X* with each *Y*. For example, in addition to CFF, you may have a memory score for the same participants. If you do enter two *Y*s, then the new variables for the second *Y* will be called 'FIT_2' and 'ERR_2'.

In addition to the new variables in the data file, SPSS will also give output with information about the bivariate regression, and the output obtained for the bivariate regression of CFF with age is shown below.

Obtained using: Analyze, Regression, Curve Estimation

Curve Fit

Model Description

Model Name		MOD_1
Dependent Variable	1	cff
Equation	1	Linear
Independent Variable		age
Constant		Included
Variable Whose Values	Label Observations in Plots	Unspecified

This table gives information about the model you requested.

Case Processing Summary

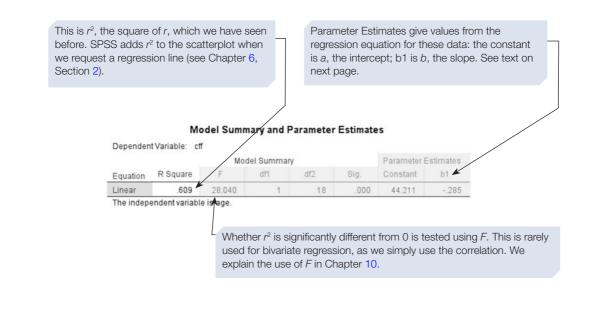
	N
Total Cases	20
Excluded Cases ^a	0
Forecasted Cases	0
Newly Created Cases	0

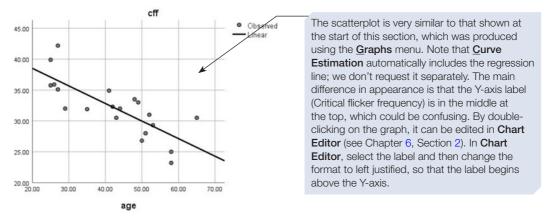
 Cases with a missing value in any variable are excluded from the analysis. The Variable Processing Summary table shows information about cases for each variable.

Variable Processing Summary

		Variables			
		Dependent cff	Independent age		
Number of Positive Valu	es	20	20		
Number of Zeros		0	0		
Number of Negative Val	Jes	0	0		
Number of Missing	User-Missing	0	0		
Values	System-Missing	0	0		

CHAPTER 6





The Parameter Estimates section of the Model Summary and Parameter Estimates table (on the previous page) shows:

- 1. The constant or a, the intercept. Thus, when the modelling process was applied to the values in the data set, the CFF value for someone who is 0 years old was estimated to be 44.21 Hz.
- 2. b1 or *b*, the slope or the regression coefficient. For these data *b* is negative; thus, if age increases by one year, the CFF reduces by 0.285 Hz. More usefully for a report, we could instead say that when people are 10 years older, their CFF would be expected to have reduced by 2.85 Hz. This is, of course, the typical reduction predicted from these data.

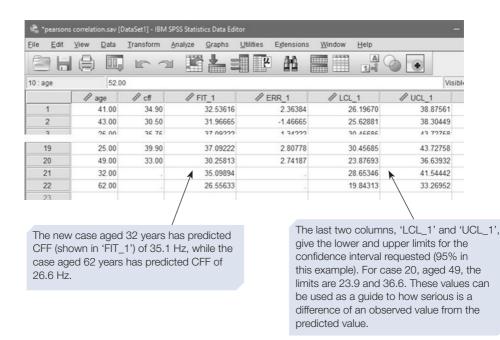
Thus, the straight line equation for these data is:

CFF = 44.21 + -.285age, or CFF = 44.21 -.285age

Using the procedure to predict Y for new cases

If you wish to predict *Y* for new cases without using the equation yourself, simply enter the new case/s into the data file and run the procedure again. As an example, we added two hypothetical new cases, with 'age' data only, to the data file and ran the procedure again. In this run, we also selected **Prediction intervals** to obtain confidence intervals for 95% CI. The section of the data file with these new cases is shown below.

An inspection of the output viewer will show you that the two new cases, without observed values for the *Y* variable, affect the excluded cases and missing values, but not the other output.



Remember that for the purposes of prediction, you should not try to extrapolate much beyond the range of values that you measured. The two new cases we have added are within the age range of the participants in the study.

Summary

- This chapter introduced you to statistical tests of correlation that will tell you if there is a significant relationship between two variables, and provide you with information on the strength and nature of this relationship.
- Your choice of test of correlation will depend on whether the data are parametric and whether the relationship between the variables is linear.
- You should first obtain a scatterplot to observe any general trend in your data and to see if any relationship is linear.
- You can add a regression line to the scatterplot.
- The value of the correlation coefficient indicates the strength of the correlation, and is a measure of effect size.
- The sign of the correlation coefficient is important as it indicates whether the relationship is positive or negative.
- It is not possible to infer a causal relationship from a correlation.
- Partial correlations were introduced. These allow you to test the strength and direction of a relationship between two variables whilst controlling for the effect of one or more other variables
- We also introduced how to compare the strength of two correlation coefficients.
- Comparing the strength of two correlation coefficients can be useful if there are grounds to hypothesise that the correlation between two variables will differ for two groups (such as men and women).
- When comparing correlations, you should check whether the groups differ on variables other than the grouping variable (gender, in our example). If they do differ on other variables, you will not know if any difference between the correlations is due to gender or those other variables.
- We also introduced bivariate regression, which allows us to predict one variable (the criterion) from another (the predictor).
- In Chapter 10, we discuss multiple regression, which allows use of more than one predictor variable, in the same model, to predict the criterion variable.
- For guidance on incorporating SPSS output into a report, or on printing the output, see Chapter 14.

Tests for nominal data

In this chapter

- Nominal data and dichotomous variables
- Chi-square test versus the chi-square distribution
- The goodness of fit chi-square
- The multidimensional chi-square
- The McNemar test for repeated measures



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Section 1: NOMINAL DATA AND DICHOTOMOUS VARIABLES

- Nominal data, also referred to as 'categorical data', are data measured using scales that categorise the observations or responses, for example being male or female.
- For convenience, each category is allocated a number when entering data into SPSS. These are numbers that cannot be put into any meaningful order; if they could, they would be ordinal data.
- With nominal data, the numbers only represent the category of which the participant belongs. That is why nominal data are sometimes called 'qualitative data', and, by contrast, the ordinal, interval and ratio levels of measurement are called 'quantitative data'. The use of these terms in this way is different from the use in qualitative and quantitative research. Quantitative research can include all levels of measurement, including nominal.
- In experimental studies the independent variable is a nominal variable, while the dependent variable is often interval or ratio (although may sometimes be ordinal). However, when both of the variables we are interested in are nominal, then we need to use a different type of test devised specifically for nominal data.

- Some nominal variables can only have two values for example, whether or not someone has ever had a head injury. These are known as 'dichotomous variables'.
- A dichotomous variable is a nominal variable that can *only* take one of two values. For example, if you classify smoking status as smoker or non-smoker, then someone who smokes only very occasionally would be classified as a smoker, whereas an ex-chain smoker would be classified as a non-smoker.
- Some nominal variables can have more than two values. For example, if you recorded the smoking status of your participants, you could use three categories: smoker, never smoked and ex-smoker. If you collected data from an online survey and recorded nationality, you might have a great many different categories among your participants.

Descriptives for nominal data

It is important to consider which summary descriptive statistics are appropriate for use with nominal data. If you have recorded your participant's sex, then calculating values such as the mean sex or median sex is meaningless.

The only summary descriptives suitable for nominal data are counts, or frequencies, and percentages. Thus, we could say that, of 20 participants, 15 (75%) are female and 5 (25%) are male. We can display those counts and percentages in a table, and can illustrate them using a bar chart. We show both of these options in this chapter. Note that histograms should be used for displaying data of at least ordinal level of measurement, and not for nominal data.

Entering nominal data into SPSS

Data entry for nominal variables is easy. For any variable measured on a nominal scale, make sure that you assign value labels to help you remember what the different values mean (see Chapter 2), then simply enter the number chosen to represent the category. As with most data in psychology, the rows in the data file represent participants, and the columns are our variables.

For example, we might have a column for the variable 'Sex', in which we enter a value of either 1 or 2 (male or female), and a column for Smoking status, which is coded as 1 for 'smoker', 2 for 'never smoked' or a 3 for 'ex-smoker'.

Section 2: CHI-SQUARE TEST VERSUS THE CHI-SQUARE DISTRIBUTION

- The chi-square test is one of the tests specifically designed to handle nominal data. There are two different forms of the chi-square test: the goodness of fit chi-square test and the multidimensional chi-square test.
- The goodness of fit chi-square is used to test whether an observed pattern of events differs significantly from what would be expected by chance alone.
- The multidimensional chi-square test can be used either as a test of association or as a test of difference between nominal variables.

The chi-square *test* makes use of the chi-square *distribution* to test for significance. The distinction between 'test' and 'distribution' is explained by Howell (2013). In this book we only cover the chi-square inferential test, and not the chi-square distribution. However, it is worth noting that, for some other statistical tests, including some for data with a level of measurement other than nominal, a chi-square distribution is used to test for significance; for example, the Kruskal–Wallis and Friedman tests covered in Chapter 8, Sections 2 and 3, respectively. Also note that the chi-square distribution is not used to test for significance in all inferential tests that are applied to nominal data. For example, for the McNemar test (Section 5), SPSS uses the binomial distribution.

Section 3: THE GOODNESS OF FIT CHI-SQUARE

In this type of chi-square test we are testing whether the observed pattern of events (or scores) differs significantly from what we might have expected to occur by chance. For example, we might ask whether a group of smokers choose brand A cigarettes more often than brand B. Here, we are effectively asking the question: 'Do significantly more than 50% of our smokers choose one brand over the other brand?' In practice, this form of the chi-square test is not often used in psychology. This example of cigarette brands relates to one of the few times the authors have ever used this form of the test. An undergraduate student undertook a project examining the effect of cigarette advertising on cigarette choice. As part of this project, she listed a series of personality characteristics that were implied by cigarette adverts. For example, some cigarette advertisements might imply a sophisticated personality. These personality statements were then presented to smokers who were asked to indicate to which of five brands of cigarettes they thought the statement best applied. The responses for each statement were analysed using the chi-square goodness of fit test to compare the observed distribution against that predicted by the null hypothesis (that the five brands would be equally often selected). This is an interesting, but rare, example of the use of this form of the chi-square test in psychology. Much more common is the second form of this test, described in the next section, which allows us to consider whether two variables are independent of one another (or conversely, whether they are associated).

To perform the goodness of fit chi-square test

Note that this type of chi-square test is accessed via the chi-square command that can be found under **Nonparametric tests** in the <u>Analyze</u> menu. However, as this form of the test is used infrequently in psychology, we will not be demonstrating it here.



A common error for novice users is to select the wrong form of the chisquare test. Make sure you know which version you want and select appropriately.

Section 4: THE MULTIDIMENSIONAL CHI-SQUARE

The multidimensional chi-square test can be thought of in two ways: as a test of association or as a test of difference between independent groups.

It can be thought of as a test of association because it allows us to test whether two variables are associated or whether they are independent of each other. For example, let's modify our cigarette example so that 50 smokers and 50 non-smokers are asked to choose which of two cigarette adverts they preferred. The multidimensional chi-square test would allow us to ask the question: 'Is the pattern of brand choice independent of whether the participant was a smoker or not?' Another example would be to determine whether receiving or not receiving a particular medical treatment was associated with living or dying. Yet another might be to see whether a person's sex was independent of their choice of favourite colour. In psychology, we often need to test whether nominal variables are, or are not, independent of each other. The experimental hypothesis would be that the two variables are not independent of each other; for example, we could hypothesise that people receiving a particular treatment are less likely to die than those not receiving the treatment.

Note that another way of phrasing that last hypothesis is that there will be differences between the number of people who die under each treatment condition. Thus, the multidimensional chi-square can be thought of as both a test of association, and a test of difference. Whichever way you think of it, the type of data and the way chisquare assesses those data are the same.

General issues for chi-square

Causal relationships

Remember, even if you find there is a significant association between your nominal variables, you cannot infer a causal relationship between them unless you have directly manipulated one of them. For example, unless you randomly allocated patients to either receive the treatment or not, then you cannot claim that the treatment reduces the rate of death. It may be that the hospital only offered the treatment to the patients who were strongest, and it may be this which gives rise to the results

you see. The only conclusion we can draw is that there is an association between the treatment a patient received and the outcome. If we wanted to test whether there was a causal relationship, we would need to undertake a randomised control trial in which patients were randomly allocated to treatment and the outcome assessed. It's important to remember this when describing the results of a chi-square test.

Type of data

In order to use chi-square, our data must satisfy the following criteria:

- The variables must be measured on a nominal level of measurement, giving frequency data. In other words, our data must be able to tell us the number of times some event has occurred. We can, of course, convert other types of data into nominal data. For example, suppose we have recorded participants' annual income – we could recode these data, scoring each participant as either 'High income' or 'Low income' (above or below the median income for the group). We could then count to give frequency data – the *number* of high- and low-income participants we have observed. (See Chapter 4, Section 5 for information on how to recode in this way.)
- 2. For multidimensional chi-square, we must have collected data of this sort on at least two variables. For example, in addition to the high-low-income data above, we might also know whether each of these participants is a smoker or not.
- 3. The categorisations of each of the variables must be mutually exclusive. In other words, each participant must be *either* a smoker *or* a non-smoker, and *either* high or low income. In other words, each participant must fall into one and only one of the cells of the contingency table (see below).
- 4. Every observation must be independent of every other observation. This will not be the case if you have more than one observation per participant. (Nominal data from a repeated measures design can be analysed by means of the McNemar test, see Section 5.)

The N*N contingency table

When we have nominal data of this form, we can best display the frequencies or counts in a contingency table. If we have two variables, each with two levels (as in the example above), then we draw what is called a 2*2 (pronounced 'two-by-two') contingency table. So, if we had 100 participants in our example data set, the contingency table might look like Table 7.1.

	High income	Low income	Row totals
Smokers	10	20	30
Non-smokers	35	35	70
Column totals	45	55	100 (Grand total)

Table 7.1	An example of a 2*2 continger	ncy table
-----------	-------------------------------	-----------

The numbers in this table represent the numbers of participants who fall into each cell of the table (and remember that each participant can be in only one cell). So we can see that, of the 30 smokers in our study, 10 are high income and 20 are low income. Similarly, we can see that, of the low-income group, 20 are smokers and 35 are non-smokers. Thus, the contingency table is a useful summary of the data.

Rationale for chi-square test

If there was no association between smoking and income, then we would expect the proportion of smokers in the high-income group to be the same as the proportion in the total sample. That is, we would expect 45/100 or 45% of the smokers to be high income. As there were 30 smokers in total, we would expect (45% of 30) = 13.5 of the smokers to be in the high-income group. In this way, we can work out the expected frequencies for each cell. The general formula is:

expected frequency = $\frac{\text{row total} \times \text{column total}}{\text{grand total}}$

The chi-square test calculates the expected frequency for each cell and then compares the expected frequencies with the observed frequencies. If the observed and expected frequencies are significantly different, it would appear that the distribution of observations across the cells is not random. We can, therefore, conclude that there is a significant association between the two variables: Income and smoking behaviour are not independent for our (fictitious) sample of participants.

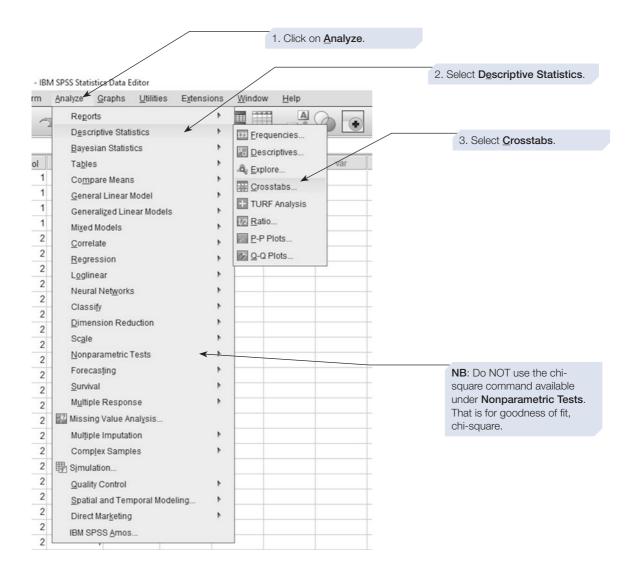
Chi-square will actually allow us to calculate whether more than two variables are independent of each other. However, as we add variables, it fast becomes very difficult to interpret the results of such an analysis, so we would recommend that you resist the temptation to add extra variables unless you are sure you know what you are doing. It is, however, perfectly reasonable to have more than two categories of each variable – for example a 3*3 chi-square is quite acceptable.

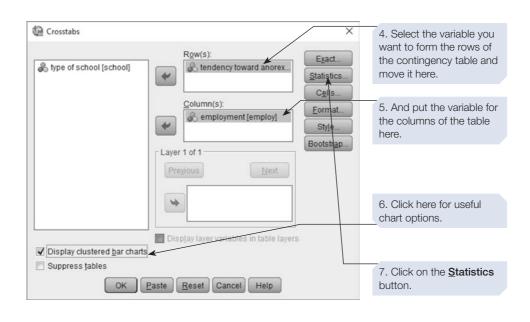
Example study: investigating tendency towards anorexia

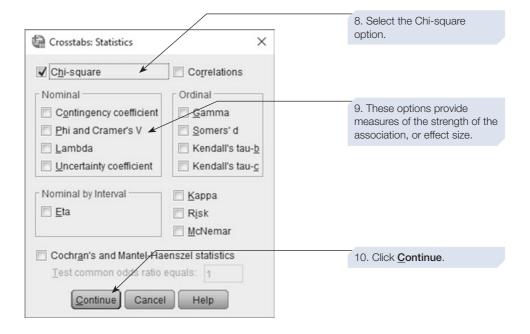
To illustrate the use of chi-square, we will use some fictitious data based on research conducted by one of our past students. Eighty young women completed an eating questionnaire, which allowed them to be classified as either high or low anorexia (participants with high scores are at greater risk of developing anorexia). In addition, the questionnaire asked for the employment status of the women's mother (full-time, part-time or unemployed) and their cultural background (Caucasian, Asian or other) and type of school they attended (private or state). Previous research has suggested that the incidence of anorexia is higher among girls attending private schools than state schools, and higher among girls whose mothers are not in full-time employment. In addition, the incidence seems to be higher in Caucasian girls than non-Caucasian girls. We therefore hypothesised that there would be an association between these factors and the classification on the eating questionnaire. To test this hypothesis, we conducted a series of chi-square analyses. (These data are available in the Appendix or from macmillanihe.com/harrison-spss-7e.)

To perform the multidimensional chi-square test

The multidimensional chi-square is accessed under the <u>C</u>rosstabs command. <u>C</u>rosstabs draws up contingency tables, and chi-square is an optional inferential statistic within this command.







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Crosstabs	Row(s):	×	
🗞 type of school [school]	Column(s): Column(s): A employment [employ] Layer 1 of 1 Preyious Next	Exact Statistics Cells Eormat Style Bootstr <u>a</u> p	11. Click C<u>e</u>lls .
Display clustered bar charts	Display layer variables in table layers	5	
Suppress tables	Paste Reset Cancel Help		

		_	12. These options control the information
Crosstabs: Cell Display	×		included in the cells of the contingency table.
Counts	z-test		Select Observed and Expected .
Expected Hide small counts	Adjust p-values (Bonferroni method)		
Less than 5			13. And <u>R</u> ow, <u>C</u> olumn and <u>T</u> otal.
Percentages	Residuals		
Row K	Unstandardized		
Column	Standardized		
<u> </u>	Adjusted standardized		
┌ Noninteger Weights —			
Round cell counts	○ Round case weights		
O Truncate cell counts	◎ Truncate case weig <u>h</u> ts		
◎ No adjust <u>m</u> ents			14. Click <u>Continue</u> , then
	Le Cancel Help		OK.

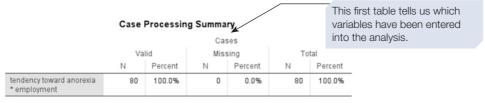
Two sets of annotated output are given below. The first is from the 2*3 chi-square exploring the association between the tendency towards anorexia and mother's employment status (the one shown above in the **Crosstabs** dialogue box). The second is from the 2*2 chi-square exploring the association between tendency towards anorexia and type of school attended.

SPSS output for chi-square without using Exact option

Obtained using menu items: Descriptive Statistics > Crosstabs

Output for first chi-square: tendency towards anorexia * employment (a variable with two levels against a variable with three levels)

CROSSTABS



This is our row variable. It has two levels (high; low).

	/			employment			
			f/t	none	p/t	Total	
tendency toward anorexia Higi	high	Count	14	13	11	38	
		Expected Count	14.7	11.9	11.4	38.0	
		% within tendency toward anorexia	36.8%	34.2%	28.9%	100.0%	
		% within employment	45.2%	52.0%	45.8%	47.5%	And this is the column variable
		% of Total	17.5%	16.3%	13.8%	47.5%	It has three levels (f/t; none;
Iov	low	Count	(17)	12	13	42	p/t) so there is a column for each level.
		Expected Count	16.3	13.1	12.6	42.0	each level.
		% within tendency toward anorexia	40.5%	28.6%	31.0%	100.0%	
		% within employment	54.8%	48.0%	54.2%	52.5%	
		% of Total	21.3%	15.0%	16.3%	52.5%	
Total		Count	31	25	24	80	·
		Expected Count	31.0	25.0	24.0	80.0	If you selected the options
		% within tendency toward anorexia	38.8%	31.3%	30.0%	100.0%	we suggested, then each cell of the table will contain five
		% within employment	100.0%	100.0%	100.0%	100.0%	values. These are explained in
		% of Total	38.8%	31.3%	30.0%	100.0%	the text below.

Within each cell of the contingency table you will see five values reported:

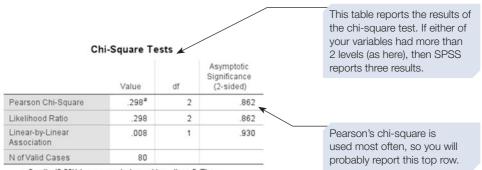
Count: This is the number of cases falling into the cell. For example, on the top left cell this is the number of cases which scored high on the anorexia scale and in full-time employment.

Expected Count: This is the number we would expect to see in this cell if there was no association between the two variables.

% within tendency toward anorexia: This is the number of cases in the cell expressed as a percentage of the row total. For example, 36.8% of girls who score high on the anorexia scale have a parent in full-time employment.

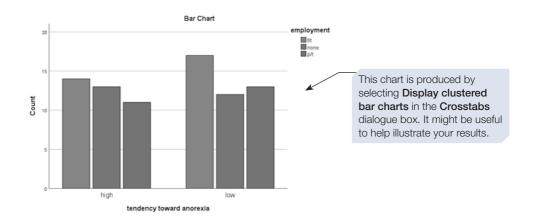
% within type of employment: This is the number of cases in the cell expressed as a percentage of the column total. For example 45.2% of the cases with a parent in full-time employment scored high on anorexia.

% of Total: This is the cases in this cell as a percentage of the total number of cases.



 a. 0 cells (0.0%) have expected count less than 5. The minimum expected count is 11.40.

If you used the **Exact** option, available on the **Crosstabs** dialogue box, the table above would have three extra columns: we will describe these later in this section.





There was no relationship between tendency towards anorexia and the employment status of the mother: $\chi^2(2, N = 80) = 0.29$, p = .862.

We will discuss reporting the outcome of chi-square in more detail, but first we explain the output for the second chi-square, which explores the association between tendency towards anorexia and type of school attended.

Output for second chi-square: tendency towards anorexia* education (two variables each with two levels)

Each row of this table gives information about each level of the column variable; e.g. this row gives information for girls in the 'high' group. It shows the figures for those in a state school separately from those attending a private school. The total is for all girls in the 'high' group. We can see that, of the total of 38 'high' girls, 4 were in a state school compared with 34 in a private school.

			type of school			
*			State	Private	Total	
tendency toward anorexia	high	Count	4	34	38	
		Expected Count	15.7	22.3	38.0	
		% within tendency toward anorexia	10.5%	89.5%	100.0%	
		% within type of school	12.1%	72.3%	47.5%	
		% of Total	5.0%	42.5%	47.5%	
	low	Count	29	12	42	
		Expected Count	17.3	24.7	42.0	
		% within tendency toward anorexia	69.0%	31.0%	100.0%	
		% within type of school	87.9%	27.7%	52.5%	
		% of Total	36.3%	16.3%	52.5%	
Total		Count	33	XX	80	
		Expected Count	33.0	47.0	80.0	
		% within tendency toward anorexia	41.3%	58.8%	100.0%	
		% within type of school	100.0%	100.0%	100.0%	
		% of Total	41.3%	58.8%	100.0%	

The columns give information about each level of the row variable. This column shows figures for girls in private schools. Figures are given separately for those in the 'high' and 'low' groups and for the total for all girls in private education. Of the total of 47 girls in private education, 34 of them were 'high' and 13 of them were 'low'. Within each cell, we are given:

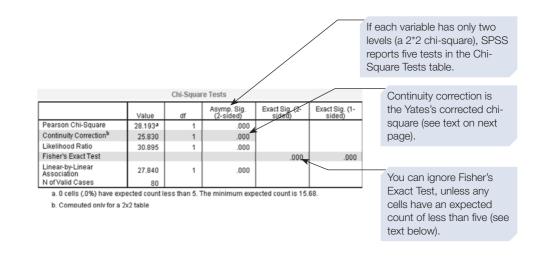
Count = the observed frequency. The number of participants falling into the cell; i.e. the number of girls who are 'low' and attended a state school.

Expected Count: the number expected for this cell assuming no association (see text).

% within tendency toward anorexia: the cases in this cell as a % of row total; i.e. % of 'low' girls who attend a state school.

% within type of school: the cases in this cell as a % of the column total; i.e. the % of girls who attend a state school who are 'low'.

% of Total: the cases in this cell as a % of the total number of participants.





If you used the **Exact** option, available on the **Crosstabs** dialogue box, the table above would have one extra column: we will describe it later in this section.

Interpreting and reporting results from chi-square

SPSS reports several different measures of p. It is probably best to use Pearson's (the chi-square test was developed by Karl Pearson). Note that, for chi-square, the value of df (degrees of freedom) is not related to the number of participants. It is the number of levels in each variable minus one multiplied together. So, for a 2*2 chi-square, df = (2-1)(2-1), which equals 1. When reporting the outcome, we also provide the value of N - the number of valid cases in our contingency table.

For a 2*2 table, SPSS also calculates the result with and without the continuity correction, or Yates's correction. This is a statistical correction used in analyses with relatively few participants or when you have reason to believe that your sample is not a very good approximation to the total population. There is disagreement about whether to use it, but the Exact test, described below, can be used instead for small samples.

It is important to understand that the chi-square result on its own cannot tell you about the pattern of your results. For that, you have to look at the contingency table. For example, when reporting the results of the second chi-square result (on previous page), you might write: 'Only a minority (12%) of those girls attending a state school scored high on the scale, whereas the majority (72%) attending a private school scored high on the scale.' If you made a specific one-tailed prediction about the direction of the relationship between the two variables (here we predicted that there will be a higher tendency towards anorexia in the private school pupils) and the pattern of results revealed by the contingency is compatible with this prediction (as here), then you can use the chi-square results to assess whether this particular association is significant.

CHAPTER 7

The strength of the association between the two variables can also be obtained in the **Crosstabs: Statistics** dialogue box by selecting Phi and Cramer's V. The Symmetric Measures table will appear after the Chi-Square Tests table.

Symmetric Measures			Approximate Significance	The value of Phi indicates the magnitude of the association. It can be considered equivalent
Nominal by Nominal	Phi	594	.000	to Pearson's <i>r</i> (see Chapter 6, Section 3).
	Cramer's V	.594	.000	,
N of Valid Cases		80		

Just as we can square *r* to give an estimate of the proportion of variation that is common to the two variables, so we can square Phi. For these data Phi = -.594. Squaring this gives us a value of .35 or 35%. This tells us that 35% of the variation in the tendency towards anorexia score is accounted for by the type of school attended. We must remember that because none of the variables in this study were manipulated a significant chi-square result does not imply a causal relationships between variables.



The relationship between tendency towards anorexia and the type of school attended was significant: $\chi^2(1, N = 80) = 28.19, p < .001$. The association was of moderate strength: Phi = -.59, and the type of school attended accounted for 35% of the variance in the score on tendency towards anorexia scale.

You could also include a table of counts and expected frequencies, bar charts and other information as appropriate.

Use of exact tests in chi-square

Look at the chi-square tables in the preceding pages. There is a note at the bottom of each, informing you of the number of cells with expected frequencies (what SPSS calls expected counts) of less than 5. It is important that you always check this note. For both chi-square analyses above, there are no cells with this problem. However, if you do undertake a chi-square analysis and SPSS reports that one or more cells has an expected frequency of less than 5, then you must take some action. If you are doing a 2*2 chi-square, you can make use of Fisher's Exact test, which SPSS reports. This test can be used when cells have low expected frequencies (Siegel and Castellan, 1988, 103–11). However, this test is only available for 2*2 tables. If you are performing something other than a 2*2 chi-square and encounter this problem, then you can use the Exact option.

To demonstrate this for you, we have undertaken a further chi-square analysis, exploring a possible association between cultural background and tendency towards anorexia. In the third SPSS output (see below), two cells have an expected frequency of less than 5.

Output for third chi-square: tendency towards anorexia* cultural background (a variable with two levels against a variable with three levels)

Chi-Square Tests

	Value	df	Asymptotic Significance (2-sided)
Pearson Chi-Square	7.866 ^a	2	.020
Likelihood Ratio	8.078	2	.018
Linear-by-Linear Association	2.016	1	.156
N of Valid Cases	80		

a. 2 cells (33.3%) have expected count less than 5. The

minimum expected count is 2.85.

A chi-square is not valid if any of the expected frequencies are less than 5. In this case two cells have this characteristic.

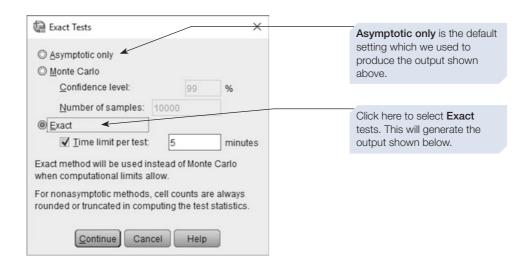
If SPSS prints a message like this you should consider using the **Exact** option (see text below).

As this was a 2^*3 chi-square, there is no Fisher's Exact test. Instead, we use the **Exact** button on the **Crosstabs** dialogue box.

Using the Exact option for chi-square

If you click on the Exact button in the Crosstabs dialogue box, then the Exact Tests dialogue box will appear (see below).

Crosstabs	×	
Row(s): type of school [school] newback Column(s): background [backgrnd] Layer 1 of 1 Previous Next Disp[ay layer variables in table layer	Statistics Cells Eormat Style Bootstrap	Click on the Exact button in the Crosstabs dialogue box.
Display clustered bar charts		
Suppress tables		
OK Paste Reset Cancel Help		



Output for third chi-square (2*3) with Exact option

		Ch	i-Square Test	s		
	Value	df	Asymptotic Significance (2-sided)	Exact Sig. (2- sided)	Exact Sig. (1- sided)	Point Probability
Pearson Chi-Square	7.866 ^a	2	.020	.015		
Likelihood Ratio	8.078	2	.018	.037		
Fisher's Exact Test	7.765			.017		
Linear-by-Linear Association	2.016 ^b	1	.156	.191	.114	.065
N of Valid Cases	80				7	
a. 2 cells (33.3%) have b. The standardized st		less than t	5. The minimum e	xpected count is 2	2.85.	
You can now report chi-square, as we sł					printed w	ee extra columi hen you select tion (see below



Analysis showed that two cells had an expected frequency of less than 5, so an exact significance test was selected for Pearson's chi-square. There was a relationship between tendency towards anorexia and cultural background: $\chi^2(2, N = 80) = 7.87$, exact p = .015).

Output for a 2*2 chi-square with Exact option

	Value	df	Asymptotic Significance (2-sided)	Exact Sig. (2- sided)	Exact Sig. (1- sided)	Point Probability
Pearson Chi-Square	7.744 ^a	1	.005	.006	.005	
Continuity Correction ^b	6.412	1	.011			
Likelihood Ratio	7.929	1	.005	.006	.005	
Fisher's Exact Test				.006	.005	
Linear-by-Linear Association	7.647°	1	.006	.006	.005	.004
N of Valid Cases	80			1		

Chi-Square Tests

a. 0 cells (0.0%) have expected count less than 5. The minimum expected count is 10.45.

b. Computed only for a 2x2 table

c. The standardized statistic is -2.765.

These two column always appear for a 2*2 chi-square. However, without the **Exact** option the only *p*-value reported was for Fisher's Exact Test. With the **Exact** option, *p*-values for other tests, including Pearson's chi-square, have been calculated and are reported here.

Performing a chi-square using the Weight Cases option

Imagine that you have an eager research assistant who has collected the data for you and diligently calculated the observed frequencies and set these out in a contingency table. Instead of entering each data point into SPSS to perform the chi-square test, you can instead enter just these observed frequencies and make use of the <u>Weight</u> Cases option. This is illustrated below.

Table 7.2 gives the observed frequencies for the data we analysed previously in this chapter. The 80 young women who completed an eating questionnaire were classified as either high or low anorexia, and the type of school they attended was a private or state school.

	High anorexia	Low anorexia
Private	34	13
State	4	35

Table 7.2 The observed frequencies	Table	7.2	The	observed	frequencies
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In order to enter these frequencies, you first need to set up a data file as shown below, creating three variables.

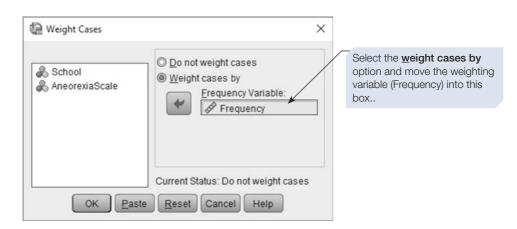
Pile Edit	[DataSet4] - IBM View Data	SPSS Statistics Transform		Graphs Ut	lities Extension	s Window	Help				
26			3 M	*=	TH M			۲			
	Name	Type	Width	Decimals	Label	Values	Missing	Columns	Align	Measure	
1	School	Numeric	8	0		(1, Private)	None	8	I Right	& Nominal	>
2	AneorexiaS	Numeric	8	0	7	{0, Low}	None	8	I Right	& Nominal	>
3	Frequency	Numeric	8	0		None	None	8	I Right	/ Scale *	1
4	1										Т
							\backslash				
e variable els (1 = F			-		/				wo levels	AnorexiaSc s (0 = Low;	

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	& S	chool	Aneorexia Scale	Frequence y	
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2		1	2	13	3
3		2	1	4	1
4		2	2	29	9
5					
-					

In Data View, enter the appropriate values for 'school' and 'anorexia', along with the corresponding observed frequency, as shown here.

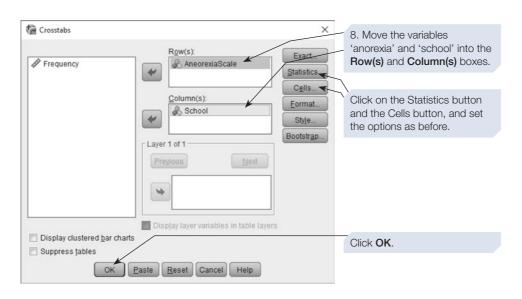
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Although no output is produced, and there is no change to the data table, you can now undertake the chi-square analysis as before.

Click on <u>Analyze</u> \Rightarrow <u>Descriptive</u> Statistics \Rightarrow <u>Crosstabs</u> to bring up the Crosstabs dialogue box (shown below).



The output produced in this way will be identical to that shown earlier in the chapter.

This is a quick and easy way to enter frequency data so it can be analysed using chi-square.

Section 5: THE MCNEMAR TEST FOR REPEATED MEASURES

- The McNemar test is used to analyse data obtained by measuring a dichotomous variable for related designs. The difference between this and the 2*2 chi-square test is that the chi-square test is for independent groups designs.
- Another difference is that, while the chi-square test can be used to test for an association between two variables, the McNemar test can only be used as a test of difference.
- Remember that a dichotomous variable, by definition, can only take one of two values (e.g. yes or no). Thus, the McNemar test is for a situation where you measure the same thing twice; for example, a 'before treatment' yes/no response, and an 'after treatment' yes/no response.

If you measure more than two values (e.g. yes/uncertain/no), SPSS will automatically apply the McNemar–Bowker test instead. We will not cover that test here.

Example study: gender and handwriting

To illustrate the use of the McNemar test, we will describe an experiment we carried out with students. It has been found that roughly two-thirds of handwriting samples can be correctly categorised as being written by a man or a woman. This is significantly above the chance level performance of 50% correct. A possible implication is that some (but not all) men and women tend to write in a gender-stereotyped manner. Very briefly, male handwriting tends to be irregular and untidy, whereas female handwriting is more rounded and neat. If this is the case, do people have a choice in their writing style? Hartley (1991) investigated this by asking children to try to imitate the handwriting of the opposite sex. We carried out a similar study, but with first-year psychology students.

The experimental hypothesis was, when participants are asked to judge the sex of the author of a sample of handwriting, they will perform more accurately when the authors were using their normal writing style compared with when they were trying to emulate the other sex. The independent variable was handwriting style with two levels: one level was the students' normal handwriting (before they knew the hypothesis) and the other level was their writing when trying to emulate the opposite sex. Participants then tried to categorise each of the two samples as either male or female. The design was, therefore, repeated measures. The order of presentation of the handwriting samples was counterbalanced across participants. The dependent variable was whether the participant's judgement of the writer's sex was correct or incorrect. Hypothetical data are available in the Appendix or from macmillanihe.com/harrisonspss-7e. Note that the writer's sex is not recorded in these data; we simply recorded, for each handwriting sample, whether their sex was judged correctly or not.

How to perform the McNemar test

There are two different ways to perform the McNemar test in SPSS, but we are going to describe the best of these which involves accessing the test from the <u>Crosstabs</u> command.

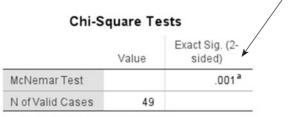
McNemar test and Crosstabs command

<u>Crosstabs</u> draws up contingency tables, and the McNemar test is an optional inferential statistic within this command. Follow steps 1 to 6 in the instructions on performing a multidimensional chi-square (Section 4). The Crosstabs: Statistics dialogue box will then appear (see below).

Crosstabs: Statistics	×	
Chi-square	Correlations	
Nominal Contingency coefficient Phi and Cramer's V	Ordinal <u>G</u> amma <u>S</u> omers' d Kendall's tau- <u>b</u>	
Uncertainty coefficient	Kendall's tau- <u>c</u>	
Nominal by Interval	 <u>K</u>appa Risk ✓ <u>McNemar</u> 	Select the <u>McNemar</u> option.
Cochran's and Mantel-Had	enszel statistics	
Test common odds ratio (equals: 1	Click on <u>Continue</u> then OK.
Continue Cancel	Help	

SPSS output for the McNemar test

Obtained using menu items: Descriptive Statistics > Crosstabs

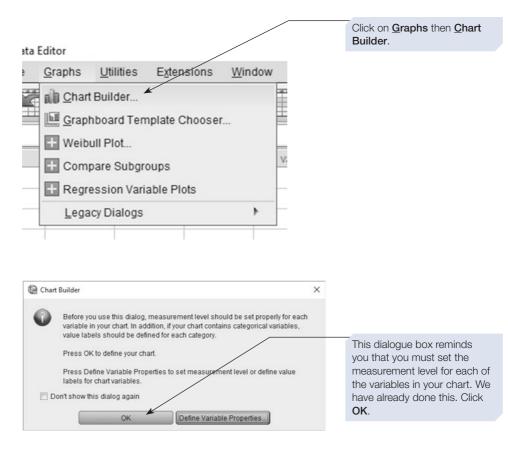


a. Binomial distribution used.

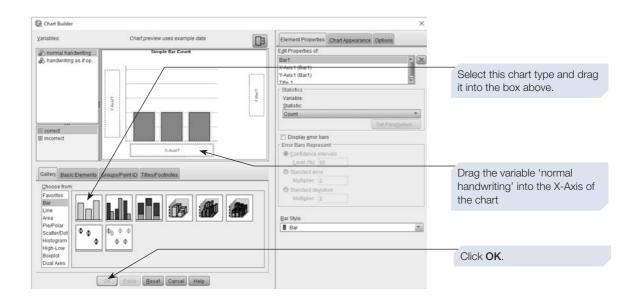
This is the output of the McNemar test. SPSS calculates the statistic using the binomial distribution, and gives the value of p and N only.

P It

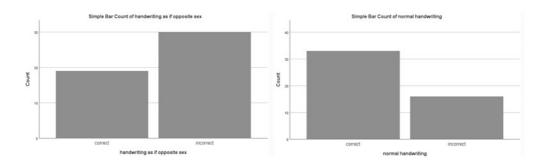
It would be useful to illustrate the results. For data from a related design, rather than using Clustered bar charts in the **Crosstabs** dialogue box, it is probably better to plot two separate bar charts using the **Graphs** menu.



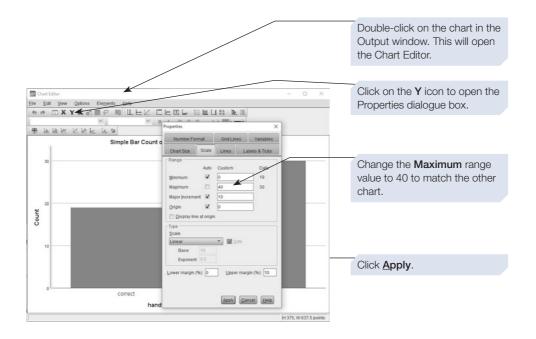
How to obtain a bar chart using Chart Builder



Now repeat this process, this time dragging the second variable 'handwriting as if opposite sex' into the X-Axis of the Chart Builder. You will now have two charts in your output. We have reproduced these alongside each other below.



There is a problem with these two charts – they are plotted using a different y-axis scale. We can easily correct this by editing the first one.



The two charts will now be plotted on the same *y*-axis scale and are ready to be used in our report.



The McNemar test using binomial distribution showed a significant difference, in the number of correct judgements, between the two conditions of handwriting style (N = 49, exact p = .001).

It would also be useful to explain the pattern of results in the following way:

Of the 49 participants, 33 (67%) correctly identified the handwriter's sex for the normal handwriting. Of those 33, 17 correctly identified the handwriter's sex for the 'opposite handwriting' and 16 incorrectly identified it. Of the 16 (33%) who were incorrect for the normal handwriting, 2 of them correctly identified the handwriter's sex for the 'opposite handwriting' and 14 of them incorrectly identified it. In brief, there were more incorrect responses when the writer had written as if they were of the opposite sex. This pattern of results is illustrated in Figure 7.1.

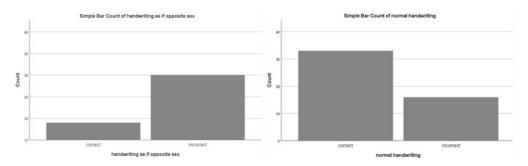


Figure 7.1 The pattern of correct and incorrect identification of handwriter's sex, when the writing was their normal handwriting and when they mimicked the writing of the opposite sex.

CHAPTER 7

Next, we show you how to create a bar chart using <u>Graphboard Template</u> Chooser.

How to obtain a bar chart using Graphboard Template Chooser

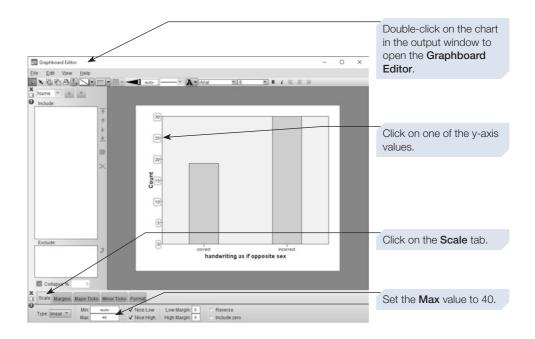
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Basic Detailed Titles Options Natural Name Upp normal handwriting handwriting as if opposite sex	Bar of Counts Choropieth of Counts	Select variable you want to plot.
Visualization of normal handwriting	Histogram Pie of Counts	Select the Bar of Counts icon.
Symmary: Count 💌		Click OK .
Manage	Easte Beset Cancel Help	

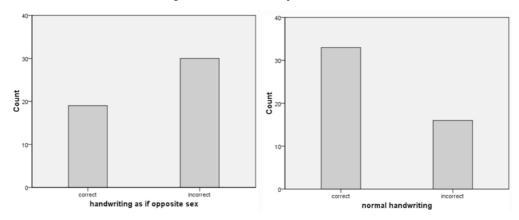
Repeat this process for the second variable.

Graphboard distinguishes between a 'bar of counts' suitable for nominal data, and 'bar' that will plot a summary descriptive for a variable of at least ordinal level of measurement.

The two charts will be in the output window. You can edit these as required by double-clicking. Double-click on the second chart and set the *y*-axis scale to a maximum value of 40.



The two charts are now plotted on the same y-axis scale.



Note about causation and the McNemar test

In related designs, you manipulate an independent variable and either collect data for both of its levels from the same participants (repeated measures) or collect data from matched participants (matched pairs). If you have carried out the normal controls for the design, then you can draw conclusions about causation from the McNemar test output.

Summary

- This chapter introduced you to nominal data and the tests that can be performed on nominal data.
- A contingency table is the best way of displaying nominal data.
- Remember our advice about descriptive statistics for nominal data; calculating the mean or median is inappropriate, as is any measure of dispersion.
- The chi-square test is a nonparametric test often used to analyse nominal data.
- The multidimensional chi-square nevertheless requires that the data satisfy certain criteria. The observations must be independent, so each participant should contribute only one data point.
- The McNemar test is used to analyse data obtained by measuring a dichotomous variable for related designs.
- For guidance on recoding values, see Chapter 4, Section 5.
- For guidance on incorporating SPSS output into a report, or on printing the output, see Chapter 14.

One-way analysis of variance

In this chapter

- An introduction to one-way analysis of variance (ANOVA)
- One-way between-subjects ANOVA, planned and unplanned comparisons, and nonparametic equivalent
- One-way within-subjects ANOVA, planned and unplanned comparisons, and nonparametric equivalent



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Section 1: AN INTRODUCTION TO ONE-WAY ANALYSIS OF VARIANCE (ANOVA)

- ANOVA is an enormously useful statistical procedure that is widely used to test for differences in a number of different experimental designs. One-way ANOVA is used when you have one independent variable (IV) with more than two groups or conditions; and Factorial ANOVA (which we will cover in Chapter 9) allows you to analyse data collected from designs with more than one IV. This chapter will focus on One-way ANOVA.
- One-way ANOVA is a parametric test that we can use to establish whether the means of our experimental conditions are different or not. In this way it can tell us whether our experimental manipulation has had an effect on the dependent variable (DV the thing we are measuring). The term 'one-way' simply refers to the fact that we use it to analyse data from experiments that have *one* IV. It may be easier to think about One-way ANOVA as an extension of the *t*-test, rather than something completely different.
- While t-tests are only able to test for differences between two groups or conditions, One-way ANOVA can allow us to compare participants' performance across multiple groups or conditions. However, while One-way ANOVA will tell us whether the scores significantly vary across our conditions (i.e. whether there has

been a significant effect of our IV on participants' scores), it won't tell us precisely which pairs of conditions are significantly different from one another (e.g. whether condition 1 is significantly different from condition 2, whether condition 2 is significantly different from condition 3, or whether condition 1 is significantly different from condition 3). Such comparisons require some additional statistical procedures called *planned* and *unplanned comparisons*.

When can we use One-way ANOVA?

As ANOVA is a parametric test, check that:

- 1. The dependent variable comprises data measured at interval or ratio level.
- 2. The data are drawn from a population that is normally distributed.
- 3. There is homogeneity of variance; that is, the samples being compared are drawn from populations that have similar variances.
- 4. In the case of independent groups designs, independent random samples must have been taken from each population.

It is not essential to have equal numbers of observations from each group or condition you are comparing across the independent variable.

How does it work?

We all know that humans vary in performance, both between individuals and within individuals over time. For these reasons, if we conduct a simple experiment comparing, say, the time it takes to learn a list of short words, medium length words and long words, we would not expect all the participants within a condition to take the same amount of time. We naturally accept that some participants will be faster than others (i.e. there will be variation between individuals). We also know that any one participant might take less or more time on one occasion than on other occasions (i.e. there will be variation within individuals). Remember that we can measure the amount of variation within a set of scores with measures of dispersion, such as the standard deviation or the *variance*.

Now let us imagine for a moment that we were robopsychologists; that is, we were interested in the psychology of robots (rather than robots interested in psychology). If we repeated our learning experiment with a group of R2D2 robots, we would expect all the robots in one condition to react in exactly the same way. Table 8.1 shows some hypothetical data for robots and for humans.

Let us first consider the data from the humans shown in this table. If we asked you to 'eyeball' the raw data and guess whether there was a difference in learning times for the three lists, you would probably have no problem saying that the difference did appear to be significant. In making this judgement, you are actually doing something quite sophisticated. What you are doing is deciding whether the natural variation between individuals within the conditions is large or small compared with the variation between individuals across the different conditions. That is, you are asking: 'OK, so not all the participants in the List A condition took the same time, and not all the participants in the List B or List C condition took the same time, but is this
 Table 8.1
 Time (in seconds) taken to learn three different lists of words for a group of robot and human participants

ROBOTS			
List A	List B	List C	
10	20	30	
10	20	30	
10	20	30	
10	20	30	
10	20	30	
10	20	30	
10	20	30	
10	20	30	
Mean = 10	Mean = 20	Mean = 30	Grand mean = 20

HUMANS			
List A	List B	List C	
30	54	68	
40	58	75	
35	45	80	
45	60	75	
38	52	85	
42	56	90	
36	65	75	
25	52	88	
Mean = 36.38	Mean = 55.25	Mean = 79.50	Grand mean = 57.04

natural variation (or noise) large or small compared with the difference in times between the three conditions?' In this case, participants within each condition might vary from each other by several seconds, but this is small compared with the larger differences between the times produced under the three different list conditions.

Let us look at the robots' data. Robots perform identically under identical conditions (or at least our robots do), so within each condition every robot has exactly the same learning time. Thus, the variance within each condition is zero. But if we compare the performance between the three conditions, it is clear that all the robots were fastest at learning the short words and all took longest to learn the long words. You might conclude that you want to switch from psychology to robopsychology, but there is also a more important point here. What we want to do is make our human participants' data more like the robots' data; that is, we want to reduce the variance within conditions down towards zero. In fact, all the practices of good experimental design, such as giving all participants the same instructions and testing under identical conditions, are designed to do just this - reduce the variance within each condition. This good experimental practice will reduce the variance but will not eliminate it – our participants will never behave exactly like the robots. So, if we cannot eliminate the variance, perhaps we can account for it. What we need is a statistical procedure that takes account of the variance within the conditions and compares this to the variance between conditions. If the variance between conditions is much larger than the variance within conditions, surely we can say that the IV is having a larger effect on the scores than the individual differences are. Clearly, for the robots, the variance within the conditions is zero, and the variance between the conditions is quite large. For our humans, the situation is not quite so clear-cut, but if we calculate the vari-

Variance between conditions > variance within conditions

ances, we will find the same basic pattern applies:

This concept of calculating the variance due to nuisance factors such as individual differences and comparing it to the variance due to our manipulation of the IV is central to ANOVA. Exactly how we calculate these variances can get rather complex for some designs, but this does not alter the basic principle that we simply want to ask whether or not the variance in the data brought about by our manipulation of the IV is larger than that brought about by the other nuisance factors such as individual differences. The variance brought about by these nuisance variables is usually referred to as the *error variance*, so we ask whether the error variance is less than the variance due to the manipulation of the IV.

A convenient way of expressing this is to calculate the ratio of the variance due to our manipulation of the IV and the error variance. This ratio is known as the *F*-ratio (named after Fisher). The *F*-ratio is:

F = Variance due to manipulation of *IV* / *Error* variance

If the error variance is small compared with the variance due to the IV, then the *F*-ratio will be a number greater than 1 (a large number divided by a smaller number always gives rise to a number greater than 1). If, on the other hand, the effect of the IV is small, and/or the error variance is large (perhaps because our participants varied considerably or because we did not adequately control the experiment), then the *F*-ratio will be a number less than 1 (a small number divided by larger number will always result in a number less than 1). Thus, we can now say that the effect of the IV is definitely not significant if the *F*-ratio is less than 1. This is because the error variance is actually larger than the variance caused by our manipulation of the IV.

So, the *F*-ratio is simply the ratio of these two estimates of variance. The larger the *F*-ratio, the greater the effect of the IV compared with the 'noise' (error variance) in

the data. An *F*-ratio equal to or less than 1 indicates a non-significant result, as it shows that the scores were equally affected or more affected by the nuisance variables (such as individual differences) as they were by the manipulation of the IV.

How do we find out if the F-ratio is significant?

Once we have calculated the value of the *F*-ratio and found it is larger than 1, we need to determine whether it is large enough to be regarded as significant. That is, we ask whether the effect of the IV is sufficiently larger than the effect of the nuisance variables to regard the result as significant. When calculating the *F*-ratio with a calculator, we consult *F* tables to discover, given the number of observations we made, what value *F* had to exceed to be considered as significant. When using SPSS to perform ANOVA, the output reports the exact *p* value for that particular *F*-ratio. This *p* value is the probability of getting this *F*-ratio by chance alone and it needs to be less than .05 for the *F*-ratio to be regarded as significant.

What about degrees of freedom?

You will remember from performing a *t*-test, another test of difference, that we need to calculate and report the degrees of freedom associated with our analysis. One complication with ANOVA is that, for each *F* value, we must report two sets of degrees of freedom. This is because we need to remember how many observations went into our calculation of the error variance and also how many went into our calculation of the variance due to the manipulation of the IV. As these are the bottom and top halves of the *F*-ratio equation, these are sometimes referred to as the 'denominator' and 'numerator' degrees of freedom, respectively. A good statistics text will explain the calculation of degrees of freedom in detail, but as SPSS calculates and reports these for you, all you need to know is to expect two values for each *F*-ratio. We will show you how to report these degrees of freedom and the *F*-ratio in subsequent sections.

What terms are used with One-way ANOVA?

Different textbooks tend to use slightly different terminologies to describe One-way ANOVA. To avoid the problems this can create, we are going to use what we consider to be the simplest terminology.

Factors

Factors are the independent variables, but as there may well be more than one of them per study, it makes sense to call them factors from now on.

Levels of factors

Levels of factors are similar to conditions. In the experiments we considered earlier, we had a single IV, which was manipulated to create two conditions. We would now describe this as a single factor with two levels. In ANOVA designs, a factor can have

as many levels as we like. For example, we might have a factor of 'Drug dose', which might be manipulated to create four levels of 0 mg, 10 mg, 20 mg and 30 mg.

Between-subjects factors

Between-subjects factors are factors whose levels vary between participants, so that each participant will experience only one level of a factor. For example, a participant can be administered either 0 mg, 10 mg, 20 mg or 30 mg. This is a factor that is manipulated using an independent groups design, which we will now refer to as a *between-subjects design*.

Within-subjects factors

Within-subjects factors are factors whose levels vary within a participant, so that each participant will experience two or more levels of a factor. For example, a participant might be administered all four different drug dosages. This is a factor that is manipulated using a repeated-measures design, which we will now refer to as a *within-subjects design*.

Main effect

The term 'main effect' is used to describe the effect a single independent variable has on a dependent variable. If our One-way ANOVA is significant, we can say that there was a significant main effect of our IV on our DV.

How is the F-ratio calculated?

You do not need to know how to calculate the *F*-ratio, as SPSS will do this for you. However, to fully appreciate the output that SPSS generates, it would be helpful to read this section and realise why the calculation is dependent on the type of factor manipulated. There are two different types of One-way ANOVA, and how they work depends on the type of experimental design you have.

Between-subjects One-way ANOVA design

Let us go back to our word list learning experiment (see Table 8.2).

This time, imagine that there are different humans taking part in each condition; that eight participants were asked to learn list A, another eight to learn list B and another eight to learn list C. There are two sources of variance of interest here:

- 1. How do the scores in one group vary from those in the other groups? We can look at how the mean of each column deviates from the grand mean. This provides us with a measure of the variance due to the factor.
- 2. How do the scores vary within each group? We can look at how each score within a column deviates from the mean for that condition. This provides us with a measure of noise.

Together, these two sources of variance must add up to the total variance (the variance between each single score and the grand mean). That is:

$$\operatorname{Var}_{(\operatorname{Total})} = \operatorname{Var}_{(\operatorname{Between Groups})} + \operatorname{Var}_{(\operatorname{Within Groups})}$$

HUMANS			
List A	List B	List C	
30	54	68	
40	58	75	
35	45	80	
45	60	75	
38	52	85	
42	56	90	
36	65	75	
25	52	88	
Mean = 36.38	Mean = 55.25	Mean = 79.50	Grand mean = 57.04

 Table 8.2
 Time (in seconds) taken to learn three different lists of words by three groups of human participants in a between-subjects design

Within-subjects One-way ANOVA design

Imagine that in our learning experiment, eight participants took part and each performed in each level of the factor. We would be able to calculate a mean score for each list and a mean score for each participant; see Table 8.3.

 Table 8.3
 Time (in seconds) taken to learn three different lists of words for the group of human participants in a within-subjects design

HUMANS				
Participant number	List A	List B	List C	Participant mean
1	35	42	64	47
2	48	60	90	66
3	36	65	75	58.67
4	40	55	70	55
5	38	52	85	58.33
6	25	42	58	41.67
7	30	42	60	44
8	42	60	90	64
	Mean = 36.75	Mean = 52.25	Mean = 74.0	Grand mean = 54.33

The calculation of F for the within-subjects design is more complicated than for the between-subjects design. Again, we want to determine the sources of variance. However, with this design, we have repeated observations of each participant as every person performs in every level of the factor. This allows us to distinguish between variation caused by individual differences and variation caused by different participants performing differently across the different conditions, and therefore separate out participant variance from error variance.

So, we have three sources of variance, and we can ask:

- 1. How do the scores in one condition vary from those in the other? We can compare overall differences between the three lists. As before, we can look at how the mean of each column deviates from the grand mean. This provides us with a measure of the variance due to our manipulation of the factor.
- 2. How do participants vary in their average scores? We can get an indication of how much individuals differ from each other by looking at how much each participant's average score deviates from the grand mean. This provides us with a measure of participant variance.
- 3. How much error variance is there? We can work this out by looking at the extent to which each score is not what we would predict from the row and column means. You can also think of this as the variance resulting from different participants responding differently to the change in the factor.

For example, with regard to the score for participant 1 in list A, we know that their mean time is 47 seconds. Participant 1 is, on average, 7.33 seconds faster compared with the overall grand mean of 54.33 seconds. The mean for the list A column is 36.75 seconds, so participants are, on average, 17.58 seconds faster at learning list A than the overall grand mean of 54.33 seconds. So, altogether, we would expect participant 1 to be 17.58 + 7.33 seconds faster than the grand mean of 54.33 seconds at learning list A, giving an expected time of 29.42 seconds. The observed score is 35 seconds, which is slower than we would expect. (Looking at participant 1's scores, we can see that they are relatively faster with lists B and C compared with list A.)

With regard to participant 2's score in list A condition, we know that their row mean is 66 seconds, which is 11.67 seconds slower than the grand mean of 54.33 seconds. So, we would expect participant 2 to be 17.58 seconds faster at learning list A, but 11.67 seconds slower because this participant is slower on average. Thus, we expect a time of 54.33 - 17.58 + 11.67 = 48.42 seconds. The observed score is 48 seconds – close to what we would expect.

The extent to which the observed scores vary from the expected scores reflects the extent to which participants are inconsistent and, as illustrated above, provides us with a measure of error variance.

Using SPSS to calculate the F-ratio

The calculation of the *F*-ratio for a within-subjects factor is tricky, and, as you will see, the SPSS output is rather more complex for a within-subjects factor compared with a between-subjects factor. Furthermore, because SPSS uses a procedure called the General Linear Model, it will give you much more information than just the *F*-ratio statistic. You will see both ANOVA and multiple regression statistics in the

SPSS output, as ANOVA can be considered to be a special case of multiple linear regression (see Chapter 10), which itself is a special case of the General Linear Model.

Effect size and ANOVA

We mentioned in Chapter 1 that it is usual to report effect size alongside the statement of whether the result is significant. This provides information about the magnitude of the finding and also may draw attention to the influence of sample size. For example, if our results show a marginal but non-significant result and a moderate or large effect size, power may be an issue, and it may be appropriate to consider a follow-up study with a larger sample. In Chapter 5, where we covered tests of differences for two samples, we introduced you to one measure of effect size, namely Cohen's d, which was calculated from the descriptive statistics provided by SPSS. You may remember that this involved calculating the size of the difference between the means relative to the standard deviation of the scores. The effect size estimates that are calculated for ANOVA are different, in that they tend to describe the proportion of the variability accounted for by each factor (or combination of factors) included in the ANOVA design (in that sense, they are similar to r^2 described in Chapter 6, which also is a measure of the proportion of variance accounted for). These effect size measures include eta squared, partial eta squared, generalized eta squared, associated omega squared and also correlational measures. Fritz, Morris and Richler (2012) explain that one can distinguish between:

- 1. Estimates that describe the effect size in the observed sample but do not consider the population from which the sample was drawn, and this is the case with eta squared and partial eta squared.
- 2. Estimates of effect size, such as omega squared measures, which try to estimate the variability in the sampled population rather than just in the observed sample, and these are therefore less likely to be inflated by chance factors.

We recommend that you consult your statistics text, or Fritz et al. (2012), for advice on how to calculate these different estimates of effect size. Field (2005) demonstrates how to calculate an omega squared measure (which involves a different equation depending on whether between-subjects or within-subjects ANOVA). Mulhern and Greer (2011) explain how to calculate eta squared and recommend against using partial eta squared, which is the only measure of effect size that SPSS will calculate for you. They explain that partial eta squared is not easy to interpret as it is an adjusted measure: the variance explained by one factor *after* taking into account the variance explained by the other factor(s). Fritz et al. (2012) suggest that partial eta squared is limited in terms of its usefulness and may only be helpful if making cross-study comparison with identical designs. Eta squared, on the other hand, is relatively straightforward to calculate by hand and will tell you the proportion of the total variability accounted for by each factor in your design. We will demonstrate how to do this in Section 2.

Whichever effect size measure you decide on, do report this for significant and non-significant effects, and identify which statistic is used.

Planned and unplanned comparisons

Imagine you have a design involving one factor with three conditions: A, B and C. If ANOVA reveals a significant effect of this factor, this suggests a difference between the conditions. However, to find out exactly where this difference is, we need to carry out further tests that compare the pairs of conditions. There are three possibilities:

- 1. All three possible comparisons are significant, so conditions A and B are significantly different from each other, as are conditions B and C, and conditions A and C.
- 2. Only two of the three comparisons are significant, for example only conditions A and C and conditions B and C differ significantly.
- 3. Only one of the three comparisons is significant, for example only conditions A and C are significantly different.

There are two types of comparisons that we can conduct:

- 1. Planned (*a priori*) comparisons. These are decided on before the data were collected. The researcher has predicted which means will differ significantly from each other.
- 2. Unplanned (*a posteriori* or *post-hoc*) comparisons. Here, differences among means are explored after the data have been collected.

Why should this matter? We need to use different tests for these two kinds of comparison because the probability of a Type I error is smaller when the comparisons are planned in advance. Type I error involves incorrectly rejecting a null hypothesis, thus concluding that there is a real effect, when in fact, the means differ due to chance. When making multiple comparisons, there is an increased risk of Type I errors. Howell (2013) gives the following example: assume that we give a group of males and a group of females 50 words and ask them to give us as many associations to these words as possible in one minute. For each word, we then test whether there is a significant difference in the number of associations given by male and female participants. We could run 50 more or less independent t tests, but we would run the risk that 2.5 of these (50*.05) will be declared 'significant' by chance.

Why is there a greater risk of making a Type I error when carrying out unplanned comparisons? Consider the following: Imagine an experiment to look at the effect of five different levels of noise on memory that employed a One-way ANOVA design. You will have five means (one for each condition) and could do a total of ten comparisons (you could compare mean 1 to mean 2; mean 1 to mean 3; mean 1 to mean 4 etc.). Assume that the null hypothesis is true, and that noise does not affect memory, but that, by chance, two of the means are far enough apart to lead us erroneously to reject the null hypothesis, thus giving rise to a Type I error. If you had planned your single comparison out of 10 that would lead us to make a Type I error. But if you look at the data before deciding which comparisons to make, you are certain to make a Type I error, since you are likely to test the largest difference you can observe.

As you will see, it is possible to ask SPSS to carry out planned and unplanned comparisons. Unplanned or *post-hoc* comparisons are easy to perform using SPSS. However, there is no need to perform these if the factor has only two levels (the

main effect is sufficient) or if the main effect is not significant. Planned comparisons, on the other hand, are less easy to perform. We demonstrate how to perform both types of comparisons in Sections 2 and 3.

Section 2: ONE-WAY BETWEEN-SUBJECTS ANOVA, PLANNED AND UNPLANNED COMPARISONS, AND NONPARAMETRIC EQUIVALENT

Example study: the effects of witness masking

To practise the use of the One-way between-subjects ANOVA, we shall consider an applied experiment which looked at the effects of masking the face of a witness. There is growing awareness that the identity of witnesses in sensitive cases should be protected, especially in light of the move towards televising live court cases. The technology to mask a witness's face is available and has been used in America. Towell, Kemp and Pike (1996) reported the results of a study investigating the effect that masking might have on jurors' memory for a witness's testimony and on jurors' perceptions of the witness's credibility. The testimony of an alleged victim of rape presented in a televised trial in America was shown to participants.

The design employed was a One-way between-subjects ANOVA design. The between-subjects factor, presentation condition, had four levels: unmasked, grey blob, pixelation and negation. These were operationalised by showing some participants the witness unmasked, so that her face was fully visible, some with her face masked by a grey blob, some with her face masked by pixelation and some with her face negated (white and black were reversed). One of the dependent variables was the percentage of facts from the testimony correctly remembered by the participants. The hypothesis was that there would be a negative effect of masking on memory. Results revealed that participants' memory for the victim's testimony was affected by presentation condition, while negating the face did not lower memory compared with the unmasked condition, masking with a grey blob and pixelation both impaired memory. For the purposes of this book, we have created a data file that will reproduce some of these findings. (These data are available from macmillanihe.com/harrison-spss-7e.)

SPSS provides two ways of carrying out a One-way between-subjects ANOVA, one using the **General Linear Model** command and one using the **One-way ANOVA** command. The first command can also be used to perform a multi-between-subjects ANOVA, as you will see in Section 3, and also has the option of including in the output the measure of effect size called partial eta squared. In Section 1, we indicated that this particular measure is not that useful; however, when the design involves just one factor, there is no difference between partial eta squared and eta squared. The second command will only permit analysis of a One-way ANOVA design, but has the advantage of a much simpler output and of providing alternative *Fs* should the assumption of homogeneity of variance be violated. Both ways allow you to do planned and unplanned comparisons to evaluate the differences between pairs of group means. See Section 1 for general guidance on these comparisons.

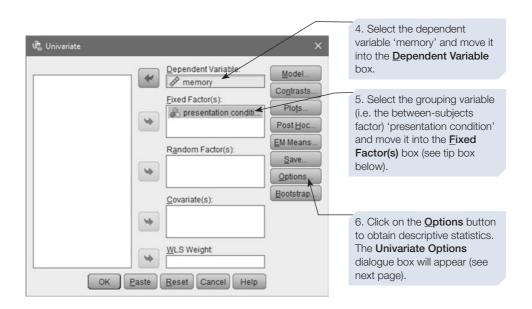
1. Click on Analyze.

CHAPTER 8

We will first describe how to use the **General Linear Model** command, followed by the **One-way ANOVA** command. We will then demonstrate how to carry out planned and unplanned comparisons and finish this section by describing a nonparametric equivalent test.

How to do it: using General Linear Model command

Eile E	dit <u>V</u> iew <u>D</u> ata	a <u>T</u> ransform	Analyze Graphs Utilities	Extensions	Window	Help		
			Re <u>p</u> orts D <u>e</u> scriptive Statistics <u>B</u> ayesian Statistics	<i>[</i> •••	2. Cl Mod		<u>G</u> eneral Li	inear
	& prescond	d 🖉 memory	Ta <u>b</u> les	/ >	ar	var	var	va
1		1 68.00	Compare Means					
2		1 75.00	General Linear Model	*	Univari	ate		
3		1 65.00	Generalized Linear Model	s 🕨	Multiva			
4		1 69.00	Mixed Models	,				
5		1 70.00	Correlate		Repeat	ed Measi	ures	
6		1 72.00	Regression		Variand	e Compo	onents	
7		65.00						
8		1 66.00	Loglinear					
9		1 58.00	Neural Networks	,				
10		1 59.00	Classify	,	3. Cl	ick on I	<u>U</u> nivariate	. The
11		2 56.00	Dimension Reduction	•	Univ	ariate	dialogue b	ox wil
12		2 58.00	Sc <u>a</u> le	*	appe	ar (see	below).	
14			Nonparametric Tests					



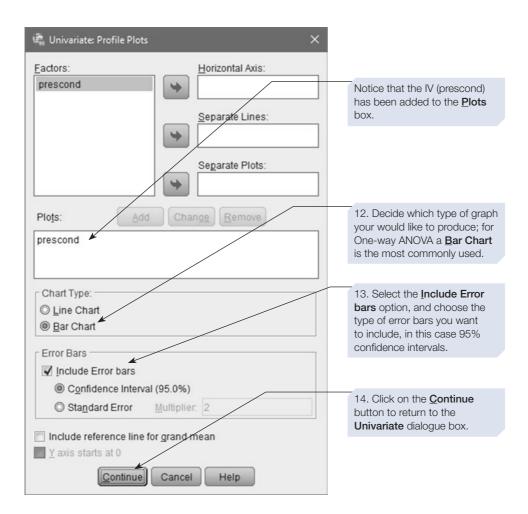


With a fixed factor, data have been collected from all the levels of the factor that are of interest to the researcher. An alternative is to choose the levels by a random procedure, but this rarely happens in psychological research.

🛱 Univariate: Options	×	7. Click here to obtain mean and standard deviation for each level of the factor.
Display Descriptive statistics	Homogeneity tests	
Observed power Parameter estimates Contrast coefficient matrix	Residual plot Lack of fit General estimable function	8. Click here to check whether the assumption of equality of variance has been violated.
Heteroskedasticity Tests Modified Breusch-Pagan test	F test	
Breusch-Pagan test Model. Parameter estimates with roby	White's test	Click here if you want to obtain partial eta squared, a measure of effect size (see guidance in Section 1).
© HC <u>0</u> © HC <u>1</u> © HC2		
© HC <u>3</u> © HC <u>4</u>		9. Click on the Continue button to return to the Univariate dialogue box.
	lence intervals are 95.0 % ncel Help	

It is often a good idea to ask SPSS to illustrate your data, so you have a graph you can use when it comes to writing up your results. SPSS allows you to do this directly from the main **Univariate** dialogue box. To produce a graph, select the **Plots**... option to open up the **Univariate**: **Profile Plots** dialogue box.

Univariate: Profile Plots Eactors: prescond	X Horizontal Axis:	10. Select the independent variable 'prescond' and move it into the <u>Horizontal Axis</u> box.
	Separate Plots:	
Plots: Add Cha	nge Remove	11. Click <u>A</u>dd to tell SPSS that you want to produce a graph for this variable.



Finally, click on the \bigcirc button, and SPSS will calculate the test for you. See below for an example of the output using the Univariate command via General Linear Model, which includes the means, standard deviations and N (number of scores) obtained by clicking on <u>Descriptive statistics</u> in the Univariate: Options dialogue box.

SPSS output for One-way between-subjects ANOVA

Obtained using menu items: General Linear Model > Univariate

Between-Sub	jects Factors
-------------	---------------

		Value Label	N
presentation condition	1	unmasked	10
	2	greyblob	10
	3	pixelated	10
	4	negated	10

Descriptive Statistics

SPSS reminds you of the factor you are analysing, what the levels of that factor are, and the number of participants in each level.

This table will appear if you requested <u>Descriptive</u> statistics in the Univariate: Options dialogue box.

Dependent Variable: memory

presentation condition	Mean	Std. Deviation	Ν	The table gives the mean and
unmasked	66.7000	5.33437	10	the standard deviation (SD) for each level of the factor.
greyblob	55.7000	3.80205	10	each level of the lactor.
pixelated	57.7000	5.41705	10	The bottom row gives the
negated	67.2000	4.58984	10	total mean and SD; that is, for
Total	61.8250	7.00142	40	all participants regardless of which condition they were in.

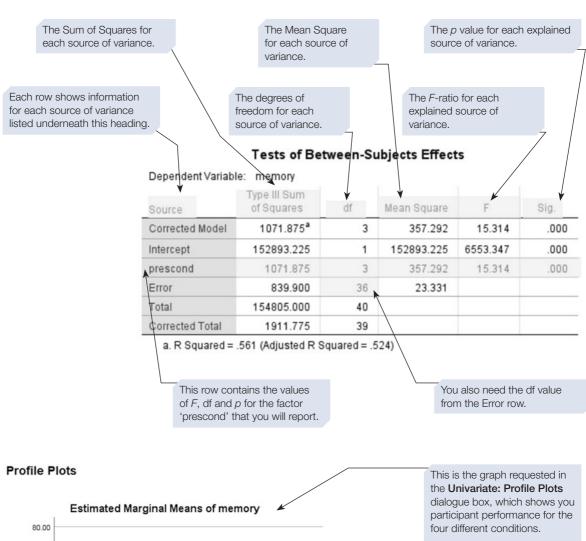
Levene's Test of Equality of Error Variances^{a,b}

		Levene Statistic	df1	df2	Sig.
memory	Based on Mean	.490	3	36	.692
	Based on Median	.499	3	36	.686
	Based on Median and with adjusted df	.499	3	34.786	.686
	Based on trimmed mean	.502	3	36	.683

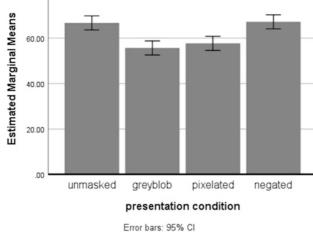
Tests the null hypothesis that the error variance of the dependent variable is equal across groups.

SPSS tests the assumption of equality of variances in a number of different ways. In this case, we are interested in the statistic **Based on Mean**.

If Levene's test is significant, this suggests that the group variances are likely to be significantly different, suggesting that the assumption of homogeneity of variance has been violated. Here, it is not significant, so the assumption has been met.



When error bars do not overlap, it is likely that the conditions are significantly different – although you need to do post-hoc tests or planned comparisons to confirm this.



Calculating eta squared: one measure of effect size

In Section 1 we provided guidance on reporting effect size when carrying out ANOVAs. Although SPSS will calculate partial eta squared, this is not easy to interpret as it is an adjusted measure: the variance explained by one factor *after* taking into account the variance explained by the other factor(s). Instead, Mulhern and Greer (2011) recommend eta squared, which is relatively straightforward to calculate by hand and will tell you the proportion of the total variability accounted for by each factor and interaction in your design. We will demonstrate how to do this now. Note, however, that there are alternative measures which may be preferable, as they attempt to estimate the effect size in the population rather than just in the sample (see Fritz, Morris and Richler, 2012 for a review and guidance on how to calculate these, and/or consult your statistics text). Eta squared (η^2) can be calculated from the ANOVA output. It is the sum of squares for your IV (prescond) divided by the corrected total sum of squares, and these values can be found in the **Tests of Between-Subjects Effects** table.

	Tests of Bet	ween-Si	ubjects Effect	ts	
Dependent Variab	le: memory				
Source	Type III Sum of Squares	df	Mean Square	F	Sig.
Corrected Model	1071.875 ^a	3	357.292	15.314	.000
Intercept	152893.225	1	152893.225	6553.347	.000
prescond	1071.875	3	357.292	15.314	.000
Error	839.900	36	23.331		
Total	154805.000	40			
Corrected Total	1911.775	39			

a. R Squared = .561 (Adjusted R Squared = .524)

For 'prescond' $(\eta^2) = 1071.875 / 1911.775 = .56$

This suggests the factor 'prescond' accounts for approximately 56% of the variance in the dependent variable.

As mentioned above, SPSS can calculate a different effect size measure for you: partial eta squared. This isn't the most helpful measure (see Section 1 for guidance), so we generally recommend using eta squared instead. However, when you have a One-way ANOVA there is actually no difference between partial eta squared and eta squared. This is because partial eta squared calculates the proportion of variance a factor (or variable) explains that is not explained by other variables in the analysis. When there are no other factors in the analysis (as is the case with One-way ANOVA), the calculation is the same as eta squared. As such, when you have a One-way ANOVA, rather than calculating eta squared by hand, you can just use the SPSS 'Estimate Effect Size' option in the Options dialogue box. Try doing this now, by downloading the dataset used in this example, and see if you get the same number as the one we just calculated by hand.



Reporting the results

In a report you might write:

A One-way between-subjects ANOVA was conducted to examine the effect of presentation condition on participants' recall of the witness testimony. This revealed a significant effect of presentation condition: F(3,36) = 15.31, p < .001, $\eta^2 = .56$.

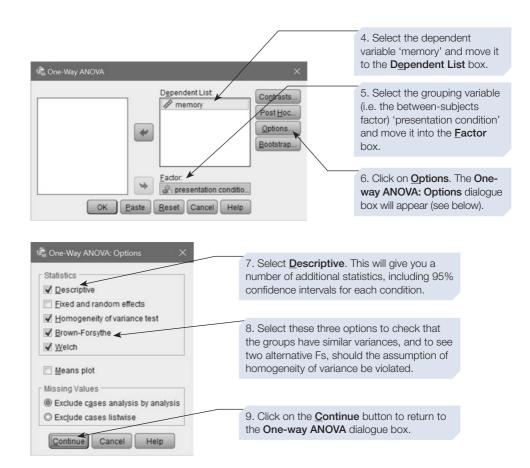


You would also want to report your descriptive statistics (i.e. means and standard deviations) and provide information regarding the confidence intervals for each condition. To identify which pair(s) of conditions significantly differed, you would carry out planned or unplanned comparisons as appropriate, and these are demonstrated later in this section.

How to do it: using One-way ANOVA command

As stated earlier, One-way between-subjects ANOVA can be carried out in two different ways in SPSS. This is the second way, which provides a simpler output and also offers alternative *F*s should the variances in the groups not be equal.

					/		1. Click	k on <u>A</u> naly	/ze.	
u n 1-	Way Be	tween-Su	bjects A	nova.sav [Datas	Set4] - IBM SPSS Statistics Data E	ditor			- 🗆	×
Eile	Edit	⊻iew	<u>D</u> ata	Transform	Analyze Graphs Utilities	Extensions	Window Help			
					Re <u>p</u> orts D <u>e</u> scriptive Statistics <u>B</u> ayesian Statistics		2. Click		pare Mea isible: 2 of 2 V	_
		& pres	scond	& memory	Ta <u>b</u> les	,	ar var	var	var	
	1	1	1	68.00	Compare Means	4	Means			*
1	2		1	75.00	General Linear Model	•	One-Sample T	Test		
:	3		1	65.00	Generalized Linear Mode	s >				
4	4		1	69.00	Mixed Models	,	Independent-Sa			
1	5		1	70.00	Correlate		Paired-Samples	s T Test		
. (6		1	72.00	Regression		One-Way ANOV	A		
1	7		1	65.00	Loglinear		1			
1	В		1	66.00	Neural Networks					
	9		1	58.00	_	ľ.	3. Click	on One-	way ANO	VA.
1	0		1	59.00	Classify				NOVA dial	
1	1		2	56.00	Dimension Reduction				see below)	0
1	2		2	58.00	Sc <u>a</u> le	,	Box Wi			
1	3	1	2	59.00	Nonparametric Tests	*				



Finally, click on the **out** button. See below for annotated output.

SPSS output for One-way between-subjects ANOVA

Obtained using menu items: Compare Means > One-way ANOVA

Oneway

This table is produced by selecting <u>Descriptive</u> in the One-way ANOVA: Options dialogue box.

Descriptives

					95% Confidence Interval for Mean				
	Ν	Mean	Std. Deviation	Std. Error	Lower Bound	Upper Bound	Minimum	Maximum	
unmasked	10	66.7000	5.33437	1.68688	62.8840	70.5160	58.00	75.00	
greyblob	10	55.7000	3.80205	1.20231	52.9802	58.4198	48.00	61.00	
pixelated	10	57.7000	5.41705	1.71302	53.8249	61.5751	51.00	68.00	
negated	10	67.2000	4.58984	1.45144	63.9166	70.4834	58.00	74.00	
Total	40	61.8250	7.00142	1.10702	59.5858	64.0642	48.00	75.00	

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		Levene Statistic	df1	df2	Sig.
memory	Based on Mean	.490	3	36	.692
	Based on Median	.499	3	36	.686
	Based on Median and with adjusted df	.499	3	34.786	.686
	Based on trimmed mean	.502	3	36	.683

This table is produced by selecting Homogeneity of variance test in the Oneway ANOVA: Options dialogue box. If Levene's test is significant, this suggests that the group variances are significantly different, suggesting that the assumption of homogeneity of variance has been violated. Here, it is not significant, so the assumption has been met.

This table shows the outcome of the analysis of variance. Each row shows information for a source of variance.



memory Sum of F Squares df Mean Square Sig. 3 15.314 Between Groups 1071.875 357.292 .000. Within Groups 839.900 36 23.331 Total 1911.775 39 You also need the df value from the Within Groups row This row contains the values of *F*, df and *p* for the factor (the Error variance row). 'prescond' that you will report.

Robust Tests of Equality of Means

memory

	Statistic ^a	df1	df2	Sig.
Welch	16.488	3	19.797	.000
Brown-Forsythe	15.314	3	33.733	.000

a. Asymptotically F distributed.

This table is produced by selecting **Brown-Forsythe** and <u>Welch</u> in the One-way ANOVA: Options dialogue box. These are alternative *F*s, which can be used should Levene's test be significant and hence the assumption of homogeneity of variance violated.

Calculating eta squared: one measure of effect size

As above, effect size can be calculated using the main ANOVA table. To calculate eta squared (η^2), again you need to take the sum of squares for your IV (in the 'Between Groups' row) and divide it by the corrected total sum of squares (in the 'Total' row), For 'prescond' (η^2) = 1071.875 / 1911.775 = .56

Reporting the results

In a report you might write:

A One-way between-subjects ANOVA was conducted to examine the effect of presentation condition on participants' recall of the witness testimony. This revealed a significant effect of presentation condition: F(3,36) = 15.31, p < .001, $\eta^2 = .56$.



You would also want to report in your results your descriptive statistics (i.e. means and standard deviations) and information regarding the confidence intervals for each condition. As graphs cannot be produced from the ANOVA dialogue boxes using this method, you may want to refer to Chapter 5, Section 3 for guidance on how to obtain an error bar chart. To identify which pair(s) of conditions significantly differed, you would carry out planned or unplanned comparisons as appropriate, and these are demonstrated next.

Planned and unplanned comparisons

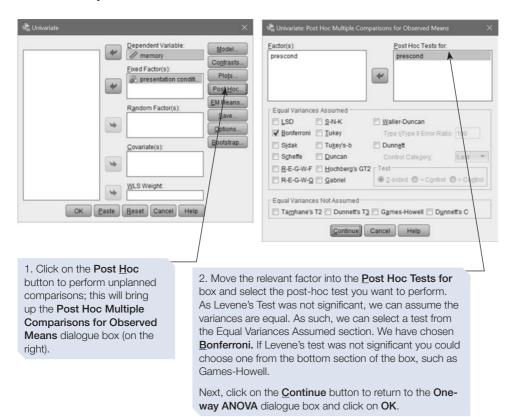
The One-way between-subjects ANOVA identified a significant effect of presentation condition, but we need to carry out further analysis to compare the pairs of conditions to pinpoint the source of this effect. In Section 1 we explained that there are two main types of comparisons: planned and unplanned. For demonstration purposes, we carry out both of these on the same data set; however, normally you would only perform one type.

Unplanned (post-hoc) comparisons in SPSS

There are a range of post-hoc tests to choose from, and you will need to consult your statistics text to select the one most suitable for your data. Which test you choose will depend on whether the assumption of homogeneity of variance has been met or not, and whether you have equal sample sizes or not (Field, 2013). We demonstrate how to select these tests for each of the two ways of conducting a One-way between-subjects ANOVA. Either way, the output is the same and is shown below.

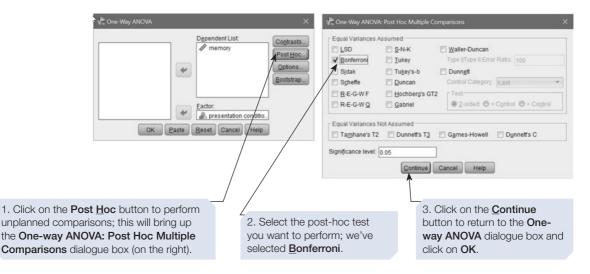
Using General Linear Model command

Click on <u>Analyze</u> \Rightarrow <u>General Linear Model</u> \Rightarrow <u>Univariate</u>.



Using One-way ANOVA command

Click on <u>A</u>nalyze \Rightarrow Compare <u>M</u>eans \Rightarrow <u>O</u>ne-way ANOVA.



SPSS output for post-hoc tests

POST HOC TESTS 🔨

This heading and the Multiple Comparisons table will appear after all the other tables in the ANOVA output.

Multiple Comparisons

Dependent Variable: memory Bonferroni

		Mean Difference (1			95% Confidence Interval	
(I) presentation condition	(J) presentation condition	Difference (I- J)	Std. Error	Sig.	Lower Bound	Upper Bound
unmasked	greyblob	11.00000	2.16012	.000	4.9690	17.0310
	pixelated	9.00000	2.16012	.001	2.9690	15.0310
	negated	50000	2.16012	1.000	-6.5310	5.5310
greyblob	unmasked	-11.00000	2.16012	.000	-17.0310	-4.9690
	pixelated	-2.00000	2.16012	1.000	-8.0310	4.0310
	negated	-11.50000	2.16012	.000	-17.5310	-5.4690
pixelated	unmasked	-9.00000	2.16012	.001	-15.0310	-2.9690
	greyblob	2.00000	2.16012	1.000	-4.0310	8.0310
	negated	-9.50000	2.16012	.001	-15.5310	-3.4690
negated	unmasked	.50000	2.16012	1.000	-5.5310	6.5310
*	greyblob	11.50000	2.16012	.000	5.4690	17.5310
	pixelated	9.50000	2.16012	.001	3.4690	15.5310

*. The mean difference is significant at the 0.05 level.

SPSS prints a complete matrix (as it does for correlations). You have to pick out the comparisons required, and ignore the repetitions.

As our factor had four levels, there are six possible comparisons. Their p values are highlighted in this column.



In a report you might write:

Employing the Bonferroni post-hoc test, significant differences were found between the unmasked and greyblob conditions (p < .001), between the unmasked and pixelated conditions (p = .001), between the greyblob and negated conditions (p < .001) and between the pixelated and negated conditions (p = .001). There was no significant difference between the unmasked and negated conditions (p = 1) or between the greyblob and pixelated conditions (p = 1).

Or, to abbreviate:

There was no significant difference between the unmasked and negated conditions or between the greyblob and pixelated conditions (for both, p = 1). The greyblob and pixelated conditions were each significantly different from each of the unmasked and negated conditions ($p \le .001$).

Planned comparisons in SPSS

Generally, for planned comparisons, the technique of linear contrasts is used, which allows us to compare one level or set of levels with another level or set of levels. The simplest way of doing this is to assign weights to each of them. These weights are known as 'coefficients'. This technique is available on SPSS, which uses the *t* statistic to test specific contrasts. Indeed, the printout will give you two t values, one for 'assume equal variances' and one for 'does not assume equal variances'. Since the variances of the groups being compared should be broadly similar (otherwise you should not be using ANOVA), you can 'assume equal variances', but check both values and their significance. A point to note here is that the overall main effect does not have to be significant for you to test for specific differences using planned comparisons.

By assigning weights (or coefficients), we can make three sorts of comparison:

- 1. We can compare one condition with one other condition.
- 2. We can compare one condition with the mean of two or more other conditions.
- 3. We can compare the mean of one set of conditions with the mean of another set of conditions.

In all three cases, we assign a weight of zero to a condition (or conditions) that we do not want to be included in the comparison. Conditions (or groups of conditions) that are to be compared with each other are assigned opposite signs (positive or negative). In all cases, the sum of the weights must be zero.

So, suppose you had four conditions, C1, C2, C3 and C4:

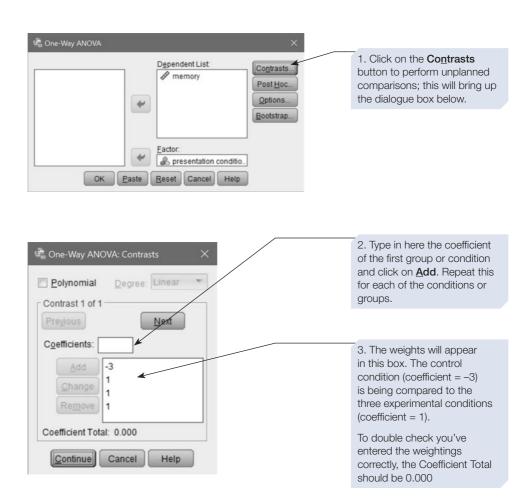
- If you wanted to compare only conditions 1 and 3, you could assign the weights: 1, 0, -1, 0.
- If you wanted to compare the average of the first two conditions with the third condition, you could assign the weights: 1, 1, -2, 0.
- If you wanted to compare the mean of the first two groups with the mean of the last two groups, you could use the weights: 1, 1, -1, -1.

If you wish to perform more than one planned comparison on the same data set, you need to check that the comparisons are independent of one another, that they are non-overlapping - these are called orthogonal comparisons. You can do this by taking each pair of comparisons and checking that the products of the coefficients assigned to each level sum to zero (see any good statistics text, for example Howell, 2013).

If you perform a planned comparison using the One-way ANOVA command, then you can design your own contrasts (as shown above), and enter the weights (coefficients) into a dialogue box. We show you how to do this next.

Using One-way ANOVA command

Click on <u>Analyze</u> \Rightarrow Compare <u>Means</u> \Rightarrow <u>One-way</u> ANOVA.



The dialogue box above shows a planned comparison, where the control group (who were shown the witness giving evidence with her face visible) is compared with the three experimental groups (who were all shown the witness giving evidence with her face masked). A linear contrast is requested, and the coefficients have been entered, first for group 1, then for groups 2, 3 and 4. The output below shows that this comparison is significant.

CHAPTER 8

SPSS output for contrasts

Obtained using menu items: Compare Means > One-way ANOVA

		ANOVA			
memory					
	Sum of Squares	df	Mean Square	F	Sig.
Between Groups	1071.875	3	357.292	15.314	.000
Within Groups	839.900	36	23.331		
Total	1911.775	39			

Contrast Coefficients

	presentation condition						
Contrast	unmasked	greyblob	pixelated	negated			
1	-3	1	1	1			

Contrast Tests Value of Contrast Contrast Std. Error t df Sig. (2-tailed) Assume equal variances -19.5000 5.29119 -3.685 36 .001 Does not assume equal -19.5000 5.66539 -3.442 13.818 .004 1 variances This row contains the values of t, df and p for the contrast you requested, and assuming equal variances.



In a report you might write:

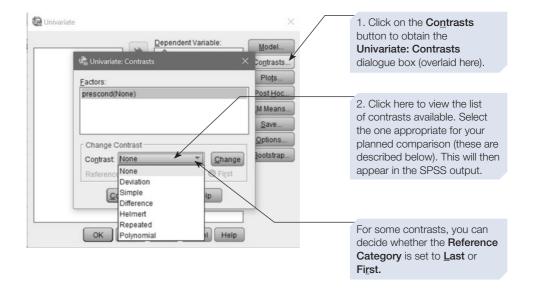
A planned comparison revealed that participants who saw the witness's face unmasked remembered significantly more of her testimony than the participants in the three masking conditions (t = 3.69, df = 36, $\rho = .001$).

Note that the contrast test can tell you whether the conditions you compared are significantly different or not, but nothing about the *direction* of the difference(s). In order to fully interpret the result, you will need to inspect the descriptive statistics for the conditions or groups being compared.

Using General Linear Model command

If you use the General Linear Model method to perform the One-way ANOVA, you have the option of choosing between a range of preset contrasts.

Click on <u>Analyze</u> \Rightarrow <u>General Linear Model</u> \Rightarrow <u>Univariate</u>.



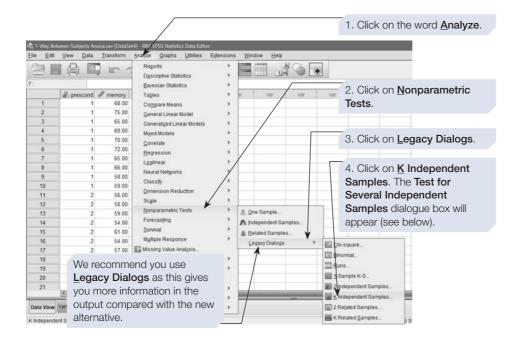
- 1. A **Deviation** contrast compares the effect for each level of the factor, except the reference category, with the overall effect. Thus, if there are three levels of the factor, two comparisons will be carried out. If the reference category is set to last, then levels 1 and 2 will be compared with the overall effect (of levels 1, 2 and 3 combined). If you select the reference category to be the first, then levels 2 and 3 will be compared with the overall effect of all three.
- 2. In a **Simple** contrast, each level of the factor is compared with the reference level, which can be set to either the first or last level. For example, if there are three levels and the reference category is set to the first level, then the second and the third levels will each be compared with the first level.
- 3. In a **Difference** contrast, each level of the factor except the first is compared with the mean effect of all previous levels. Thus, if there are three levels, level 3 is compared with the combined effects of level 2 and level 1, and level 2 is compared with level 1.
- 4. A **Helmert** contrast is the reverse of the **Difference** contrast, and the effect of each level of the factor except the last is compared with the mean effect of subsequent levels. If there are three levels, level 1 is compared with levels 2 and 3 combined, and level 2 is compared with level 3.
- 5. In a **Repeated** contrast, each level is compared with the previous level, so if there are three levels, level 2 is compared with level 1, and level 3 with level 2.
- 6. The **Polynomial** contrast should be used to look for trends in the data; for example, it can be used to look for a linear trend.

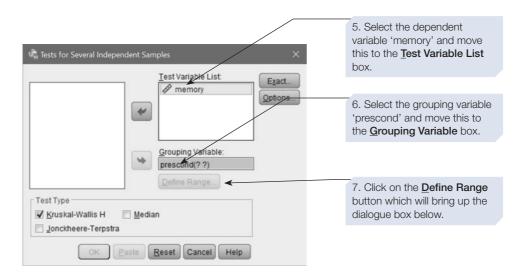
CHAPTER 8

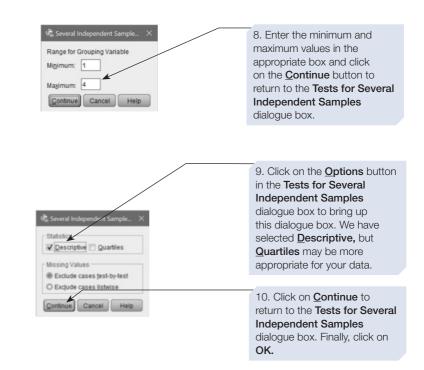
The Kruskal–Wallis test

To end this section, we introduce you to the nonparametric equivalent of the One-way between-subjects ANOVA, which can be used if your data fail to meet the assumptions for the ANOVA. For demonstration purposes, we perform the Kruskal–Wallis test on the same data used throughout this chapter.

How to do it

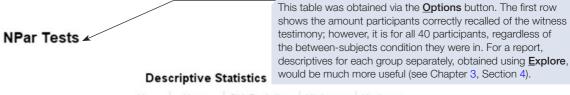






SPSS output for Kruskal–Wallis test

Obtained by using menu items: Nonparametric Tests > K Independent Samples



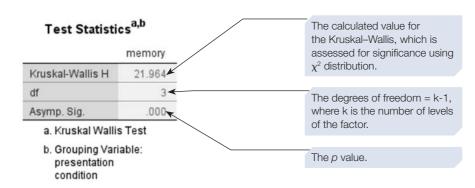
	Ν	Mean	Std. Deviation	Minimum	Maximum
memory	40	61.8250	7.00142	48.00	75.00
presentation condition	40	2.50	1.132	1	4

Kruskal-Wallis Test +

	Ranks		
	presentation condition	Ν	Mean Rank
memory	unmasked	10	28.70
	greyblob	10	10.40
	pixelated	10	13.45
	negated	10	29.45
	Total	40	

This table provides information about the calculations for the Kruskal–Wallis, which can be considered as an extension of the Mann–Whitney *U* test. Therefore, you might like to refer to the annotated output of that test for an explanation (Chapter 5, Section 5).







Reporting the results

In a report you might write:

A Kruskal–Wallis test was conducted to examine the effect of presentation condition on participants' recall of the witness testimony. This revealed a significant effect of presentation condition: $x^2(3, N = 40) = 21.96, p < .001$.

Section 3: ONE-WAY WITHIN-SUBJECTS ANOVA, PLANNED AND UNPLANNED COMPARISONS AND NONPARAMETRIC EQUIVALENT

Example study: the Stroop effect

Many experiments have been conducted to investigate the Stroop effect (Stroop, 1935). The most common way of demonstrating this effect is to show participants the names of colours printed in an incongruous colour (e.g. the word 'red' written in green ink) and ask them to name the colour of the ink. Results show that this is not an easy task because of our tendency to read the word, which then interferes with the task of naming the colour of the ink. In one experiment conducted with undergraduate students, we devised three lists: One list was incongruent and contained four words with strong colour associations (grass, coal, blood and sky) repeated three times in a random order, each time in a different incongruent colour ink (e.g. 'grass' printed in black, red and blue ink). The second list was congruent and contained the same four words repeated three times in a random order, each time in their congruent colour ink (e.g. 'grass' printed in green ink). The third list was neutral and contained four new words, matched in word length to the original words, and repeated three times. These words were not associated with any particular colour and were printed in one of the four different colour inks (e.g. 'table' written in green). These three lists constituted the different experimental conditions, and all participants completed all three lists. The order of the lists was counterbalanced across participants.

The design employed was a One-way within-subjects ANOVA design. The withinsubjects factor, the type of list, had three levels: incongruent, congruent and neutral. The dependent variable was the total time taken in seconds to name the colour of the ink of the 12 words in the list. The hypothesis was that there would be an effect of type of list on performance, with the shortest naming time for the congruent list and the longest naming time for the incongruent list. (These data are available from mac-millanihe.com/harrison-spss-7e.)

Understanding the output

The within-subjects ANOVA output contains several additional sections not seen previously in the between-subjects output. We describe these here and draw attention to them again when working through the output:

- 1. Mauchly's test of sphericity: this is important. If you have two or more levels of a within-subjects factor, SPSS will produce a test called the Mauchly's test of sphericity. The Mauchly's test of sphericity is a statistical test to determine whether the data entered into the within-subjects ANOVA meet certain assumptions, rather like the Levene's equality of variance test. The assumption of sphericity is that the variances of the differences between all possible pairs of levels of the factor are equal. Because the same participants are performing in each level of the within-subjects factor, you would assume that the correlations between all possible combinations of levels are roughly the same. If there are only two levels, there will be only one correlation, so this test is only valuable when there are three or more levels. A chi value is estimated to test the significance of the Mauchly's test of sphericity procedure (hence the output reports 'Approx. Chi-square'). The significance of this value of chi is reported. If it is significant (i.e. less than .05), then the assumption of sphericity has been violated. When this occurs there are two things you can do: corrections using epsilon or multivariate tests, described next.
- 2. Corrections using epsilon: SPSS provides three estimates of a statistic called epsilon that can be used to correct for a violation revealed by Mauchly's test of sphericity. The greater the violation, the smaller the value of epsilon. All you need to do is decide which of the three estimates of epsilon to use. Greenhouse–Geisser epsilon is probably the most appropriate value to use, but if you have relatively few participants, this can be rather too conservative (i.e. its use will decrease the chances of finding a significant result) in these cases the Huynh–Feldt epsilon may be preferable. The third estimate, the lower-bound epsilon, is a minimum value for epsilon that will give the most conservative correction. SPSS gives corrected values in the table. When reporting any result, make it clear which one you have used.
- 3. **Multivariate tests**: these make fewer assumptions about the data and hence are more appropriate when the Mauchly's test of sphericity is significant. In the Multivariate Tests table, SPSS reports four different multivariate statistics: Pilliai's Trace, Wilks' Lambda, Hotelling's Trace and Roy's Largest Root. Each of these tests reports a value of *F* with associated degrees of freedom and a significance value. You will probably find that there is little difference between the significance of *F* reported by these four procedures, so we suggest you pick one of them and report it. The multivariate values of *F* are always lower than the univariate values; hence, if a result is not significant by the univariate method, it cannot be significant by

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the multivariate method. For this reason, SPSS does not report the multivariate estimates when the univariate test is non-significant.

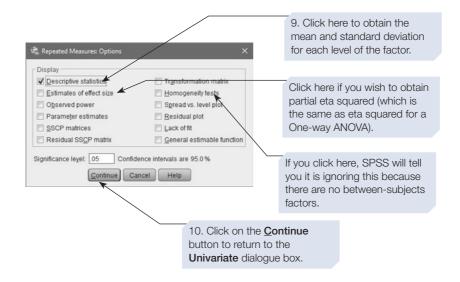
4. Finally, it is worth mentioning the **Test of between-subjects effects** table that appears towards the end of the output even when your design does not include a between-subjects factor. This can be ignored when you only have within-subjects factors in your design, as is the case here with the One-way within-subjects ANOVA. What SPSS is doing is assuming that 'participant' is an additional between-subjects factor in the analysis. One way to think of this is to say that the part of the output reporting the between-subjects effects is asking 'Did all participants perform the same?' It is in the nature of psychology that participants are variable in almost all tasks, and hence you will find that the *F*-ratio is invariably very high and highly significant. As we are not normally interested in this question of whether the participants are all performing in the same way (we usually want to know about general trends across groups of participants), we can ignore this section of the output. Indeed, you will rarely see this result reported in psychology papers.

How to do it

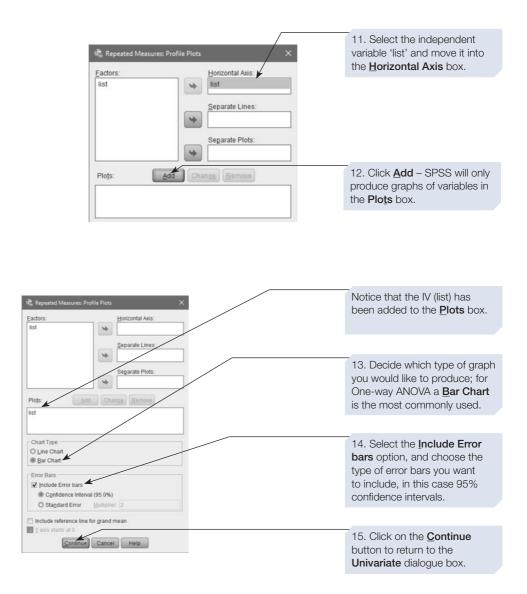
						1. Click on the word <u>A</u>nalyze .
un 1	-Way Wi	thin-Subjects A	nova.sav (DataSel	t2] - IBM SPSS Statistics Data Edito	or	
<u>F</u> ile	<u>E</u> dit	<u>V</u> iew <u>D</u> ata	Transform	Analyze Graphs Utilities	Extensions	Window Help
E				Reports Descriptive Statistics Bayesian Statistics		2. Click on <u>General Linear</u> Model.
	1	Incong 13.00		Ta <u>b</u> les Compare Means	/:	
	2	13.00 16.00		General Linear Model	* ,	Univariate
	4	13.00	8.00	Generalized Linear Models Mixed Models	,	Multivariate
	5 6	14.00 15.00		<u>C</u> orrelate Regression	,	Variance Components
	7 8	14.00 13.00		Loglinear Neural Networks	, ,	
	9	16.00 17.00		Classify Dimension Reduction	•	
	12			Sc <u>a</u> le Nonparametric Tests	,	3. Click on <u>R</u> epeated
	13 14			Forecasting	,	Measures. The Repeated Measures Define Factor(s)
	15 16			<u>S</u> urvival M <u>u</u> ltiple Response	,	dialogue box will appear (see below).
	17 18			Missing Value Analysis Multiple Imputation	,	
	19			Comp <u>l</u> ex Samples	•	

	$\hat{w}_{\rm m}$ Repeated Measures Define Factor(s) $~~ imes~$	$\hat{w}_{\rm m}$ Repeated Measures Define Factor(s) $~~ imes$
4. Replace the factor name	Within-Subject Factor Name:	Within-Subject Factor Name:
suggested by SPSS by	▲ factor1	list
highlighting factor1 and	Number of Levels:	Number of Levels: 3
typing the word 'list'.		
'list' represents your IV:	Add	Add list(3)
the type of list, the factor	Chinge	Change
manipulated in this Stroop		
study, which was the lists of words in different colour inks.	Remove	Remove
	Measure Name:	Measure Name:
	Add	Add
	Change	Change
5. There	e were three different	Remove
	he study, so type the	Itemore
	'3' in the Number of box and click on the ancel Help	Define Reset Cancel Help
	tton. The dialogue	Denne Reset Cancer Help
	now show list(3)	
(see rig	nt).	
		6. Click on the Define button,
		the Repeated Measures
		dialogue box will appear (see below).
		50000).
		7. Move the variables into the
		Within-Subjects Variables
🖓 Repeated Measu	res X	box in a sensible order (see tip
	Within-Subjects Variables Model	box below).
incongruent lis incongruent lis	trall Contrasts	
	7 * _?_(2) Plojs	
	?(3)	
	EM Means	
	Save	
	Options	
	Between-Subjects Factor(s):	<u></u>
	*	8. Click on the Options button
	Covariates:	to obtain descriptive statistics.
		The Repeated Measures:
	*	Options dialogue box will
(OK Paste Reset Cancel Help	appear (see below).
	Lase Peace Cancer Helb	

As SPSS does a trend test, we entered the variables in line with the hypothesis: first, congruent list, then neutral list and then incongruent list. We hypothesised that the time taken to name the ink colour would be shortest for the congruent list, longer for the neutral list and longest for the incongruent list.



Next, we can ask SPSS to produce a graph of the data using the **Plots**... in the main **Univariate** dialogue box. This opens up the **Univariate**: **Profile Plots** dialogue box.



Click on **ok** to obtain the SPSS output shown below. You will find that there is a significant effect of type of list, and you may wish to include in your results section a table displaying the mean for each condition and the 95% confidence intervals (see Chapter 3, Section 5).

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SPSS output for One-way within-subjects ANOVA

Obtained using menu items: General Linear Model > **Repeated Measures**

Wit	hin-Subjects Factors	These are the names for each level of the factor 'list'. If these were entered into the Repeated Measures dialogue box in a meaningful order,
Meas	ure: MEASURE_1	they will be displayed here in that order, and the table Tests of Within-Subjects Contrasts (shown
list	Dependent Variable	on the next page) will also be relevant.
1	cong	
2	neutral	
3	incong	
	Decerintive Statistics K	Useful descriptives that you can incorporate into your report, obtained by ticking

Descriptive Statistics

	Mean	Std. Deviation	Ν
congruent list	8.9000	.73786	10
neutral list	11.1000	1.19722	10
incongruent list	14.4000	1.50555	10

you report, obtained by ticking Descriptive statistics in the **Repeated Measures: Options** dialogue box.

Multivariate Tests^a

Effect		Value	F	Hypothesis df	Error df	Sig.
list	Pillai's Trace	.920	45.993 ^b	2.000	8.000	.000
	Wilks' Lambda	.080	45.993 ^b	2.000	8.000	.000
	Hotelling's Trace 🔻	11.498	45.993 ^b	2.000	8.000	.000
	Roy's Largest Root	11.498	45.993 ^b	2.000	8.000	.000

a. Design: Intercept Within Subjects Design: list

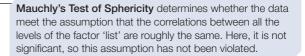
b. Exact statistic

If Mauchly's Test of Sphericity (shown

below) is significant, you might consider using one of the four multivariate statistics shown in this table, as these make fewer assumptions about the data. Guidance on these statistics was given towards the start of this section.



Remember that, should Mauchly's test be significant, there are two options: either adopt a multivariate approach and report one of the four statistics given in the **Multivariate Tests** table above, or use the values for your chosen epsilon from the **Tests of Within-Subjects Effects** table (shown below).



Mauchly's Test of Sphericity^a

Measure: MEASURE_1

						Epsilon ^b	
Within Subjects Effect	Mauchly's W	Approx. Chi- Square	df	Sig.	Greenhouse- Geisser	Huynh-Feldt	Lower-bound
list	.892	.914	2	.633	.903	1.000	.500

Tests the null hypothesis that the error covariance matrix of the orthonormalized transformed dependent variables is proportiona to an identity matrix.

a. Design: Intercept

Measure: MEASURE_1

Within Subjects Design: list

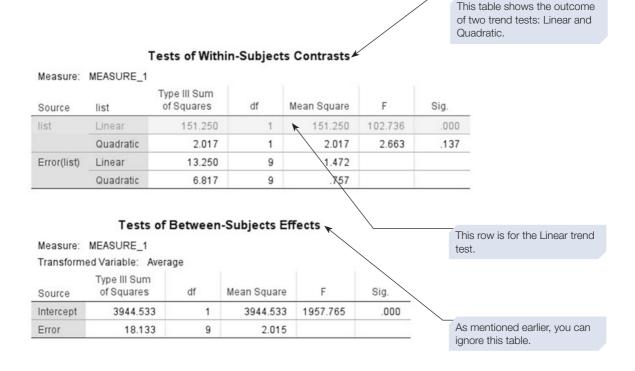
Than oubjects Besign hat

b. May be used to adjust the degrees of freedom for the averaged tests of significance. Corrected tests are displayed in the Tests of Within-Subjects Effects table.

Tests of Within-Subjects Effects

Source		Type III Sum of Squares	df	Mean Square	F	Sig.	
list	Sphericity Assumed	153.267	2	76.633	68.741	.000	
Ť	Greenhouse-Geisser	153.267	1.805	84.906	68.741	.000	You also need the
	Huynh-Feldt	153.267	2.000	76.633	68.741	.000	df for the Error
	Lower-bound	153.267	1.000	153.267	68.741	.000	term.
Error(list)	Sphericity Assumed	20.067	18	1.115			
1	Greenhouse-Geisser	20.067	16.246	1.235			
	Huynh-Feldt	20.067	18.000	1.115			
	Lower-bound	20.067	9.000	2.230			

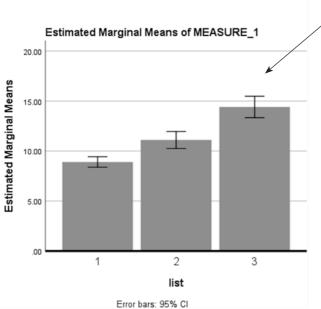
This row is the one you use when Mauchly's test is not significant. It gives the values for the within-subjects factor 'list' that has three levels (the three types of list). Underneath are the three estimates of epsilon. Because Mauchly's test was not significant, no correction is needed and the four different entries for F in this table are identical. For more information, see the start of this section.



For these data, as we predicted, there is a significant linear trend, F(1,9) = 102.74, p < .001. The means suggest that participants took the shortest time to name the ink colour for the congruent list and took the longest time to name the colour of the ink in the incongruent list, with the mean for the neutral list falling between these two values.

Furthermore, for these data, there is no significant quadratic trend, F(1,9) = 2.66, p = .137. A linear trend test is used to see if the points tend to fall onto a straight line (as here). A quadratic trend test looks for a U-shaped or inverted U-shaped trend. If you entered the three levels in the order 'cong', 'incong' and 'neutral', then the quadratic trend would be significant. You might like to try this.





This is the graph requested in the **Profile Plots** dialogue box, which illustrates participant performance in the three different conditions.

When error bars do not overlap, it is likely that the conditions are significantly different – although you need to do post-hoc tests or planned comparisons to confirm this.

Calculating eta squared: one measure of effect size

As with One-way between-subjects ANOVA, we can calculate eta squared to measure effect size. Again, eta squared (η^2) is calculated by dividing the sum of squares for the IV (list) by the total sum of squares. As the total sum of squares isn't given in the **Tests of Within-Subjects Effects**, we need to calculate it by hand by simply adding the IV sum of squares to the Error sum of squares.

Source		Type III Sum of Squares	df	Mean Square	F	Sig.
list	Sphericity Assumed	153.267	2	76.633	68.741	.000
	Greenhouse-Geisser	153.267	1.805	84.906	68.741	.000
	Huynh-Feldt	153.267	2.000	76.633	68.741	.000
	Lower-bound	153.267	1.000	153.267	68.741	.000
Error(list)	Sphericity Assumed	20.067	18	1.115		
	Greenhouse-Geisser	20.067	16.246	1.235		
	Huynh-Feldt	20.067	18.000	1.115		
	Lower-bound	20.067	9.000	2.230		

Tests of Within-Subjects Effects

For 'list' $(\eta^2) = 153.267 / (153.267 + 20.067)$

= 153.267 / 173.334



A One-way within-subjects ANOVA was conducted on response times to a Stroop task. There was a significant effect of the type of list: F(2,18) = 68.74, p < .001, $\eta^2 = .88$. A significant linear trend emerged, F(1,9) = 102.74, p < .001, with naming times increasing across congruent, neutral and incongruent lists.



You would also want to report in your results section your descriptive statistics and information regarding the confidence intervals for each condition.

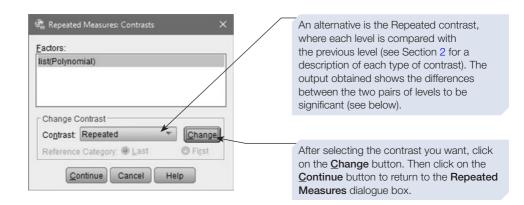
Note that the **Tests of Within-Subjects Contrasts** table shows only whether a trend is significant or not. It is not a test of whether the individual conditions significantly differ from one another. However, ANOVA commands via the <u>General Linear</u> **Model** allow you to choose between a range of preset contrasts, described in Section 2. It is also possible to carry out some unplanned (post-hoc) comparisons. We illustrate how to do each of these next for a within-subjects factor, using the data analysed above. Consult Section 1 for general guidance on planned versus unplanned comparisons.

Planned comparisons: more contrasts for within-subjects factor

Click on <u>Analyze</u> \Rightarrow <u>General Linear Model</u> \Rightarrow <u>Repeated Measures</u>

Within-Subjects Variables Model	in the Repeated Measures dialogue box.
(list): Contrasts Plots Post Hoc	
List(Polynomial) EM Means Save Qptions Contrast Contrast Contrast Polynomial Change Deviation Simple Contiference Helmert Repeated Polynomial	Click here to view the list of contrasts available. The default is Polynomial contrasts, which we saw in the output described above, showing the linear and quadratic trends.

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Tests of Within-Subjects Contrasts

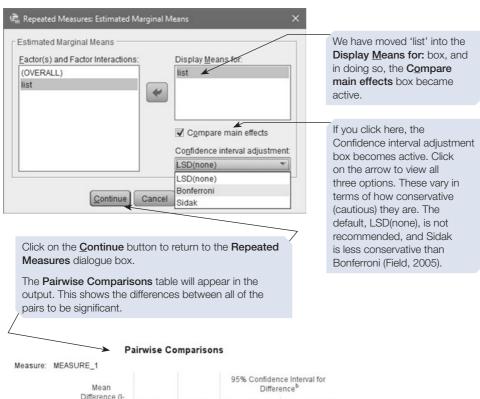
MEASURE_1

Source	list	Type III Sum of Squares	df	Mean Square	F	Sig.
list	Level 1 vs. Level 2	48.400	1	48.400	27.923	.001
	Level 2 vs. Level 3	108.900	1	108.900	54.149	.000
Error(list)	Level 1 vs. Level 2	15.600	9	1.733		
	Level 2 vs. Level 3	18.100	9	2.011		

The **Tests of Within-Subjects Contrasts** table will appear in the output. This shows the contrasts to be significant.

Unplanned comparisons for within-subjects factor ANOVA

In terms of unplanned (post-hoc) comparisons, the choice for a within-subjects factor is more limited than for a between-subjects factor. If you click on the **Post Hoc** button in the **Repeated Measures** dialogue box, you will not see the within-subjects factor listed. However, a smaller selection of post-hoc tests are available if you select the <u>EM</u> **Means...** in the **Repeated Measures** dialogue box and open up the **Estimated Marginal Means** box.



		Difference (I-				ence
(I) list	(J) list	J)	Std. Error	Sig. ^b	Lower Bound	Upper Bound
1	2	-2.200	.416	.002	-3.421	979
	3	-5.500	.543	.000	-7.092	-3.908
2	1	2.200	.416	.002	.979	3.421
	3	-3.300	.448	.000	-4.615	-1.985
3	1	5.500	.543	.000	3.908	7.092
	2	3.300	.448	.000	1.985	4.615

Based on estimated marginal means

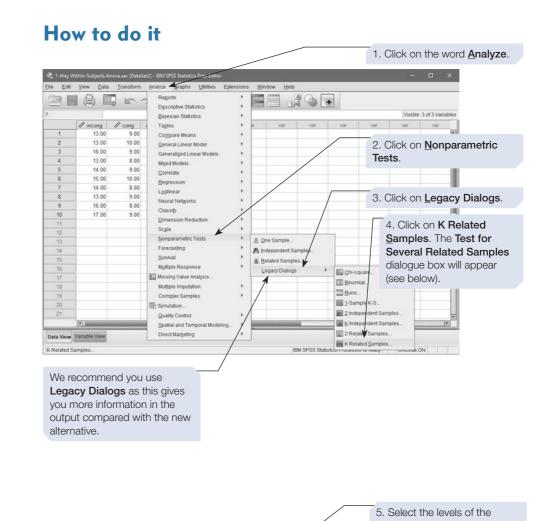
*. The mean difference is significant at the .05 level.

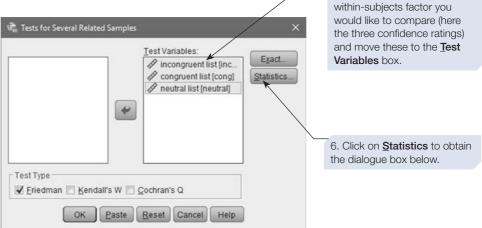
b. Adjustment for multiple comparisons: Bonferroni.

Finally, it is worth noting that, as with the One-way between-subjects ANOVA, there is a nonparametric version of the One-way within-subjects ANOVA, and this is described in the last part of this section.

The Friedman test

The Friedman test is the nonparametric equivalent of the One-way within-subjects analysis of variance. Confusingly, the Friedman test is sometimes referred to as the 'Friedman two-way ANOVA' (this is because for a within-subjects analysis of variance, the participants are also considered to be a factor). For demonstration purposes, we perform this test using the same data as for the One-way within-subjects ANOVA.







7. Select **Descriptive** or **Quartiles**, depending on what is most appropriate for your data. We demonstrate the second

8. Click on the Continue button to return to the Tests for Several Related Samples dialogue box. Finally, click on OK.

SPSS output for Friedman test

Obtained by using menu items: Nonparametric Tests > K **Related Samples**

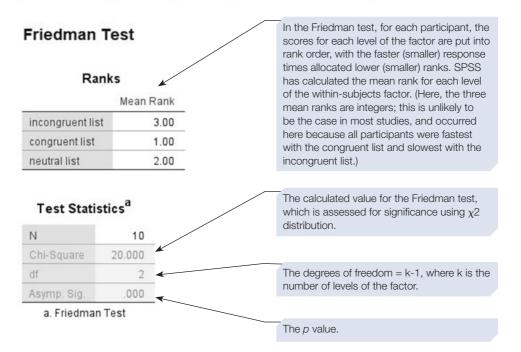




Descriptives for each level of the within-subjects factor, here in the form of quartiles, were obtained via the Statistics button in the dialogue box.

Descriptive Statistics

	Ν	25th	50th (Median)	75th
incongruent list	10	13.0000	14.0000	16.0000
congruent list	10	8.0000	9.0000	9.2500
neutral list	10	10.0000	11.0000	12.0000





Reporting the results

In a report you might write:

A Friedman test revealed that response times to a Stroop task varied significantly for incongruent, congruent and neutral lists: x^2 (2, N = 10) = 20.00, p < .001.

Summary

- This chapter introduced you to One-way ANOVA, planned and unplanned comparisons and nonparametric equivalents of the One-way ANOVA.
- These tests of differences are used for experimental designs involving more than two groups or conditions.
- Effect size (eta squared) can be calculated by hand by dividing the sum of squares for the IV by the total sum of squares. For One-way ANOVA, this can also be produced selecting 'Estimates of effect size' in the Options dialogue box.
- Appropriate descriptive statistics, the mean and standard deviation, can be obtained either by following the advice in Chapter 3 or by selecting the appropriate options on the ANOVA Options dialogue box.
- Error bar charts are often used to display statistically significant findings, and can be produced using the ANOVA options in some cases.
- If your dependent variable is a total score calculated from several raw scores, which have already been entered into your data file, then see Chapter 4 for guidance on how to create such a total score in SPSS.
- For guidance on incorporating SPSS output into a report, or on printing the output, see Chapter 14.

9 Factorial analysis of variance

In this chapter

- An introduction to factorial analysis of variance (ANOVA)
- Two-way between-subjects ANOVA
- Two-way within-subjects ANOVA
- Mixed ANOVA



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Section 1: AN INTRODUCTION TO FACTORIAL ANALYSIS OF VARIANCE (ANOVA)

- The previous chapter introduced you to the simplest type of ANOVA: the One-way ANOVA. While this allows you to analyse data collected from experimental designs with more than two conditions, it only allows you to explore the influence of one independent variable at a time. However, human beings are complicated creatures, and psychologists often want to investigate how multiple factors (or variables) can impact people's behaviour. In these situations, researchers can use what is called a factorial design to investigate their hypotheses or research questions. Factorial simply means that their experimental design have more than one IV.
- Factorial ANOVA allows us to analyse the data from these types of factorial design, by allowing us to investigate the effect of more than one IV at once. For example, we could examine the effect of participants' sex as well as their age on their memory for a list of words. Here, we have two IVs (sex and age) and one dependent variable (DV) (memory score). A single ANOVA test will allow us to examine simultaneously the effect of these two IVs. In fact, ANOVA can handle any number of IVs in a single experiment, but, in practice, we rarely include more than three or four for reasons that will become apparent shortly.

- A major advantage of ANOVA is that it allows us to investigate how these IVs combine to affect the DV. For example, we can ask questions about how the sex *and* the age of a participant *combine* to affect memory score it might be that male participants' performances decline with age but that female participants' performances improve with age. Such an *interaction* between these two variables is of theoretical importance, and it is only by investigating both variables in one design that we can discover this interaction.
- For more information about the assumptions of ANOVA, and how it works, please refer back to Chapter 8. The basic principles underlying One-way and Factorial ANOVA are the same.

Different types of Factorial ANOVA?

As with One-way ANOVAs and *t*-tests, Factorial ANOVAs come in a number of different varieties depending on your experimental design. These are:

Between-subjects ANOVA

Between-subjects (or Independent) ANOVA is used when *all* of your factors (or independent variables) are between subjects. That is, when all have different participants taking part in each of the different levels (or conditions).

Within-subjects ANOVA

Within-subjects (or Repeated Measures) ANOVA is used when *all* of your factors (or independent variables) are within subjects. In this case, this means that the same participants take part in *all* of the conditions in your experiment.

Mixed ANOVA

Mixed ANOVA is used when you have an experimental design that includes one or more within-subjects factor(s) and one or more between-subjects factor(s).

Factorial ANOVAs

Exactly which type of Factorial ANOVA you use will vary according to the experimental design you have (i.e. between-subjects, within-subjects or mixed), but also on the number of factors you have in your study. When describing an ANOVA design, we need to specify three things:

- 1. The number of factors involved in the design.
- 2. How many levels there are of each factor.
- 3. Whether each factor is a within- or between-subjects factor.

Just as a one-way ANOVA has one factor, we can use the number of factors to describe a two-way ANOVA (which has two factors), or a three-way ANOVA (which has three factors) and so on (e.g. a six-way ANOVA would have six factors). What this does not tell you is how many levels (or conditions) each factor has. You could describe

CHAPTER 9

this in longhand, but there is an easier convention. For example, a three-way ANOVA, in which the first factor, 'Sex', had two levels, the second factor, 'Age', had three levels and the third factor, 'Drug dose', had five levels, could be described more simply as a 2*3*5 ANOVA design. Note that, in this terminology, the number of numerals (three in this case) describes the number of factors, and the values of the numerals indicate the number of levels of each of these factors. Using this terminology, we just need to make it clear whether the factors were within- or between-subject factors. We could do this by writing:

A 2*3*5 (Sex*Age*Drug dose) mixed ANOVA design was employed, where Sex and Age were between-subjects factors and Drug dose was a within-subjects factor.

Main effects and interactions

Using ANOVA, we can analyse data from studies that incorporate more than one factor. We can assess the effect of each of these factors on their own and the interaction between the factors. The term 'main effect' is used to describe the independent effect of a factor. For example, in the 2*3*5 ANOVA described above, three main effects will be reported. The main effect of 'Sex' will tell us whether men performed significantly differently from women, irrespective of their age or drug dose. The main effect of 'Age' will tell us whether age affects performance, irrespective of sex or drug dose. Finally, the main effect of 'Drug dose' will tell us whether drug dosage affects performance, irrespective of the sex or age of the participants. These main effects simply compare the mean for one level of a factor with the mean of the other level(s) of that factor; for example, comparing mean male performance levels to mean female performance levels. Interactions, on the other hand, assess the combined effect of the factors. An interaction that assesses how two factors combine to affect performance is called a *two-way interaction*.

Interactions and moderation

Interactions are important, as they can tell us how one variable might moderate the effects of another. Moderation is an important statistical concept in psychological research, as psychologists are often interested in how the influence of an IV on a DV might be moderated by a third variable.

When we carry out a standard Factorial ANOVA, we are usually just looking to see whether two variables (A and B) interact in terms of their effects on a criterion variable C. We don't usually make any distinction in the role of A and B; they are both just considered to be independent variables (or factors). If an interaction occurs, this means that the effect of A on C is different at different levels of B; and likewise, the effect of B on C is different at different levels of A.

In contrast, if you are interested in the moderating effects of a specific variable, then you assign your factors differently: One remains an IV and the other becomes the Moderator. In this case, you would specifically investigate the effect the Moderator, B, has on the relationship between A and C.

Mathematically, there is no difference between interaction and moderation, and they are analysed in the same way. However, the concept of moderation affects the way you interpret the results.

Understanding ANOVA output

When attempting to understand the output from the ANOVA command in SPSS, it is helpful if you know in advance how many results you are looking for:

- 1. *A one-way ANOVA*, where the single factor is called A, will give rise to just a single main effect of A.
- 2. *A two-way ANOVA*, where the factors are called A and B, will give rise to two main effects (main effect of A and main effect of B), and a single two-way interaction (A*B). This is a total of three results (3 *F*-ratios).
- 3. A three-way ANOVA, where the factors are called A, B and C, will give rise to three main effects (main effect of A, main effect of B and main effect of C), three two-way interactions (A*B, A*C and B*C) and a single three-way interaction (A*B*C). This is a total of seven results.
- 4. A four-way ANOVA, where the factors are called A, B, C and D, will give rise to four main effects (main effect of A, main effect of B, main effect of C and main effect of D), six two-way interactions (A*B, A*C, A*D, B*C, B*D and C*D), four three-way interactions (A*B*C, A*B*D, A*C*D and B*C*D), and a single four-way interaction (A*B*C*D). This is a total of 15 results.

You can now see why it is unusual to include more than four factors in a design. The number of possible interactions rises steeply as the number of factors increases. Furthermore, it is unlikely that you hypothesised about the shape of these higher-level interactions, and if they are significant, they can be hard to describe and/or explain. Using SPSS, it is easy to undertake a four- or even five-way ANOVA, but rather more difficult to explain the results. Our advice is to try to limit yourself to a maximum of three factors.

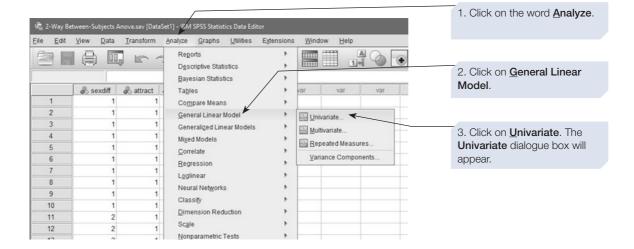
The next three sections will walk you through how to carry out between-subjects, within-subjects and mixed ANOVA using SPSS.

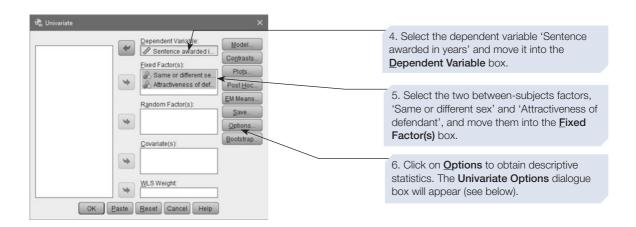
Section 2: TWO-WAY BETWEEN-SUBJECTS ANOVA

Example study: the effect of defendant's attractiveness and sex on sentencing

To practise how to analyse data with a two-way between-subjects ANOVA design, we will consider the possible effects of attractiveness and also the gender of the defendant in a mock trial. In the study described here, conducted by one of our past students, the testimony of a hypothetical defendant describing a murder and admitting guilt was presented as written text to 60 participants: 20 participants simply received the written text with no photo attached, 20 participants received the text and a photo of an attractive defendant, and 20 participants received the text and a photo of an unattractive defendant. The photo was of either a man or a woman. Participants were asked to indicate how many years in jail the defendant should receive as punishment. The design employed was a 3*2 between-subjects ANOVA design. The first between-subjects factor was the knowledge about attractiveness, which had three levels; the factor is operationalised as showing no photo of the defendant (so no knowledge about attractiveness is provided, a control condition), a photo of an attractive defendant or a photo of an unattractive defendant. The second between-subjects factor was same or different sex, operationalised by showing a photo of the defendant of the same or opposite sex as the participant. Sex of the defendant was stated in the written text for the participants who received no photo. The dependent variable was the sentence given, operationalised as how many years the defendant should spend in prison, ranging from a minimum of 3 to a maximum of 25. The hypothesis tested was that the unattractive defendant would be sentenced more harshly and that the length of sentence given might also depend on the sex of the participant. (These data are available from macmillanihe.com/harrison-spss-7e.)

How to do it

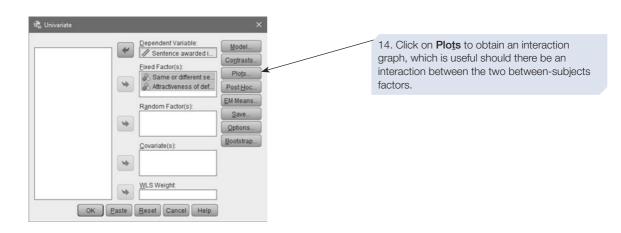




Display Descriptive statistics	7. Click here to check for assumption of homogeneity of variance.
Estimates of effect size Spread vs. level plot	
Observed power Contrast coefficient matrix General estimable function Heteroskedasticity Tests	8. Click here to obtain the mean and standard deviation for each level of the factor.
Modified Breusch-Pagan test Model Breusch-Pagan test Wodel White's test Model	9. Click here if you want to obtain partial eta squared, a measure of effect size, but see guidance provided in Chapter 8.
Parameter estimates with robust standard errors HC0 HC1 HC2	10. Click on the Continue button to return to the Univariate dialogue box.
● HC3	
© HC4	



Univariate: Post Hoc Multiple Comparisons for Observed Means X	
Eactor(s): Post Hoc Tests for:	12. You only need to do post-hoc tests for factors (or variables) that have more than two levels.
attract	As sexdiff only has two levels, we have only moved attract across to the Post Hoc Tests for: box. Doing this makes a number of post-hoc options below become active.
Equal Variances Assumed LSD S-N-K Sonferroni Tukey Type II Error Ratio: 100	In this case, we have selected Bonferroni (a relatively conservative test), but consult your statistics textbook for more about the other options.
Sidak Tukey's-b Dunngtt Sigheffe Duncan Control Category Last	
R-E-G-W-F Hochberg's GT2 R-E-G-W-Q Gabriel Image: State of the s	13. Click on Continue to return to the Univariate dialogue box
Equal Variances Not Assumed	
Cancel Help	



'attract' into the <u>H</u> the <u>Separate Line</u> with the fewest lev box.)	etween-subjects factor, moving lorizontal Axis box and 'sexdiff' into es box. (As a general rule, the factor vels should go in the Separate Lines vill become active. Click on this to add Plots: box.
Image: Second structure Image: Second structure <td>Univariate: Profile Plots X Eactors: Horizontal Axis: sexdiff Separate Lines: attract Separate Plots:</td>	Univariate: Profile Plots X Eactors: Horizontal Axis: sexdiff Separate Lines: attract Separate Plots:
Plots: Add Charles Remove	Plots: Add Change Remove
Chart Type: <u>Line Chart</u> <u>Bar Chart</u>	Chart Type: Line Chart Bar Chart
 Error Bars Confidence Interval (95.0%) 16. Decide which type of graph you would like to produce; for Factorial ANOVA a Line Chart can be really helpful when of interpreting your interaction. Select the Include Error bars option, and 	Error Bars Confidence Interval (95.0%) Standard Error Multiplier: 2 Include reference line for grand mean Y axis starts at 0 Continue Cancel Help
choose the type of error bars you want to include, in this case 95% confidence intervals.	
	17. Click on Continue to return to the Univariate dialogue box

Finally, click on the or button in the Univariate dialogue box to display the output. The annotated output is shown below.

CHAPTER 9

SPSS output for two-way between-subjects ANOVA

Obtained using menu items: General Linear Model > Univariate

Univariate Analysis of Variance

Betweer	n-Subiec	ts Factors			SPSS reminds you of the factors you are analysing, the levels of each factor are
		Value Label	Ν		and the number of particip
Same or different sex	1	Same sex as defendant	30		in each level.
	2	Opposite sex to defendant	30		
Attractiveness of	1	Attractive	20		
defendant	2	Unattractive	20		
	3	No picture	20		This table was produced
Dependent Variable: Ser	ntence awa	escriptive Stat rded in years reness of	istics 🖌		statistics in the Univariate Options dialogue box.
Same or different sex	defend		Mean	Std. Deviation	N
Same sex as defendant	Attractiv	e k	7.50	1.780	10
	Unattra	ctive	11.20	2.300	10
	No pict	ure k	14.50	1.269	10
	Total		11.07	3.403	30
Opposite sex to defendant	Attractiv		7.50	2.415	10
	Unattra	ctive	10.30	2.058	10
	No pict	ure	13.50	1.650	10
	Total		10.43	3.191	30
Total	Attractiv		7.50	2.065	20
	Unattra	///	10.75	2.173	20
	No pict	ure	14.00	1.522	20
/_///	Total		10.75	3.286	60
	ctor 'attrac	descriptives for east', collapsing acro ; 'sexdiff'.			
/			VV E	ach of these six ro	ws shows the descriptives for o
se two Total rows show o ach level of the factor 'so psing across the levels o	exdiff',		pi pi	f the conditions of articipants who we	the study. Thus, the first row is f ere given a photo of an attractive the same sex as that defendan

СО factor, 'attract'. SPSS tests the assumption of equality of variances in a number of different ways. In this case, we are interested in the statistic **based on the mean**.

If Levene's test is significant, this suggests that the group variances are likely to be significantly different, suggesting that the assumption of homogeneity of variance has been violated. Here, it is not significant, so the assumption has been met.

Levene's Test of Equality of Error Variances^{a,b}

		Levene Statistic	df1	df2	Sig.
Sentence awarded in years	Based on Mean	1.509	5	54	.202
	Based on Median	1.195	5	54	.324
	Based on Median and with adjusted df	1.195	5	43.110	.328
	Based on trimmed mean	1.497	5	54	.206

Tests the null hypothesis that the error variance of the dependent variable is equal across groups.

a. Dependent variable: Sentence awarded in years

b. Design: Intercept + sexdiff + attract + sexdiff * attract

Tests of Between-Subjects Effects

Dependent Variable: Sentence awarded in years

This table shows the outcome of the analysis of variance. Each row shows information for a source of variance.

Source	Type III Sum of Squares	df	Mean Square	F	Sig.
Corrected Model	431.550 ^a	5	86.310	22.658	.000
Intercept	6933.750	1	6933.750	1820.236	.000
sexdiff	6.017	1	6.017	1.579	.214
attract	422.500	2	211.250	55.457	.000
sexdiff * attract	3.033	2	1.517	.398	.674
Error	205.700	54	3.809		
Total	7571.000	60			
Corrected Total	637 250	59			

This row shows information about the main effect of the factor 'sexdiff'. This row shows information about the main effect of the factor 'attract'. This row shows information about the interaction between the factors 'sexdiff' and 'attract'.

CHAPTER 9

Post Hoc Tests

Attractiveness of defendant

SPSS prints out a matrix (as it does for correlations), and you have to pick out the three possible comparisons.

Multiple Comparisons

Dependent Variable:	Sentence awarded in years
Bonferroni	

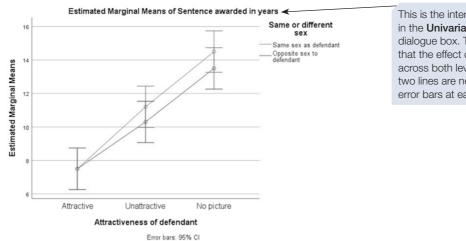
(I) Attractiveness of	(J) Attractiveness of	Mean Difference (I-			95% Confidence Interval	
defendant	defendant	J)	Std. Error	Sig.	Lower Bound	Upper Bound
Attractive	Unattractive	-3.25	.617	.000	-4.77	-1.73
	No picture	-6.50	.617	.000	-8.02	-4.98
Unattractive	Attractive	3.25	.617	.000	1.73	4.77
	No picture	-3.25	.617	.000	-4.77	-1.73
No picture	Attractive	6.50	.617	.000	4.98	8.02
	Unattractive	3.25	.617 /	.000	1.73	4.77

The error term is Mean Square(Error) = 3.809.

*. The mean difference is significant at the .05 level.

As our factor 'attract' had three levels, there are three possible comparisons. Their p values are highlighted.

Profile Plots



This is the interaction graph requested in the **Univariate: Profile Plots** dialogue box. The graph suggests that the effect of 'attract' is similar across both levels of 'sexdiff' – the two lines are nearly parallel, and the error bars at each point overlap.

Calculating eta squared: one measure of effect size

In Chapter 8, Section 1 we provided guidance on reporting effect size when carrying out ANOVAs. Although SPSS will calculate partial eta squared, this is not easy to interpret as it is an adjusted measure: the variance explained by one factor *after* taking into account the variance explained by the other factor(s). Instead, Mulhern and Greer (2011) recommend eta squared, which is relatively straightforward to calculate by hand and will tell you the proportion of the total variability accounted for by each factor and interaction in your design. We will demonstrate how to do this now. Note, however, that there are alternative measures which may be preferable, as they attempt to estimate the effect size in the population rather than just in the sample (see Fritz, Morris and Richler, 2012 for a review and guidance on how to calculate these, and/or consult your statistics text). Eta squared (η^2) can be calculated from the ANOVA output. It is the sum of squares for the factor or interaction divided by the corrected total sum of squares, and these values can be found in the **Tests of Between-Subjects** Effects table.

	Tests of Bet	ween-Su	ubjects Effect	s	
Dependent Variabl	le: Sentence awar	ded in yea	rs		
Source	Type III Sum of Squares	df	Mean Square	F	Sig.
Corrected Model	431.550 ^a	5	86.310	22.658	.000
Intercept	6933.750	1	6933.750	1820.236	.000
sexdiff	6.017	1	6.017	1.579	.214
attract	422.500	2	211.250	55.457	.000
sexdiff * attract	3.033	2	1.517	.398	.674
Error	205.700	54	3.809		
Total	7571.000	60			
Corrected Total	637.250	59			

a. R Squared = .677 (Adjusted R Squared = .647)

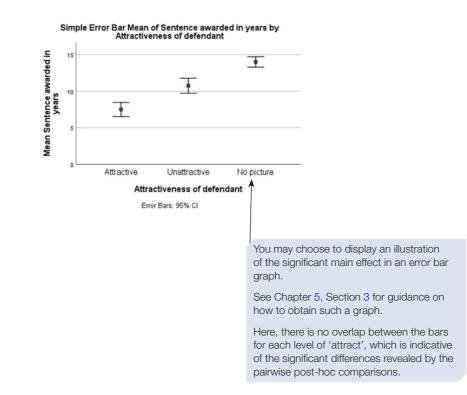
- For the factor 'sexdiff' $(\eta^2) = 6.02 / 637.25 = .009$
- For the factor 'attract' $(\eta^2) = 422.50 / 637.25 = .663$
- For the interaction 'sexdiff*attract' (η^2) = 3.03 / 637.25 = .005

The factor 'attract' accounts for approximately 66% of the variance in the dependent variable, whereas the other factor, 'sexdiff', and the interaction account for only a tiny proportion. The remainder of the variance is accounted for by error (205.7 / 637.25 = .32).



A two-way between-subjects ANOVA was conducted on sentencing judgements. Whether the sex of the defendant was the same as or different from the sex of the participant did not affect the length of sentence given (*F* (1,54) = 1.58, *p* = .214, η^2 = .009). However, information about the attractiveness of the defendant did influence sentencing judgements (*F* (2,54) = 55.46, *p* < .001. η^2 = .663), see graph below. There was no significant interaction between these two factors (*F* (2,54) = 0.40, *p* = .674, η^2 = .005).

Bonferroni post-hoc comparisons showed significant differences between no picture and attractive conditions (p < .001), no picture and unattractive conditions (p < .001), and attractive and unattractive conditions (p < .001).



Section 3: TWO-WAY WITHIN-SUBJECTS ANOVA

Example study: the effects of two memory tasks on finger tapping performance

To practise a two-way within-subjects ANOVA, we shall look at an experiment examining the effects of two memory tasks on tapping performance. Research has identified that right index finger tapping is largely controlled by the left hemisphere, and left index finger tapping by the right hemisphere. If a cognitive task is performed at the same time as this finger tapping task, then the way in which the cognitive task interferes with such tapping could reflect the extent to which either hemisphere is involved in controlling the cognitive task. Many studies that required participants to tap as fast as possible with their index finger while also performing a verbal task found that right-hand tapping was disrupted more than left-hand tapping. This result is compatible with the notion that the left side of the brain, more so than the right side, is involved in controlling both right-hand tapping and many verbal tasks. In a study by Towell, Burton and Burton (1994), participants were asked to tap with each hand while memorising either the words presented to them on a screen (a verbal memory task) or the position of the words on the screen (a visuospatial memory task). As before, memorising the words should disrupt right-hand tapping more than lefthand tapping due to the language processing involved. However, as the right side of the brain is thought to be involved in visuospatial tasks, memorising the positions of words should disrupt left-hand tapping more.

The design employed was a 2*2 within-subjects ANOVA. Each factor had two levels: the first was the tapping hand (left or right hand) and the second was the memory task (memorising the words or memorising the positions). All participants were tested under each possible combination of the two factors, making it a within-subjects design. The dependent variable was a percentage change score, showing the extent to which tapping is slowed down by the concurrent performance of the memory task. The hypothesis tested was that there would be an interaction between tapping hand and memory task. This hypothesis was supported, and for the purposes of this book, we have created a data file that will reproduce some of the findings reported by Towell et al. (1994). (These data are available from macmillanihe.com/ harrison-spss-7e.)

Labelling within-subjects factors

Consider the factors and levels in this example; they could be set out as in Table 9.1. As each factor has two levels, there are four conditions in total, each with one level of one factor and one level of the other factor. When entering data from a within-subjects ANOVA, you will need to have a column for each condition that each participant takes part in. In this case, a $2^{*}2$ within-subjects ANOVA has $2 \ge 4$ conditions (just as a $2^{*}3$ within-subjects ANOVA will have 6 conditions, or a $2^{*}2^{*}3$ within-subjects ANOVA will have 12).

Factor 1	Tapping h	nand			
Levels	Right		Left		
Factor 2	Memory 1	Memory task Me		lemory task	
Levels	Words	Position	Words	Position	
Column name, SPSS data file, for conditions	h1s1	h1s2	h2s1	h2s2	

 Table 9.1
 The numbering system for within-subjects factors

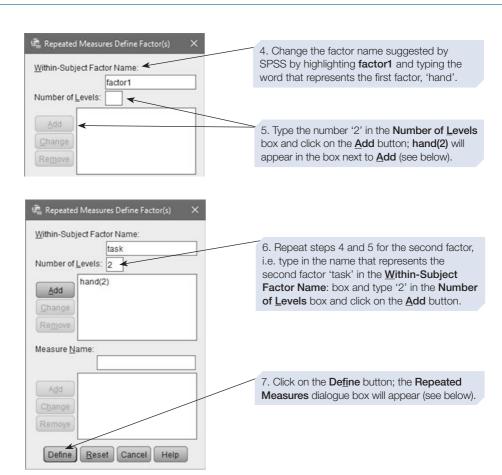
In this example, the name that was given in the SPSS data file to each column containing the data for each condition incorporates a number for each level of each factor, as shown in the bottom row of Table 8.4. For example:

- 'h1s1' means tapping hand 1 (right) and stimulus for task 1 (memorising words).
- 'h2s2' means tapping hand 2 (left) and stimulus for task 2 (memorising positions).

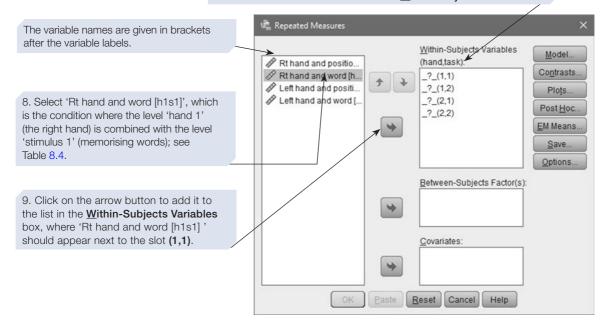
You should jot down a rough table such as this before entering the data for any design with two or more within-subjects factors. This will help you when you define the within-subjects factors, because you will find that the numbers you have used for the column names will match the numbers that SPSS uses when requesting variable selection.

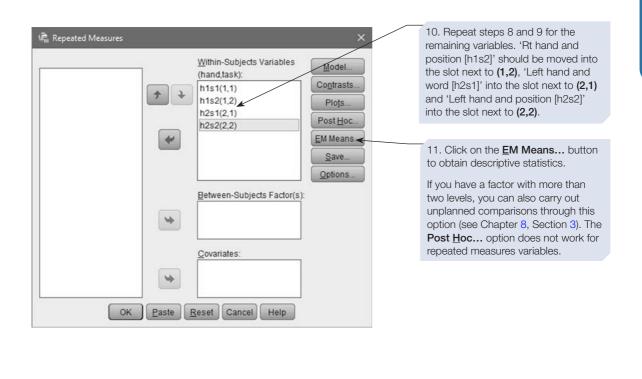
How to do it

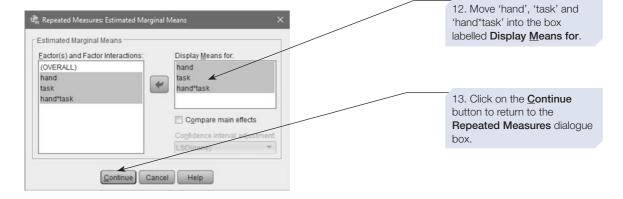
💼 2-Way W	ithin-Subjects An	ova.sav [DataSet	2] - IBM SPSS Statistics Data Editor					
Eile <u>E</u> dit	⊻iew <u>D</u> ata	Transform	Analyz <u>y G</u> raphs <u>U</u> tilities	Extensions	Windo	w <u>H</u> elp		1. Click on the word Analyze
			Reports Descriptive Statistics Bayesian Statistics	* * *		 		
	@ h1s2	A h1s1	Tables	•	var	var	var	var var
1	-1.10	-5.52	Compare Means	*				2. Click on <u>G</u> eneral Linear
2	11.11	1.48	General Linear Model	*	Univariate			Model.
3	-4.19	2.40	Generalized Linear Models		Multivariate			
4	6.74	13.78	Mixed Models					
5	-4.91	61	Correlate		Variance Components		5	
6	11.11	8.11	Regression				ents	3. Click on <u>R</u>epeated
7	7.49	8.14	Loglinear					Measures. The Repeated
8	1.74	4.07	Neural Networks					Measures Define Factor(s)
9	-3.47	-4.10	Classify	í.				dialogue box will appear (see
10	6.04	7.09	-					below).
11	6.51	9.47	Dimension Reduction					0010 10/01.
12	7.02	19.06	Sc <u>a</u> le					
13	5.91	15.68	Nonparametric Tests	,	-			



This key helps explain the numbers in the brackets below: 'hand' refers to the first number and 'task' to the second number. This will help you enter the variables in the correct order into the <u>Within-Subjects Variables</u> box.

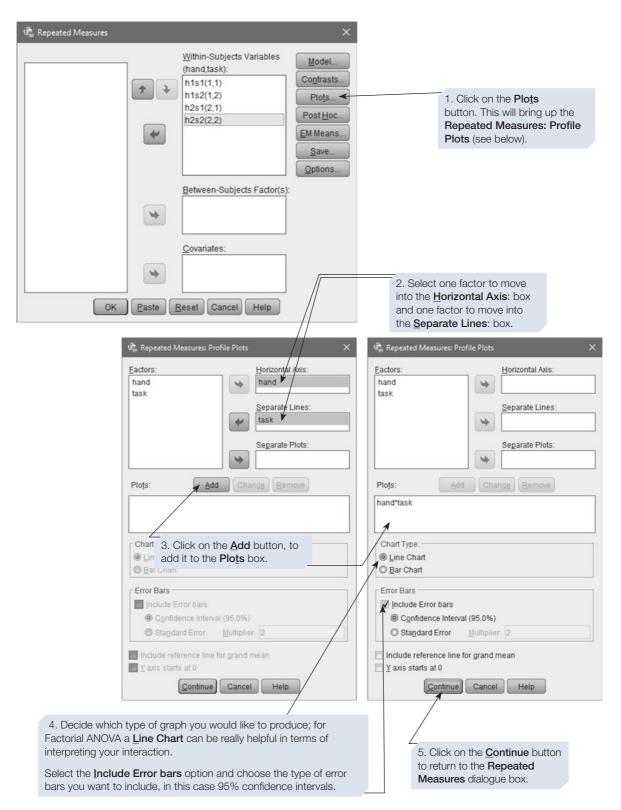






Click on **SPSS** will perform the calculations. You may wish to obtain an interaction graph should the analysis reveal a significant interaction – this is an option available on the **Repeated Measures** dialogue box; see the steps outlined below.

How to obtain an interaction graph



Click on **or** to obtain the ANOVA output, shown below, with the interaction graph at the end of the output.

The output for a within-subjects ANOVA is lengthier than that for a between-subjects ANOVA, and we explain the purpose of some of the tables at the start of Chapter 8, Section 3. If you have not read this already, now would be a good time to do so.

SPSS output for two-way within-subjects ANOVA

Obtained using menu items: General Linear Model > Repeated Measures

Within-Subjects Factors

Measure:	MEASURE_1

hand	task	Dependent Variable		
1	1	h1s1		
	2	h1s2		
2	1	h2s1		
	2	h2s2		

This table is important when there are more than two levels of a within-subjects factor and when **Mauchly's Test of Sphericity** (next table) is significant.

See the guidance provided at start of Chapter 8, Section 3 on understanding the output of within-subjects ANOVAs.

Multivariate Tests^a

Effect		Value	F	Hypothesis df	Error df	Sig.
hand	Pillai's Trace	.006	.133 ^b	1.000	23.000	.719
	Wilks' Lambda	.994	.133 ^b	1.000	23.000	.719
	Hotelling's Trace	.006	.133 ^b	1.000	23.000	.719
	Roy's Largest Root	.006	.133 ^b	1.000	23.000	.719
task	Pillai's Trace	.078	1.955 ^b	1.000	23.000	.175
	Wilks' Lambda	.922	1.955 ^b	1.000	23.000	.175
	Hotelling's Trace	.085	1.955 ^b	1.000	23.000	.175
	Roy's Largest Root	.085	1.955 ^b	1.000	23.000	.175
hand * task	Pillai's Trace	.173	4.807 ^b	1.000	23.000	.039
	Wilks' Lambda	.827	4.807 ^b	1.000	23.000	.039
	Hotelling's Trace	.209	4.807 ^b	1.000	23.000	.039
	Roy's Largest Root	.209	4.807 ^b	1.000	23.000	.039

Mauchly's Test of Sphericity^a

Measure: MEASURE_1

					Epsilon ^b		
Within Subjects Effect	Mauchly's W	Approx. Chi- Square	df	Sig.	Greenhouse- Geisser	Huynh-Feldt	Lower-bound
hand	1.000	.000	0		1.000	1.000	1.000
task	1.000	.000	0		1.000	1.000	1.000
hand * task	1.000	.000	0		1.000	1.000	1.000

Tests the null hypothesis that the error covariance matrix of the orthonormalized transformed dependent variables is proportional to an identity matrix.

This table is important when there are more than two levels. Here, both of the withinsubjects factors have only two levels, so we don't need to use it. This row gives the values for the factor 'hand', the withinsubjects factor with two levels (right and left). This table shows the outcome of the analysis of variance. For information about the nonhighlighted rows, see Chapter 8, Section 3.

Tests of Within-Subjects Effects

Source		Type III Sum of Squares	df	Mean Square	F	Sig.
hand	Sphericity Assumed	2.295	1	2.295	.133	.719
	Greenhouse-Geisser	2.295	1.000	2.295	.133	.719
	Huynh-Feldt	2.295	1.000	2.295	.133	.719
	Lower-bound	2.295	1.000	2.295	.133	.719
Error(hand)	Sphericity Assumed	398.267	23	17.316		
	Greenhouse-Geisser	398.267	23.000	17.316		
	Huynh-Feldt	398.267	23.000	17.316		
	Lower-bound	398.267	23.000	17.316		
task	Sphericity Assumed	70.906	1	70.906	1.955	.175
1	Greenhouse-Geisser	70.906	1.000	70.906	1.955	.175
	Huynh-Feldt	70.906	1.000	70.906	1.955	.175
	Lower-bound	70.906	1.000	70.906	1.955	.175
Error(task)	Sphericity Assumed	834.213	23	36.270		
	Greenhouse-Geisser	834.213	23.000	36.270		
	Huynh-Feldt	834.213	23.000	36,270		
	Lower-bound	834.213	23.000	36.270		
hand * task	Sphericity Assumed	21.441	1	21.441	4.807	.039
×	Greenhouse-Geisser	21.441	1.000	21.44	4.807	.039
$\langle \rangle$	Huynh-Feldt	21.441	1.000	21.441	4.807	.039
\backslash	Lower-bound	21.441	1.000	21.441	4.807	.039
Error(hand*task)	Sphericity Assumed	102.585	23	4.460	\	
	Greenhouse-Geisser	102.585	23.000	4.460		
	Huynh-Feldt	102.585	23.000	4.460		
	Lower-bound	102.585	23.000	4,460		

This row gives the values for the factor 'task', the withinsubjects factor with two levels (memorising the word or memorising its position). This row gives the values for the interaction between the factor 'hand' and the factor 'task'. You also need the degrees of freedom for the error associated with each main effect or interaction. In this example, the error df is the same for all three. This is not always the case. This table shows the outcome of trend tests. Each factor only has two levels, and so:

1. Only linear tests can be carried out, and not quadratic.

2. The values are simply those for the analysis of variance.

If, however, you have at least one factor with three or more levels, this table would be useful, as

shown in the one-way within-subjects ANOVA example in Chapter 8.

Tests of Within-Subjects Contrasts

Measure: MEASURE_1

Source	hand	task	Type III Sum of Squares	df	Mean Square	F	Sig.
hand	Linear		2.295	1	2.295	.133	.719
Error(hand)	Linear		398.267	23	17.316		
task		Linear	70.906	1	70.906	1.955	.175
Error(task)		Linear	834.213	23	36.270		
hand * task	Linear	Linear	21.441	1	21.441	4.807	.039
Error(hand*task)	Linear	Linear	102.585	23	4.460		

	Tests of I	Between	-Subjects Effe	ects		
Measure:	MEASURE_1					
Transform	ed Variable: Avera	age				You can ignore this table; see
Source	Type III Sum of Squares	df	Mean Square	F	Sig.	guidance in Chapter 8, Section 3.
Intercept	2544.862	1	2544.862	33.553	.000	0.
Error	1744.438	23	75.845			

These three tables give the descriptives requested in the EM Means... box: 'hand',

Estimated Marginal Means -

1. hand

Measure: MEASURE_1

			95% Confidence Interval			
hand	Mean	Std. Error	Lower Bound	Upper Bound		
1	4.994	1.017	2.890	7.098		
2	5.303	.952	3.334	7.273		

2. task

			95% Confidence Interval			
task	Mean	Std. Error	Lower Bound	Upper Bound		
1	6.008	1.137	3.655	8.361		
2	4.289	1.021	2.178	6.401		

3. hand * task 🗲

				95% Confid	ence Interval
hand	task	Mean	Std. Error	Lower Bound	Upper Bound
1	1	6.326	1.352	3.530	9.123
	2	3.662	.967	1.662	5.662
2	1	5.690	1.122	3.369	8.011
	2	4.916	1.248	2.334	7.499

This table shows descriptives for each of the conditions of the study. Thus, the first row gives details of performance when participants were tapping with their right hand while memorising words. The bottom row gives details of performance when participants were tapping with their left hand while memorising the positions of the words.

'task' and 'hand*task' were moved into the

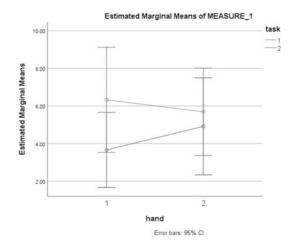
box labelled **Display Means for**.

This table shows descriptives for each level of the factor 'hand', collapsed across the other factor 'task'. We used the code 1 = right and 2 =left, so the first row is for the right hand, and the second row is for the left hand.

This table shows descriptives for each level of the factor 'task' collapsed across the two levels of the factor 'hand'.

The interaction graph shown below was obtained by clicking on the **Plots** button at the bottom of the **Repeated Measures** dialogue box.

The labels and title are not helpful, so double-click on the graph in the SPSS output to go into the **Chart Editor** dialogue box. Now double-click on the labels and title to edit them.



The lines on the graph can help you interpret your interaction. If the lines are converging or diverging (as is the case in this example), it suggests there may be an interaction between the factors. This is because it illustrates that there is a different pattern of scores across the levels of one IV, for each level of another (you will need to look at the ANOVA statistics to find out whether this interaction is significant or not). If the lines are parallel, it is unlikely that the factors are interacting.

When reporting a significant interaction, try to put what you see in the graph into words, linking the patterns you see in the different DV scores to the different factors and levels. For example, start by looking at the lines. Each line represents one condition of one IV (in this case, task type). If a line is sloping upwards it means that the scores in that condition are increasing across the levels of the other IV (in this case, increasing across right and left hand conditions); if a line is sloping downwards, then the opposite is true. As increased scores mean more disruption to tapping, in this case we could say something like:

- Right-hand tapping (hand 1) was more disrupted by memorising words (task 1) than was left-hand tapping (hand 2).
- Left-hand tapping was more disrupted by memorising the position of the words than was right-hand tapping.

Next, look at the groups of points on the graph: Are they closer together for some conditions than others? The degree of overlap between the error bars give an impression of how different (or not) participant performance was on tasks 1 and 2 (i.e. remembering the word vs the position) according to the different levels of the second variable (i.e. whether they were tapping with their right or left hand). This time we could say something like:

- Right-hand tapping (hand 1) was more disrupted by memorising words (task 1) than remembering their position (task 2).
- Left-hand tapping was similarly disrupted by both tasks.

Calculating eta squared: one measure of effect size

As with one-way within-subjects ANOVA, we can calculate eta squared to measure effect size. We need to do this for each factor, and for the interaction term. Again, eta squared (η^2) is calculated by dividing the sum of squares for the IV (list) by the total sum of squares. However, as the total sum of squares isn't given in the **Tests of Within-Subjects Effects**, we need to calculate it by hand for each IV and interaction by adding the appropriate sum of squares to the related Error sum of squares.

Source		Type III Sum of Squares	df	Mean Square	F	Sig.
hand	Sphericity Assumed	2.295	1	2.295	.133	.719
	Greenhouse-Geisser	2.295	1.000	2.295	.133	.719
	Huynh-Feldt	2.295	1.000	2.295	.133	.719
	Lower-bound	2.295	1.000	2.295	.133	.719
Error(hand)	Sphericity Assumed	398.267	23	17.316		
	Greenhouse-Geisser	398.267	23.000	17.316		
	Huynh-Feldt	398.267	23.000	17.316		
	Lower-bound	398.267	23.000	17.316		
task	Sphericity Assumed	70.906	1	70.906	1.955	.175
	Greenhouse-Geisser	70.906	1.000	70.906	1.955	.175
	Huynh-Feldt	70.906	1.000	70.906	1.955	.175
	Lower-bound	70.906	1.000	70.906	1.955	.175
Error(task)	Sphericity Assumed	834.213	23	36.270		
	Greenhouse-Geisser	834.213	23.000	36.270		
	Huynh-Feldt	834.213	23.000	36.270		
	Lower-bound	834.213	23.000	36.270		
hand * task	Sphericity Assumed	21.441	1	21.441	4.807	.039
	Greenhouse-Geisser	21.441	1.000	21.441	4.807	.039
	Huynh-Feldt	21.441	1.000	21.441	4.807	.039
	Lower-bound	21.441	1.000	21.441	4.807	.039
Error(hand*task)	Sphericity Assumed	102.585	23	4.460		
	Greenhouse-Geisser	102.585	23.000	4.460		
	Huynh-Feldt	102.585	23.000	4.460		
	Lower-bound	102.585	23.000	4.460		

Tests of Within-Subjects Effects

For 'hand' $(\eta^2) = 2.295 / (2.295 + 398.267)$

= 2.295 / 400.562= .006 For 'hand' (η^2) = 70.906 / (70.906 + 834.213) = 70.906 / 905.119 = .08 For 'hand' (η^2) = 21.441 / (21.441 + 102.585) = 21.441 / 124.026 = .17



Reporting the results

In a report you might write:

A two-way within-subjects ANOVA was conducted on decrement of finger tapping speed. The main effect of tapping hand was not significant: F(1,23) = 0.13, p = .719, $\eta^2 = .006$. The main effect of type of task was not significant: F(1,23) = 1.96, p = .175, $\eta^2 = .08$. There was a significant interaction between tapping hand and type of task: F(1,23) = 4.81, p = .039, $\eta^2 = .17$. This interaction is displayed in the graph above, suggesting memorising words disrupted right-hand tapping more than left-hand tapping. The opposite pattern was observed when participants memorised the position of the words. While memorising the words was generally more disruptive to hand tapping than memorising position, this was more pronounced in the right hand condition.

You would also want to report your descriptive statistics (e.g. means and standard deviations) and information regarding the confidence intervals for each condition in your results section.

If you wanted to break down the interaction to find out exactly where the significant differences lie, you could do this using follow-up paired *t*-tests (see Chapter 5). To find out whether performance on the two tasks was significantly different for the right-hand tapping condition, you would need to compare conditions h1s1 and h1s2. You could then do the same for the left hand (h2s1 vs h2s2). You could also explore whether there was a difference in the performance of the right and left hand for (a) the memorising word condition (h1s1 vs h2s1), and (b) the memorising position task (h1s2 vs h2s2). However, when reporting multiple follow-up *t*-tests you would need to carry out a correction to your alpha level (i.e. the level at which we accept significance, in this case p < .05) to avoid Type 1 error. The simplest way to do this is to manually apply a Bonferroni correction by dividing the alpha level by the number of comparisons you are carrying out (.05/4 = .0125). However, there are other methods that may be more suitable, so consult your statistics textbook about how best to achieve this.

Section 4: MIXED ANOVA

Here, we show you how to perform an ANOVA that involves between- and withinsubjects factors in the same experiment. To do this, we will use a study employing a three-way mixed design, which will also give you an idea of how to interpret any ANOVA with more than two IVs.

Example study: perceptual expertise in the ownage bias

As adults, we are experts at face recognition. However, there are some faces which we are more expert at recognising than others. For example, over 50 years of research has demonstrated that we are better at recognising faces belonging to our own racial or ethnic group compared with those of a different, less familiar race or ethnicity. More recently, research has suggested that we may also display an 'own-age bias'; being better at recognising faces belonging to our own age group compared with others. Why these biases occur is unclear, but there is some evidence to suggest differential contact may have a role to play. In other words, it may be that we are better at recognising own-group faces as a result of the increased experience we tend to have with those faces. Exactly what perceptual/cognitive mechanisms may be responsible for producing such an effect is unclear, however, one prominent explanation suggests our perceptual face-processing mechanisms may become better tuned for the groups of faces with which we have more experience. Specifically, it has been suggested that this increased in-group contact may lead to an increased ability to extract the configural (or holistic) information from in-group faces, while out-group faces are processed in a more piecemeal featural manner.

The following example explores the role of contact and perceptual expertise in the own-age bias using a paradigm that has been repeatedly used to illustrate the relationship between configural processing and expertise: the inversion effect. That is, the finding that the more expertise we have with a group of stimuli, the more inversion (turning it upside-down) disrupts our ability to process that stimuli configurally (or 'expertly'). Therefore, if the own-age bias is the result of greater expertise with own-age faces, one would predict a greater disruption of own-age compared with other-age face recognition following inversion.

In the following study, Harrison (2011) investigated the role of expertise in the own-age bias by looking at the effect of inversion on adult faces and children's faces with children (child face experts, adult novices), 'low contact' young adults (adult experts, child face novices) and 'high contact' teachers (adult and child face experts). Specifically, they showed participants 64 faces (32 own-age; 32 other-age) and measured how accurate participants were at recognising them. Half were presented upright, and half were upside-down. If perceptual expertise is responsible for own-age bias, one would expect novices (in this case, children and 'low contact' adults) to display the classic own-age bias for upright faces, but for this difference to disappear when the faces were inverted. The inversion cost (decrease in accuracy following inversion) would be expected to affect own-age faces more than other-age faces. For the teacher 'high contact' expert group, no own-age bias would be expected in upright faces (as they are experts with both own-age and children faces), and inversion should affect both groups of faces equally.

A 3*2*2 mixed design was used, with one between-subjects factor: group (three levels: children, 'low contact' adults and 'high contact' teachers); and two repeated measures factors: face age in photograph (two levels: child, 8–11 years old, and young adult, 18–25 years old) and orientation (two levels: upright and inverted faces). The dependent variable was percentage accuracy (measuring how many faces participants correctly recognised).

Factor 1	Face age			
Levels	Adult		Child	
Factor 2	Orientation		Orientation	
Levels	Upright	Inverted	Upright	Inverted
Column name, SPSS data file, for conditions	fa1o1	fa1o2	fa2o1	fa2o2

Table 9.2 The numbering system for within-subjects factors

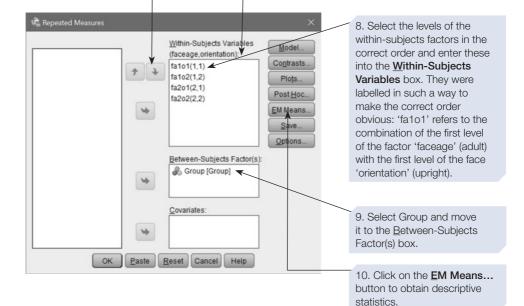
For the purposes of this book, we have created a data file that will reproduce some of the findings of this study. The data file is available from macmillanihe.com/harrison-spss-7e. In the data file, the columns holding the data for the combination of levels of the two within-subjects factors have been named using the numbering systems described in the previous Section (see Table 9.2).

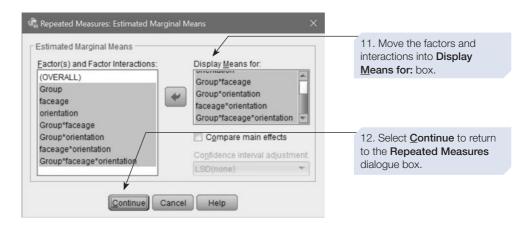
How to do it

ur. In	version.	sav [Data	Set1] - IB	M SPSS Statisti	cs Data Ed	litor			1	. Click on the	word <u>Analyze</u> .		
Eile	<u>E</u> dit	⊻iew	<u>D</u> ata	Transform	Analyze	<u>G</u> raphs	<u>U</u> tilities	Extensions	Window	<u>H</u> elp			
		*	Group	ADUAC	<u>B</u> aye Ta <u>b</u> le	criptive Stat sian Statis	tics	> > > >		Click on <u>G</u> er odel.	eral Linear		
	1 1 68.7 2 1 81.2 3 1 75.0			eral Linear	1	*	Univaria	ata					
3			Generalized Linear Models				ultivariate						
	4			81.2	- <u>m</u> R			Repeat	Repeated Measures				
	5	1		75.0	1.1.1	Regression			Varianc	variance Components			
)	6			75.0									
	7	1		87.5		Loglinear 🕨							
	8			68.7	Neur	Neural Networks			- 3	3. Click on Repeated			
	9			81.2	Clas					— .			
10 1 87.5		Dime	Dimension Reduction			Measures. The Repeated							
1	11			81.2	Sc <u>a</u> le	Scale •			Measures Define Factor(s)				
	10	Î		02.7	Nonr	parametric	Tests	•		dialogue box will appear (see below).			

🖷 Repeated Measures Define Factor(s) 🛛 🗙	
Within-Subject Factor Name: faceage	4. Change the factor name suggested by SPSS by highlighting factor1 and typing the word that represents the first within-subjects factor,
	'faceage'.
Add Change Remove	5. Type the number '2' in the Number of Levels box and click on the Add button, faceage(2) will appear in the box next to Add (see below).
Measure Name:	
Add Change Remove	
Define Reset Cancel Help	
🕼 Repeated Measures Define Factor(s) 🛛 🗡	
Within-Subject Factor Name: orientation Number of Levels: 2 faceage(2)	6. Repeat steps 4 and 5 for the second within-subjects factor, i.e. type in the name that represents the second factor 'orientation' in the <u>Within-Subject Factor Name</u> : box and type '2' in the Number of Levels box and click on the <u>A</u> dd button.
Within-Subject Factor Name: orientation Number of Levels: 2	second within-subjects factor, i.e. type in the name that represents the second factor 'orientation' in the <u>Within-</u> Subject Factor Name: box and type '2' in the Number of Levels box and click on the
Within-Subject Factor Name: orientation Number of Levels: 2 Add Change	second within-subjects factor, i.e. type in the name that represents the second factor 'orientation' in the <u>Within-</u> Subject Factor Name: box and type '2' in the Number of Levels box and click on the
Within-Subject Factor Name: orientation Number of Levels: 2 Add Change Remove	second within-subjects factor, i.e. type in the name that represents the second factor 'orientation' in the <u>Within-</u> Subject Factor Name: box and type '2' in the Number of Levels box and click on the

If, when following Step 8, you put one of the levels in the wrong place, you can move it by highlighting it and then clicking on the up or down arrow as appropriate. This key helps explain the numbers in the brackets below: 'faceage' refers to the first number and 'orientation' to the second number. This will help you enter the variables in the correct order into the <u>Within-Subjects</u> Variables box.





💼 Repeated Measures	Within-Subjects Variables	×	
*+	(faceage,orientation): fa1o1(1,1) fa1o2(1,2) fa2o1(2,1) fa2o2(2,2)	Model Contrasts Plots Post Hoc	13. Click on the Options button to obtain more statistics.
*	Between-Subjects Factor(s):	EM Means Save Options	
	& Group [Group]		
OK Paste E	Reset Cancel Help		

Repeated Measures: Options ×	14. Click here to check for assumption of homogeneity of variance.
Descriptive statistics Tansformation matrix Estimates of effect size Homogeneity tests Observed power Spread vs. level plot Parameter estimates Residual plot SSCP matrices Lack of fit Residual SSCP matrix General estimable function	Vanance.
Significance level: 05 Centidence intervals are 95.0 %	15. Click on the <u>Continue</u> button to return to the Repeated Measures dialogue box.
If you want partial eta squared, a measure of effect size, select <u>Estimates of effect size</u> .	

Click on \fbox{ok} , SPSS will calculate the ANOVA and produce the output that is explained below.

SPSS output for three-way mixed ANOVA Obtained using menu items: General Linear Model > Repeated Measures

General Linear Model

Within-Subjects Factors

Measure: MEASURE_1

faceage	orientation	Dependent Variable
1	1	fa1o1
	2	fa1o2
2	1	fa2o1
	2	fa2o2

If you requested **Descriptive statistics** in the **Repeated Measures: Options** dialogue box, the Descriptive Statistics table would appear after these first two tables. Instead, we moved all factors and interactions into the box labelled **Display Means for** in the **Repeated Measures Options** dialogue box, which produces tables that appear at the end of the output.

Between-Subjects Factors

		Value Label	Ν
Group	1	High Contact	32
	2	Low Contact	32
	3	Children	32

Box's Test of Equality of Covariance Matrices^a

Box's M	27.826
F	1.305
df1	20
df2	31046.084
Sig.	.163

Tests the null hypothesis that the observed covariance matrices of the dependent variables are equal across groups.

> a. Design: Intercept + Group Within Subjects Design: faceage + orientation + faceage * orientation

This table appears because we selected Homogeneity tests in the Repeated Measures Options dialogue box. This is checking that for each level of the betweensubjects factor, the pattern of correlations among the levels of the within-subjects factor are the same. Check that this is not significant.

6
<u>m</u>
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Effect		Value	F	Hypothesis df	Error df	Sig.
faceage	Pillai's Trace	.006	.600 ^b	1.000	93.000	.441
	Wilks' Lambda	.994	.600 ^b	1.000	93.000	.441
	Hotelling's Trace	.006	.600 ^b	1.000	93.000	.441
	Roy's Largest Root	.006	.600 ^b	1.000	93.000	.441
faceage * Group	Pillai's Trace	.230	13.881 ^b	2.000	93.000	.000
	Wilks' Lambda	.770	13.881 ^b	2.000	93.000	.000
	Hotelling's Trace	.299	13.881 ^b	2.000	93.000	.000
	Roy's Largest Root	.299	13.881 ^b	2.000	93.000	.000
orientation	Pillai's Trace	.798	366.353 ^b	1.000	93.000	.000
	Wilks' Lambda	.202	366.353 ^b	1.000	93.000	.000
	Hotelling's Trace	3.939	366.353 ^b	1.000	93.000	.000
	Roy's Largest Root	3.939	366.353 ^b	1.000	93.000	.000
orientation * Group	Pillai's Trace	.126	6.674 ^b	2.000	93.000	.002
	Wilks' Lambda	.874	6.674 ^b	2.000	93.000	.002
	Hotelling's Trace	.144	6.674 ^b	2.000	93.000	.002
	Roy's Largest Root	.144	6.674 ^b	2.000	93.000	.002
faceage * orientation	Pillai's Trace	.005	.433 ^b	1.000	93.000	.512
	Wilks' Lambda	.995	.433 ^b	1.000	93.000	.512
	Hotelling's Trace	.005	.433 ^b	1.000	93.000	.512
	Roy's Largest Root	.005	.433 ^b	1.000	93.000	.512
faceage * orientation *	Pillai's Trace	.139	7.498 ^b	2.000	93.000	.001
Group	Wilks' Lambda	.861	7.498 ^b	2.000	93.000	.001
	Hotelling's Trace	.161	7.498 ^b	2.000	93.000	.001
	Roy's Largest Root	.161	7.498 ^b	2.000	93.000	.001

Multivariate Tests^a

a. Design: Intercept + Group

Within Subjects Design: faceage + orientation + faceage * orientation

b. Exact statistic

For information about these two tables, see Chapter 8 Section 3. The **Mauchly's Test of Sphericity** is only important if there are more than two levels of either within-subjects factors. If this is not significant, the assumption of sphericity has not been violated.

Mauchly's Test of Sphericity^a

Measure: MEASURE_1

						Epsilon ^b			
Within Subjects Effect	Mauchly's W	Approx. Chi- Square	df	Sig.	Greenhouse- Geisser	Huynh-Feldt	Lower-bound		
faceage	1.000	.000	0		1.000	1.000	1.000		
orientation	1.000	.000	0		1.000	1.000	1.000		
faceage * orientation	1.000	.000	0	52	1.000	1.000	1.000		

Tests the null hypothesis that the error covariance matrix of the orthonormalized transformed dependent variables is proportional to an identity matrix.

a. Design: Intercept + Group

Within Subjects Design: faceage + orientation + faceage * orientation

b. May be used to adjust the degrees of freedom for the averaged tests of significance. Corrected tests are displayed in the Tests of Within-Subjects Effects table.

Tests of Within-Subjects Effects

Source		Type III Sum of Squares	df	Mean Square	F	Sig.
faceage	Sphericity Assumed	49.235	1	49.235	.600	.441
	Greenhouse-Geisser	49.235	1.000	49.235	.600	.441
	Huynh-Feldt	49.235	ares df Mean Square F 9.235 1 49.235 6.600 9.235 1.000 49.235 6.600 9.235 1.000 49.235 6.600 9.235 1.000 49.235 6.600 9.235 1.000 49.235 6.600 9.235 1.000 49.235 6.600 9.235 1.000 1139.323 13.881 8.646 2.000 1139.323 13.881 8.646 2.000 1139.323 13.881 8.646 2.000 82.076 1 3.057 93.000 82.076 1 3.057 93.000 82.076 1 8.923 1.000 14228.923 366.353 8.923 1.000 14228.923 366.353 8.923 1.000 14228.923 366.353 8.923 1.000 259.196 6.674 8.923 2.000 259.196 6.674 <t< td=""><td>.600</td><td>.441</td></t<>	.600	.441	
	Lower-bound	49.235	1.000	49.235	.600	.441
faceage * Group	Sphericity Assumed	2278.646	2	1139.323	35 .600 35 .600 35 .600 35 .600 35 .600 35 .600 35 .600 35 .600 35 .600 23 13.881 23 13.881 23 13.881 23 13.881 23 13.881 23 366.353 23 366.353 23 366.353 23 366.353 23 366.353 23 366.353 23 366.353 23 366.353 23 366.3674 96 6.674 96 6.674 97	.000
	Greenhouse-Geisser	2278.646	2.000	1139.323	13.881	.000
	Huynh-Feldt	2278.646	2.000	1139.323	13.881	.000
	Lower-bound	2278.646	2.000	1139.323	13.881	.000
Error(faceage)	Sphericity Assumed	7633.057	93	82.076		
	Greenhouse-Geisser	7633.057	93.000	82.076		
	Huynh-Feldt	7633.057	93.000	82.076		
	Lower-bound	7633.057	93.000	82.076		
orientation	Sphericity Assumed	14228.923	1	14228.923	366.353	.000
	Greenhouse-Geisser	14228.923	1.000	14228.923	366.353	.000
	Huynh-Feldt	14228.923	1.000	14228.923	366.353	.000
	Lower-bound	14228.923	1.000	14228.923	366.353	.000
orientation * Group	Sphericity Assumed	518.392	2	259.196	6.674	.002
	Greenhouse-Geisser	518.392	2.000	259.196	6.674	.002
	Huynh-Feldt	518.392	2.000	259.196	6.674	.002
	Lower-bound	518.392	2.000	259.196	6.674	.002
Error(orientation)	Sphericity Assumed	3612.061	93	38.839		
	Greenhouse-Geisser	3612.061	93.000	38.839		
	Huynh-Feldt	3612.061	93.000	38.839		
	Lower-bound	3612.061	93.000	38.839		
faceage * orientation	Sphericity Assumed	14.648	1	14.648	.433	.512
	Greenhouse-Geisser	14.648	1.000	14.648	.433	.512
	Huynh-Feldt	14.648	1.000	14.648	.433	.512
	Lower-bound	14.648	1.000	14.648	.433	.512
faceage * orientation *	Sphericity Assumed	507.813	2	253.906	7.498	.001
Group	Greenhouse-Geisser	507.813	2.000	253.906	7.498	.001
	Huynh-Feldt	507.813	2.000	253.906	7.498	.001
	Lower-bound	507.813	2.000	253.906	7.498	.001
Error	Sphericity Assumed	3149.414	93	33.865		
(faceage*orientation)	Greenhouse-Geisser	3149.414	93.000	33.865		
	Huynh-Feldt	3149.414	93.000	33.865		
	Lower-bound	3149.414	93.000	33.865		

Measure: MEASURE_1

This table shows the outcome of trend tests (see Chapter 8, Section 3 for more information about these). In this case, as each within-subjects factor only has two levels, only linear tests can be carried out, and not quadratic. If one of your within-subjects factors has three or more levels, quadratic tests could also be done.

Tests of Within-Subjects Contrasts

Source	faceage	orientation	Type III Sum of Squares	df	Mean Square	F	Sig.
faceage	Linear		49.235	1	49.235	.600	.441
faceage * Group	Linear		2278.646	2	1139.323	13.881	.000
Error(faceage)	Linear		7633.057	93	82.076		
orientation		Linear	14228.923	1	14228.923	366.353	.000
orientation * Group		Linear	518.392	2	259.196	6.674	.002
Error(orientation)		Linear	3612.061	93	38.839		
faceage * orientation	Linear	Linear	14.648	1	14.648	.433	.512
faceage * orientation * Group	Linear	Linear	507.813	2	253.906	7.498	.001
Error (faceage*orientation)	Linear	Linear	3149.414	93	33.865		

Levene's Test of Equality of Error Variances^a

		Levene Statistic	df1	df2	Sig.
Jpright Adult Faces	Based on Mean	.177	2	93	.838
	Based on Median	.200	2	93	.819
	Based on Median and with adjusted df	.200	2	92.067	.819
	Based on trimmed mean	.241	2	93	.786
nverted Adult Faces	Based on Mean	.909	2	93	.406
	Based on Median	.721	2	93	.489
	Based on Median and with adjusted df	.721	2	91.954	.489
	Based on trimmed mean	.949	2	93	.391
pright Child Faces	Based on Mean	1.186	2	93	.310
	Based on Median	.758	2	93	.471
	Based on Median and with adjusted df	.758	2	88.168	.471
	Based on trimmed mean	1.059	2	93	.351
nverted Child Faces	Based on Mean	1.090	2	93	.341
	Based on Median	1.301	2	93	.277
	Based on Median and with adjusted df	1.301	2	91.664	.277
	Based on trimmed mean	1.087	2	93	.341

Tests the null hypothesis that the error variance of the dependent variable is equal across groups.

a. Design: Intercept + Group

Measure: MEASURE_1

Within Subjects Design: faceage + orientation + faceage * orientation

This table appears because we selected <u>Homogeneity tests</u> in the **Repeated Measures Options** dialogue box. If tests based **on the mean** are non-significant for all levels of the within-subjects factors, the assumption of homogeneity has been met. If any were significant, this would have implications for the accuracy of *F* for the between-subjects factor.

Estimated Marginal Means -

1. Group

			95% Confidence Interval			
Group	Mean	Std. Error	Lower Bound	Upper Bound		
High Contact	73.730	.987	71.771	75.690		
Low Contact	69.531	.987	67.572	71.491		
Children	54.590	.987	52.630	56.549		

2. faceage

			95% Confidence Interval				
faceage	Mean	Std. Error	Lower Bound	Upper Bound			
1	66.309	.718	64.882	67.735			
2	65.592	.749	64.106	67.079			

3. orientation

Measure: N	IEASURE_1		95% Confid	ence Interval
orientation	Mean	Std. Error	Lower Bound	Upper Bound
1	72.038	.609	70.829	73.246
2	59.863	.693	58.486	61.240

4. Group * faceage

Measure: MEASURE_1

MARANNA MEAGUIDE 4

				95% Confidence Interval			
Group	faceage	Mean	Std. Error	Lower Bound	Upper Bound		
High Contact	1	73.438	1.244	70.967	75.908		
	2	74.023	1.297	71.448	76.599		
Low Contact	1	73.145	1.244	70.674	75.615		
	2	65.918	1.297	63.343	68.493		
Children	1	52.344	1.244	49.873	54.814		
	2	56.836	1.297	54.261	59.411		

5. Group * orientation

Measure: MEASURE_1

				95% Confidence Interval			
Group	orientation	Mean	Std. Error	Lower Bound	Upper Bound		
High Contact	1	80.957	1.054	78.864	83.050		
	2	66.504	1.201	64.119	68.889		
Low Contact	1	76.074	1.054	73.981	78.168		
	2	62.988	1.201	60.603	65.373		
Children	1	59.082	1.054	56.989	61.175		
	2	50.098	1.201	47.712	52.483		

These tables appear because we moved all factors and interactions into the box labelled **Display Means for** in the **Repeated Measures Options** dialogue box.

The descriptives in the first three tables are for each level of each factor: 'group', 'faceage' and 'orientation'.

The next three tables show descriptives for each level of each two-way interaction, i.e. for each combination of levels of two of the factors (ignoring the third factor).

Should any main effect or interaction be significant, you would look here at the relevant table to interpret the results.

6. faceage * orientation

Measure: MEASURE_1

				95% Confidence Interval			
faceage	orientation	Mean	Std. Error	Lower Bound	Upper Bound		
1	1	72.591	.767	71.069	74.114		
	2	60.026	.882	58.274	61.778		
2	1	71.484	.823	69.850	73.119		
2	2	59.701	.931	57.851	61.550		

7. Group * faceage * orientation -

					95% Confid	ence Interval
Group	faceage	orientation	Mean	Std. Error	Lower Bound	Upper Bound
High Contact	1	1	80.469	1.328	77.832	83.106
		2	66.406	1.528	63.372	69.441
	2	1	81.445	1.425	78.615	84.276
		2	66.602	1.613	63.399	69.804
Low Contact	1	1	81.445	1.328	78.808	84.082
		2	64.844	1.528	61.809	67.878
	2	1	70.703	1.425	67.872	73.534
		2	61.133	1.613	57.930	64.335
Children	1	1	55.859	1.328	53.222	58.496
		2	48.828	1.528	45.794	51.862
	2	1	62.305	1.425	59.474	65.135
		2	51.367	1.613	48.165	54.570

The final matrix table provides descriptives for each level of the three-way interaction and hence for each of the twelve conditions.

Post Hoc Tests

Group

The final section is for our Post Hoc tests. This table displays the **Multiple Comparisons** matrix table. You have to pick out the comparisons required, and ignore the repetitions.

Multiple Comparisons

Measure: MEASURE_1

Bonferroni		

		Mean Difference (I-			95% Confid	ence Interval
(I) Group	(J) Group	J)	Std. Error	Sig.	Lower Bound	Upper Bound
High Contact Low Conta		4.1992	1.39540	.010	.7972	7.6013
	Children	19.1406	1.39540	.000	15.7386	22.5427
Low Contact	High Contact	-4.1992	1.39540	.010	-7.6013	7972
	Children	14.9414	1.39540	.000	11.5393	18.3435
Children	High Contact	-19.1406	1.39540	.000	-22.5427	-15.7386
	Low Contact	-14.9414	1.39540	.000	-18.3435	-11.5393
Desident and a					1	

Based on observed means.

The error term is Mean Square(Error) = 31.154.

*. The mean difference is significant at the .05 level.

As our factor had three levels, there are three possible comparisons:

• High Contact vs Low Contact

- High Contact vs Children
- Low Contact vs Children

All comparisons are significant.



A three-way mixed ANOVA was performed on participants' face recognition accuracy scores. While there was no significant main effect of face age (*F* < 1), there was a significant main effect of orientation (*F* (1,93) = 366.35, p < .001). There was also a significant main effect of group on accuracy scores (*F* (2,93) = 103.96, p < .001). Post-hoc Bonferroni tests revealed that this was a result of the children performing significantly worse than both of the adult groups, and the high-contact adult group performing significantly more accurately than the low-contact adult group (p < .001 in all cases).

There were also a couple of higher-order interactions. While the face age * orientation interaction was not significant (*F* < 1), there was a significant interaction between group and orientation (*F* (2,93) = 6.67, *p* = .002). Furthermore, face age significantly interacted with group (*F* (2,93) = 13.88, p < .001) with those in the low-contact group performing more accurately for young adult faces than the children's faces, while children and the high-contact group showed the opposite pattern.

A significant three-way interaction between group, face age and orientation was also observed (F(2, 93) = 7.50, p = .001). The low-contact group showed a larger inversion deficit for young adult faces than children faces; the children showed the opposite; and the high-contact group was similarly affected by inversion for both face age groups.

You would also want to report your descriptive statistics and information regarding the confidence intervals for each condition. Interactions are often best illustrated with Graphs (see Chapter 3).

If you want to break down the significant interactions to find out what is driving them, you can do this using follow-up *t*-tests (see Chapter 5). For example, for each group you might want to run paired *t*-tests for each face age to see whether inversion significantly impacts recognition. As one of the hypotheses predicted that children and low-contact adults should show a typical own-age bias for upright faces, while 'high contact' teachers may not, you might also want to run a separate two-way ANOVA on upright trials only. However, when reporting multiple follow-up tests you might need to carry out a correction to your alpha level (i.e. the level at which we accept significance, in this case p < .05) to avoid Type 1 error. See tip box at the end of Section 3 above and consult your statistics textbook about how best to achieve this.

Summary

- This chapter introduced you to the different types of Factorial ANOVA for between-subjects, within-subjects and mixed designs.
- These tests of difference are used for experimental designs involving more than one IV.
- Appropriate descriptive statistics, the mean and standard deviation, can be obtained either by following the advice in Chapter 3 or by selecting the appropriate options on the ANOVA Options dialogue box (see Chapter 8).
- Error bar charts and interaction graphs are often used to display the statistically significant findings.
- For guidance on incorporating SPSS output into a report, or on printing the output, see Chapter 14.

Multiple regression

In this chapter

- An introduction to multiple regression
- Standard or simultaneous method of multiple regression
- Sequential or hierarchical method of multiple regression
- Statistical methods of multiple regression



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Section 1: AN INTRODUCTION TO MULTIPLE REGRESSION

- Regression was introduced in Chapter 6, Sections 7 and 8, and we advise you to reread those sections before starting this chapter.
- Bivariate regression, also covered in Chapter 6, involves one dependent variable, which we term the *criterion variable*, and only one independent variable, which we refer to as the *predictor variable*.
- Multiple regression involves a criterion variable and two or more predictor variables.
- The predictor variables can be measured using a range of scales (although ideally at interval or ratio level), but the criterion variable should be measured using a ratio or interval scale.
- Although it is not possible to produce totally accurate predictions, multiple regression allows us to identify which set of the predictor variables together provides the best prediction of the score on the criterion variable.
- As with bivariate correlation and regression, multiple regression does not imply causal relationships unless variables have been manipulated.
- Regression can be used to model and simplify data we have collected, in a way that allows prediction of future cases.

From bivariate to multiple

Bivariate correlation measures the strength of association between two variables, and bivariate regression allows one variable to be predicted by one other variable. We can extend this. Multiple correlation measures the strength of association between one variable and a set of other variables, and multiple regression allows prediction of one variable (the criterion variable) from the set of other variables (the predictor variables). In practice, the term *multiple regression* usually includes multiple correlation.

In multiple regression, the criterion variable is still represented by *Y*, and the predictor variables by *X*. Now that there are two or more predictors, we use subscripts to identify them: X_1 , X_2 , X_3 , etc., up to X_k , where k = number of predictors.

Having more than one predictor variable is useful when predicting human behaviour, as our actions, thoughts and emotions are all likely to be influenced by some combination of several variables. The advantage of applying multiple regression instead of several bivariate correlations (between the criterion variable and each of the predictor variables) is that multiple regression corrects for the correlations among the predictor variables. The bivariate correlation between a particular variable (X_1) and the criterion variable may be partly, or only, due to the correlation between X_1 and another predictor variable (X_2) (see the ice cream and temperature example in Chapter 6, Section 1). Using multiple regression we can test theories (or models) about precisely which set of variables best predict the criterion variable. There are a few different methods for conducting multiple regression, which can be used to apply different models. We will discuss these methods below. An excellent text is that by Miles and Shevlin (2001).

An example

Consider our example in Chapter 6, Section 7, of predicting how confident in using SPSS students will be after completing a module (we shall refer to this as 'SPSS confidence'). For bivariate regression, we suggested that time spent practising could be a predictor. Other variables can also be considered; attendance at lectures, attendance at practical sessions, time spent reading SPSS books, anxiety about using computers before the module and total study time might all contribute towards SPSS confidence. We can propose a model that suggests that this set of predictor variables will predict SPSS confidence. Next, we collect data, perhaps by surveying a module with 200 or 300 students, in order to test the model in terms of how many and which variables give rise to the most accurate predicted by time spent practising, attendance at practical sessions, time spent reading SPSS books and computer anxiety before the module, with the other variables we measured not helping us to predict SPSS confidence. Note that we would predict a positive effect for most of those variables, but a negative effect for computer anxiety before the module.

How does multiple regression relate to analysis of variance?

What we are doing in both ANOVA and multiple regression is seeking to account for the variance in the scores we observe. Thus, in the example above, people might vary greatly in their levels of SPSS confidence. Some of this variance will be accounted for by the variables we have identified. For example, we might be able to say that time spent practising accounts for a fairly large percentage of the variance in SPSS confidence; hence, it is useful to know how much time a student spends practising when trying to predict their SPSS confidence. These concepts are rather similar to those underlying ANOVA. In ANOVA, we determine how much of the variance is accounted for by the IVs (relative to the percentage of the variance we cannot account for). In multiple regression, we are not restricted to just a few levels of the IVs; instead we measure the variables and test how well they predict the score on the dependent variable (or criterion variable). Thus, ANOVA is actually a rather specific and restricted example of the general approach adopted in multiple regression.

Another way in which ANOVA is more restricted than multiple regression relates to whether the predictors/IVs are *orthogonal* or not. Two variables are orthogonal if they are uncorrelated. ANOVA designs are orthogonal if all the IVs are categorical and there are equal numbers of observations in each cell of the design. This design, combined with good experimental control, has the great advantage of allowing clear causal statements about effects on the DV. It cannot, however, be achieved with many variables of interest in psychology. Multiple regression allows for the common, *nonorthogonal*, situation, in which predictors/IVs are correlated with each other. We return to this issue below, when discussing the various multiple regression methods.

The similarities between multiple regression and ANOVA (and also between bivariate correlation and the *t*-test) stem from the fact that these statistical techniques all estimate how much of the variance in one variable is explained by one or more other variables. These other variables might be manipulated directly in the case of controlled experiments (allowing causal relationships to be tested), or be naturally occurring in non-experimental studies, but the underlying principle is the same. Each of these procedures is usually treated separately, but they are fundamentally all the same procedure. This underlying single approach is called the General Linear Model, a term we first introduced when undertaking ANOVA in Chapter 8, Section 1.

Causation

Whether we can draw conclusions about cause-and-effect relationships depends on whether variables were manipulated or simply measured, and whether there was random allocation of participants to conditions. In ANOVA, we often directly manipulate the factors, meaning that random allocation is possible, and then we measure the resulting change in the dependent variable. Sometimes, however, the levels of the factors are chosen from existing groups (for example, if we compare students educated to degree level or not), and so random allocation is not feasible and causal relationships cannot be assumed. In multiple regression, we normally measure the naturally occurring scores on a number of predictor variables and try to establish which set of the observed variables gives rise to the best prediction of the criterion variable. In such studies, a significant outcome does not imply causation. If one can, however, manipulate a predictor variable used in a multiple regression, then conclusions about causation could be drawn (provided all the normal controls for experimental research are applied). If no variable has been directly manipulated, then we might speculate that a causal relationship exists, but we would need to design a new study to test that hypothesis.

CHAPTER 10

When should I use multiple regression?

- 1. You can use this statistical technique when exploring linear (straight line) relationships between the predictor and criterion variables. If any relationship is nonlinear it might be necessary to transform one or more of the variables.
- 2. The criterion variable should be measured on a scale at either interval or ratio level. If the criterion variable is nominal then other procedures must be employed (see Chapter 12).
- 3. The predictor variables you select should be measured on a ratio, interval or ordinal scale. A nominal predictor variable is legitimate, but only if it is dichotomous. For example, if we are interested in whether people live in a city or not, then that variable has two values and can be used in a multiple regression. If, however, we classify their environment as city, town or village, then that could not be entered into multiple regression as a single variable. Instead, you would create three different variables, each with two categories (city/not city; town/not town and village/ not village). The term 'dummy variable' is used to describe this type of dichotomous variable.
- 4. Multiple regression requires a large number of observations. The estimate of R (the multiple correlation, see below) depends on the number of predictor variables and the number of participants (N). One rule of thumb is to have at least ten times as many participants as predictor variables. Tabachnick and Fidell (2014) suggest that, to test both overall correlation and the effect of individual predictors, N should equal the greater of the following: either the number of predictors times 8 plus 50 or the number of predictors plus 104.
- 5. You should screen your data for outliers, normality and homoscedasticity of residuals. [See Tabachnick and Fidell (2014) for guidelines.] SPSS provides you with a means of checking residuals, which we describe below.
- 6. You should check for multicollinearity. When choosing a predictor variable, you should select one that might be correlated with the criterion variable, but that is not strongly correlated with the other predictor variables. Non-orthogonality, with correlations among the predictor variables, is not unusual. The term 'multi-collinearity' (or collinearity) is used to describe the situation when a high correlation is detected between two or more predictor variables. Such high correlations cause problems when trying to draw inferences about the relative contribution of each predictor variable to the success of the model. SPSS provides you with means of checking for collinearity, which we describe below.

The multiple regression equation

This equation allows us to predict the criterion variable *Y* from the set of predictor variables X_1 , X_2 , X_3 , X_4 , etc. It is an extension of the bivariate regression equation:

$$Y' = A + B_1 X_1 + B_2 X_2 + B_3 X_3 + \ldots + B_k X_k$$

where:

Y' is the predicted value of the criterion variable.

- B is the regression weight, or regression coefficient, for each predictor variable; B indicates how much Y' will change if that X changes by one unit. We return to regression coefficients below.
- k is the number of predictor variables.

When solving the equation, values of A and B are set so that:

- 1. The least squares criterion is used; as with bivariate regression, this means that $\Sigma(Y-Y')^2$ is at a minimum;
- 2. The correlation of *Y* with *Y*' is maximised; the multiple correlation, *R*, is the Pearson's correlation coefficient for *Y* with *Y*'. This is explained further below.

Regression coefficients: *B* (unstandardised) and beta (standardised)

Regression coefficients (or regression weights) are measures of how strongly each predictor variable influences the criterion variable if all the other predictor variables were held constant. B indicates the change in the measured units of the criterion variable for a change in one unit on the predictor variable (if all other predictors are held constant). With two or more predictors, we will be interested in which of them has most effect on the criterion variable; that is, which is the strongest predictor. As B is unstandardised (measured in the original units of its X), it can be difficult to interpret. For example, the effect on Y' from a change of one unit in 'course anxiety' cannot be directly compared with the effect of a change of one unit in 'time spent practising'. In addition, the variability of each variable will differ, and this will also affect the coefficients. A useful alternative is beta (β) , the *standardised* regression coefficient, which is measured in units of standard deviation, allowing us to more easily compare the influence of several predictors. For example, if $\beta = 2.5$ for one of the predictor variables, a change of one standard deviation in that variable will result in a change of 2.5 standard deviations in the criterion variable. Thus, a higher β value for one predictor variable indicates a greater impact of that predictor variable on the criterion variable; the β values for the different predictor variables are directly comparable.

A regression coefficient is either negative or positive, indicating whether an increase in the predictor will result in a decrease or increase in the criterion variable.

With only one predictor variable in the model, β is equivalent to the correlation coefficient between the predictor and the criterion variable. This equivalence makes sense, as this situation is a correlation between two variables. When you have more than one predictor variable, however, you cannot compare the contribution of each predictor variable by simply comparing the bivariate correlation coefficients. Each bivariate relationship may be affected by one or more of the other predictor variables. Multiple regression takes account of all the relationships, and β allows you to compare the strength of the relationships between predictor variables and the criterion variable.

The significance of each predictor in explaining the variance in the criterion variable is assessed using t, and we will highlight this in the annotated output in Section 2.

R, R-squared and adjusted R-squared

R is a measure of the correlation between the observed values of the criterion variable and its predicted values. In our example, this would be the correlation between the observed SPSS confidence score (reported by participants) and the SPSS confidence score predicted for them from the predictor variables. Unlike *r* or r_s for bivariate correlation, *R* can only take positive values even if all the individual correlations or β s are negative; this is because *R* is *not* the relationship between the criterion variable and predictor variables – instead *R* is the relationship between the observed and the predicted scores for the criterion variable.

R-squared (R^2) indicates the proportion of the variance in the criterion variable that is accounted for by the model; in our example, the proportion of the variance in the SPSS confidence accounted for by the set of predictor variables (time spent practising etc.). In essence, this is a measure of how well we can predict the criterion variable if we know the predictor variables.

However, R and R^2 tend to somewhat overestimate the success of the model. One reason for this is that R, unlike r, can never be negative (because it is the correlation of Y and Y'), and as all chance variations will be positive, they tend to increase R. Such random chance variations are greater in small samples. Thus, an adjusted R^2 value is calculated, which takes into account the number of predictor variables in the model and the number of observations (participants) that the model is based on. This adjusted R^2 value gives the most useful measure of the success of the model. If, for example, we have an adjusted R^2 value of .75, we can say that our model has accounted for 75% of the variance in the criterion variable. (This would be considered a large proportion of variance explained.) If there are few participants, then adjusted R^2 can be quite a bit smaller than R^2 (as chance variations are greater in small samples). The value of adjusted R^2 can even be negative, but this is an artefact and the convention is to report a negative adjusted R^2 as 0.

Whether R^2 is significant or not is assessed using *F*, and we will highlight this in the annotated output in Section 2.

Regression methods

The relative contribution of each predictor variable to explaining the variance in *Y* can be assessed by different methods that give rise to different models. For each method, the success of the model/s in predicting the criterion variable is assessed. The methods vary in the following ways:

- 1. Shared variance is treated in different ways; shared variance is an issue in nonorthogonal situations in which predictors share variance with each other as well as with the criterion variable.
- 2. Predictors can be entered as a single block or separately.

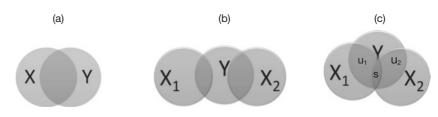


Figure 10.1 Venn diagrams

Notes: The overlapping areas illustrate:

- (a) Proportion of variance in Y explained by X in a bivariate correlation.
- (b) Proportion of variance in Y explained, in a multiple correlation, by each of two Xs, which are orthogonal with each other. All the explained variance is unique to either X₁ or X₂.
- (c) Proportion of variance in Y explained, in a multiple correlation, by each of two Xs, which are non-orthogonal; 'u' indicates the variance explained in Y that is unique to each X, and 's' indicates the variance explained in Y that is shared by X₁ and X₂.
- 3. When predictors are entered separately, the order in which they are entered into the regression equation can be specified in advance by selecting a specific method.

Unique and shared variance

Venn diagrams are commonly used to illustrate unique and shared variance:

- Figure 10.1(a) shows the situation in a bivariate correlation; the overlapping area represents the proportion of variance in *Y* explained by *X*; this is a simple situation, but adding even one more predictor often complicates the issue.
- Figure 10.1(b) illustrates a multiple correlation, in which each X explains some variance in Y, but does not share variance with the other. Thus, each X explains unique variance, and there will be no problem whichever regression method is used.
- Figure 10.1(c) illustrates a multiple correlation, in which each X explains some unique variance, but also shares some of the variance explained in Y with the other X. This is the non-orthogonal situation. See figure legend for further explanation

The way in which the different regression methods treat shared variance is explained next.

Simultaneous or standard method

In the simultaneous method, the researcher specifies the set of predictor variables in the model, and they are all entered into the model at the same time. This method is performed in SPSS by selecting all the predictor variables in one single block and using the option Enter. Each predictor is assessed on the variance it explains that is additional to the variance explained by all the other predictors combined. It is probably the safest method to use if you have no theoretical rationale that certain predictor/s may be more important than others. It has the drawback, however, that when the regression coefficients are calculated, only unique variance is attributed to each predictor variable. Shared variance, indicated by 's' in Fig. 10.1(c), is not attributed to

any predictor when the regression coefficients are calculated. Thus, a predictor that is, in fact, important may appear to account for only a small proportion of the variance because it shares a fair amount of variance with other predictors. All the variance explained in the criterion variable does, however, contribute to *R* and other statistics that summarise the whole model. So, the model as a whole may be quite strong even if some or all the predictor variables appear quite weak.

Sequential or hierarchical method

In the sequential or hierarchical method, the researcher enters predictor variables in a particular sequence or hierarchy determined by theoretical or empirical considerations. Unless you have reason to specify a particular order of entry, you should not use this method. Each predictor can be entered singly, or as part of a set of predictors, which means entering in several blocks. Note that SPSS calls each entry a 'block' whether it contains only one or more than one predictor variable. Each block specifies a different model, which is assessed for how well it predicts the criterion variable. First, let us consider single predictors. The first predictor to be entered is assigned all the variance that it explains in the criterion variable. Each subsequent predictor is only assigned its additional unique variance. Any shared variance remains with predictors higher up the hierarchy. As each predictor is entered into the model, its contribution is assessed. If the predictive power of the model is not significantly increased relative to the previous model, then that predictor is dropped. The process moves on to the next predictor and so on until all have been considered. The models can then be compared.

> Whether SPSS assigns shared variance as we have just outlined depends on which SPSS option you use when entering each predictor. Indeed, one option is to use **Enter** so that, as each predictor is added, the set of predictors in each successive model is analysed using the standard method; as described previously, shared variance will not be attributed to any predictor. This is explained further in Section 3.

In addition to entering predictors singly, we can enter two or more predictors at the same stage. The variance assigned to each block should follow the same hierarchical principle outlined above for predictors entered one at a time, and again this will depend on the SPSS option that you use. Within a block containing more than one predictor, the variance assigned to each of those predictors will follow the rules of the SPSS option that you choose for that block. That may be the **Enter** option (which follows the standard 'rules' explained above), but could be one of the statistical options we explain below. Using the **Forward** option seems to treat individually entered predictors in a way closest to the spirit of the sequential method, as it leaves shared variance with a predictor you entered earlier in the hierarchy. Alternatively, with the **Remove** option, you first **Enter** all the predictors and then **Remove** them from the model in blocks according to your rationale; the use of this option will not be further described here.

Statistical (or stepwise) methods

There are several statistical (also known as 'stepwise') regression methods. In all of them, the researcher specifies the set of predictor variables, but then the order in which they are entered into (or taken out of) the model is determined according to their strength and not according to any theoretical rationale. Thus, these methods are considered controversial (Tabachnick and Fidell, 2014, 174) and 'unwise regression' (see Howell, 2013, 540), although they may be useful in exploratory research. Guidance on validating results from these methods is given below. SPSS offers the following statistical options. The description here is for the situation when you include all the predictors in a single block:

- 1. In Forward (stepwise) selection, SPSS enters the predictors into the model one at a time in an order determined by the strength of their correlation with the criterion variable (the strongest is entered first). As each predictor is added, its effect is assessed, and if it does not significantly add to the success of the model, it is excluded.
- 2. In **Backward** (stepwise) selection, SPSS first enters all the predictors into the model. The weakest predictor (based on partial correlation with the criterion variable) is then removed and the regression recalculated. If the model was significantly weakened, then the predictor is re-entered otherwise it is deleted. This procedure is repeated until only useful predictors remain in the model.
- 3. The **Stepwise** option is the most sophisticated of the statistical methods. SPSS enters the predictors into the model one at a time in an order determined by the strength of their correlation with the criterion variable, and the strength of the predictor in adding to the predictive power is assessed. If adding the predictor contributes to the model, it is retained. So far, this is like the **Forward** option, but then all other predictors in the model are retested for their contributes significantly is removed. Thus, with this method, you end up with the smallest possible set of predictors included in your model.

Validating results from statistical regression methods

Statistical methods for multiple regression should be used with caution and only when you have a large number of cases. This is because minor variations in the data due to sampling error can have a large effect on the order in which variables are entered and therefore the likelihood of them being retained. If you decide to select a statistical method, then you should validate your results with a second independent set of data. This can be done either by conducting a second study, or, if you have sufficient data, by first randomly splitting your data set into two halves. Use **Select Cases** with **Random sample of cases** method (see Chapter 4, Section 4) and set sample size at 50%. A variable called 'filter_\$' will be added to your data file. To ensure that you run a separate analysis on each randomly selected half, proceed as follows:

- 1. Reset Select Cases to All cases.
- Use Select <u>Cases</u> with If <u>condition</u> is satisfied method to select those cases for which filter_\$ = 0 and run your analysis on the selected cases.
- 3. Use Select Cases with If condition is satisfied method to select those cases for which filter_\$ = 1 and run your analysis on those cases. Only results that are common to both analyses should be reported.

Section 2: STANDARD OR SIMULTANEOUS METHOD OF MULTIPLE REGRESSION

Example study: state anxiety

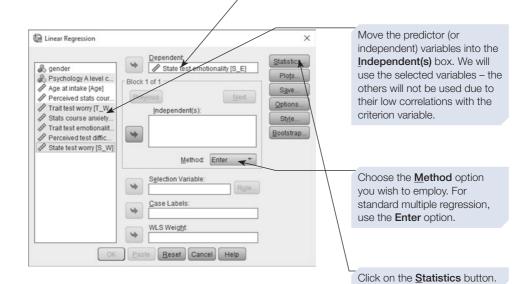
Rosemary Snelgar and colleagues Alan Porter and Tina Cartwright conducted a study into the the common phenomenon that students are anxious both about statistics and about tests, which might affect their performance on modules in which research methods and statistics are taught. We will use a subset of the data from this study to illustrate three methods of conducting a multiple regression: standard, sequential and stepwise. The reserchers measured both trait and state anxiety for two components of anxiety. The worry component includes such things as concern about the consequences of doing badly in the statistics course. The emotionality component of anxiety includes physiological responses, such as feeling 'jittery'. They also measured other variables suggested to be related to statistics and test anxiety. Trait test worry ('tw'), trait test emotionality ('te'), perceived statistics course difficulty ('pcd') and statistics course anxiety ('ca') were measured early in the module. Later, perceived test difficulty ('ptd') was measured just before a test, and state test worry ('sw') and state test emotionality ('se') were measured just after the test. Other variables included were whether or not students had psychology as one of their 'A' levels ('psycha'), because students without this often assume that they will struggle, age and gender. State test emotionality is the criterion variable. These data are available from macmillanihe.com/harrison-spss-7e.

The researchers first checked whether the data were skewed: age was positively skewed, as would be expected with a student population, and 'ca' just failed the skewness test (see Tabachnick and Fidell, 2014, 113), but and the other variables were not skewed. They also obtained correlation coefficients between state test emotionality and the other variables. The relationship of 'se' with gender, age, maths experience and perceived statistics course difficulty were all low and not significant; for our current purpose, these variables were omitted from the multiple regression.

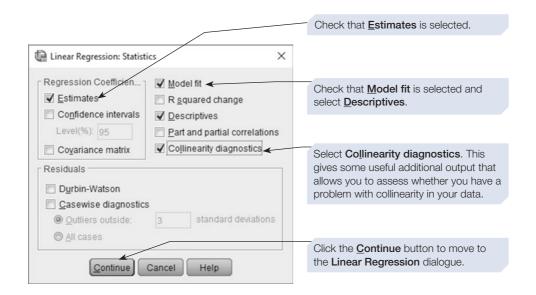
This analysis is illustrated below and the relevant output is annotated.

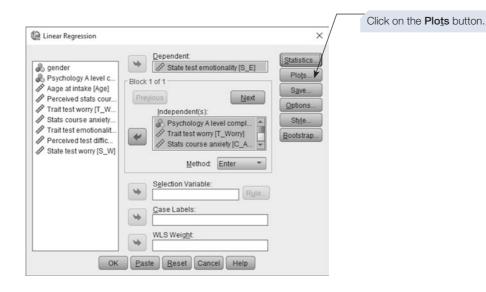
Analyze Graphs Utilities Exte	ensions	Window	<u>H</u> elp		
Reports Descriptive Statistics	;		14	۲	
Bayesian Statistics					
Tables		C Anx	T_Emot	PTD	
Compare Means	,		v ·	V 1 10	
General Linear Model		2.50	3.25	3	
Generalized Linear Models	,	3.75	4.00	3	 Click on <u>A</u> nalyze⇒ <u>R</u> egression⇒ <u>L</u> inea
Mixed Models	,	2.25	2.25	4	
Correlate		2.25	2.25	2	
Regression		200	1.00	3	
Loglinear			atic Linear Modeling	-4	
Neural Networks		Linear.		4	
Classify	,		Estimation	4	
Dimension Reduction		Partial	Lea <u>s</u> t Squares	2	
Scale	,	Binary I	Logistic	3	
Nonparametric Tests	,	Multino	mial Logistic	2	
Forecasting	,	Ordinal	I	3	
Survival		Probit.		2	
– Multiple Response	,	Nonline	ear	3	
Missing Value Analysis		Weight	Estimation	1	
Multiple Imputation			e Least Squares	2	
Complex Samples		Quantil		3	
Bimulation				3	
Quality Control	,	Optima	I Scaling (CATREG)	2	
Spatial and Temporal Modeling	,	2.50	2.25	3	
Direct Marketing		2.50	2.75	2	
IBM CBCC Amon		1.50	1.75	2	
IBM SP35 Amos		1.00	1.25	2	

Move the criterion (or dependent) variable into the Dependent box.



We are using 'State test emotionality' as criterion or dependent variable. Our predictor variables are 'Psychology 'A' level', 'Trait test worry', 'Trait test emotionality', 'Stats course anxiety', 'Perceived test difficulty' and 'State test worry'. Note that 'Psychology 'A' level' is a nominal variable, but it is dichotomous (completed or not completed), and so it is a suitable predictor for multiple regression.





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Linear Regression: Plots	Move *ZRESID (standardised residuals) into the <u>Y</u> box and *ZPRED (standardised predicted values) into the <u>X</u>
CEPENDNT Scatter 1 of 1 Previous Next	box to produce a scatterplot of these variables.
*ZRESID *DRESID *ADJPRED *SRESID	
*SDRESID	Select Normal probability plot for a graph that allows you to check normality.
 ☐ Histogram ☑ Normal probability plot 	
Continue Cancel Help	Click on <u>Continue</u> and then OK.

The SPSS multiple regression default is to **Exclude cases listwise**. Unless you change this (using the Options button), SPSS will only analyse data from participants who have no missing values. The alternatives are pairwise or replace with the mean, but the latter should be used with caution. This data set is complete with no missing values.

SPSS output for standard multiple regression

Obtained using menu items: Regression > Linear (method = enter)

> This first table is produced by the **Descriptives** option, and is useful for reporting your results.

Descriptive Statistics

	Mean	Std. Deviation	Ν
State test emotionality	2.5556	1.03775	90
Psychology A level completed	1.5333	.50168	90
Trait test worry	2.3139	.78572	90
Stats course anxiety	2.1889	.63744	90
Trait test emotionality	2.5722	.87510	90
Perceived test difficulty	2.8611	.71118	90
State test worry	2.7639	.75174	90

Remember that, for any nominal variables ('psycha' here), you should not report mean and SD, but counts (or frequencies) instead.

This second table gives details of the correlation between each pair of variables. We do not want strong correlations between the predictor variables. The values here are acceptable.

Correlations

	K	State test emotionality	Psychology A level completed	Trait test worry	Stats course anxiety	Trait test emotionality	Perceived test difficulty	State test worry
Pearson Correlation	State test emotionality	1.000	.212	.493	.500	.609	.562	.705
	Psychology A level completed	.212	1.000	.148	.173	.218	.155	.129
	Trait test worry	.493	.148	1.000	.504	.644	.201	.403
	Stats course anxiety	.500	.173	.504	1.000	.532	.387	.406
	Trait test emotionality	.609	.218	.644	.532	1.000	.267	.346
	Perceived test difficulty	.562	.155	.201	.387	.267	1.000	.529
	State test worry	.705	.129	.403	.406	.346	.529	1.000
Sig. (1-tailed)	State test emotionality		.022	.000	.000	.000	.000	.000
	Psychology A level completed	.022		.082	.051	.019	.073	.113
	Trait test worry	.000	.082		.000	.000	.029	.000
	Stats course anxiety	.000	.051	.000		.000	.000	.000
	Trait test emotionality	.000	.019	.000	.000		.006	.000
	Perceived test difficulty	.000	.073	.029	.000	.006		.000
	State test worry	.000	.113	.000	.000	.000	.000	
N	State test emotionality	90	90	90	90	90	90	90
	Psychology A level completed	90	90	90	90	90	90	90
	Trait test worry	90	90	90	90	90	90	90
	Stats course anxiety	90	90	90	90	90	90	90
	Trait test emotionality	90	90	90	90	90	90	90
	Perceived test difficulty	90	90	90	90	90	90	90
	State test worry	90	90	90	90	90	90	90

a. Dependent Variable: State test emotionality

b. Predictors: (Constant), State test worry, Psychology A level completed, Trait test emotionality, Perceived test difficulty, Stats course anxiety, Trait test worry

Va	ariables Ent	tered/Re	moved	a ^a			
Model	Variables Entered	Variak Remo		Method			This third table tells us ab
	State test worry, Psychology A level completed, Trait test emotionality, Perceived test difficulty, States course anxiety, Trait test worry ^b			Enter			the predictor variables an method used. Here we ca that all our predictor varia were entered simultaneou that is because we used Enter option.
	endent Variabl equested varia			onality			This table and the next ar important. R ² indicates th overall explanatory power
		Model S	Summ	ary ^b			the model. The Adjusted Square value tells us that model accounts for 66.05
Model	R	R Square		ljusted R Square	Std. Err the Esti		variance in the 'se' score
1	.827 ^a	.683		.660		60492	
cor Sta	edictors: (Cor mpleted, Trait ats course an	test emot xiety, Trait	tionality test wo	, Perceived te rry			This table reports an AN that assesses significance
b. De	pendent Vari			notionality			R^2 . The first df is the num
			NOVA ^a	*			predictors (<i>m</i>); the residu <i>N-m</i> -1 (here 90-6-1). <i>p</i> <
	S	um of	df	Mean Square	F	Sig.	so the model is significan
Model	S	quares					
	gression	65.475	6	10.913	29.822	.000 ^b	
1 Re			6 83	10.913 .366	29.822	.000 ^b	

CHAPTER 10

B (the unstandardised coefficient) for each predictor variable shows the predicted increase in the value of the criterion variable for a 1 unit increase in that predictor (while controlling for the other predictors). The next column gives the standard error of B.

		Unstandardize	d Coefficients	Standardized Coefficients		
Model		► B	Std. Error	Beta	t	Sig.
1	(Constant)	-1.394	.344		-4.050	.000
	Psychology A level completed	.073	.132	.035	.553	.582
	Trait test worry	.021	.113	.016	.190	.850
	Stats course anxiety	.049	.129	.030	.381	.704
	Trait test emotionality	.433	.102	.365	4.239	.000
	Perceived test difficulty	.304	.110	.208	2.757	.007
	State test worry	.614	.108	.445	5.671	.000

Coefficients^a

a. Dependent Variable: State test emotionality

The standardised coefficient, beta (β) , gives a measure of the contribution of the variable to the model in terms of standard deviations. β is the predicted change in SD of the criterion variable, for a change of 1 SD in the predictor (while controlling for the other predictors). Thus, if trait test emotionality increased by 1 SD, we can predict that state test emotionality would increase by .37 SD; and if state test worry increased by 1SD, we can predict that state test emotionality would increase by .45 SD.

The *t*-test values indicate whether the predictor's regression coefficient is significant. For the standard method, however, it only tests the unique variance explained by the predictor. Thus, a predictor that is correlated with the criterion variable but shares variance explained with another predictor may have a non-significant β .

If you request **Collinearity diagnostics** (in the **Linear Regression: Statistics** dialogue box), two additional columns appear on the right of the Coefficients table. These extra columns are usually sufficient to check whether your data meet this assumption.

The tolerance values are a measure of the correlation between the predictor variables and can vary between 0 and 1. The closer to zero the tolerance value is for a variable, the stronger the relationship between this and the other predictor variables. You should worry about variables that have a very low tolerance. SPSS will not include a predictor variable in a model if it has a tolerance of less than .0001. However, you may want to set your own criteria rather higher, perhaps excluding any variable that has a tolerance level of less than .01.

Collinearity Statistics Sig. Tolerance VIF t -4.050.000 .553 .582 941 1.063 .190 850 518 1.931 .704 .381 .605 1.654 4.239 .000 516 1.940 2.757 .007 .668 1.498 .000 1.610 5.671 621

VIF is an alternative measure of collinearity (in fact, it is the reciprocal of tolerance); a large value indicates a strong relationship between predictor variables.

				-	Collinearity D	iagnostics ^a					
				Variance Proportions							
Model	Dimension	Eigenvalue	Condition Index	(Constant)	Psychology A level completed	Trait test worry	Stats course anxiety	Trait test emotionality	Perceived test difficulty	State test worry	
1	1	6.686	1.000	.00	.00	.00	.00	.00	.00	.00	
	2	.104	8.025	.02	.37	.13	.01	.10	.02	.00	
	3	.078	9.251	.01	.42	.04	.01	.07	.11	.11	
	4	.042	12.622	.00	.01	.32	.56	.06	.01	.18	
	5	.036	13.610	.00	.01	.24	.34	.72	.08	.00	
	6	.030	14.930	.60	.15	.05	.07	.00	.03	.47	
	7	.024	16.663	.38	.04	.22	.01	.05	.74	.23	

a. Dependent Variable: State test emotionality

Residuals Statistics^a

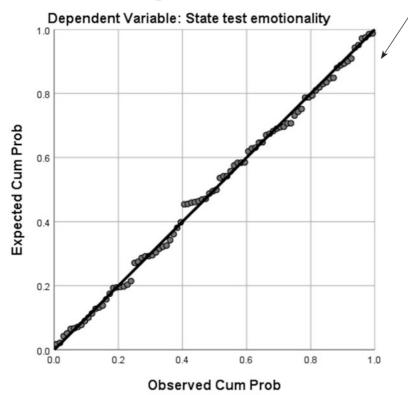
The Residuals Statistics table is produced if you request any Plots.

	Minimum	Maximum	Mean	Std. Deviation	Ν
Predicted Value	1.0614	4.3618	2.5556	.85772	90
Residual	-1.27640	1.35944	.00000	.58417	90
Std. Predicted Value	-1.742	2.106	.000	1.000	90
Std. Residual	-2.110	2.247	.000	.966	90

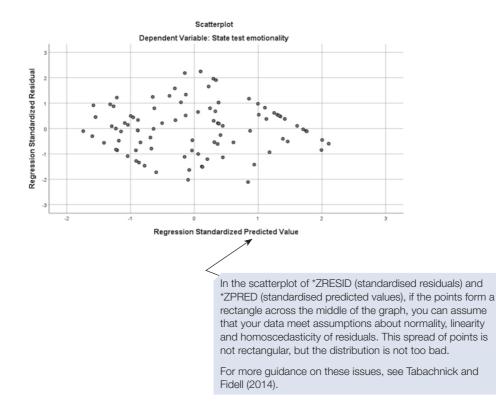
a. Dependent Variable: State test emotionality

Charts

Normal P-P Plot of Regression Standardized Residual



This graph is produced by the **Normal probability plot** option. If the points, which represent the cumulative expected and observed probabilities, are reasonably close to the straight line, then you can assume normality of the residuals.



Next, we give some guidance on reporting a multiple regression.

We haven't shown this here, but you can save various new variables produced by the multiple regression as we showed you for bivariate regression (Chapter 6 Section 8). Click on **Save** in the **Linear Regression** dialogue box to get the **Linear Regression: Save** dialogue box. You could explore the new variables that are available through the options there.

Reporting the results

When you describe your data analysis, in a methods subsection or in the results section, you should report assumption checks that you carried out, whether the data met these checks and any transformations you made as a consequence (for example to correct for skewed data).

When reporting the results of a multiple regression analysis, you should inform the reader about the proportion of variance accounted for by the model, the significance of the model and the significance of the predictor variables. You should also include a table of the beta values for each predictor. In addition, if the regression coefficient of any predictor is negative (unlike in this example), you must point out that direction of impact on the criterion variable. Note that some psychological constructs are scored in opposite directions; for example, a high score on a self-esteem questionnaire may indicate either high self-esteem or low self-esteem. In addition to being clear in your methods subsection what a high score indicates for each of your constructs, you also need to think carefully about the direction of the relationship between each predictor and the criterion variable in order to describe these correctly in the results section. You would also include other information, such as summary descriptives for all variables, and the correlation matrix for all the variables you included in the multiple regression analysis.

For the multiple regression itself, in a report you might write:

Those variables that were significantly correlated with the criterion variable, state test emotionality, were entered as predictors into a multiple regression using the standard method. A significant model emerged: F(6,83) = 29.82, p < .001. The model explains 66.0% of the variance in state test emotionality (adjusted $R^2 = .660$).

Table 10.1 gives information about regression coefficients for the predictor variables entered into the model. Trait test emotionality, Perceived test difficulty, and State test worry were significant predictors, with a positive relationship to State test emotionality. Psychology 'A' level, Trait test worry, and Statistics course anxiety were not significant predictors.

 Table 10.1
 The unstandardised and standardised regression coefficients for the variables entered into the model

Variable	В	SE B	β	р
State test worry	.61	.11	.45	< .001
Trait test emotionality	.43	.10	.37	< .001
Perceived test difficulty	.30	.11	.21	.007
Psychology A level completed	.07	.13	.04	.582
Stats course anxiety	.05	.13	.03	.704
Trait test worry	.02	.11	.02	.850

CHAPTER 10

Section 3: SEQUENTIAL OR HIERARCHICAL METHOD OF MULTIPLE REGRESSION

Only use the sequential or hierarchical method if you have a rationale for entering predictors in a particular sequence, as described in Section 1. Remember that first you should run a standard multiple regression, requesting collinearity diagnostics and residual plots in order to check those assumptions.

m <u>Analyze G</u> raphs <u>U</u> tilities E <u>x</u> ter	
Reports Descriptive Statistics	
Typ Bayesian Statistics	abel Values M
eric Ta <u>b</u> les	evel completed {1.00, yes} None
eric Compare Means	None None
General Linear Model	course difficulty None None
eric Generalized Linear Models	None None
eric Mixed Models	nxiety None None
Correlate	onality None None Click on
eric Regression	Analyze⇒Regression⇒Linear
eric Loglinear) III i i i i i i i i i i i i i i i i i
Neural Networks	Linear ne
Classify	Curve Estimation
Dimension Reduction	Partial Least Squares
Scale	Binary Logistic
Nonparametric Tests	Multinomial Logistic
Forecasting	M Orginal
Survival	Probit
Multiple Response	Nonlinear
Missing Value Analysis	Weight Estimation
Multiple Imputation	2-Stage Least Squares
Comp <u>l</u> ex Samples	Quantile
聞 Simulation	Optimal Scaling (CATREG)
Quality Control	>
Spatial and Temporal Modeling	N

In the Linear Regression dialogue box, we enter predictors in blocks (as shown below).

	Note this is Block 1 of 1.
Gender Segender Sychology A level c Age at intake [Age] Perceived stats cour Trait test worty [T_W Stats course anxiety Trait test emotionality [S_E] Previous Independent(s): Stats course anxiety Stats course anxiety	We are only entering psycha ('Psychology A level completed') as a predictor in block 1. This is because it is the only predictor (of those correlated with the criterion) that was established before students began the course.
Selection Variable: Selection Variable: Rule Case Labels: WLS Weight:	Note that we are using the Enter option. We will discuss options for each block below.
OK Paste Reset Cancel Help	When you have entered predictors for this block, click on <u>Next</u> .

Linear Regression		×	
gender Psychology A level c Age at intake [Age] Perceived stats cour Trait test worry [T_W Stats course anxiety Trait test emotionalit Perceived test diffic	Dependent. State test emotionality [S_E] Block 2 of 2 Preyious Independent(s): Trait test worry [T_Worry] Stats course anxiety [C_Anx] Trait test emotionality [T_Em.]	Statistics Plots Save Options Style Bootstrap	Note this should say "Block 2 of 2". This is a bug in our version of SPSS. Note that the Previous button is now available for you to move between blocks.
State test worry [S_W]	Method: Enter Selection Variable: Rule		The second block of predictors are those that were measured early in the course. Now click the Next button
ОК	Case Labels: WLS Weight Paste Reset Cancel Help		again to set up the third block.

 Linear Regression gender Psychology A level c Age at intake [Age] Perceived stats cour Trait test worry [T_W Stats course anxiety Trait test emotionalit Perceived test diffic State test worry [S_W] 	Dependent. State fest emotionality [S_E] Block 3 of 3 Previous Next Block 3 of 3 Perceived test difficulty [PTD] State test worry [S_W] Method: Enter *	X Statistics Plots. Save Qptions Style Bootstrap	We are now in Block 3 of 3. We have moved the third and final block of predictors (that were measured just before or after the test) into the Independents box, and are ready to move on.
	Selection Variable: Que. Q		Click on the <u>S</u>tatistics button, for the next step.
ОК	Paste Reset Cancel Help		



When you have entered two or more blocks of predictors, you can use the <u>Mext</u> and **Previous** buttons to move between blocks to check the predictor varaibles in each block.

D	ics X		When blocks are entered, SPSS will produce output for different models. Selecting R gquared change will test for
Estimates Confidence intervals Level(%): 95 Covariance matrix	Model fit R squared change Descriptives Part and partial correlations Collinearity diagnostics	s r ((((significant difference between successive models.We have not selected Collinearity diagnostics because we checked that output from the standard multiple regression see Section 2).
Residuals	3 standard deviations		
Continue	Cancel Help	(Click on <u>Continue,</u> then OK.

The output is illustrated below. We have only shown the tables that are different from the output for the standard or simultaneous method shown in Section 2.

SPSS output for sequential multiple regression

Obtained using menu items: Regression > Linear (method = enter; three blocks have been entered)

Model	Variables Entered	Variables Removed	Method
1	Psychology A level completed ^b		Enter
2	Trait test worry, Stats course anxiety, Trait test emotionality ^b		Enter
3	Perceived test difficulty, State test worry ^b		Enter

Each time a block is introduced, a new model is formed, and a multiple regression is conducted on each successive model. Thus, with three blocks, there are three models.

This tells us which variables were entered in each block. None were removed because we used **Enter** for each block.

a. Dependent Variable: State test emotionality

b. All requested variables entered.

....a

Model Summary

						Change Statistics				
Model	R	R Square	Adjusted R Square	Std. Error of the Estimate	R Square Change	F Change	df1	df2	Sig. F Change	
1	.212ª	.045	.034	1.01986	.045	4.150	1	88	.045	
2	.652 ^b	.425	.398	.80509	.380	18.738	3	85	.000	
3	.827°	.683	.660	.60492	.258	33.781	2	83	.000	

a. Predictors: (Constant), Psychology A level completed

b. Predictors: (Constant), Psychology A level completed, Trait test worry, Stats course anxiety, Trait test emotionality

c. Predictors: (Constant), Psychology A level completed, Trait test worry, Stats course anxiety, Trait test emotionality, Perceived test difficulty, State test worry

We can see that Model 1, which included only 'Psychology A level', accounted for just 3.4% of the variance (Adjusted R^2 = .034). The inclusion of the second block of predictors into Model 2 resulted in an additional 38% of the variance being explained (R^2 change = .380). Model 3 explained an additional 26% of the variance (R^2 change = .258), and in total accounts for 66% of the variance (Adjusted R^2 = .660).

SPSS provides a test of whether each model is significantly different from the preceding model in terms of proportion of variance explained in the criterion variable.

For Model 1, this comparison is against a 'model' with no predictors (that is, when $R^2 = 0$), and the F and p values are therefore the same as for the model alone (shown in the next table).

The next row indicates that Model 2 explains significantly more variance than Model 1 [F (3,85) = 18.74, p < .001], and the third row indicates that Model 3 explains significantly more variance than Model 2 [F (2,83) = 33.78, p < .001].

The degrees of freedom here are: first, the number of predictors entered in that block and, second, N - m - 1, where *m* is the total number of predictors in that model.

		-
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	101	v A

Model		Sum of Squares	df	Mean Square	F	Sig.
1	Regression	4.317	1	4.317	4.150	.045 ^b
	Residual	91.531	88	1.040		
	Total	95.847	89			
2	Regression	40.753	4	10.188	15.718	.000°
	Residual	55.095	85	.648		
	Total	95.847	89			
3	Regression	65.475	6	10.913	29.822	.000 ^d
	Residual	30.372	83	.366		
	Total	95.847	89			

This table reports the ANOVA result for the three models. They are all significant, although Model 1 (with 'Psychology A level' as the sole predictor) only just reaches significance.

a. Dependent Variable: State test emotionality

b. Predictors: (Constant), Psychology A level completed

 Predictors: (Constant), Psychology A level completed, Trait test worry, Stats course anxiety, Trait test emotionality

d. Predictors: (Constant), Psychology A level completed, Trait test worry, Stats course anxiety, Trait test emotionality, Perceived test difficulty, State test worry

Co	effic	ien	tsa

		Unstandardized Coefficients		Standardized Coefficients		/
Model		В	Std. Error	Beta	t	Sig
1	(Constant)	1.882	.347		5.418	.000
	Psychology A level completed	.439	.215	.212	2.037	.045
2	(Constant)	015	.388		038	.970
	Psychology A level completed	.144	.175	.069	.823	.413
	Trait test worry	.147	.147	.112	1.005	.318
	Stats course anxiety	.350	.164	.215	2.138	.035
	Trait test emotionality	.483	.135	.407	3.568	.001
3	(Constant)	-1.394	.344		-4.050	.000
	Psychology A level completed	.073	.132	.035	.553	.582
	Trait test worry	.021	.113	.016	.190	.850
	Stats course anxiety	.049	.129	.030	.381	.704
	Trait test emotionality	.433	.102	.365	4.239	.000
	Perceived test difficulty	.304	.110	.208	2.757	.007
	State test worry	.614	.108	.445	5.671	.000

See Section 2 for an explanation of this output.Note, for predictors in more than one model (such as 'Psychology A level completed'), the regression coefficients change in successive models (see text below).

a. Dependent Variable: State test emotionality



For the predictor 'psycha' in Model 1, β is identical to its correlation coefficient with 'se'; as explained in Section 1, with one predictor, the regression coefficient and correlation coefficient are the same.

Here, SPSS gives some information for the predictors that were excluded from each model. When the **Enter** option is used in the **Linear Regression** dialogue box, the variables excluded for each model are simply the predictors that were entered in later blocks.

Excluded Variables ^a						
Model	_	Beta In	t	Sig.	Partial Correlation	Collinearity Statistics Tolerance
1	Trait test worry	.471 ^b	5.064	.000	.477	.978
	Stats course anxiety	.477 ^b	5.120	.000	.481	.970
	Trait test emotionality	.591 ^b	6.811	.000	.590	.952
	Perceived test difficulty	.542 ^b	6.104	.000	.548	.976
	State test worry	.689 ^b	9.117	.000	.699	.983
2	Perceived test difficulty	.402°	5.082	.000	.485	.838
	State test worry	.542°	7.456	.000	.631	.779

a. Dependent Variable: State test emotionality

b. Predictors in the Model: (Constant), Psychology A level completed

c. Predictors in the Model: (Constant), Psychology A level completed, Trait test worry, Stats course anxiety, Trait test emotionality

A note on sequential method and regression coefficients

As we explained earlier, in the sequential or hierarchical method, the researcher uses a theoretical or empirical rationale to decide the order in which predictors are entered into the model. This contrasts with the standard or simultaneous method, in which all predictors are entered at once. As we demonstrated above, in SPSS we use successive blocks to enter predictors in the sequence we decided on, and we are given information about whether each additional block of predictors significantly improves the explanatory power (that is, whether R^2 is significantly increased).

However, the regression coefficients above do not follow the sequential 'rule' that shared variance remains with predictors entered earlier. This is because, for each block, we used the SPSS **Enter** option; thus, for each model all the predictors in that model are analysed with the standard method. In Model 1 the predictor 'Psychology A level completed' had a β of .21, whereas in Model 2 the β was much lower at .06. This reduction demonstrates that 'Psychology A level completed' shares explained variance with at least one of the predictors which we added in Model 2, because the standard method excludes shared variance when calculating regression coefficients (although it is taken into account for R^2).

You may be specifically interested in the regression coefficients for predictors entered according to your rationale and wish the predictors higher up the hierarchy to 'keep' the variance they explain in the criterion variable, with variables further down the hierarchy only allocated any additional variance they explain. In this case, instead of using the **Enter** option for each predictor when you add it as a block, you would use another SPSS option such as **Forward** (see Section 1). Note, however, that if you enter two or more predictors in a block (as we did above), they will all be treated in that block by the **Forward** statistical criterion, and you might not want that.



Reporting the results

When reporting the results of a sequential multiple regression analysis, you should inform the reader about the proportion of the variance accounted for by each model, the significance of each model and whether a later model explained significantly more variance than an earlier model. In addition, you can report on the significance of the predictor variables for one or more of the models, as appropriate for the purpose of your report. It is also a good idea to include a table of the beta values for each predictor, as seen on Page 315. (See also the general guidance on reporting the results provided in Section 2.) In a report you might write:

Model 1, with Psychology 'A' level as the only predictor, explained 3.4% of variance and was significant [F(1,88) = 4.15, p < .045]. Model 2, in which Stats course anxiety, Trait test worry and Trait test emotionality were added, explained significantly more variance [R^2 change = .38, F(3,85) = 18.74, p < .001]. The model explains 40% of the variance in State test emotionality (adjusted $R^2 = .398$) and was significant [F(4,85) = 15.72, p < .001]. Model 3, in which Perceived test difficulty and State test worry were added, explained yet more variance, and this increase was also significant [R^2 change = .26, F(2,83) = 33.78, p < .001]. Model 3 explained 66.0% of the variance in State test emotionality (adjusted $R^2 = .660$) and was significant [F(6,83) = 29.82, p < .001]. The significant predictors in Model 3 were Trait test emotionality, Perceived test difficulty and State test worry.

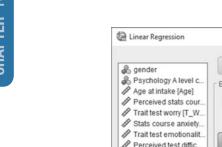
Section 4: STATISTICAL METHODS OF MULTIPLE REGRESSION

As explained in Section 1, **Stepwise** is one of the statistical methods that should be used with caution; results must be validated by cross-checking on two separate samples. Stepwise checks all predictors as it enters each one and removes any that no longer meet its statistical criterion. Thus, it results in a model with the smallest number of predictors. See Regression methods in Section 1 for more explanation.

How to perform multiple regression using the stepwise method

Remember to run a standard multiple regression first, requesting collinearity diagnostics and residual plots to check those assumptions.

Click on <u>Analyze⇒Regression⇒Linear</u>.



 Linear Regression gender Psychology A level c Age at intake [Age] Perceived stats cour Trait test worry [T_W Stats course anxiety Trait test emotionalit Perceived test diffic State test worry [S_W] 	State test emotionality [S_E] Block 3 of 3 Previous Independent(s): Psychology A level completion Trait test worry [T_Worry] State course anxiety [C_A	× Xots Save Style otstrap	We use as predictors those variables that had some correlation with the criterion. Those were: Psychology A level; Trait test worry; Trait test emotionality; Statistics course anxiety; Perceived test difficulty; State test worry.
	Method: Enter Selection Variable: Stepwise		
	Case Labels: Backward Forward		Select Stepwise as the Method.
	WLS Weight:		
ОК	Paste Reset Cancel Help		Click on the <u>Statistics</u> button.

🚇 Linear Regression: Statisti	cs	×	
Regression Coefficien Estimates Confidence intervals Level(%): 95 Coyariance matrix Residuals Durbin-Watson Casewise diagnostics	Model fit R gquared change Descriptives Part and partial correlatio Collinearity diagnostics	ns	Select R <u>s</u>quared change .
Outliers outside: O All cases	3 standed deviation	IS	Click <u>Continue</u> , then OK.

Below are the parts of the output which are different from those we have already explained in Sections 2 and 3.

Model	Variables Entered	Variables Removed	Method
1	State test worry	ŀ	Stepwise (Criteria: Probability-of- F-to-enter <= . 050, Probability-of- F-to-remove >= .100).
2	Trait test emotionality	ŀ	Stepwise (Criteria: Probability-of- F-to-enter <= . 050, Probability-of- F-to-remove >= .100).
3	Perceived test difficulty		Stepwise (Criteria: Probability-of- F-to-enter <= . 050, Probability-of- F-to-remove >= .100).

Variables Entered/Removed^a

a. Dependent Variable: State test emotionality

SPSS output for a statistical multiple regression

Obtained using menu items: Regression > Linear (method = stepwise)

This table shows us the order in which predictors were entered and removed. In this run, three predictors were added in successive models, and none were removed.

In stepwise, predictors are considered in order of magnitude of their correlation with the criterion variable. They are all reconsidered whenever the next predictor is entered, on the basis of a significant contribution to R^2 , and may be removed if they do not meet that statistical criterion. The criteria for entry and removal are shown in the right-hand column of this table. They can be amended, if you wish, in the **Linear Regression: Options** dialogue box.

				Model	Summary				
						Cha	nge Statistic	s	
Model	R	R Square	Adjusted R Square	Std. Error of the Estimate	R Square Change	F Change	df1	df2	Sig. F Change
1	.705ª	.497	.491	.74043	.497	86.830	1	88	.000
2	.805 ^b	.648	.640	.62299	.151	37.303	1	87	.000
3	.825°	.681	.670	.59623	.033	8.985	1	86	.004

a. Predictors: (Constant), State test worry

b. Predictors: (Constant), State test worry, Trait test emotionality

c. Predictors: (Constant), State test worry, Trait test emotionality, Perceived test difficulty

Model		Sum of Squares	df	Mean Square	F	Sig.
1	Regression	47.603	1	47.603	86.830	.000 ^b
	Residual	48.244	88	.548		
	Total	95.847	89			
2	Regression	62.081	2	31.040	79.977	.000°
	Residual	33.766	87	.388		
	Total	95.847	89			
3	Regression	65.275	3	21.758	61.206	.000 ^d
	Residual	30.572	86	.355		
	Total	95.847	89			

ANOVA

a. Dependent Variable: State test emotionality

b. Predictors: (Constant), State test worry

c. Predictors: (Constant), State test worry, Trait test emotionality

d. Predictors: (Constant), State test worry, Trait test emotionality, Perceived test difficulty

The Model Summary table shows Adj. R^2 and R^2 change, and whether the change is significant for each model.

The ANOVA table shows the test of significance for each model considered independently.

You will extract values from these two tables for your report; see detailed annotations on the sequential multiple regression above.

Coefficients ^a							
		Unstandardize	d Coefficients	Standardized Coefficients			
Model		В	B Std. Error		t	Sig.	
1	(Constant)	133	.299		446	.657	
	State test worry	.973	.104	.705	9.318	.000	
2	(Constant)	849	.278		-3.061	.003	
	State test worry	.775	.094	.561	8.273	.000	
	Trait test emotionality	.491	.080	.414	6.108	.000	
3	(Constant)	-1.280	.302		-4.240	.000	
	State test worry	.627	.102	.454	6.121	.000	
	Trait test emotionality	.467	.077	.394	6.032	.000	
	Perceived test difficulty	.316	.105	.216	2.997	.004	

As for the sequential multiple regression, the Coefficients table shows the regression coefficients and whether they are significant, for each model. See annotations on the previous two multiple regressions for detailed comments.

a. Dependent Variable: State test emotionality

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Reporting the results

See the general guidance on reporting the results provided for the standard method in Section 2. Results for the models from a statistical method can be reported in a similar way to that shown for the sequential method, and with a table showing coefficient information for the final model (as seen on Page 315). If you use a statistical method, however, you should validate your results in the way described at the end of Section 1 and report both multiple regressions. Only results that are common to both analyses should be relied on.

In a report you might write:

State test worry, in Model 1, explained 49% of variance and was significant [*F* (1,88) = 86.83, p < .001]. Model 2, in which Trait test emotionality was added, explained significantly more variance [*R*² change = .15, *F* (1,87) = 37.30, p < .001]. The model explains 64% of the variance in State test emotionality (adjusted *R*² = .640) and was significant [*F* (2,87) = 79.98, p < .001]. Model 3, in which Perceived test difficulty was added, explained another 3.3% of variance, and this increase was significant [*R*² change = .033, *F* (1,86) = 8.96, *p* = .004]. Model 3 explained 67.0% of the variance in State test emotionality (adjusted *R*² = .670) and was significant [*F* (3,86) = 61.21, p < .001].

Summary

- This chapter introduced you to multiple regression, a statistical technique that allows us to predict someone's score on one variable (the criterion variable) on the basis of their scores on other variables (the predictor variables).
- Remember that multiple regression requires a large number of observations.
- There are different ways that the predictor variables can be assessed, and unless you have a theoretical model in mind, it is safest to use the Enter option, which applies the simultaneous method.
- Multiple regression makes certain assumptions about the data; if your criterion variable is measured at nominal level, see Chapter 12.
- For guidance on recoding values, see Chapter 4, Section 5.
- For guidance on incorporating SPSS output into a report, or on printing the output, see Chapter 14.

11 Analysis of covariance and multivariate analysis of variance

In this chapter

- An introduction to analysis of covariance
- Performing analysis of covariance in SPSS
- An introduction to multivariate analysis of variance
- Performing multivariate analysis of variance in SPSS



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Section 1: AN INTRODUCTION TO ANALYSIS OF COVARIANCE

- Analysis of covariance (ANCOVA) is a statistical procedure that allows us to look at the effect of one or more factors on a dependent variable, while partialling out or removing the effect of another variable. ANCOVA can be considered as a cross between ANOVA, partial correlation and multiple regression.
- With ANCOVA, we look at the effect of one or more factors on a dependent variable (as does ANOVA), but in addition, we acknowledge the influence of another variable – a covariate. ANCOVA partials out (or removes) the effect of the covariate by using the regression equation to measure its influence (as in multiple regression).
- In this chapter we consider a one-way ANCOVA, looking at how one factor, the independent variable, affects the dependent variable once the variance of the covariate is removed.

What does ANCOVA do?

In Chapter 8 we explained that ANOVA essentially works by comparing the amount of variance brought about by our experimental manipulation (i.e. the systematic variance in the data that is due to our IV) against the amount of error variance in the data

(i.e. the unsystematic variance that is brought about by other 'nuisance variables' such as individual differences). This comparison produces the *F*-ratio; and as this statistic increases, so does the relative amount of variance our IV can explain compared with the error variance. If we can reduce the error variance in some way, this can help to make the *F*-ratio more sensitive, increasing the likelihood of finding a significant result if our manipulation has been successful. ANCOVA allows us to effected

tively reduce the unexplained error variance by including extraneous variables (or covariates) that may be influencing our DV in some way. By getting rid of the effects due to a covariate, ANCOVA allows us to reduce the error variance, which in turn leads to a larger *F*-value. The inclusion of covariates can therefore increase statistical power.

You may realise from your own experience of conducting research that it is not always possible to control for all possible confounding variables. However, where you can identify and measure a variable that you know will influence the dependent variable, then ANCOVA may be beneficial as it will allow you to get rid of these unwanted, confounding effects.

When should I use ANCOVA?

ANCOVA can be used when you have one between-subjects factor (IV) – and this is what we will demonstrate in this chapter – or it can also be used when you have multiple between-subjects factors. It is also possible to have more than one covariate. It is best to use ANCOVA when participants are randomly assigned to one of the levels of the independent variable, and ideally, the covariate should be measured beforehand rather than afterwards, as exposure to the independent variable may affect the covariate.

Broadly speaking, there are two reasons to use an ANCOVA (and the two are not mutually exclusive):

- 1. When you want to reduce the amount of error variance in your experimental design.
- 2. When you want to partial out the effects of potentially confounding variables.

An example

Imagine that we want to investigate how quickly students could learn how to use different computerised statistical packages (the IBM SPSS Statistics software referred to here as SPSS is not the only one available). You might choose three different packages and then randomly assign them to three different groups, and measure how quickly it takes the students to learn how to carry out an ANCOVA using their designated package. Our hypothesis may be that the group given SPSS to learn would be the fastest (as this package is very user-friendly). However, there are a number of other extraneous variables that might affect our DV (the students' learning speed), such as familiarity with computer software. If we ran a correlation between our covariate and our DV, we would probably find that there was a positive correlation between them, so that as familiarity increases, so does speed of learning. ANCOVA examines the association between the covariate(s) and DV (e.g. learning speed) and removes the variance due to these associations. So if we measured these variables before exposing students to one of the statistical packages, we could then partial out (or remove) their effects on the DV so that we can obtain a clearer insight into the differences in learning speed for the different packages.

You might be thinking that as we randomly assigned our participants to the three groups, there probably wouldn't be any systematic differences in software familiarity across the three groups. Therefore, the means on this covariate would not differ too much between the three experimental conditions. This is actually the ideal situation, and represents a fundamental assumption of ANCOVA: the assumption that the covariate is independent from the factor of interest (i.e. your IV, or experimental manipulation). This is important to consider, because when your conditions differ in terms of the covariate (i.e. when the covariate and IV are correlated), removing the variance associated with the effect of your IV; and this can lead to spurious results (Derks et al., 2018).

Despite this, there are many examples in the literature where psychologists have used ANCOVA in situations where the means on the covariate differ significantly between experimental conditions. This is most common in two types of situations: (1) where randomisation to experimental conditions has not been possible, resulting in systematic differences between experimental groups on a confounding variable that is not the IV; (2) in pretest-posttest designs, where measures are taken before and after exposure to the experimental manipulation. For example, a researcher exploring the effect of a new antidepressant might do so by assigning participants to one of two conditions: the experimental group (who would take the new drug for 6 weeks), and the control group (who would take sugar pills for 6 weeks). To test the effectiveness of the drug, participants would be asked to fill out a depression questionnaire before and after the 6 week period. While this isn't a problematic design in itself, it becomes a problem if the scores on the pretest (e.g. the baseline depression scores) reveal any pre-existing, systematic differences between the groups. In these situations, (where experimental conditions vary according to a confounding variable – such as pre-existing depression) researchers often mistakenly include this confound into the analysis in an attempt to 'control for' the differing baseline scores using ANCOVA. However, the inclusion of the confound as the covariate is not valid in these particular cases. This is because the covariate is related to the IV (i.e. the value of the covariate changes according to the experimental condition), so the removal of the covariate's effect partials out some of the variance that would ordinarily be attributed to the experimental manipulation, reducing the experimental effect and rendering it unreliable.

So what can you do instead? In situations where your covariate differs according to the different experimental conditions, it would be more appropriate to include the covariable as an extra factor (or IV) into a Factorial ANOVA. While this won't allow you to partial out the effects of the variable, it will enable you to tell whether it is moderating the effect of the IV on the DV (see Chapter 9, Section 1), which can be just as informative.

Checklist for choosing one or more covariates

When including covariates into your analysis, it is important to consider whether their inclusion is appropriate, and whether or not you are violating any key assumptions of the test.

- 1. A covariate should be chosen on the basis of existing theory and research.
- 2. A covariate should ideally be measured using a scale at ratio, interval or ordinal level. According to Howell (2002), it can even be nominal, but like the predictor variables in multiple regression, if nominal, it has to be a dichotomous variable.
- 3. Ideally, a covariate should be measured before the experimental manipulation takes place.
- 4. The covariate should be independent from the experimental manipulation (i.e. experimental conditions should not differ on covariate scores).
- 5. A covariate should be measured reliably, that is, if it were measured several times across a time period, there would be a high correlation between the scores.
- 6. The relationship between a covariate and the dependent variable must be linear (straight line). You can check this by looking at the scatterplots for each group, and if there is more than one covariate, then they should not be strongly correlated with each other.
- 7. There should be homogeneity of regression slopes. The relationship between the dependent variable and the covariate should be similar for all experimental groups, so that the regression lines are parallel. So, using our first example, the relationship between learning speed of statistical package and familiarity with computer software should be similar for each of the three groups tested.
- 8. In addition, ANCOVA makes the same assumptions as ANOVA (see Chapter 8, Section 1).

Section 2: PERFORMING ANALYSIS OF COVARIANCE IN SPSS

Example study: exposure to low levels of organophosphates

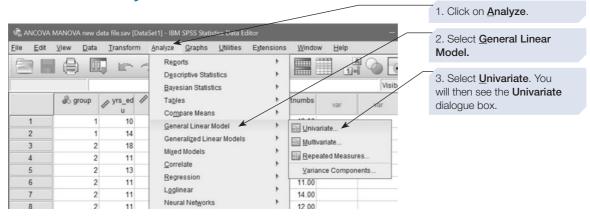
This example is taken from one of Virginia's papers investigating the effects of lowlevel exposure to certain toxic chemicals on neuropsychological functioning (Mackenzie Ross et al., 2010). The chemicals under consideration were organophosphates (OPs), which had previously been found to cause neuropsychological or psychiatric impairment in high doses. Here, the performance of two groups of participants was compared on a number of different psychometric tasks designed to measure neuropsychological functioning, and the two groups differed in terms of their exposure to particular chemicals. Participants were either farmers who had long-term, low-level occupational exposure to the OPs (as they are commonly used as insecticides and pesticides for agricultural purposes), or controls who had not been exposed to the chemicals. The researchers also measured other factors known to affect performance on the psychometric tests, such as years in education. This allowed the researchers to compare the performance of the two groups on these tests after removing the effects of the potentially confounding variables.

One particular finding that Mackenzie Ross et al. report is that the group of farmers exposed to low levels of organophosphate pesticides scored lower overall on the Wechsler Adult Intelligence Scale (WAIS) than the control group. Further analysis demonstrated that this was a result of impaired performance on certain WAIS subtests, indicating neurotoxic damage.

For the purposes of demonstrating ANCOVA here, the researchers have provided us with part of their data set so that we can show you how to tease out the effects of a covariate, namely the number of years in education, from overall performance on the WAIS. Their data were carefully screened for outliers and transformed where necessary to ensure they met the assumptions for parametric tests (the data file is available from macmillanihe.com/harrison-spss-7e).

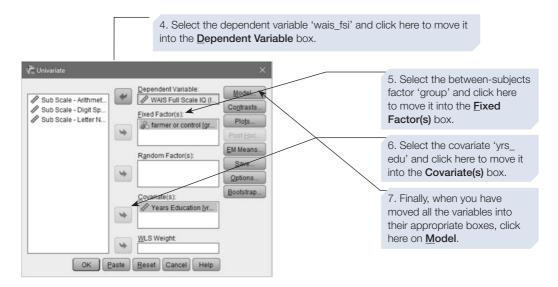
The analysis shown here involves one between-subjects factor called 'group', which had two levels. Group 1 involved 127 sheep famers (working and retired) who had been exposed to low levels of organophosphate pesticides. Group 2 involved 78 controls (working and retired) who had not been exposed to these chemicals. The dependent variable we will look at here is overall performance on the WAIS ('wais_fsi' in the data file). The covariate consists of the number of years in education ('yrs_edu' in the data file). For the purposes of this demonstration, the hypothesis tested is that there will be a detrimental effect of low-level exposure to organophosphate pesticides on overall performance on the WAIS. Note that we are not replicating any analysis that Mackenzie Ross et al. undertook, but using a subset of their data to demonstrate how you would perform ANCOVA.

We have chosen to show you the ANCOVA procedure with a one-way betweensubjects design, so that we can also show you how to check for homogeneity of regression slopes statistically. It is much harder to check for this assumption statistically with more complex designs, so we shall also show you how you can look for this graphically and at the same time check that there is a linear relationship between your covariate and dependent variable. We will do this first, and after that, we show you how to perform the ANCOVA test itself.



How to check for homogeneity of regression slopes

If you have read Chapter 8 or performed ANOVA in SPSS, you will have already seen the dialogue boxes that appear next. This is because SPSS has incorporated ANCOVA as an option in the main ANOVA dialogue box.



When you click on the <u>Model</u> button, the following Univariate: Model dialogue box will appear.

Chivariate: Model Specify Model Full factorial Build terms Build custern terms	8. Click on <u>Build terms</u> . The variables in the left-hand box then become active. Check that Interaction is showing in the Build Term(s) box.
Eactors & Covariates:	9. Select the factor 'group' and click on the arrow under Build Term(s) so that it appears in the Model box. Then repeat with the covariate 'yrs_edu'.
Build Jerm: Sum of sguares: Type III Clear Term Add Remove Build Jerm: Clear Term Add Remove Build Jerm: Clear Term Add Remove Help	10. Next, select both 'group' and 'yrs_edu' in the left-hand box by clicking on one after the other so that both become highlighted, and click on the arrow so that the interaction 'group*yrs_edu' appears in the right-hand box. Click on the <u>Continue</u> button.

Once you click on the <u>Continue</u> button, you will return to the Univariate dialogue box. Now click on the <u>ok</u> button. The output is shown below.

SPSS output from procedure to check for homogeneity of regression slopes

Univariate Analysis of Variance

Note that N for each group is reduced, and this is because there are missing scores in the data file.

Between-Subjects Factors

		Value Label	Ν
farmer or control	1	farmer	122
	2	control	74

Tests of Between-Subjects Effects

Dependent Variable: WAIS Full Scale IQ (IQ indexed score)

Source	Type III Sum of Squares	df	Mean Square	F	Sig.
Corrected Model	5585.739 ^a	3	1861.913	17.030	.000
Intercept	24315.619	1	24315.619	222.405	.000
group	54.513	1	54.513	.499	.481
yrs_edu	3199.170	1	3199.170	29.261	.000
group * yrs_edu	10.549	1	10.549	.096	.756
Error	20991.465	192	109.331		
Total	2260748.000	196			
Corrected Total	26577.204	195			

SPSS reminds you that 'WAIS Full Scale IQ' is the dependent variable. Here, you are interested in looking to see if there is an interaction between the covariate, 'yrs_edu', and the independent variable, 'group'.

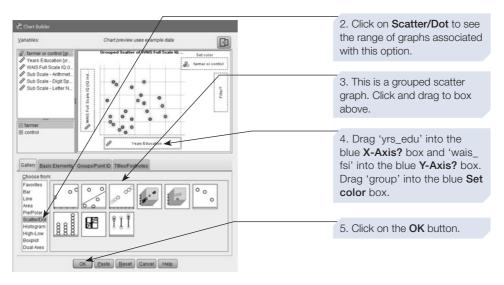
a. R Squared = .210 (Adjusted P

This is the only row that you are interested in. If this interaction is significant, then the data violate the assumption of homogeneity of regression slopes. Here, SPSS reports the interaction to be non-significant, so this assumption has *not* been violated.

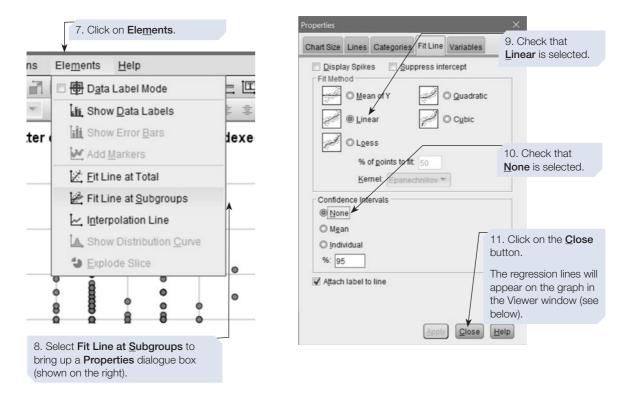
Now that we have checked for homogeneity of regression slopes, we can perform the ANCOVA test. First, however, we show you how to inspect the relationship between the covariate and the dependent variable graphically, using scatterplots. This procedure can be used to check that there is a linear relationship between the covariate and the dependent variable for both levels of the factor, and also that there is homogeneity of regression slopes.

How to check for linear relationship between covariate and dependent variable

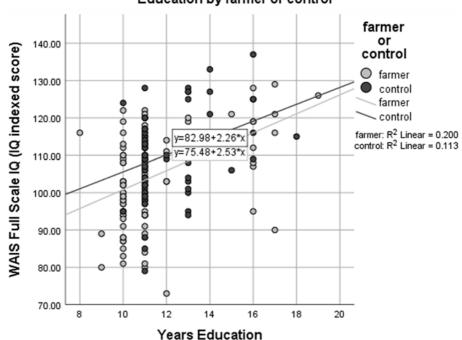
1. Click on <u>G</u>raphs and Chart Builder (see Chapter 3, Section 8 if you are unfamiliar with Chart Builder).



6. Double-click on the scattergram in the SPSS output to bring up the **Chart Editor** window (shown below on the left).



SPSS output for graph



Grouped Scatter of WAIS Full Scale IQ (IQ indexed score) by Years Education by farmer or control

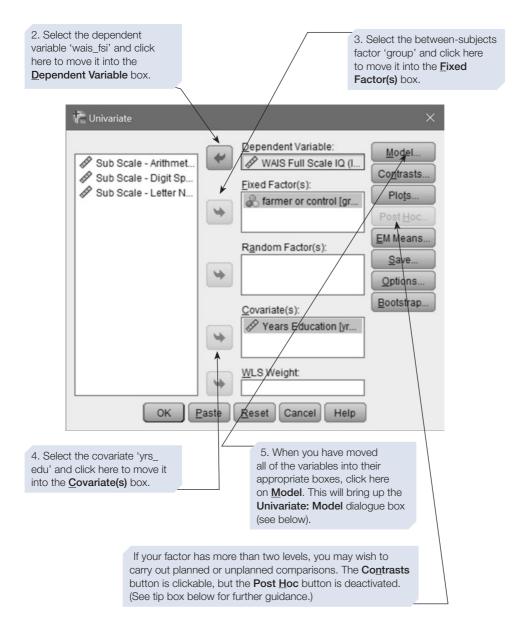
Inspect your graph to see if there is a linear relationship between the covariate and the dependent variable, one of the assumptions for ANCOVA. Here, the relationship for each group ('farmer' and 'control') looks only moderately linear. In fact, calculating the square root of the R² Linear values gives you Pearson's r, which can indicate the strength of the linear relationship between the variables; in this case r=.45 for farmers, and r=.34 for controls (see Chapter 6, Section 1). In practice, if the linear relationships look relatively weak (or non-existent), you may need to transform your data or remove outliers to improve linearity (see Tabachnick and Fidell, 2007, Chapter 4, for details).

Remember that the slopes of the regression lines should be roughly parallel, that is, the relationship between the covariate and the dependent variable should be similar for all groups (the assumption of homogeneity of regression slopes). This is important because ANCOVA assumes that the overall relationship between the dependent variable and the covariate is true for each group. We already know that this assumption has not been violated by our earlier check, and this is confirmed here by the fact that the slopes are almost parallel. The *R*-squared values can also be used to directly indicate how strong the relationship is between the dependent variable and the covariate (see Chapter 10, Section 1).

We have shown you here how to check for a linear relationship between the covariate and the dependent variable, and how to check if there is homogeneity of regression slopes. Next, we show you how to perform the ANCOVA test.

How to perform ANCOVA

 Click on <u>Analyze</u> ⇒ <u>General Linear Model</u> ⇒ <u>Univariate</u>. You will then see the Univariate dialogue box. You will find that you have already performed actions 2–4 shown for this dialogue box when checking for homogeneity of regression slopes.



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Inivariate: Model X Specify Hodel Specify Hodel © Full factorial © Build terms	6. In the Specify Model options, select Full f<u>a</u>ctorial .
Eactors & Covariates: Model:	
Type: Interaction	7. Click on the Continue button.
→ By * (Within) Clear Term Add Remove Build Ierm.	

You will return to the **Univariate** dialogue box. Click on the <u>EM Means...</u> button to bring up the Estimated Marginal Means dialogue box (shown below).

Univariate: Estimated Marginal Means Estimated Marginal Means Eactor(s) and Factor Interactions: OVERALL) group G	8. Move 'group' into the Display Means for box to obtain the means for the levels of 'group' after they have been adjusted for the effect of the covariate.
Compare main effects Confidence interval adjustment: LSD(none)	If you have more than two levels to your factor, you can perform group comparisons here (see Chapter 8).
9. Click on the Continue button. You will return to the Univariate dialogue box.	

You will return to the **Univariate** dialogue box. Click on the **Options** button to bring up the **Univariate: Options** dialogue box (shown below).

While partial eta squared isn't usually the best measure of effect size, it can be useful in the case of ANCOVA. This is because it calculates the proportion of variance a factor (or variable) explains that is not explained by other variables in the analysis. In this case, it allows us to calculate the proportion of variance in our DV (WAIS scores) that is explained by our IV (group membership: exposed vs non-exposed), that is not related to our covariate (years education).

Display Descriptive statistics Estimates of effect size Deserved power Parameter estimates	Homogeneity tests Spread vs. level plot Residual plot Lack of fit	10. Click here to check for assumption of homogeneity; remember that ANCOVA must also meet the assumptions underlying ANOVA.
Contrast coefficient matrix Heteroskedasticity Tests	General estimable function	
Modified Breusch-Pagan test Model Breusch-Pagan test Model Parameter estimates with rob	Model	11. Select the Estimates of effect size option. This will produce partial eta squared, which can be appropriate for ANCOVA (see below).
© нс <u>о</u> © нс <u>1</u> © нс <u>2</u>		12. Click on the <u>Continue</u> button. You will return to the Univariate dialogue box.
HC3 HC4 Significance level: 05 Confid	dence intervals are 95.0 %	

Click on or and SPSS will calculate the ANCOVA. The output is shown next.

If a covariate is selected, the **Post <u>Hoc</u>** button in the **Univariate** dialogue box is deactivated; however unplanned comparisons can be carried out by clicking in the box next to **Compare main effects** in the **Univariate: Estimated Marginal Means** dialogue box above (see Chapter 8, Section 3). Planned comparisons can be carried out by clicking the **Contrasts** button in the **Univariate** dialogue box (see Chapter 8, Section 2).

SPSS output for ANCOVA

Univariate Analysis of Variance

Between-Subjects Factors

		Value Label	Ν
farmer or control	1	farmer	122
	2	control	74

Levene's Test of Equality of Error Variances^a

Dependen	t Variable:	WAIS Full Sc	ale IQ (IQ indexed score)
F	df1	df2	Sig.

.005 1 194 .943

Tests the null hypothesis that the error variance of the dependent variable is equal across groups.

a. Design: Intercept + yrs_edu + group

The first line highlighted shows that the covariate is significantly related to the dependent variable. The next line shows the main effect of group is significant when the effects of the covariate 'yrs_edu' are partialled out. To check that the assumption of equality of variance was not violated, we clicked on **Homogeneity tests** in the **Univariate: Options** dialogue box. If this is not significant, as here, then this assumption has not been violated.

The effect size estimates are given in the **Partial Eta Squared** column.

Source	Type III Sum of Squares	df	Mean Square	F	Sig.	Partial Eta Squared
Corrected Model	5575.190 ^a	2	2787.595	25.617	.000	.210
Intercept	30770.937	1	30770.937	282.772	.000	.594
yrs_edu	4310.391	1	4310.391	39.611	.000	.170
group	818.381	1	818.381	7.521	.007	.038
Error	21002.014	193	108.819			
Total	2260748.000	196				
Corrected Total	26577.204	195				

a. R Squared = .210 (Adjusted R Squared = .202)

Estimated Marginal Means

	far	mer or cor	ntrol	
Dependent Variab	le: WAIS Fu	II Scale IQ (II	indexed score)	
	V		95% Confid	ence Interval
farmer or control	Mean	Std. Error	Lower Bound	Upper Bound
farmer	105.165 ^a	.946	103.299	107.032
control	109.403 ^a	1.217	107.003	111.803

 a. Covariates appearing in the model are evaluated at the following values: Years Education = 11.72.



Reporting the results

In a report you might write:

A one-way between-subjects analysis of covariance was carried out to assess the impact of exposure to low levels to organophosphate pesticides on overall performance on the WAIS. Checks were carried out to confirm homogeneity of regression slopes and linear relationship between covariate and dependent variable. The between-subjects factor comprised two groups: farmers who had been exposed to the pesticides and a control group who had not. The covariate comprised number of years in education, and this was significantly related to the overall WAIS scores: F(1,193) = 39.61, p < .001, partial $\eta^2 = .17$. Adjusting for this covariate resulted in a significant effect of the between-subjects factor group: F(1,193) = 7.52, p = .007, partial $\eta^2 = .04$. The adjusted mean WAIS score for those exposed to the chemicals was 105.17 compared with 109.40 for the control group.

Y ir

You would also want to report your descriptive statistics in your results, and information regarding the confidence intervals for each condition.

Section 3: AN INTRODUCTION TO MULTIVARIATE ANALYSIS OF VARIANCE

- Multivariate analysis of variance (MANOVA) is a statistical procedure similar to ANOVA and is used when the design of the experiment involves more than one dependent variable.
- Like ANOVA, MANOVA can be used to explore the effects of one or more independent variables and interactions between independent variables. However,

These are the adjusted means, i.e. the effect of the covariate has been statistically removed. To obtain these, we clicked on **Display Means for** in the **Univariate: Options** dialogue box. whereas ANOVA can only be used when there is one dependent variable (hence it is described as a 'univariate' test), MANOVA can handle several dependent variables all together (hence it is described as a 'multivariate' test).

- This type of design is quite common in clinical research when evaluating the impact of an intervention programme, as several different outcome measures can be explored simultaneously, for example cognitive and behavioural measures.
- MANOVA is useful when seeking to measure a variable that is complex to operationalise, and when a single dependent variable fails to capture all the elements of this complex variable. Instead, several different dependent variables could be measured, each of which represents a different element.

An example

Imagine you wanted to investigate whether being exposed to certain chemicals, for example those used to spray crops, has a detrimental effect on human health. You could select different groups of people who had different levels of exposure to such chemicals and compare them with a control group of people who had no or minimal exposure. You might then measure a number of different aspects of health, including cognitive processes. For example, you might choose to use different subtests of the Wechsler Adult Intelligence Scale (WAIS) that have been found to be reliable and valid measures of working memory. You could perform one-way ANOVAs on the scores from each different measure of working memory, but as you may remember from Chapter 8, Section 1, by performing multiple tests, you would run an increased risk of making a Type 1 error (that is, incorrectly rejecting a null hypothesis). To avoid this, you could use MANOVA.

What does MANOVA do?

MANOVA allows you to not only look at the effect of different independent variables and see if these interact, but also tells you if there is any relationship between the different dependent variables. Because they are all analysed simultaneously, MANOVA can check whether the different levels of the factors not only differ from one another on one dependent variable but also whether they differ along a combination of several dependent variables. It does this by creating a new dependent variable, which is the linear combination of each of the original dependent variables. MANOVA will tell you if the mean differences among groups on the combined dependent variable are larger than expected by chance.

There are other tests that also combine variables. In Chapter 10, a model containing a combination of predictor variables sought to predict the scores on a criterion variable. In Chapter 12 you will read about how variables are combined together to predict category membership in a type of analysis called 'discriminant analysis'.

You may remember that for ANOVA the statistic calculated is the *F*-ratio, which is the ratio of the variance due to the manipulation of the IV and the error variance. Conceptually, MANOVA does something similar, but this is statistically far more complicated and it will provide you with a choice of four different statistics to choose from, all of which indicate whether there are significant differences among the levels of the independent variable on the linear combination of the dependent variables. They are Pillai's Trace, Hotelling's Trace, Wilks' Lambda and Roy's Largest Root. SPSS will report a value for each of these, along with the *F* tests for each. If your factor has only two levels, then the *F* tests reported will be identical. This is because, when the factor has only two levels, and hence one degree of freedom, there is only one way of combining the different dependent variables to separate the levels or the groups. However, when your factor has more than two levels, then the *F* tests reported for the four test statistics are usually different and it is possible that some may be significant and some not. Most researchers report the values for the Wilks' Lambda, so we suggest you report these too. However, Pillai's Trace is considered to be the most robust (although all four are reasonably robust), so you might consider reporting the values

Following up a significant result

for Pillai's Trace when your sample size is small.

If you find a significant result, you will then want to follow it up. One possibility is to look at the univariate ANOVAs that are included in the SPSS output, after the section that presents the MANOVA test statistics. This will tell you which of the individual dependent variables are contributing to the significant overall result. If you do this, you need to consider something we mentioned earlier, the possibility of committing a Type I error. The analyses you carry out following a significant MANOVA are considered to be 'protected', because if the multivariate test is non-significant, then any subsequent tests are ignored. However, this notion of 'protection' is a little misleading because a significant MANOVA often reflects a significant difference for one rather than all dependent variables. Therefore, it is probably best to ensure against a Type I error, and there are several ways of doing so. One conservative way is to apply the Bonferroni correction. Normally, a result is regarded as 'significant' if the p value is less than .05. If our design involves two dependent variables and we want to look at the two ANOVAs performed on these, we apply the following correction: $.05 \div 2 =$.025, and for our result to be significant, p now has to be less than .025. If our design involves three dependent variables and we want to look at the three ANOVAs performed on these, we apply the following correction: $.05 \div 3 = .017$, and for our result to be significant, p now has to be less than .017. So, .05 is divided by the number of dependent variables in the study.

Another possibility is to explore a significant MANOVA result by conducting discriminant analysis (described in Chapter 12); see the following paragraph for more guidance on this.

When should I use MANOVA?

MANOVA can be used when your design is a simple one-way design – as demonstrated in this chapter – or with more complex designs where you have more than one independent variable or factor. There should be some conceptual reason for considering several dependent variables together in the same analysis. Adding dependent variables may decrease the power of the test, so MANOVA should only be used when there is a reason to measure several dependent variables.

There is some controversy over the extent to which the dependent variables can or should be correlated (see Cole, Maxwell, Arvey and Salas, 1994). Multicollinearity should be avoided (see Chapter 10, Section 1), so check that the correlation coefficients for any pair of dependent variables do not exceed .9, and correlations around .8 are also cause for concern. Tabachnick and Fidell (2014, 291) suggest that, in terms of the research design, one should avoid selecting dependent variables that are correlated because they essentially measure the same thing, albeit in a slightly different way. They write: 'MANOVA works best with highly negatively correlated DVs and acceptably well with moderately correlated DVs in either direction (about |.6|)' (p. 310). Certainly, if the dependent variables are highly positively correlated, and MANOVA shows a significant result, it is difficult to tease apart the contribution of each of the individual dependent variables to this overall effect. Rather than looking at the univariate ANOVAs, you would need to explore your data using a discriminant analysis, as this will allow you to explore the relationship between the dependent variables. We recommend that you perform tests of correlation to check the strength of the correlations between your dependent variables to help you decide whether or not to use MANOVA or how best to follow up any significant MANOVA result.

Checklist for using MANOVA

- 1. There should be a theoretical or empirical basis underpinning your choice of dependent variables.
- 2. The dependent variables should be measured using an interval or ratio scale, and any relationship between them should be linear (straight line). You can check this by looking at the scatterplots between pairs of dependent variables for each level of your factor. (If you are not sure how to generate scatterplots on SPSS, see Chapter 6, Section 2.)
- 3. You must ensure that the number of cases in each cell is greater than the number of dependent variables.
- 4. There should be *homogeneity of variance–covariance matrices*, and this is similar to the assumption of homogeneity of variance, mentioned previously in relation to parametric tests. SPSS can check this assumption for you, and we will show you how to do this in Section 4.
- 5. There should be univariate and multivariate *normality of distributions*. Assessment of multivariate normality is difficult in practice, and cannot be checked using SPSS; however, you should at least check that each dependent variable is normally distributed, that is, that univariate normality holds, as this is likely to reflect multivariate normality. Giles (2002) points to two ways in which normality may be violated. The first is *platykurtosis*, and this is evident when the distribution curve looks like a low plateau. You can check for this by generating histograms of each dependent variable. The second is the presence of *outliers*; these are data points far outside the area covered by the normal distribution. (See Chapter 3, Section 6; and/ or Tabachnick and Fidell, 2007, Chapter 4, for advice on screening for outliers.)

Generally, if you have equal sample sizes and a reasonable number of participants in each group, and you've checked for outliers before conducting your analysis, MANOVA will still be a valid test even with modest violations of these assumptions.

Section 4: PERFORMING MULTIVARIATE ANALYSIS OF VARIANCE IN SPSS

Example study: exposure to low levels of organophosphates

Let us return to the example from Section 2 and the study by Mackenzie Ross et al. (2010), which looked at whether low-level exposure to organophosphate pesticides brought about neuropsychological or psychiatric impairment. The researchers had asked working and retired sheep farmers to complete a range of cognitive and mood tests, and compared their performance to that of a matched control group, and to published test norms.

Their study examined two groups. Group 1 involved 127 sheep famers (working and retired) who had been exposed to low levels of organophosphate pesticides. Group 2 involved 78 controls (working and retired) who had not been exposed to these chemicals. The dependent variables we will look at here are scores on three subtests of the WAIS designed to measure working memory: the arithmetic test, the digit span test and the letter number substitution test.

The hypothesis we shall be testing is that there will be a difference between the groups in terms of their working memory. Analysis conducted by Mackenzie Ross et al. (2010) revealed that farmers were significantly impaired on measures of memory, including working memory, visual memory and auditory memory. Note that we are not replicating any analysis that Mackenzie Ross et al. undertook, and their data were carefully screened for outliers and transformed where necessary to ensure they met the requirements for parametric tests. Instead, we are using a subset of their data provided by the researchers to demonstrate how you would perform MANOVA. (The data file is available from macmillanihe.com/harrison-spss-7e.)

Before conducting the MANOVA procedure, we first check the correlations among the dependent variables. To obtain such correlations, click on <u>Analyze</u> \Rightarrow <u>Correlate</u> \Rightarrow <u>Bivariate</u> and then select your dependent variables (see Chapter 6 for more detail on obtaining correlations). Below is the SPSS output. Scatterplots would confirm that the relationships between the dependent variables are linear.

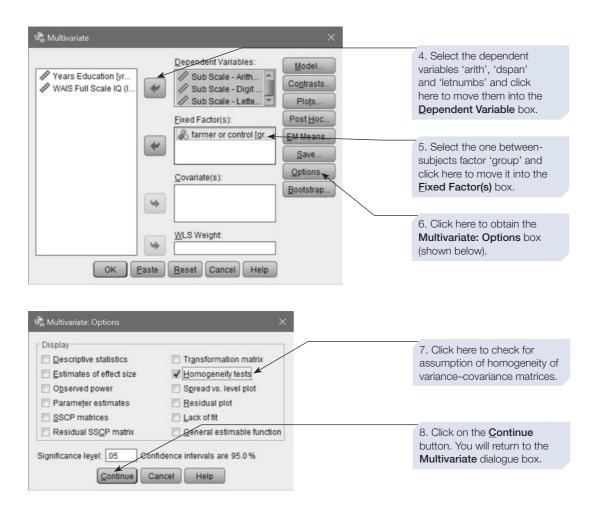
	Correl	ations		
		Sub Scale - Arithmetic	Sub Scale - Digit Span	Sub Scale - Letter Number Substitution
Sub Scale - Arithmetic	Pearson Correlation	1	.356	.374
	Sig. (2-tailed)		.000	.000
	N	197	197	197
Sub Scale - Digit Span	Pearson Correlation	.356	1	.577
	Sig. (2-tailed)	.000		.000
	Ν	197	201	201
Sub Scale - Letter	Pearson Correlation	.374	.577	1
Number Substitution	Sig. (2-tailed)	.000	.000	
	N	197	201	205

This correlation matrix suggests that there are moderate correlations between the different dependent variables.

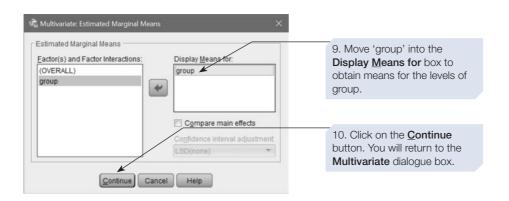
**. Correlation is significant at the 0.01 level (2-tailed).

How to perform MANOVA

🖷 ANCOVA	MANOVA new c	lata file.sav [Da	taSet1] - IBM SPSS Statistics Data Edito	- <u> </u>	
<u>Eile Edit</u>	<u>V</u> iew <u>D</u> ata	Transform	Analyze	xtensions <u>W</u> indow <u>H</u> elp	1. Click on <u>Analyze</u> .
8			Reports Descriptive Statistics Bayesian Statistics	Visibl	2. Select <u>General</u> Linear Model.
	& group		Compare Means	tnumbs var var	Linear Wodel.
1	1	10	General Linear Model		
2	1	14	Generalized Linear Models	Univariate	
3	2	18	-	Multivariate.	3. Select <u>Multivariate</u> .
4	2	11	Mixed Models	Repeated Measures	You will then see the
5	2	13	<u>C</u> orrelate	Variance Components	Multivariate dialogue
6	2	11	Regression	11.00	box.
7	2	11	Loglinear	14.00	
8	2	11	Neural Networks	12.00	
9	2	11	Classify	10.00	
10	2	11	Dimension Reduction	11.00	
11	2	11	Sc <u>a</u> le	▶ 10.00	



You will return to the **Multivariate** dialogue box. Click on the **<u>EM</u> Means** button to bring up the **Estimated Marginal Means** dialogue box (shown below).



Click on or and SPSS will calculate the MANOVA. The output is shown next.

SPSS output for MANOVA

General Linear Model

Between-Subjects Factors

		Value Label	Ν
farmer or control	1	farmer	122
	2	control	75

Box's Test of Equality of Covariance Matrices^a

8.416
1.377
6
163196.327
.219

that the observed covariance matrices of the dependent variables are equal across groups.

> a. Design: Intercept + group

Box's test checks whether your data violate the assumption of homogeneity of variance–covariance matrices and was obtained by clicking on <u>Homogeneity tests</u> in the **Multivariate: Options** dialogue box. If this is significant, you have violated this assumption. As this test is very sensitive, it is most useful when your sample size is small and unequal. Here, p (sig. value) = .219, so we have not violated this assumption.

Effect		Value	F	Hypothesis df	Error df	Sig.
Intercept	Pillai's Trace	.985	4162.385 ^b	3.000	193.000	.000
	Wilks' Lambda	.015	4162.385 ^b	3.000	193.000	.000
	Hotelling's Trace	64.700	4162.385 ^b	3.000	193.000	.000
	Roy's Largest Root	64.700	4162.385 ^b	3.000	193.000	.000
group	Pillai's Trace	.204	16.469 ^b	3.000	193.000	.000
	Wilks' Lambda	.796	16.469 ^b	3.000	193.000	.000
	Hotelling's Trace	.256	16.469 ^b	3.000	193.000	.000
	Roy's Largest Root	.256	16.469 ^b	3.000	193.000	.000

Multivariate Tests^a

a. Design: Intercept + group

b. Exact statistic

In this table, we are only interested in the results for the variable 'group', and we ignore those reported for the Intercept. SPSS reports the four MANOVA test statistics that tell us whether the new combined dependent variable, 'working memory', is different across the two groups of participants. Here, *p* is smaller than .05 for each test statistic, so all are significant. We are going to report the values for Wilks' Lambda, the row highlighted, citing the *F* value, the two sets of degrees of freedom and *p*.

These statistics were obtained by clicking on <u>Homogeneity tests</u> in the **Multivariate: Options** dialogue box. If Levene's test **Based on Mean** p > .05, as here, then there is equality of variance. This is important in terms of the reliability of the results below and in supporting the robustness of the multivariate statistics.

Levene's Test of Equality of Error Variances^a

\checkmark		Levene Statistic	df1	df2	Sig.
Sub Scale - Arithmetic	Based on Mean	.030	1	195	.862
	Based on Median	.038	1	195	.845
	Based on Median and with adjusted df	.038	1	194.594	.845
	Based on trimmed mean	.043	1	195	.836
Sub Scale - Digit Span	Based on Mean	.005	1	195	.943
	Based on Median	.004	1	195	.949
	Based on Median and with adjusted df	.004	1	191.941	.949
	Based on trimmed mean	.002	.030 1 .038 1 .038 1 .038 1 .038 1 .038 1 .038 1 .038 1 .043 1 .005 1 .004 1 .004 1 .004 1 .004 1 .002 1 .553 1 .423 1	195	.964
Sub Scale - Letter	Based on Mean	.553	1	195	.458
Number Substitution	Based on Median	.423	1	195	.516
	Based on Median and with adjusted df	.423	1	192.766	.516
	Based on trimmed mean	.533	1	195	.466

Tests the null hypothesis that the error variance of the dependent variable is equal across groups.

a. Design: Intercept + group

Source	Dependent Variable	Type III Sum of Squares	df	Mean Square	F	Sig.	
Corrected Model	Sub Scale - Arithmetic	.096 ^a	1	.096	.012	.913	
	Sub Scale - Digit Span	3.853 ^b	1	3.853	25.349	.000	
ntercept group	Sub Scale - Letter Number Substitution	223.241°	1	223.241	34.855	.000	
group Error	Sub Scale - Arithmetic	24387.812	1	24387.812	3058.878	.000	
	Sub Scale - Digit Span	1884.565	1	1884.565	12397.739	.000	
	Sub Scale - Letter Number Substitution	21021.617	1	21021.617	3282.098	.000	
group	Sub Scale - Arithmetic	.096	1	.096	.012	.913	
	Sub Scale - Digit Span	3.853	1	3.853	25.349	.000	
	Sub Scale - Letter Number Substitution	223.241	1	223.241	34.855	.000	
Error	Sub Scale - Arithmetic	1554.695	195	7.973			
	Sub Stale - Digit Span	29.642	195	.152			
	Sub Scale - Letter Number Substitution	1248.962	195	6.405			
Total	Sub Scale - Arithmetic	27390.000	197	The row	labelled	'aroun'	
	Sub Scale - Digit Span	1988.921	197		 The rows labelled 'group each of your dependent the univariate ANOVAs, a 		
	Sub Scale - Letter Number Substitution	22680.000	197				
Corrected Total	Sub Scale - Arithmetic	1554.792	196	these a	s you woul	d for a d	
	Sub Scale - Digit Span	33.495	196	there ar	e three de	pendent	
	Cub Caple Letter	4 4 7 2 2 2 2	400	Development		land lave of	

Tests of Between-Subjects Effects

 Sub Scale - Arithmetic
 1554.792
 196

 Sub Scale - Digit Span
 33.495
 196

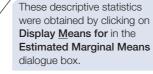
 Sub Scale - Letter Number Substitution
 1472.203
 196

 a. R Squared = .000 (Adjusted R Squared = .005)

b. R Squared = .115 (Adjusted R Squared = .111)

c. R Squared = .152 (Adjusted R Squared = .147)

The rows labelled 'group' provide values for each of your dependent variables. These are the univariate ANOVAs, and you interpret these as you would for a one-way ANOVA. As there are three dependent variables, we apply Bonferroni correction by dividing 0.05 by 3, so sig. values need to be smaller than 0.017 for results to be significant. This is the case for two of the three dependent variables, digit span and letter number substitution.



Estimated Marginal Means

farmer or control

Dependent Variable	farmer or control	Mean	Std. Error	95% Confidence Interval	
				Lower Bound	Upper Bound
Sub Scale - Arithmetic	farmer	11.434	.256	10.930	11.939
	control	11.480	.326	10.837	12.123
Sub Scale - Digit Span	farmer	3.041	.035	2.971	3.111
	control	3.329	.045	3.240	3.418
Sub Scale - Letter Number Substitution	farmer	9.541	.229	9.089	9.993
	control	11.733	.292	11.157	12.310



A one-way between-subjects multivariate analysis of variance was carried out to assess the impact of exposure to low levels of organophosphate pesticides on working memory. The between-subjects factor comprised two groups: farmers who had been exposed to the pesticides and a control group, who had not. The dependent variables comprised scores on three different measures of working memory, all of which were subtests of the WAIS. Assumptions of homogeneity of variance-covariance matrices and equality of variance were confirmed, and moderate correlations were found among the dependent variables. There was a significant difference between the two groups on the combined dependent variable 'working memory', F(3,193) = 16.47, p < .001; Wilks' Lambda = .8. Analysis of each individual dependent variable, using a Bonferroni adjusted alpha level of .017, showed that there was no significant contribution of the subtest arithmetic, F(1, 195)= 0.01, p = .913. The two groups differed significantly on the other two subtests: digit span, F(1,195) = 25.35, p < .001 and letter number substitution, F(1,195) = 34.86, p < .001. The mean scores for those exposed to the chemicals was lower for both subtests (M = 3.04 and M = 9.54, respectively) compared with those of the control group (M = 3.33 and M = 11.73, respectively).

You would also want to report a measure of effect size in your results section (see Chapter 8, Section 1), your descriptive statistics and information regarding the confidence intervals for each condition. Furthermore, the example shown here involved a between-subjects factor that involved only two levels. If your design involves a factor with more than two levels, and significant differences emerge, you will need to follow these up to find out where the significant differences lie. The easiest way would be to conduct one-way between-subjects ANOVAs on any dependent variable that emerges as significant in the MANOVA, and conduct unplanned (or planned) comparisons, as shown in Chapter 8, Section 2 for one-way between-subjects design and Chapter 8, Section 3 for a one-way within-subjects design.

Next, we finish this section by looking at how to perform MANOVA with one within-subjects factor.

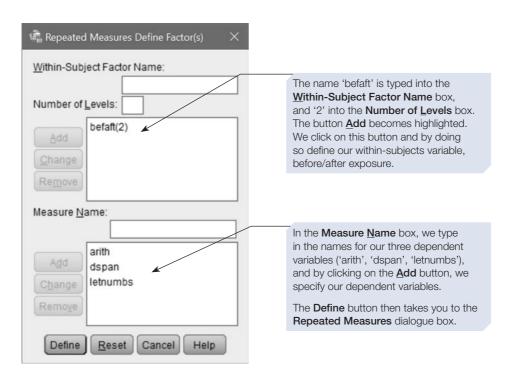
https://cochrana.ir/SPSS

A note on within-subjects designs

As mentioned previously, MANOVA can be used when your design is a simple oneway design or with more complex designs where you have more than one independent variable or factor. We have shown you how to perform this test when the design involves one between-subjects factor, and we will end this chapter by providing a note on how to conduct this test when the design involves one within-subjects factor.

Imagine our design involved a within-subjects factor, before and after being exposed to chemicals, and three dependent variables each measuring working memory. Our data file would contain six SPSS variables representing the three 'before' scores and the three 'after' scores. In the dialogue box shown below, we name the variables as follows: 'befaft' for the within-subjects factor (before and after exposure to chemicals), and the three dependent variables are 'arith', 'dspan' and 'letnumbs' (the arithmetic test, the digit span test and the letter number substitution test).

We would start analysing the data in the following way: Click on <u>Analyze</u> \Rightarrow <u>General Linear Model</u> \Rightarrow <u>Repeated Measures</u>.



We would then get the dialogue box shown in Chapter 8, Section 3, for the oneway within-subjects ANOVA.

Within-Subjects Variables (befaft): Contrasts aft_arith(1,arith) aft_arith(2,arith)	Move the six variables across into the appropriate slots in the <u>Within-Subjects Variables</u> box.
bef_dspan(1,dspan) aft_dspan(2,dspan) bef_letnumbs(1,letnum aft_letnumbs(2,letnumbs) Between-Subjects Factor(s):	Click on EM Means to produce descriptive statistics. As there are no between- subjects factors in this design, we do not need to test the assumption of homogeneity of variance through the Options box.
Covariates:	Finally, click on OK .

The output will look similar to that shown previously, and will provide values for the combined dependent variable followed by univariate ANOVA test statistics.

Summary

- This chapter introduced you to two different statistical procedures related to ANOVA.
- ANCOVA allows you to look at the effect of one or more factors on a dependent variable, while removing the effect of a covariate.
- MANOVA allows you to look at the effect of one or more factors when there is more than one dependent variable.
- We have shown you how to conduct both tests using SPSS when there is one between-subjects factor.
- Both make certain assumptions about the data, listed here, and we have shown you how to check for these.
- If your dependent variable is a total score from several raw scores that have already been entered into your data file, see Chapter 4 for guidance on how to create such a total score in SPSS.
- For guidance on incorporating SPSS output into a report, or on printing the output, see Chapter 14.

12 Discriminant analysis and logistic regression

In this chapter

- Discriminant analysis and logistic regression
- An introduction to discriminant analysis
- Performing discriminant analysis in SPSS
- An introduction to logistic regression
- Performing logistic regression in SPSS

SPSS for Psychologists online

Visit macmillanihe.com/harrison-spss-7e for data sets, online tutorials and exercises.

Section 1: DISCRIMINANT ANALYSIS AND LOGISTIC REGRESSION

- Discriminant analysis and logistic regression are statistical procedures that can be used to predict category membership from a number of predictor variables.
- Like multiple regression, these procedures predict an outcome. In multiple regression, the outcome variable is a continuous variable, such as score on a test. In contrast, discriminant analysis and logistic regression procedures can be employed when the outcome variable is categorical, for example pass vs fail on a test.
- The two procedures are designed for use when we want to predict a categorical outcome on the basis of several other variables. For example, we might try to predict whether someone is going to pass or fail their driving test on the basis of a number of predictor variables, such as their age, sex, educational achievement and employment history.
- Discriminant analysis makes various assumptions about the predictor variables, including that they should be normally distributed. Logistic regression makes no such assumptions about the distribution of the predictor variables.

An example

Consider a forensic psychologist working with prison inmates. The psychologist might be interested in identifying factors that predict whether or not a particular inmate will be reconvicted after release from prison. To undertake this analysis, the psychologist would first collect data from the records of past prisoners. These data would include the outcome (whether or not the prisoner was reconvicted after release) in addition to a number of other variables that the psychologist thinks may influence the chances of reconviction. These predictor variables might include the length of the sentence served, a measure of the behaviour of the prisoner when in custody (such as the number of days 'on report'), a measure of the level of drug use and a measure of social support (such as number of hours of prison visits per month). The psychologist could use either discriminant analysis or logistic regression to analyse these data. Although these two procedures are technically very different, their outcomes are similar, in that each allows the psychologist to identify which combination of these variables (if any) is useful in predicting reconviction. The psychologist could then apply this knowledge to all new prisoners entering the jail in order to identify those who are most likely to be reconvicted after their release and to target resources towards these high-risk prisoners in an attempt to reduce recidivism.

Similarities and differences between discriminant analysis and logistic regression

Both discriminant analysis and logistic regression allow us to predict a categorical dependent variable on the basis of a number of predictor or independent variables. These independent variables are normally continuous variables, but logistic regression can also handle categorical independent variables. In general, logistic regression can be used in a wider range of situations than discriminant analysis. For example, discriminant analysis makes various assumptions about the predictor variables, including that they should be normally distributed. Logistic regression makes no such assumptions about the distribution of the predictor variables.

Both logistic regression and discriminant analysis can be used in situations where the outcome variable has more than two categories. However, in cases with more than two categories, the output becomes more difficult to interpret. For the purpose of this book, we are only going to consider cases with dichotomous outcome variables.

An important distinction lies in the interpretation of the output. The output of discriminant analysis is rather more difficult to interpret than that of logistic regression. Discriminant analysis results in the calculation of a discriminant function – a formula that combines the predictor variables to calculate a value that is then used to predict category membership. The value of the discriminant function is somewhat arbitrary and tells us relatively little about the basis on which the prediction is being made. Although logistic regression is a more complex procedure, the output is rather easier to interpret. Logistic regression computes the probability (actually the log odds – see Section 4 for more details) that a case will belong to a particular category.

As a result of these advantages, many researchers now recommend the use of logistic regression over discriminant analysis. However, we describe discriminant analysis because it is still commonly employed in some fields of psychology.

Section 2: AN INTRODUCTION TO DISCRIMINANT ANALYSIS

Tabachnick and Fidell (2014) note that discriminant analysis is rather like reverse MANOVA. Indeed, discriminant analysis could be described as a cross between backwards analysis of variance and multiple regression. In ANOVA designs, we manipulate membership of some group or category (the levels of the IV or factor) and investigate the effect on the DV. If we find a significant effect of one of our IVs, then we could say that we can partially predict a participant's DV score if we know which category they belong to. A similar attempt to predict category membership is at the heart of discriminant analysis.

Discriminant analysis is similar to multiple regression in that both techniques allow us to make predictions on the basis of several predictor variables. The difference is that, while multiple regression is used to predict a participant's score on the DV, discriminant analysis is used to predict which of a number of categories or groups the participant will belong to.

In discriminant analysis, the dependent variable is sometimes also referred to as the 'category variable', the 'criterion variable' or the 'grouping variable'. The other variables are independent variables or predictor variables.

An example

Let us return to our earlier example of a forensic psychologist working with prison inmates. It would be useful to predict which of these individuals are at the greatest risk of being reconvicted following release. This would allow us to target treatment towards these high-risk individuals and produce a cost-effective programme to reduce reconviction rates. First, we will need to collect data from each of the prisoners on a number of likely variables such as age at release, number of previous convictions and the level of drug use. You can now investigate whether some weighted combination of these variables reliably discriminates between those prisoners who are reconvicted following release, and those who are not reconvicted. This combination of the predictor variables is called a *discriminant function* (it is a mathematical function that discriminates between the categories). Having calculated this discriminant function, you could apply it to all new inmates entering the prison, and, based on their score on each of the predictor variables, make a prediction regarding their chances of reconviction following release. You might then go on to develop a treatment programme that seeks to reduce reconviction by directly tackling some of the variables you have shown to significantly predict reconviction. For example, if drug use was a variable that predicted reconviction, you might consider including an intensive drug and alcohol module as part of the programme.

Two steps in discriminant analysis

As the above example shows, there are often two distinct steps involved in discriminant analysis:

1. Use cases where category membership is already known to develop a discriminant function that can reliably predict category membership in these cases.

2. Use this discriminant function to predict category membership for a new group of cases for whom category membership is not known.

A discriminant function is a mathematical formula that combines a set of predictor variables to predict the value of a categorical variable. For example, in the output shown in Section 3, the discriminant function (the variable named 'Dis1_1') has a high positive value for prisoners who we predict will be reconvicted and a high negative value for prisoners who we predict will not be reconvicted. The higher the value, the stronger the prediction. Once derived, the function can be applied to new prisoners, allowing us to make an informed prediction of their behaviour at the end of their sentence.

Assumptions

The category variable can have two or more distinct levels. Category membership must be mutually exclusive (each case must be classified into no more than one category) and collectively exhaustive (every case must be a member of a category). The requirements for predictor or independent variables are similar to those for dependent variables in MANOVA, but some violation of the rules regarding the distribution of these variables may be acceptable, especially if the analysis results in accurate classification. If the rate of prediction of classification is not acceptable, this might be due to the violation of some of these assumptions, especially those regarding outliers and homogeneity of variance (see Tabachnick and Fidell, 2014).

Methods in discriminant analysis

As in multiple regression, there are different methods that can be adopted. SPSS allows you to choose between the 'enter' method (also known as the 'simultaneous' method) and the stepwise method (also called 'statistical').

Choosing a method to adopt

- 1. Unless you have some good reason to do otherwise, you should use the enter (simultaneous) method.
- 2. The stepwise (statistical) method can be used to produce a discriminant function that includes the minimum number of predictor variables.

What does each method tell us?

Suppose we were seeking to predict reconviction based on age, previous convictions and drug use. The two types of analysis would give us slightly different information about the data. The enter (simultaneous) method would tell us how good a prediction we can make on the basis of these three predictor variables together. In addition, we would be able to see how much each of the predictor variables contributes to the discriminant function (and hence the accuracy of our prediction). In contrast stepwise or statistical methods would allow us to identify the best combination of predictor variables to use to predict category membership.

How does each method work?

- 1. Enter (simultaneous): All the variables are entered simultaneously and the predictive power of the combination of all the variables is considered.
- 2. Stepwise (statistical): If there are no theoretical grounds for predicting the relative importance of the variables, then stepwise can be used to determine the smallest useful set of predictor variables. In stepwise, the variables are entered and/or removed based on statistical assessments of their importance. However, just like stepwise multiple regression, this approach can be dangerous. The variables adopted, and hence the predictions made, can be influenced by minor variation in the predictor variables. Just as with multiple regression, if you choose to adopt a statistical method in discriminant analysis, then you should double-check the validity of your discriminant function by using cross-validation procedures. Discriminant analysis (especially when you use statistical methods) tends to overestimate the success of the discriminant function. Cross-validation reduces this overestimation by checking the validity of the discriminant function derived. There are two basic approaches to cross-validation:
 - a. We can calculate a discriminant function based on one half of our data and test it out on the other half (a little like split-half reliability).
 - b. We can test the ability of the discriminant function to classify the same cases measured at some second time interval. This is a little like test-retest reliability.

Within the statistical methods, there are a variety of statistical criteria to adopt. These are the criteria by which SPSS decides which predictor variables to enter and/or remove from the discriminant function. This is a complex subject area that is covered in some detail by Tabachnick and Fidell (2014). If in doubt, we would advise you to use the default settings in SPSS.

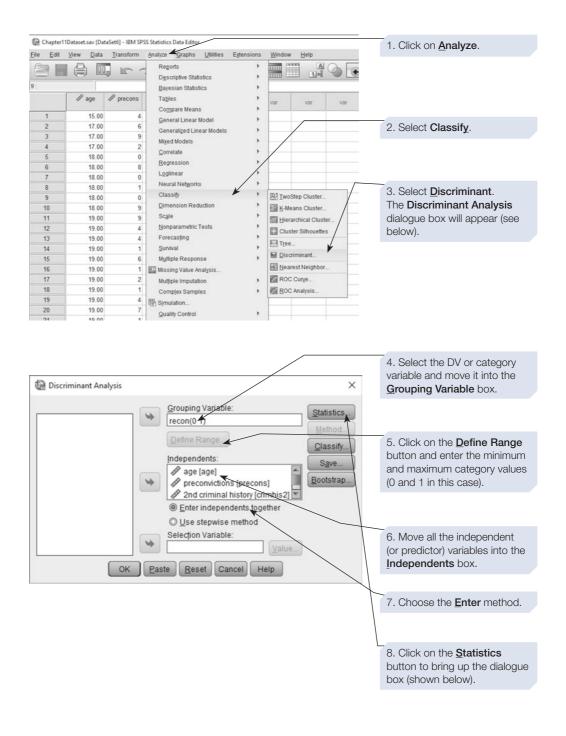
Section 3: PERFORMING DISCRIMINANT ANALYSIS IN SPSS

Example study: reconviction among offenders

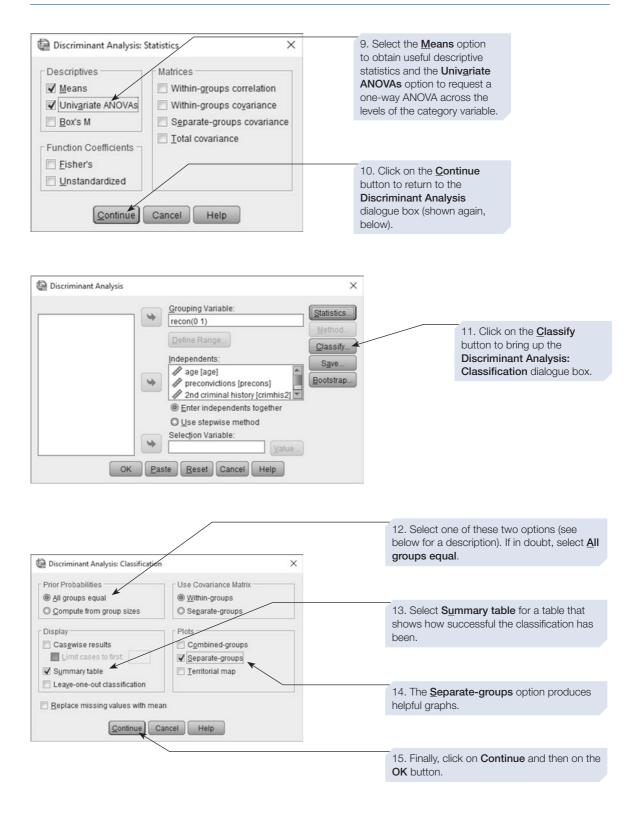
Hollin, Palmer and Clark (2003) collected data from 221 prison inmates in an attempt to identify the variables that predict reconviction among offenders. The data file we have constructed for this chapter contains just a few of the variables they measured. The variable 'age' records the prisoner's age in years, and 'precons' the number of previous convictions. The variable 'recon' is a categorical variable that records whether the prisoner was reconvicted after release. The value 1 indicates reconviction, while 0 indicates that the prisoner was not reconvicted. The variables 'crimhist2' and 'educemp2' are derived from a scale called the LSI-R (Level of Service Inventory – Revised), which is used by psychologists working in prisons. The LSI-R measures several aspects of an offender's life, including previous offending, drug use, behaviour in prison, family history and education and employment. In the data file we will be using here, we have only included the measures of criminal history (crimhist2) and education and employment history (educemp2). We will use discriminant analysis to

determine whether age, previous convictions, criminal history and education and employment history can be used to predict reconviction. These data are available from macmillanihe.com/harrison-spss-7e.

To perform a simultaneous (enter method) discriminant analysis







The annotated output is shown on the next page.



In the **Discriminant Analysis: Classification** dialogue box, you can set the **Prior Probabilities** to either **All groups equal** or **Compute from group sizes**. The default setting is **All groups equal**. In this case, the analysis assumes that, all other things being equal, the probability of a case falling into any particular category is equal. Thus, if you have two groups, it assumes that the prior probability is 0.5. For three groups, it would be 0.333 and so on. There may be occasions when the probability of each outcome might not be the same. For example, you might decide that it is rather more likely that a patient will live rather than die regardless of the treatment you offer. In cases like this, it might be more appropriate to ask SPSS to compute the prior probabilities on the basis of the number of cases that fall into each outcome category. Thus, if you have 100 cases with 60 falling into one category and 40 into another, then the prior probabilities would be set at 0.6 and 0.4, respectively. If in doubt, it is safer to leave this option at the default setting of **All groups equal**.

SPSS output for discriminant analysis using enter method

Obtained using menu items: Classify > Discriminant (enter independents together)

Discriminant

Analysis Case Processing Summary

Unweighte	d Cases	Ν	Percent
Valid		221	100.0
Excluded	Excluded Missing or out-of-range group codes	0	.0
	At least one missing discriminating variable	0	.0
	Both missing or out-of- range group codes and at least one missing discriminating variable	0	.0
	Total	0	.0
Total		221	100.0

This table tells you that 100% of the 221 cases in the data file have been included in the analysis.

If any case had a missing value for one of the IVs (the predictor variables), the case would have been dropped from the analysis and this would have been reported in this table.

•	N.
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Group Statistics Valid N (listwise) Unweighted Weighted Mean Std. Deviation reconvicted 157.000 31.9873 12.86966 157 no age preconvictions 157 000 4 9809 6.44451 157 2nd criminal history 5.5796 3.17063 157 157.000 2nd education and 4.6115 2.86144 157 157.000 employment yes age 26.0000 8.32285 64 64,000 preconvictions 7.7656 6.55802 64 64.000 2nd criminal history 2.88383 64.000 7.5313 64 2nd education and 6.0938 2.58026 64 64.000 employment Total age 30.2534 12.02872 221 221.000 preconvictions 5.7873 6.58545 221 221.000 6.1448 3.20891 2nd criminal history 221 221.000 2nd education and 2.85771 221.000 5.0407 221 employment

This is the table of means we requested in the Discriminant Analysis: Statistics dialogue box. It gives the mean and SD for each of our IVs broken down by category membership. For example, we can see the mean age of those not reconvicted was 31.99 years compared with 26 years for those who were reconvicted.

1	ests of Equality	of Group	Means		
	Wilks' Lambda	F	df1	df2	Sig.
	.949	11.818	1	219	.001
onvictions	.963	8.403	1	219	.004
criminal history	.924	18.127	1	219	.000
education and	.944	12.894	1	219	.000

This table was produced because we requested Univariate ANOVAs in the **Discriminant Analysis: Statistics** dialogue box. It shows whether there is a significant effect of category for each of the predictor variables. For example, here we can see that there is a significant difference in the age of those reconvicted and not reconvicted F(1, 219) = 11.818, p = 0.001. In addition, SPSS gives Wilks' Lambda, a multivariate test of significance, which varies between 0 and 1. Values very close to 1 indicate that the differences are not significant.

Analysis 1

age

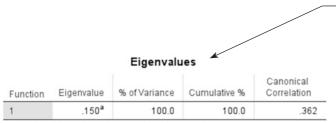
prec

2nd

2nd employment

Summary of Canonical Discriminant Functions

Canonical functions are used to discriminate between pairs of categories, or groups. For each possible orthogonal (independent) contrast, a discriminant function is calculated that best discriminates between the categories. The number of canonical functions is either one less than the number of categories, or equal to the number of predictor variables, whichever is the smaller. The following tables give details of each of the discriminant functions calculated. In this example, there are only two categories, so only one function is calculated.



a. First 1 canonical discriminant functions were used in the analysis.

The eigenvalue is a measure of how well the discriminant function discriminates between the categories (the larger the value, the better the discrimination).

The % of Variance column allows you to compare the relative success of the functions. When there is only one function (as here), this column and the Cumulative % column tell us nothing useful. Where there are several functions, you will probably find that only the first few usefully discriminate among groups.

Wilks' L	ambda 🖌			This table provides a test of the null hypothesis that the value of the discriminant function is the same for the reconvicted and non-reconvicted cases. As p is less than 0.05, we can reject the null hypothesis.
ambda	Chi-square	df	Sig.	

.000

Standardized Canonical
Discriminant Function
Coefficients
Function

Wilks' Lambda

.869

30.421

4

Test of Function(s)

	Function	
	1	
age	601	
preconvictions	.249	
2nd criminal history	.547	
2nd education and employment	.174	

This table allows you to see the extent to which each of the predictor variables is contributing to the ability to discriminate between the categories. The coefficients have been standardised so that you can compare the contribution of each regardless of the units in which it was measured. Rather like correlation coefficients, the values range from -1 to +1. In this case, 'age' and '2nd criminal history' are making a larger contribution than the other predictor variables.

Structure Matrix

	Function	
	1	
2nd criminal history	.742	
2nd education and employment	.625	
age	599	
preconvictions	.505	

Pooled within-groups correlations between discriminating variables and standardized canonical discriminant functions

Variables ordered by absolute size of correlation within function.

The Structure Matrix table gives a different measure of the contribution that each variable is making to the discriminant function. Here, the variables are ordered by the magnitude of their contribution. The negative value for the variable 'age' tells us that age correlates negatively with the value of the function, whereas '2nd criminal history' correlates positively. This is because older prisoners are less likely to be reconvicted, whereas prisoners with a higher criminal history score are more likely to be reconvicted.

If you have more than two categories, and hence more than one function calculated, this table also allows you to see which of the functions each variable is contributing most to (for example, is a particular variable helping you predict between membership of category a and b or between b and c?).

Functions at K Group Centroids

	Function
reconvicted	1

reconvicted	
no	247
yes	.605
Unstandardiz	ed
canonical dis	criminant
functions eval	luated at

This table gives the mean value of the discriminant function for each of the categories. Note that here the mean value of the function is positive for reconvicted prisoners but negative for non-reconvicted prisoners. In this way, the function is discriminating between the two categories of prisoners.

Classification Statistics

group means

Classi	fication Processing Sun	nmary
Processed	I	221
Excluded	Missing or out-of-range group codes	0
	At least one missing discriminating variable	0
Used in Ou	utput	221

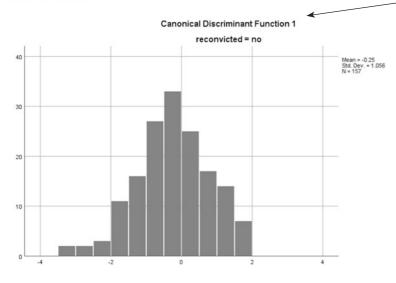
This table informs you of the total number of cases processed, the number excluded and the number used in the output.

Prior Probabilities for Groups 🛩

		Cases Used in Analysis		
reconvicted	Prior	Unweighted	Weighted	
no	.500	157	157.000	
yes	.500	64	64.000	
Total	1.000	221	221.000	

Prior probability is the assumed probability that a particular case will belong to a particular category. In this case, in the **Discriminant Analysis: Classification** dialogue box (see above), under Prior Probabilities, we selected 'All groups equal', so the probability is equal for all groups – i.e. 0.5 or 50%.

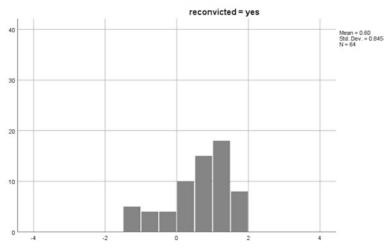
Separate-Groups Graphs



These are the Separate-Groups Plots we requested in the **Discriminant Analysis: Classification** dialogue box (see above). Note we have edited the plots so that the Y-axis is drawn to the same scale on both.

If the discriminate function is discriminating between the two groups of inmates, then the distribution will be different in these two plots. It is apparent that the distribution for those who were reconvicted is slightly shifted to the right, relative to those who were not reconvicted, but there is a lot of overlap between the two groups, indicating the discrimination is far from perfect.

Canonical Discriminant Function 1





Predicted Group Member

		reconvicted	no	yes	Total
Original	Count	no	104	53	157
		yes	16	48	64
	%	no	66.2	33.8	100.0
		yes	25.0	75.0	100.0

a. 68.8% of original grouped cases correctly classified.

This is the Summary table we requested in the Discriminant Analysis: Classification dialogue box (see above). It provides a particularly useful summary of the success (or otherwise) of our discriminant function. It shows a cross-tabulation of the category membership (reconvicted or not) against what we would have predicted using our discriminant function.

We can see that in 104 cases the discriminant function correctly predicted that the offender would not be reconvicted, and in 48 cases it correctly predicted that they would be reconvicted. Thus, 152 (104 + 48) of our 221 cases were correctly classified - a success rate of 68.8% (as noted in the footnote to the table). However, the table also shows us that 25% of the prisoners we predicted would not be reconvicted were reconvicted, and that 33.8% of the cases we predicted would be reconvicted were not. It is up to you to interpret these failures of prediction; in some cases it may be more important to avoid one type of error than another. For example, here you may feel that it is more important to avoid erroneously predicting that someone will not be reconvicted than erroneously predicting that they will.



A discriminant analysis was performed with reconviction as the DV and age, number of previous convictions, and the criminal history and education and employment subscales of the LSI-R as predictor variables. A total of 221 cases were analysed. Univariate ANOVAs revealed that the reconvicted and non-reconvicted prisoners differed significantly on each of the four predictor variables. A single discriminant function was calculated. The value of this function was significantly different for reconvicted and non-reconvicted prisoners (chi-square = 30.42, df = 4, p < .001). The correlations between predictor variables and the discriminant function suggested that age and criminal history were the best predictors of future convictions. Age was negatively correlated with the discriminant function value, suggesting that older prisoners were less likely to be reconvicted. Criminal history was positively correlated with the discriminant function value, suggesting that prisoners with higher numbers of previous convictions were more likely to be reconvicted. Overall, the discriminant function successfully predicted the outcome for 68.8% of cases, with accurate predictions being made for 66.2% of the prisoners who did not go on to be reconvicted and 75% of the prisoners who were reconvicted.

To perform a stepwise (or statistical) discriminant analysis

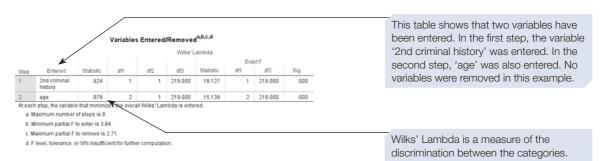
To perform a stepwise discriminant analysis, follow the procedure outlined for the enter method except that, at step 7 (see above, page 358), select <u>Use stepwise method</u>, then click on the <u>Method</u> button to bring up the Discriminant Analysis: Stepwise Method dialogue box (see below).

Discriminant Analysis: Stepwise Method X	statistical rules used to determine when a variable is
Method Criteria © Wilks' lambda © Unexplained variance © Mahalanobis distance	entered into the equation. If in doubt, select <u>W</u> ilks' lambda, the default setting.
Smallest F ratio Use grobability of F Rao's V Entry: 05 V-to-enter: 0	The criteria for entry and removal from the equation can be based on either <i>F</i> values or probability values, and the values for both can be adjusted. Tabachnick and Fidell (2014) suggest that the Entry probability
Display Summary of steps F for pairwise distances Continue Cancel	value could be changed from .05 to .15. This is a more liberal rule that will ensure that any important variable gets entered into the equation.

If in doubt, leave the Method set to Wilks' lambda, and the Criteria set to Use <u>F</u> Value, with the <u>Entry</u> and <u>Removal</u> values set to 3.84 and 2.71, respectively (see above). Then click on the <u>Continue</u> button and complete steps 8–15 (see above, pages 358–359).

The output produced by the stepwise discriminant analysis is similar to that for the simultaneous discriminant analysis, and so only tables that differ are shown below. This output was produced using the default settings.

> A smaller value indicates better discrimination. The value of Lambda falls as the second variable, 'age', is entered.



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	Variab	les in the <i>l</i>	Analysis	
р		Tolerance	F to Remove	Wilks' Lambda
	2nd criminal history	1.000	18.127	
	2nd criminal history	1.000	17.566	.949
	age	1.000	11.299	.924

Step

1

2

This table shows variables included in the analysis at each step. In this example, there are only two steps. The other variables were not added because they did not meet the entry criteria we set in the **Discriminant Analysis: Stepwise Method** dialogue box (see above).

Variables Not in the Analysis 🔫

Step		Tolerance	Min. Tolerance	F to Enter	Wilks' Lambda
0	age	1.000	1.000	11.818	.949
	preconvictions	1.000	1.000	8.403	.963
	2nd criminal history	1.000	1.000	18.127	.924
	2nd education and employment	1.000	1.000	12.894	.944
1	age	1.000	1.000	11.299	.878
	preconvictions	.794	.794	1.080	.919
	2nd education and employment	.724	.724	2.326	.914
2	preconvictions	.789	.789	1.609	.872
	2nd education and employment	.700	.700	.826	.875

This table lists all the variables not included in the equation at each step.

This table indicates how successful the discriminant function is at each step. A smaller value of Lambda indicates a better discrimination between categories. Here, the function significantly distinguishes between the two categories at both steps 1 and 2.

		_		Wilks' La	mbda				
	Number of						Exa	ct F	
Step	Variables	Lambda	df1	df2	df3	Statistic	df1	df2	Sig.
1	1	.924	1	1	219	18.127	1	219.000	.000
2	2	.878	2	1	219	15.139	2	218.000	.000

This is the Classification Results table for the stepwise discriminant analysis. It is interesting to compare this with the equivalent table produced using the enter method (see above). Which method results in the most successful prediction?

Classification Results^a

		F	Predicted Group Membership				
		reconvicted	no	yes	Total		
Original	Count	no	102	55	157		
		yes	19	45	64		
	%	no	65.0	35.0	100.0		
		yes	29.7	70.3	100.0		

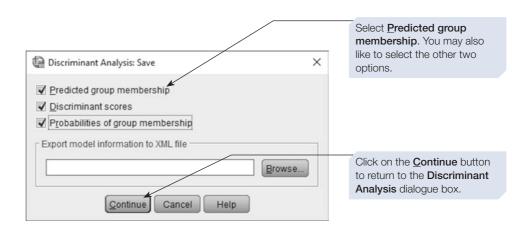
a. 66.5% of original grouped cases correctly classified.

Using discriminant function analysis to predict group membership

So far we have been trying to develop discriminant functions that can predict category membership with a reasonable level of accuracy. Having produced these functions, the next step is to use them to try to make real predictions. For example, if we collect data from another group of prisoners before they are released, we can apply our discriminant function to predict which of these prisoners will be reconvicted following release.

To do this we need to compute the discriminant function for the new cases. The easiest way to do this is to add the new cases to the existing data file. Because we do not know whether these individuals will be reconvicted, we will have to enter a missing value for the variable 'recon'. As a result, these new cases will not be included when the discriminant function is calculated, and therefore, the result will be identical to that found before these cases were added.

Follow steps 1–15 listed above (pages 358–359), selecting <u>Use stepwise method</u> at step 7, but then click on the <u>Save</u> button in the <u>Discriminant Analysis</u> dialogue box. The <u>Discriminant Analysis</u>: <u>Save</u> dialogue box will be revealed (see below).



Now click on the own button. In addition to the output described above, several new variables will be computed and added to your data file (see below).

	variable Dis _ predicted ca hip.		N		the value of the discriminant fu that the value i	nction. Notice s higher for are predicted to
Statistics Dat	ta Editor aphs Utilitie	es E <u>x</u> tension	s Window	Help		
B recon	& crimhis2	educemp	& Dis_1	Dis1_1	Ø Dis1_2	Dis2_2
0	8.00	8.00	1	1.29516	.28525	.71475
0	2.00	4.00	0	33125	.60150	.39850
1	4.00	7.00	1	.17479	.49945	.50055
1	10.00	10.00	1	1.69294	.22375	.77625
0	4.00	1.00	0	.12066	.51052	.48948
1	4.00	9.00	0	.12066	.51052	.48948
0	11.00	10.00	1	1.89183	.1967/	.80323
1	4.00	8.00	0	.12066	.51052	.48948
0	1.00	2.00	0	63841	.65992	.34008
1	4.00	3.00	0	.12066	51052	.48948
1	10.00	9.00	1	1.58467	.23951	.76049
1	6.00	9.00	1	.\$7257	.41883	.58117
1	10.00	6.00	1	1.58467	.23951	.76049
1	1.00	3.00	0	69255	.66979	.33021
1	9.00	5.00	1	1.33164	.27920	.72080
0	3.00	3.00	0	18650	.57281	.42719
0	6.00	6.00	1	.57257	.41883	.58117
0	4.00	3.00	0	.06652	.52158	.47842
0	8.00	6.00	1	1.07862	.32268	.67732
	variable is th group 1 (not r	e predicted pr reconvicted).	robability of	pre	s new variable g dicted probabilit up 2 (reconvicte	y of being in



You may want to rename the new variables to help you remember what they represent.

In this way, we can first compute the best discriminant function and then use this to make real predictions. For example, we can predict that participant 1 will be reconvicted on release, and that our estimate of the probability of reconviction is 71.4%. Based on this relatively high risk of reconviction, we might decide to target some extra resources at this individual prior to their release in the hope of reducing their likelihood of reconviction.

Section 4: AN INTRODUCTION TO LOGISTIC REGRESSION

Logistic regression differs from discriminant analysis in that, whereas discriminant analysis computes a function that best discriminates between two categories, logistic regression computes the log odds that a particular outcome will occur. For example, we can use logistic regression to compute the odds that a particular prisoner will be reconvicted after release.

The odds of an event occurring are given by the ratio of the probability of it occurring to the probability of it not occurring. For example, if four horses are running in a race and we pick one of them at random, the odds of our horse winning will be 0.25/(1 - 0.25) = 0.333. The odds for any event lie between the values of 0 and +infinity. This is problematic for the mathematics involved in logistic regression, and to overcome this problem, the log of the odds are calculated (natural log or log_e). The log odds of an event will vary between -infinity and +infinity, with a high value indicating an increased probability of occurrence. A positive value indicates that the event is more likely to occur than not (odds are in favour), while a negative value indicates that the event is more likely not to occur (odds are against). To illustrate the difference between odds and log odds, consider the example of the four-horse race. We have already seen that the odds of us picking the correct horse are 0.333. The odds of us picking the wrong horse are given by 0.75/(1-0.75) = 3. If we now take the log of each of these values, we will see that the log odds of us picking the correct horse are $\log_{e}(0.333) = -1.1$ and the log odds of us picking the wrong horse are $\log_{e}(3) =$ 1.1. The advantage of log odds over odds is clear from this example: unlike odds, log odds are symmetric about zero. The term 'logistic' in the name logistic regression derives from this use of log odds.

It is possible to use logistic regression in situations when there are two or more categories of the grouping variable. In cases where there are just two categories of the grouping variable, the SPSS binary logistic regression command should be employed. Where there are more than two categories, the multinomial logistic regression command should be used. The multinomial command is not described here, but is similar to the binary logistic regression command.

As for discriminant analysis, we will show how the analysis can be used to predict category membership for cases where it is not known.

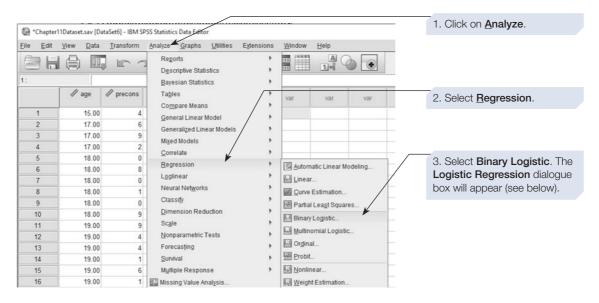
CHAPTER 12

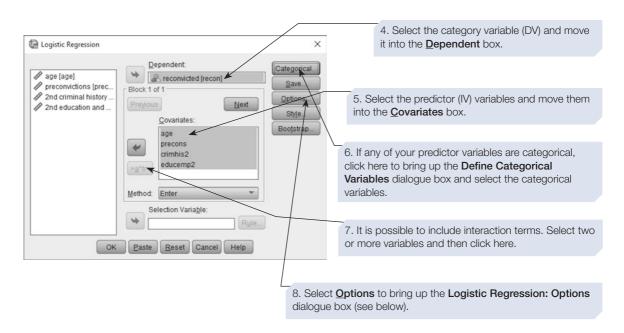
Section 5: PERFORMING LOGISTIC REGRESSION ON SPSS

Example study: reconviction among offenders

We will demonstrate binary logistic regression using the prisoner data set used to demonstrate discriminant analysis (see Section 3). This will allow us to compare the output of these two commands.

To perform a binary logistic regression





Logistic Regression: Options X	9. Select Classification plots
Statistics and Plots	······
Classification plots	
✓ Hosmer-Lemeshow goodness-of-fit Iteration history	10. Select Iteration history.
Cl for exp(B): 95 %	
Quttiers outside 2 std. dev	
© All cases	11. Select CI for exp(B).
Display At each step At last step	_1 ()
Probability for Stepwise Classification cutoff: 0.5 Entry: 0.05 Removal: 0.10 Maximum Iterations: 20 Conserve memory for complex analyses or large datasets	12. Select the <u>Hosmer-</u> Lemeshow goodness-of- fit test; this will help you to evaluate the model.
Cancel Help	
	13. Click <u>Continue</u> to return to the Logistic Regression dialogue box.

Finally, click on the or button. The annotated output is shown below.

SPSS output for logistic regression using Enter method

Obtained using menu items: Regression > Binary Logistic

Logistic Reg	gression ase Processing Sum		This table tells you that 100% of the 221 cases have been processed.	
Unweighted Case	esa	N	Percent	
Selected Cases	Included in Analysis	221	100.0	
	Missing Cases	0	.0	
	Total	221	100.0	
Unselected Case	S	0	.0	
Total		221	100.0	
a. If weight is in number of c	n effect, see classification ases.	table for th	ie total	

Dependent Variable Encoding

Original Value	Internal Value
no	0
yes	1

This table tells you how the two outcomes (reconvicted and not reconvicted) have been coded; this is important when interpreting the output.

Block 0: Beginning Block

lteration History^{a,b,c}

Iteration		-2 Log likelihood	Coefficients Constant		cases an the most
Step 0	1	266.132	842	-	
	2	265.990	897	_	
	3	265.990	897		

a. Constant is included in the model.

- b. Initial -2 Log Likelihood: 265.990
- c. Estimation terminated at iteration number 3 because parameter estimates changed by less than .001.

Classification Table^{a,b}

		Predicted					
		reconvi	icted	Percentage			
Observed		no	yes	Correct			
reconvicted no		157	0	100.0			
	yes	64	0	.0			
Overall Perce	entage			71.0			
	reconvicted	reconvicted no	Observed no reconvicted no 157 yes 64	no yes Observed no yes reconvicted no 157 0 yes 64 0			

a. Constant is included in the model.

b. The cut value is .500

This section of the output headed Block 0: Beginning Block reports the results of the most basic attempt to predict outcome; one in which all cases are predicted to result in the most common outcome.

This shows the iteration history.

This table reports the results of this simple prediction. As most of our prisoners are not reconvicted after release, the predicted outcome for all has been set to 'not reconvicted'. This crude method results in an accurate prediction for 71% of cases. Hopefully, our logistic regression will do better than this.

			ind bie 5 in	the Lqu	acion					
		В	S.E.	Wald	df	5	Sig.	Exp(B)		
Step 0	Constan	t897	.148	36.612		1	.000	.408		
		Variables r	ot in the E	quation				\checkmark		
				Score	df	Sig.				 These two tables tell us the
Step 0	Variables	age		11.316	1	.001				so far no variables have be
		preconvictions		8.167	1	.004				entered into the equation.
		2nd criminal hi	story	16.894	1	.000				
		2nd education employment	and	12.288	1	.000	•		_	We can ignore both these tables.
	Overall Stat	istics		28.909	4	.000				

Variables in the Equation

Block 1: Method = Enter

This block of output reports the results of the logistic regression analysis. This analysis should result in a more accurate prediction than that reported in Block 0.

Iteration History^{a,b,c,d} Coefficients 2nd education 2nd criminal history preconviction and -2 Log Constant age employmen Iteration likelihood s Step 1 237 530 - 888 - 031 024 109 038 1 233.181 -.797 -.052 .033 .155 .033 233.002 -.727 -.058 .036 .165 .029

-.058

-.058

.036

.036

.166

.166

.029

.029

a. Method: Enter

b. Constant is included in the model.

233.001

233.001

-.722

-.722

c. Initial -2 Log Likelihood: 265.990

d. Estimation terminated at iteration number 5 because parameter estimates changed by less than .001.

Logistic regression employs a process known as iteration. In an iterative process, we attempt to arrive at the best answer to a problem through a series of approximations. Each iteration results in a slightly more accurate approximation than the previous iteration. This table reports this iterative process. The statistic -2 log likelihood is used in logistic regression to measure the success of the model. A high value indicates that the model poorly predicts the outcome. With each iteration we can see the value falling; however, the benefit derived at each iteration decreases until after four iterations there is no change in the value and SPSS terminates this process. You can also see how the coefficients of each of the predictor variables are adjusted at each iteration.

Omnibus Tests of Model Coefficients

		Chi-square	df	Sig.
Step 1	Step	32.988	4	.000
	Block	32.988	4	.000
	Model	32.988	4	.000

Omnibus tests are general tests of how well the model performs. In a report, we can use either this or the Hosmer–Lemeshow test (shown below). When the Enter method has been employed (as here), there is only one step and so the Step, Block and Model rows in this table will be identical.

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	Mod	el Summary 🔫	
Step	-2 Log likelihood	Cox & Snell R Square	Nagelkerke R Square
1	233.001 ^a	.139	.198

 a. Estimation terminated at iteration number 5 because parameter estimates changed by less than .001.

Hosmer and Lemeshow Test 🔨

Step	Chi-square	df	Sig.
1	5.876	8	.661

This table gives useful statistics that are equivalent to R^2 in multiple regression. It is not possible to compute an exact R^2 in logistic regression, but these two statistics are useful approximations. Here, we can see that our model accounts for between 13.9% and 19.8% of the variance.

This table gives the results of the Hosmer– Lemeshow test we requested. It gives a measure of the agreement between the observed outcomes and the predicted outcomes. This statistic is a test of the null hypothesis that the model is good, hence a good model is indicated by a high p value, as in this example, where p = .661. If the pvalue is less than .05, then the model does not adequately fit the data.

Contingency Table for Hosmer and Lemeshow Test 🔶

		reconvic	ted = no	reconvict	ed = yes		
		Observed	Expected	Observed	Expected	Total	
Step 1	1	21	20.940	1	1.060	22	
	2	18	19.824	4	2.176	22	
	3	20	18.573	2	3.427	22	
	4	19	17.698	3	4.302	22	
	5	18	16.845	4	5.155	22	
	6	17	15.747	5	6.253	22	
	7	12	14.333	10	7.667	22	
	8	12	12.867	10	9.133	22	
	9	9	10.798	13	11.202	22	
	10	11	9.376	12	13.624	23	

This table is used in the calculation of the Hosmer–Lemeshow statistic reported in the previous table. The cases are ranked by estimated probability on the criterion variable and then divided into 10 'deciles of risk'. Within each decile, the numbers of observed and expected positive (reconviction) and negative (no reconviction) outcomes are calculated. The Hosmer–Lemeshow statistic (above) is then calculated from this 10*2 contingency table. Note that a high proportion of the participants in decile 1 are not reconvicted, whereas the majority of those in decile 10 are reconvicted. This pattern indicates that our model is good.

Classification Table^a

Predicted

		reconv	icted	Percentage	
Observed		no	yes	Correct	
reconvicted	no	138	19	87.9	
	yes	44	20	31.3	
Overall Perce	entage			71.5	
	reconvicted	reconvicted no	Observed no reconvicted no 138 yes 44	reconvicted no 138 19 yes 44 20	

a. The cut value is .500

This table summarises the results of our prediction and should be compared with the equivalent table in Block 0. Our model correctly predicts the outcome for 71.5% of cases. Although, overall, this is not much better than the situation reported in Block 0, in this new classification table we correctly predict reconviction in 31.3% of prisoners who are reconvicted.

Compare this with the outcome of the discriminant analysis reported earlier in this chapter.

Variables in the Equation	Variab	les i	in th	ie Ee	quati	on
---------------------------	--------	-------	-------	-------	-------	----

		В	S.E.	Wald	df	Sig.	Exp(B)
Step 1 ^a	age	058	.018	10.221	1	.001	.944
	preconvictions	.036	.026	1.871	1	.171	1.037
	2nd criminal history	.166	.068	5.860	1	.015	1.180
	2nd education and employment	.029	.070	.166	1	.684	1.029
	Constant	722	.642	1.267	1	.260	.486

a. Variable(s) entered on step 1: age, preconvictions, 2nd criminal history, 2nd education and employment.

This table contains some of the most critical information.

The first column gives the coefficients for each predictor variable in the model. The negative coefficient for 'age' indicates that the odds of reconviction declines with increasing age.

The Wald statistic and associated Sig. values indicate how useful each predictor variable is. In this case, only 'age' and 'crimhis2' are significant; the analysis could be rerun with only these variables included.

The Exp(B) column gives an indication of the change in the predicted odds of reconviction for each unit change in the predictor variable. Values less than 1 indicate that an increase in the value of the predictor variable is associated with a decrease in the odds of the event. Thus, for every extra year of age, the odds of the prisoner being reconvicted on release decrease by a factor of 0.944.

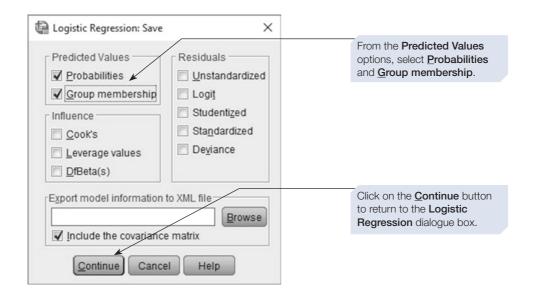


A logistic regression analysis was performed with reconviction as the DV, and age, number of previous convictions, and the LSI-R criminal history, education and employment subscales as predictor variables. A total of 221 cases were analysed, and the full model significantly predicted reconviction status (omnibus chi-square = 32.99, df = 4, p < .001). The model accounted for between 13.9% and 19.8% of the variance in reconviction status, with 87.9% of the non-reconvicted prisoners successfully predicted. However, only 31.3% of predictions for the reconvicted group were accurate. Overall, 71.5% of predictions were accurate. Table 11.1 (not included here, but this would be your table of betas similar to that at the top of this page) gives coefficients and the Wald statistic and associated degrees of freedom and probability values for each of the predictor variables. This shows that only age and criminal history reliably predicted reconviction. The values of the coefficients reveal that an increase of one year of age is associated with a decrease in the odds of conviction by a factor of 0.94, and that each unit increase in criminal history score is associated with an increase in the odds of reconviction by a factor of 1.18.

CHAPTER 12

Using logistic regression to predict group membership

We can now use the results of our logistic regression analysis to calculate the level of risk of reconviction for each new inmate who enters the prison. To do this, we need to add to the data file the data from each new prisoner as they arrive to start their sentence. For each of these new cases, enter a missing value for the category variable ('recon'). Now repeat the analysis as before, but in the Logistic Regression dialogue box, click on the Save button. This will bring up the Logistic Regression: Save dialogue box (see below).



Now click on the **ox** button. In addition to the output described above, two new variables will be added to the data file. The first of these gives the predicted probability of reconviction for each case, and the second gives the predicted group (reconvicted or not). We could use these predicted probabilities to make decisions about treatment or early release. However, in doing so, it would be important to remember the classification table (see output above), which showed that, when we predict reconviction, we will often be wrong.

ile Edit	View Data	Iransform	Analyze <u>G</u>	raphs Utilitie	Extensions	s Window Help			the predicted probability of reconviction. Note that participant 1 is predicted to b
3:	l age	& precons	& recon		P educemp	PRE_1	& PGR_1	v	at higher risk of reconviction than participant 2.
1	15.00	4	0	8.00	8.00	.52662	1		
2	17.00	6	0	2.00	4.00	.26018	0		
3	17.00	9	1	4.00	7.00	.37297	0		
4	17.00	2	1	10.00	10.00	.57609	λ		
5	18.00	0	0	4.00	1.00	.25450	0		
6	18.00	8	1	4.00	9.00	.36440	0	\backslash	
7	18.00	0	0	11.00	10.00	.58469	1	\langle	
8	18.00	1	1	4.00	8.00	.30193	0	\backslash	
9	18.00	0	0	1.00	2.00	.17612	0		This new variable gives the
10	18.00	9	1	4.00	3.00	.33359	0		predicted outcome with
11	19.00	9	1	10.00	9.00	.60237	1		· ·
12	19.00	4	1	6.00	9.00	.39465	0		regards to reconviction. It is
13	19.00	4	1	10.00	6.00	.53708	1		predicted that participant 1
14	19.00	1	1	1.00	3.00	.17712	0		will be reconvicted but that
15	19.00	6	1	9.00	5.00	.50668	1		participant 2 will not.
16	19.00	1	0	3.00	3.00	.23062	0		

Summary

- This chapter introduced two statistical procedures, discriminant analysis and logistic regression, which can be used to make a prediction about category membership. See Chapter 10 for a test that makes a prediction about an outcome variable that is a continuous score.
- Logistic regression and discriminant analysis differ in the assumptions they make about the predictor variables.
- If you need to recode any of your predictor variables, see Chapter 4 for guidance on recoding values or computing new variables.
- For guidance on incorporating SPSS output into a report, or on printing the output, see Chapter 14.

13 Factor analysis, and reliability and dimensionality of scales

In this chapter

- An introduction to factor analysis
- Performing a basic factor analysis using SPSS
- Other aspects of factor analysis
- Reliability analysis for scales and questionnaires
- Dimensionality of scales and questionnaires

SPSS for Psychologists online

Visit macmillanihe.com/harrison-spss-7e for data sets, online tutorials and exercises.

Section 1: AN INTRODUCTION TO FACTOR ANALYSIS

- Factor analysis enables us to look at the relationship between a large number of variables (for example, questions on a questionnaire) to investigate whether there is an underlying structure in the pattern of correlations between them. This effectively allows us to see whether our original variables can be grouped or summarised using a smaller number of dimensions, known as factors.
- An example is the factor structure underlying human abilities and aptitudes (e.g. Kline, 1994). To illustrate this, imagine that you gave people a large battery of diverse tests designed to establish their general mental ability (or intelligence). These types of test batteries can often take as long as 3 hours to complete, and produce a vast amount of data, comprising scores from each individual test or question completed. Rather than taking all of these raw data and somehow using this as a measure of aptitude (which can make it difficult to see the wood for the trees), it can be more meaningful to try to summarise the findings in terms of the patterns observed in the scores; and look for different latent factors that may be underlying the data. For example, while all of the questions and tasks in the test battery may have been different, you might find that some particular scores are highly correlated with one another, while others are not; which suggests that some tests may be tapping into (or measuring) common underlying abilities, while

others may be representing something different. For example, you might find that people who perform well on tasks involving drawing complex shapes also do well at putting together puzzle pieces (as they both involve visual–spatial abilities); and those who do well on reading comprehension tests also score well on measures of vocabulary (as both involve language skills). However, tests tapping into visual– spatial tasks may not be strongly related to measures of language ability, suggesting these are two distinct underlying factors involved in general intelligence.

- Thus, factor analysis allows us to establish whether any latent factors drive the observed scores on our individual tests (or variables); And if so, it can tell us how many factors are likely to be underlying our data, and how our individual variables relate to each factor. When the factors that are identified are psychological in nature, they can often be referred to as *psychological constructs* (like visual spatial abilities, or language skills), but the more general term is *factors*.
- Factor analysis is sometimes called a 'data reduction technique', because it aims to reduce large amounts of data down to more meaningful groupings. Furthermore, you can use the outcome of a factor analysis to choose a smaller set of variables than those initially measured, for use in future analysis or studies.
- There are two key types of factor analysis:
 - (1) Exploratory factor analysis. This is the type of factor analysis that we will cover in this chapter. As the name implies, exploratory factor analysis is often used as an exploratory tool to investigate whether participants' scores on large batteries of tests or large amounts of questions (such as those on a question-naire) can be explained by variations in a smaller number of underlying factors. In other words, it allows you to identify the underlying factor structure of a measure. It is often used in the development of new measures and questionnaires. For example, it can be used to help determine whether a questionnaire is really measuring the underlying factors (or psychological constructs) that it is intended to.

Exploratory factor analysis generally does not test hypotheses (except for the number of factors, see below). Instead, it explores the possibility of a factor structure underlying the variables. The analysis provides a large amount of information, from which the researcher can make inferences about factors. Hypotheses about the number of factors can be tested in IBM SPSS statistics software (SPSS); this will be briefly mentioned in Section 3.

(2) Confirmatory factor analysis is usually used in a strictly confirmatory fashion, where you have a well-defined hypothesis regarding the structure and relationships in your data. It is a statistical technique used to verify the factor structure of a set of observed variables. In this case, hypothesised underlying factors are set out in advance, and a factor analysis is used to test that hypothesis. While it is a useful and important tool, its use is more complex and specialised than exploratory factor analysis, and is better analysed using a structural equation modelling package such as AMOS (now available from IBM and known as IBM SPSS Amos) or LISREL, rather than SPSS. As such, it will not be covered in this book.

How this chapter is organised

This chapter is organised somewhat differently from other chapters, in that later in this section we will be using tables produced by SPSS to explain some aspects of factor analysis. We will not at that point explain how to produce those tables; instead, we just use them to talk about factor analysis. We will also explain how these tables can give an indication of whether a factor analysis is likely to be useful with the variables you have measured. For this section, the tables are *not* derived from data suitable for a real factor analysis. Instead, the data are simply convenient for the purpose of this demonstration. If you are already familiar with factor analysis, you could omit this section. If you are going to use factor analysis in real research, then you will need to understand more about factor analysis than can be explained in any book on how to use SPSS. We recommend the following: Cooper (2010, Chapters 16 and 17); Giles (2002); Kline (1994, who introduces the calculations for factor analysis in detail); and Tabachnick and Fidell (2014).

Section 2 returns to our normal style, showing how to perform a basic factor analysis and then how to interpret the output. Section 3 discusses some other aspects of factor analysis. Finally, in Sections 4 and 5, we provide some guidelines on using SPSS to check the reliability and dimensionality of scales/questionnaires. First, we continue this section with an introduction to factor analysis.

How factor analysis relates to other statistical tests

Correlation and covariance

Bivariate correlation identifies the level of linear relationship between two variables (see Chapter 6). The covariance of two variables, the extent to which they vary together, is the unstandardised equivalent. Factor analysis identifies variables that all relate to a single factor by exploring the relationships between the variables. It does so by carrying out calculations based on the matrix of correlation coefficients, or the matrix of covariances, between each variable with each of the other variables. As a default, SPSS uses the correlation matrix in factor analysis.

Multiple regression

In multiple regression, we are interested in which of a number of predictor variables may be useful in predicting scores on another variable, the criterion or outcome variable. Factor analysis could be used in predicting what score someone might get on a variable; for example, if you find out that a new test can be accounted for by the factor verbal ability, then someone who scores highly on other tests that tap verbal ability should also score highly on the new test.

In multiple regression, correlations between the predictor variables should not be too high; if they are, then those predictor variables may actually be measuring the same thing and so do not add to our understanding of what predicts the criterion variable. By contrast, in factor analysis, we are interested in relatively high correlations, because we want to investigate which variables are measuring aspects of the same thing (the same dimension, factor or psychological construct). See Tabachnick and Fidell (2014), however, for limitations on multicollinearity in factor analysis.

Analysis of variance

It is important that you recognise the difference between the meaning of the term *fac-tor* in ANOVA and its meaning in factor analysis. In a true experimental ANOVA design, the experimenter manipulates the levels of each factor in order to explore cause-and-effect relationships between the factors and the dependent variable. In an ANOVA design incorporating any factor with levels from a natural group variable, the levels are not manipulated but instead they are chosen from existing groups, such as male and female, or high extroversion and low extroversion, so causal relationships cannot be assumed. In either case, in ANOVA terminology, the word 'factor' is simply a more convenient name for an independent variable, whereas in factor analysis, it means a dimension (or a psychological construct) that underlies several measured variables. So, in factor analysis, the term has a more profound but less concrete meaning than it does in ANOVA. Of course, in any specific study in which an ANOVA design is used, one or more of the factors may represent a psychological construct.

Discriminant analysis and logistic regression

Discriminant analysis and logistic regression are methods for determining which variables will predict membership of, or discriminate between, different categories of another variable. The example given in Chapter 12 was an investigation of which variables would discriminate between those offenders who had been reconvicted and those who had not. In factor analysis, by contrast, we are usually not interested in the scores of any individual participant. Instead, we want to explain the pattern of correlations between variables, and identify factors (dimensions) underlying those variables. (Although we may be interested in an individuals' score on those factors, once they have been established).

Correlation matrix and other matrices in factor analysis

Here, we consider certain aspects of factor analysis, including how it makes use of the matrix of correlations. For our explanations, we will use the data previously used for Spearman's r_s correlation coefficient (Chapter 6, Section 4). Those data were for only three variables, and thus are *not* appropriate for factor analysis. Three variables, however, give rise to small matrices that will be much more useful for illustration than the large matrices produced by the number of variables typically entered into a real factor analysis.

One point to note is that factor analysis in SPSS makes use of Pearson's, not Spearman's, correlation coefficients. This is acceptable for our data because factor analysis can be used with continuous and discrete data. For a continuous scale the data points can have fractional values, whereas for a discrete scale the data points have to be integers (whole numbers). Both continuous and discrete data, however, can be of ratio, interval or ordinal level of measurement. For the study on the relationships between attractiveness, believability and confidence, the three variables, 'confdt' (confidence in the woman's testimony), 'believ' (how believable the woman was) and 'attrct' (how attractive the woman was), were all entered into a Pearson's r analysis. The table below is from the SPSS output. Details of the study are in Chapter 6, Section 4, and how to obtain and interpret the output for Pearson's correlation coefficient is described in Chapter 6, Section 3.

We show the Pearson's r output itself for comparison purposes. Next, we will show various matrices that can be produced by means of the factor analysis command. Details of how to obtain these tables are given in Section 2 of this chapter; for now we simply use them to explain aspects of factor analysis. They can also help to indicate whether our variables have *factorability*; that is, whether it is likely that there are any factors underlying the variables we measured (Tabachnick and Fidell, 2014, 667).

Pearson's r output

		confdt	believ	attrct
confdt	Pearson Correlation	1	.278	.073
	Sig. (2-tailed)		.008	.498
	N	89	89	89
believ	Pearson Correlation	.278**	1	.429
	Sig. (2-tailed)	.008		.000
	N	89	89	89
attrct	Pearson Correlation	.073	.429	1
	Sig. (2-tailed)	.498	.000	
	N	89	89	89

Convolations

**. Correlation is significant at the 0.01 level (2-tailed).

In Pearson's r output, each cell of the matrix contains the values of r, p and N, allowing a quick assessment of the strength and significance of each individual bivariate correlation. In the equivalent table produced by factor analysis, shown next, the matrix of coefficients is in the upper part of the table, and the matrix of p values, if you request it, is in the lower part of the table. This separation reflects the fact that the coefficients are used in factor analysis calculations, whereas the p values are just for information. These coefficient values are also known as the *observed correlations*. In a matrix, there is usually a distinction between the on-diagonal values and the off-diagonal values: see annotations below.

Correlation matrix from factor analysis output

1(a). The upper part of this table is a complete matrix of the correlation coefficients for each variable with each of the other variables and with itself. 1(b). The diagonal of a matrix is usually worth noting in factor analysis. The diagonal in this matrix holds the correlation coefficient for each variable with itself (and therefore these on-diagonal values are all equal to 1.0).

	ooner	ation Mat		
•		confdt	believ	attrct
Correlation	confdt	1.000	.278	.073
	believ	.278	1.000	.429
	attrct	.073	.429	1.000
Sig. (1-tailed)	confdt		.004	.249
×	believ	.004		.000
	attrct	.249	.000	

1 (c). These off-diagonal values are the three possible correlation coefficients between the different pairs of variables. These values are called the *observed* correlations. They are mirrored below left of the diagonal (see 1(a)).

a. Determinant = .751

2. The lower half of the table is a matrix of the *p* values corresponding to the correlation coefficients. The diagonal is left blank in this matrix.



For an indication of factorability, look at the sizes of the correlation coefficients between the different variables. If the coefficients are mostly small (less than .3), there is little likelihood that a factor structure underlies the variables.

Partial correlations from factor analysis output

The anti-image table (shown on the next page) contains two matrices. The upper matrix is automatically printed if you request the anti-image matrices, but for now we are most interested in the lower matrix.

In the lower matrix, the off-diagonal values are the *partial correlations* with the signs reversed, known as the *negative partial correlations*. (This term does not mean that the value has to be negative; simply that the original sign is reversed.) A partial correlation is the correlation between two variables when the effects of any other variables are controlled for or 'partialled out' (see Chapter 6, Section 5). For example, if there is a correlation between the number of ice creams sold and the number of

drownings that occur, it is unlikely that one causes the other. Instead, it is more likely that this relationship is explained by a third variable that is related to both, e.g. temperature. To investigate this possibility, one could partial out the effect of temperature from both variables and then find the partial correlation between them; it is likely to be very small in that example.

Anti-image Matrices

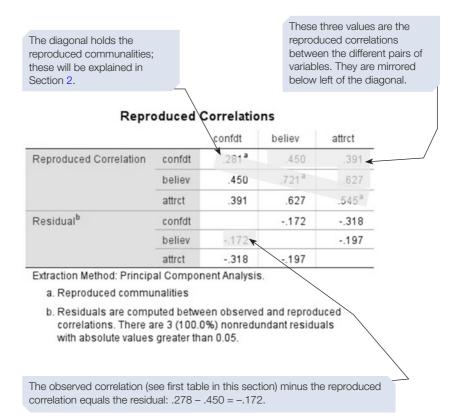
	in mag				
		confdt	believ	attrct	
Anti-image Covariance	confdt	.920	228	.046	
	believ	228	.755	335	
	attrct	.046	335	.813	
Anti-image Correlation	confdt	.515 ^a	273	.053	
	believ	273	.504ª	427 🔫	
	attrct	.053	427	.506 ^a	
a. Measures of Samp		Jacy(MSA)			
The diagonal holds the val of the Kaiser–Meyer–Olkin KMO) measure of samplir adequacy for each variable The KMO will be discusse Section 2.	ng e.			negative parti (see text): the indicate whet	her there is a re underlying the

As stated above, if the correlation coefficients are mostly small, then the variables are unlikely to have factorability. However, if the correlation coefficients are mostly large, then we cannot conclude that the variables definitely do have factorability. The partial correlations provide the next check, as explained here. If there is a factor structure underlying a number of variables, then they should all be fairly well correlated with each other. So, we would expect the correlation between any two variables to become weaker once the effects of the other variables have been partialled out. If the partial correlation between two variables is not weaker than the correlation between them, then those two variables have a strong relationship to each other and little relationship to any of the other variables we measured.

Use of the partial correlations as an indication of factorability depends on their absolute size; thus, the fact that SPSS reverses their sign and prints the negative partial correlations is irrelevant. If the partial correlations are mostly large, then there is little likelihood of a factor structure underlying the variables.

Reproduced correlations and residuals from factor analysis output

The *reproduced correlations* (the off-diagonal values in the upper matrix of the table below) are the values predicted for the correlations between the variables under the assumption that the factor analysis is correct. So, if the solution to the factor analysis represents the actual state of things, then these are the correlations that would be expected. Factor analysis then compares these predicted correlations with those that were actually obtained (the observed correlations); each *residual* value, in the lower matrix, is equal to the reproduced correlation subtracted from the observed correlation. If residuals are small, then the observed values are close to those predicted by the factors, which suggests a good factor fit (see tip box).





Factorability: the smaller the residuals are (unlike here), the more likely it is that the factor analysis gives a good explanation of the data. That is, the more likely it is that the identified factors do explain the actual state of things in the real world.

So far we have introduced you to some concepts that will be useful in understanding factor analysis. For that purpose, we used data that were simply convenient. In Section 2 we will show you how to carry out and interpret a real factor analysis. First, some issues about when factor analysis can be used and more terminology.

When should I use factor analysis?

The following criteria are usually suggested:

- 1. The variables should be of at least ordinal level of measurement.
- 2. The variables should be normally distributed: if they do not meet the criteria for normal distribution, you should consider a transformation (see Tabachnick and Fidell, 2014).
- 3. The relationships between variables should be reasonably linear.
- 4. It is usually considered that for a successful factor analysis, at least 100 participants should have provided data, and some say that 200 or more are required. Two pointers are:
 - a. There should be more participants than variables. Kline (1994) suggests a minimum ratio of 2:1, but the more the better. Thus, if you wish to explore the factor structure underlying a questionnaire that contains 60 items, you should test a minimum of 120 participants.
 - b. There should be more participants than extracted factors. Kline suggests a minimum ratio of 20:1. In truly exploratory factor analysis, however, we do not know how many factors there will be.

In general, the more participants you test, the more likely it is that any factors that do underlie the measured variables will be revealed; thus, a sample size of 200 is a sensible minimum target.

Usefulness/validity of a factor analysis

Any data set containing a number of variables can be factor analysed, but the outcome may be invalid or simply useless, or meaningless. While SPSS can tell you which variables are related, as a researcher you have to *qualitatively* interpret the quantitative output. SPSS cannot tell you whether or not the relationships you find make sense; that is up to you to decide.

A number of methods for assessing whether the variables entered have factorability have already been described. Other pointers to factorability, and to whether the extracted factors are a good solution to explaining the variables, will be described below. These pointers should be inspected before considering the factors that may have been extracted. One should then consider whether the solution provided by the factor analysis makes sense.

Terminology

In addition to the terms introduced above, the following terms are required to follow the output from even a fairly basic factor analysis. Don't worry if you find them difficult to follow at this stage – they should start to make more sense when we walk though the analysis, and put the terms in context.

Component and factor

Kline (1994, 36) states: 'components are real factors because they are derived directly from the correlation matrix. Common factors of factor analysis are hypothetical because they are estimated from the data.' Common factors are therefore an example of what Skinner (1965) referred to as 'explanatory fictions'. This does not mean, however, that they are not useful, but it is important that you use caution when interpreting any factors you derive from a factor analysis. The simplest type of factor analysis, principal component analysis, extracts components as its name suggests, whereas other types extract common factors. Thus, principal component analysis and factor analysis are somewhat different things; we give more detail in Section 3.

In this book, in common with many others, the terms 'factor' and 'component' are used interchangeably when we discuss principal component analysis, but the distinction between them should be borne in mind throughout.

Extraction

Extraction is the name for the process by which the important factors are identified; those factors are then said to have been *extracted*. When performing a factor analysis, there are several statistical methods that can be used to extract the underlying factors. We illustrate the principal component method of extraction, and mention other methods in Section 3. Principal components analysis (PCA) aims to reduce a set of variables by looking for clusters of variables that appear to be related to one another (and therefore may be tapping into the same underlying factor). As such, PCA is primarily concerned with identifying variables that share variance with one another. Extraction is not an exact process, as should become clear as you work through Section 2.

Communality

Communality is a measure of how much variance in the data from a particular variable is explained by the analysis. The values can range from 0 to 1. Initially, all the factors (or components) in the analysis (equal to the number of variables entered) are used to calculate the communalities, and all the variance is accounted for. Thus, in principal component analysis, the *initial* communalities are all equal to 1 (indicating that all the variance is explained). After the factor or component extraction, *extraction* communalities for each variable are calculated based on the extracted factors only. The higher the value of the extraction communality is for a particular variable, the more of its variance has been explained by the extracted factors. The communality is calculated from factor loadings (see below).

Eigenvalue

Eigenvalue is a measure of how much variance in the data is explained by a single factor. Remember that the analysis initially considers all the possible factors, and allocates the same number of possible factors as there are variables. The higher the value of the eigenvalue, the more of the variance in the data is explained by that factor. The magnitude of the eigenvalue can be used to determine whether a factor explains

CHAPTER 13

sufficient variance for it to be considered a useful factor. The default value for this in SPSS is 1.0, but it is possible to request SPSS to extract factors using a reference eigenvalue other than 1.0, as explained in Section 3.

Scree plot

The scree plot is a useful graph of the eigenvalues of all the factors initially considered. It can be used to decide on the number of factors that should be extracted. An annotated example is shown in Section 2.

Factor loadings

A factor loading is calculated for each combination of variable and extracted factor, and it tells you how much each variable loads on to (or is related to) each factor. These values are useful for seeing the pattern of which variables are likely to be explained by which factor. The factor loading can be thought of as the coefficient of the correlation between the component (or factor) and the variable; thus, the larger the number, the more likely it is that the component underlies that variable. Loadings may be positive or negative; this issue is considered further in Section 3.

The initial factor loadings can be inspected for patterns. Almost always, however, a rotation will be used (see below), and the rotation factor loadings are a more useful indicator of patterns. SPSS uses the before-rotation factor loadings to calculate the extraction communalities; we calculate an example in our annotations of the SPSS output in Section 2.

Rotation

A factor analysis prior to rotation provides an explanation of how many factors underlie the variables; for some purposes this is sufficient. In psychology, however, we normally wish to understand what it all means; so we want to establish whether any psychological constructs might underlie the variables. Rotation is a mathematical technique available in factor analysis that arrives at the simplest pattern of factor loadings, and helps us to interpret what our factors represent.

Sometimes, the initial unrotated factor solution results in strong correlations of a variable with multiple factors – making it difficult to know which factor the variable best belongs to. Rotation helps to determine the best fit for the variables onto the different factors. To explain what rotation essentially does, consider Figure 13.1. On the left-hand side of the figure, unrotated factor loadings from a fictitious example are plotted onto a graph showing loadings onto factor 1 (on the *x* axis), against loadings onto factor 2 (on the *y* axis). In this figure, some of the individual variables load equally well onto both factors. Rotating the points on the graph shifts the loadings around the axes in a way that lines them up more closely with one of the two factors, as can be seen on the right-hand side of the graph. Rotating the loadings in this way tends to maximise the variables' loadings onto the factors. Importantly though, rotation does not change the fundamental relationships between the factor loadings; it just maximises their alignment to the factors.

There are different methods of rotation, with a broad distinction between orthogonal rotation methods and oblique rotation methods. Which rotation option you

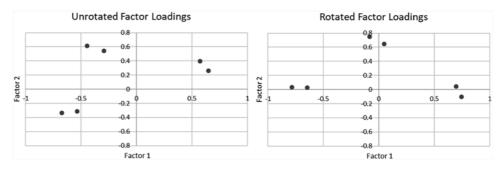


Figure 13.1 Scatterplots illustrating factor rotation

choose will depend on whether you believe the factors to be related or not. Orthogonal methods produce solutions that assume the factors produced are not correlated, whereas oblique methods produce factors that may have some intercorrelations between the factors. Kline (1994) points out that psychological constructs may well be correlated with one another, or non-orthogonal (see Chapter 10, Section 1, where we discuss this in relation to multiple regression). Kline recommends using the **Direct Oblimin** method (one of the list of oblique rotation methods available in SPSS).

Section 2: PERFORMING A BASIC FACTOR ANALYSIS USING SPSS

Hypothetical study

To illustrate the use of this analysis, we shall suggest a hypothetical survey. Suppose that a psychologist who is interested in aesthetic values wanted to investigate people's appreciation of different types of plants. For example, they wished to study whether there were any underlying dimensions to a liking for plants typically found in different places, such as 'wild' English countryside, cottage gardens and formal gardens. They might construct a questionnaire listing various plants to which people are asked to give a score from 1 (extreme dislike) to 7 (extreme liking). The plants could be wild flowers (e.g. celandine, primrose, bluebell, buttercup, daisy, speedwell), cottage gardentype flowers (such as cornflower, poppy, sweetpea, lavender, aster) and formal garden type plants (e.g. rose, wisteria, delphinium, hellebore). A data set of 50 cases is used to illustrate factor analysis, but, as explained above, many more participants are required for the results of a factor analysis to be valid. The data are available from macmillanihe.com/harrison-spss-7e. Note that these hypothetical data are skewed, which might be problematic for a real factor analysis (see Tabachnick and Fidell, 2014). The data file includes the sex of the participants, with 23 men and 27 women, so that you can try the option of selecting members of a group, if you wish. Remember that the survey described is hypothetical; we do not know what you would find if you tried this study yourself.

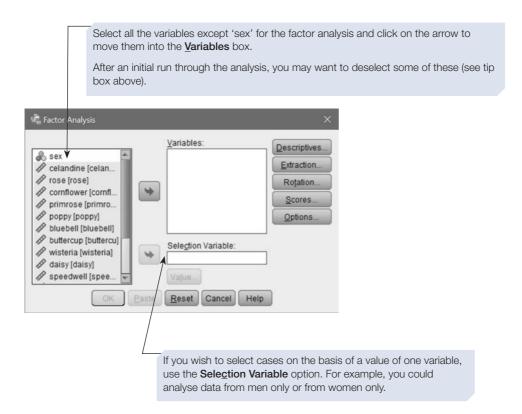
How to perform the analysis

Here, we show you how to carry out and interpret a principal component analysis; other extraction methods are described in Section 3.

The first time we run a factor analysis, we can include all the variables (of ordinal, interval or ratio level of measurement) we have measured. We should then inspect the indicators of factorability in the factor analysis output and decide:

- 1. Whether there is any factor structure at all if not, then give up.
- 2. Whether all the variables are useful if any are not, then run a new factor analysis in which you only include the useful variables.

Click on <u>Analyze</u> \Rightarrow <u>Dimension Reduction</u> \Rightarrow <u>Factor</u>. In the Factor Analysis dialogue box (shown below), select all the variables you want to enter into the factor analysis, and move them across into the <u>V</u>ariables box.



Next click on the **Descriptives** button, and the **Factor Analysis: Descriptives** dialogue box will appear (as shown below).

🛱 Factor Analysis: Descriptives $ imes$	
Statistics	If you select Univariate descriptives , SPSS will print the mean, standard deviation and number of cases for each of the variables you enter into the factor analysis.
Correlation Matrix	
 ✓ <u>C</u>oefficients Inverse ✓ <u>S</u>ignificance levels ✓ <u>R</u>eproduced ■ Dateminant 	The Initial solution option is normally preselected.
☐ Determinant ▲ Anti-image ✓ KMO and Bartlett's test of sphericity	
Cancel Help	Select the Correlation Matrix options as shown here. Don't select Inverse or Determinant for now; we will explain those options in Section 3.

Output from some of the **Correlation Matrix** options were described in Section 1. All those we have selected here will be described below. We will cover the **Inverse** and **Determinant** options in Section 3. When you have selected all the options you require, click on the **Continue** button to return to the **Factor Analysis** dialogue box. Now click on the **Extraction** button; the **Factor Analysis: Extraction** dialogue box (shown below) will appear.

🛱 Factor Analysis: Extraction X	Select the Method of extraction you want to use: for this example, we are using the principal components method.
Method: Principal components Analyze Correlation matrix Covariance matrix Scree plo	Select Scree plot to obtain a useful graph (annotated below; see page ??).
Extract Based on Eigenvalue Eigenvalues greater than: 1 Fixed number of factors Factors to extract:	Ensure that Correlation matrix is selected. (SPSS can make use of the covariance matrix instead: extra output is produced. We do not cover it in this book.)
Maximum Iterations for Convergence: 25	The Extract options are explained in Section 3. For this example, we used the default.

Some of the other options in the Factor Analysis: Extraction dialogue box will be described in Section 3. When you have selected all the options you require, click on the Continue button to return to the Factor Analysis dialogue box. Now click on the Rotation button; the Factor Analysis: Rotation dialogue box will appear (see below).

Factor Analysis: Rotation X	
Method None Quartimax Yarimax Equamax Direct Oblimin Promax Delta: Kappa Display Display	Select the Direct Oblimin rotation method. The Delta option will become available; it affects how correlated the factors will be. The default value, zero, gives the most oblique solution. Once you have selected a method, all the Display options will be highlighted. Normally, you will leave those settings at their default values, as shown.
Rotated solution Loading plot(s)	Note that other rotation methods can be used: we mention them in Section 3.
Maximum Iterations for Convergence: 25	

Other options in the Factor Analysis: Rotation dialogue box will be described in Section 3. Click on the <u>Continue</u> button to return to the Factor Analysis dialogue box. Clicking on the <u>Options</u> button brings up a dialogue box that allows you to specify how to treat missing values. It also allows you to control the appearance of part of the output; we will describe that in Section 3. Finally, click on <u>ok</u>. The annotated output is shown below.

Output from factor analysis using principal component extraction and direct oblimin rotation

Obtained using menu items: Analyze > Dimension Reduction > Factor

	Mean	Std. Deviation	Analysis N	
celandine	5.2200	1.56870	50	This table is produced by
rose	5.2200	1.51577	50	the Univariate Descriptives
cornflower	5.4600	1.66856	50	option in the Factor Analysis Descriptives dialogue box.
primrose	5.3800	1.39810	50	
рорру	6.1200	1.25584	50	
bluebell	5.7200	1.29426	50	
buttercup	5.3400	1.63645	50 🖌	Remember that, in a real factor
wisteria	5.3200	1.42055	50	analysis, you should aim to te a minimum of 200 participants
daisy	5.6200	1.29189	50	
speedwell	5.9600	1.21151	50	
sweetpea	5.2800	1.57843	50	
delphinium	4.3800	1.71298	50	
aster	5.4000	1.55183	50	
lavender	5.9400	1.50387	50	
helebore	3.9000	1.85439	50	

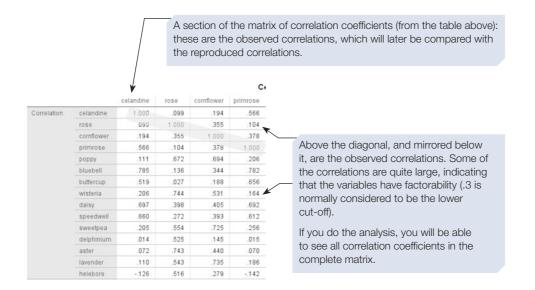
Descriptive Statistics



The descriptives table above is the only part of the output with a note of the number of participants you entered into the factor analysis, so we recommend that you always request this table, for reference purposes.

		celandine	rose	comflower	primrose	DODDY	bluebell	buttercup	wisteria	daisy	speedwell	sweetpea	delphinium	aster	lavender	helebore
orrelation	celandine	1.000	.099	.194	.566	.111	.785	.519	.206	.697	.660	.205	.014	.072	.110	126
oneiauon	rose	.099	1.000	.194	.104	672	.136	027	.200	398	.000	554	.014	.072	543	120
	comflower	.194	.355	1.000	.378	.694	.136	.188	.531	.405	.272	.725	.020	440	.735	.510
	primrose	.134	.104	.378	1.000	206	.782	.100	.164	.405	.612	.725	.015	.070	.186	142
	poppy	.111	.672	.694	.206	1.000	.235	.019	.699	.406	.379	.724	.453	.572	.868	.461
	bluebell	.785	.136	344	.200	.235	1.000	653	.000	667	.373	319	- 015	.189	.243	122
	buttercup	.519	.027	.188	.656	.019	.653	1.000	.154	.468	.511	.089	.048	.106	.025	- 204
	wisteria	.206	744	.100	.164	.699	.000	.154	1.000	.524	.458	.687	.603	.811	.678	.562
	daisy	.697	.398	.405	.692	.406	.667	468	.524	1.000	.851	.584	.334	.352	.366	.214
	speedwell	.660	272	.403	.612	379	.722	.511	.458	.851	1.000	.582	253	.323	301	.180
	sweetpea	.205	.554	.725	.256	.724	.319	.089	.687	584	582	1.000	.458	.620	.721	.512
	delphinium	.014	.525	.145	.015	.453	015	.048	.603	334	253	.458	1.000	.556	.397	.719
	aster	.072	.743	.440	.070	.572	.189	.106	.811	352	323	.620	.556	1.000	649	532
	lavender	.110	.543	.735	.186	.868	.243	.025	.678	.366	.301	.721	.397	.649	1.000	.444
	helebore	126	.516	279	- 142	461	122	- 204	.562	214	180	.512	.719	.532	.444	1.000
Sig. (1-tailed)	celandine		.246	.088	.000	222	.000	.000	.076	.000	.000	.076	.462	.309	.224	.193
	1059	.246		.006	.236	.000	.173	.427	.000	.002	.028	.000	.000	.000	.000	.000
	comflower	.088	.006		.003	.000	.007	.095	.000	.002	.002	.000	.158	.001	.000	.025
	primrose	.000	.236	.003		.076	.000	.000	.128	.000	.000	.036	.458	.315	.098	.162
	рорру	.222	.000	.000	.076		.051	.447	.000	.002	.003	.000	.000	.000	.000	.000
	bluebell	.000	.173	.007	.000	.051		.000	.056	.000	.000	.012	.458	.094	.045	.198
	buttercup	.000	.427	.095	.000	.447	.000		.143	.000	.000	.270	.371	.232	.431	.078
	wisteria	.076	.000	.000	.128	.000	.056	.143		.000	.000	.000	.000	.000	.000	.000
	daisy	.000	.002	.002	.000	.002	.000	.000	.000		.000	.000	.009	.006	.004	.068
	speedwell	.000	.028	.002	.000	.003	.000	.000	.000	.000		.000	.038	.011	.017	.106
	sweetpea	.076	.000	.000	.036	.000	.012	.270	.000	.000	.000		.000	.000	.000	.000
	delphinium	.462	.000	.158	.458	.000	.458	.371	.000	.009	.038	.000		.000	.002	.000
	aster	.309	.000	.001	.315	.000	.094	.232	.000	.006	.011	.000	.000		.000	.000
	lavender	.224	.000	.000	.098	.000	.045	.431	.000	.004	.017	.000	.002	.000		.001
	helebore	.192	.000	.025	.162	.000	.198	.078	.000	.068	.106	.000	.000	.000	.001	

The contents of this table are produced by the <u>Coefficients</u> and <u>Significance levels</u> options in the **Factor Analysis: Descriptives** dialogue box. Here, the whole table is shown shrunk to fit. Next, there is an annotated section of the upper matrix.



If you selected the **Inverse** option in the **Factor Analysis: Descriptives** dialogue box, then that table would appear here. It is not required for principal components analysis, so we will show it in Section 3.

KMO	and	Bartlett's	Test	*
-----	-----	------------	------	---

Kaiser-Meyer-Olkin Me	.754	
Bartlett's Test of	Approx. Chi-Square	659.906
Sphericity	df	105
	Sig.	.000

The KMO measure of sampling adequacy is a test of the amount of variance within the data that could be explained by factors. As a measure of factorability, a KMO value of .5 is poor; .6 is acceptable; a value closer to 1 is better.

The KMO value here is the mean of individual values shown in the next table.

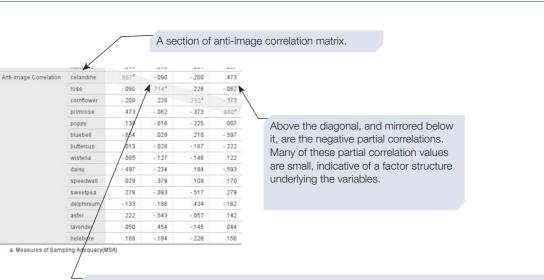
This table is produced by the <u>KMO and</u> Bartlett's Test of Sphericity option in the Factor Analysis: Descriptives dialogue box. These tests give some information about the factorability of your data, in addition to the information from the various matrices as seen above and described in Section 1.

Bartlett's test indicates that the data are probably factorable if p < .05; however, it is considered to be a sensitive test, so it is better to use it in reverse: if p > .05, do not continue; but if p < .05, check other indicators of factorability before proceeding.

This table is produced by the <u>Anti-image</u> option in the Factor Analysis: Descriptives dialogue box. The upper matrix contains negative partial covariances, and the lower matrix contains negative partial correlations. This is the whole table shrunk to fit. Below is an annotated section of the lower matrix.

contribute -200 226 752* -373 -225 210 -187 -146 118 109 -517 4.34 -057 -145 primose 4.47 -002 -3.37 6.60° 007 -527 -222 122 -503 1.10 2.79 -162 1.44 popy -138 -616 -007 -727 -022 1.50 -508 -0.77 2.79 -162 -142 -0.44 popy -138 -616 -027 -027 1.02 -508 -0.77 -2.99 -3.26 0.07 -7.27 -0.22 1.05 -508 -0.37 -2.99 -514 -7.20 bluebelt -684 0.28 2.10 -597 -0.22 7.07 1.22 3.01 -3.33 -0.57 -2.29 -2.24 -1.05 bluebelt -684 0.28 2.10 -597 -0.22 7.17 7.12 3.01 -3.33 -157 -2.29					*	Anti-image	Matrices	s									
Inss -019 214 0.69 -0.92 -0.96 -0.05 -0.06 -0.07 -0.06 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.06 -0.07 -0.07 -0.07 -0.07 -0.07 -0.07 -0.07 -0.07 -0.07 -0.07 -0.07 -0.07 -0.07 -0.07 -0.07 -0.0			celandine	rose	comflower	primrose	poppy	bluebell	buttercup	wisteria	daisy	speedwell	sweetpea	delphinium	aster	lavender	helebore
conflower -0.43 0.90 2.25 -0.74 -0.36 0.90 -0.56 -0.20 <t< td=""><td>Anti-image Covariance</td><td>celandine</td><td>.210</td><td>019</td><td>043</td><td>.091</td><td>.021</td><td>110</td><td>.004</td><td>017</td><td>076</td><td>.005</td><td>.055</td><td>033</td><td>.043</td><td>008</td><td>.044</td></t<>	Anti-image Covariance	celandine	.210	019	043	.091	.021	110	.004	017	076	.005	.055	033	.043	008	.044
pimose 0.91 -0.20 -0.90 0.91 -0.90 0.90 -0.90 0.90 -0.90 0.90		1059	019	.214	.050	012	096	.005	008	026	036	.068	019	.047	106	.076	049
popy 0.21 040 0.05 1.11 -0.00 1.03 -0.02 0.23 0.21 0.44 0.04 -0.04 0.00 buideed 1.10 0.05 0.06 1.00 1.00 1.04 0.04 0.04 0.06 0.00 0.04 0.06 0.00 0.06 0.00		comfower	043	.050	.225	074	036	.036	056	030	.029	.020	- 106	.112	011	025	061
blackeli -110 0.00 0.00 -0.00 133 -0.41 0.21 0.44 -0.40 -0.20 -0.04 0.01 0.01 0.02 0.04 0.01 0.00 0.00 0.00 0.00 unstria 0.017 0.026 0.00 0.020 0.02 0.02 0.01 0.02 0.00		primrose	.091	012	074	.174	.001	091	058	.022	083	.028	.050	- 037	.025	.007	.037
buttercup 0.04 -0.09 -0.09 -0.20		poppy	.021	096	036	.001	.112	008	.033	023	.023	047	.004	049	.073	087	.035
wisteria -0.07 -0.26 -0.09 0.022 -0.20 0.21 -0.26 1.09 -0.25 -0.09 0.012 -0.07 0.001 dairy -0.06 -0.05 -0.08 0.02 -0.01 0.02 0.01 0.02 0.01 0.01 0.02 0.01 dairy -0.05 -0.08 0.03 0.01 0.02 0.01 0.02 0.01 1.02 0.01 0.03 0.01 0.02 0.01 0.02 0.03 0.01 1.02 0.00 0.01 1.02 0.03 0.01 1.02 0.03 0.01 0.02 0.03 0.01 0.03 0.01 0.02 0.03 0.01 0.03 0.01		bluebell	110	.005	.036	091	- 008	.133	041	.021	.044	045	025	.046	035	014	016
dainy -0.76 -0.36 0.29 -0.49 0.20 0.44 0.24 -0.25 112 -0.71 -0.39 -0.11 0.21 -0.24 seeedwall 0.05 0.06 0.02 0.07 0.44 0.05 0.07 0.47 0.05 0.07 0.407 0.05 0.07 0.407 0.00 0.07 0.407 0.00 0.407 0.00 0.407 0.00 0.407 0.00 0.407 0.00 0.407 0.00 0.404 0.00 0.017 0.00 0.404 0.00 0.017 0.017 0.00 0.404 0.00 0.017 0.017 0.00 0.404 0.00 0.017 0.017 0.00 0.00 0.04 0.01 0.01 0.00 0.04 0.01 0.01 0.00 0.04 0.01 0.01 0.00 0.04 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01		buttercup	.004	008	056	058	.033	041	.398	036	.024	049	.050	098	007	.004	.098
speedwell 0.05 0.08 0.00 0.02 -0.47 0.04 -0.09 -0.71 1.15 0.47 0.00 -0.29 0.55 swedpus 0.05 0.01 0.02 0.02 0.04 -0.02 0.00 0.01 0.02 0.05 0.04 0.02 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.00 0.01 0.00 0.01 0.00 0.00 0.01 0.00 0.00 0.01 0.00 0.00 0.01 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00		wisteria	017	026	030	.022	023	.021	036	.190	025	+.009	.012	027	080	.001	011
sweetyee 0.55 0.19 0.10 0.09 0.01 0.09 0.10 0.09 0.01 diginnum 0.033 0.47 0.112 0.001 0.012 0.09 0.017 0.010 0.044 0.00 0.017 aster 0.033 0.47 0.112 0.037 0.48 0.40 0.00 0.017 0.010 0.011 0.00 0.044 0.06 0.07 0.07 0.017 0.010 0.011 0.021 0.09 0.011 0.01 0.021 0.09 0.011 0.02 0.05 0.07 0.01 0.01 0.05 0.017 0.01 0.01 0.05 0.017 0.01 0.01 0.05 0.017 0.01 0.01 0.05 0.017 0.01 0.01 0.05 0.017 0.01 0.01 0.05 0.01 0.01 0.05 0.01 0.01 0.05 0.017 0.011 0.01 0.015 0.01 0.011 0.01 0.015 0.01 <td></td> <td>daisy</td> <td>076</td> <td>036</td> <td>.029</td> <td>083</td> <td>.023</td> <td>.044</td> <td>.024</td> <td>025</td> <td>.112</td> <td>071</td> <td>039</td> <td>011</td> <td>.021</td> <td>024</td> <td>+.003</td>		daisy	076	036	.029	083	.023	.044	.024	025	.112	071	039	011	.021	024	+.003
deptinium -0.33 9.47 112 -0.37 -0.49 9.46 -0.90 -0.27 -0.11 0.30 -0.44 2.96 -0.49 0.18 atter 0.43 -0.06 -0.11 0.20 0.73 -0.90 0.70 0.00 0.20 -0.00 0.70 0.01 0.20 0.00 0.01 0.71 0.70 0.71 0.70 <td></td> <td>speedwell</td> <td>.005</td> <td>.068</td> <td>.020</td> <td>.028</td> <td>047</td> <td>045</td> <td>049</td> <td>009</td> <td>071</td> <td>.152</td> <td>047</td> <td>.030</td> <td>029</td> <td>.055</td> <td>035</td>		speedwell	.005	.068	.020	.028	047	045	049	009	071	.152	047	.030	029	.055	035
ster 0.43 -1.06 0.21 0.25 0.20 -0.20 -0.21 -0.20 -0.20 -0.07 1.07 Izende -0.00 0.76 -0.25 0.07 -0.01 -0.01 -0.25 -0.07 -0.01 -0.21 -0.25 -0.07 -0.01 -0.02 -0.05 -0.07 -0.01 -0.02 -0.05 -0.07 -0.01 -0.02 -0.05 -0.07 0.05 -0.01 -0.05 -0.07 0.05 -0.01 -0.05 -0.07 -0.01 -0.01 -0.05 -0.07 0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.01 -0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.01 -0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.05 -0.01 -0.01 -0.01 -0.01		sweetpea	.055	019	- 106	.050	.004	- 025	.050	.012	039	- 047	.187	044	009	017	.005
Izender -0.00 0.07 -0.07 -0.07 -0.01 <t< td=""><td></td><td>delphinium</td><td>033</td><td>.047</td><td>.112</td><td>037</td><td>049</td><td>.046</td><td>098</td><td>027</td><td>011</td><td>.030</td><td>044</td><td>.296</td><td>046</td><td>.018</td><td>187</td></t<>		delphinium	033	.047	.112	037	049	.046	098	027	011	.030	044	.296	046	.018	187
helebox 0.44 -0.49 -0.40 0.03 -0.31 -0.40 -0.03 -0.31 -0.30 -0.31 0.00 -1.41 0.00 -1.41 0.00 -1.41 0.00 -1.41 0.00 -0.41 0.00 -0.31 0.00 -0.41 -0.41		aster	.043	106	011	.025	.073	035	007	080	.021	029	009	046	.179	071	.020
Activity Setandine 567* -0.90 -2.00 4.73 1.38 -6.54 0.10 -0.90 -2.00 2.73 1.38 -6.54 0.10 -0.90 -2.00 2.26 -0.50 -0.90 -2.00 2.26 -0.90 -0.71 -0.90 -0.71 -2.27 -2.34 -0.79 -0.90 1.16 -1.44 -0.97 -1.42 -1.44 -0.97 -1.42 -1.47 -1.29 -2.27 1.20 -1.30 -2.07 -2.07 -2.07 -2.07 -2.07 -2.07 -2.07 -2.07 -2.		lavender	008	.076	+.025	.007	087	014	.004	.001	024	.055	017	.018	071	.129	+.019
rose 090 .714 ^a .226 618 .028 029 127 234 279 090 186 454 commover 200 .226 757 ^b 373 225 107 187 184 109 517 444 607 1445 primrose 203 602 373 .660 ^a 007 197 122 593 109 517 444 607 1445 primrose 473 602 073 600 007 122 518 120 513 117 121 115 120 513 172 143 117 229 121 115 123 115 <		helebore		049	061	.037	.035	- 016	.098	011	003	035	.005	187	.020	019	.329
conflower 200 2.28 7.52* 27* 210 187 184 1.18 1.19 51* A.34 51* 145 primose A.73 02 373 .600* .007 527 .222 .53 .10* .279 .162 .145 primose A.73 60* .007 .50* .222 .123 .50* .729 .122 .53 .170 .279 .162 .142 .044 popy .138 .618 .225 .007 .72* .022 .158 .10* .158 .20* .29* .24* .10* .11* <td>Anti-image Correlation</td> <td>celandine</td> <td>.667*</td> <td>090</td> <td>- 200</td> <td>.473</td> <td>.138</td> <td>654</td> <td>.013</td> <td>085</td> <td>497</td> <td>.029</td> <td>.278</td> <td>133</td> <td>.222</td> <td>050</td> <td>.168</td>	Anti-image Correlation	celandine	.667*	090	- 200	.473	.138	654	.013	085	497	.029	.278	133	.222	050	.168
primose 4.73 062 373 .060 ⁴ 0.07 597 222 1.22 593 1.70 2.79 162 1.42 .044 popp 1.18 618 225 .007 7.12 ⁴ 602 1.56 .016 2.08 618 .027 290 113 161 120 117 116 117 116 117 117 120 117 1117 1117 <td></td> <td>rose</td> <td>090</td> <td>.714*</td> <td>.226</td> <td>062</td> <td>618</td> <td>.028</td> <td>028</td> <td>127</td> <td>- 234</td> <td>.379</td> <td>093</td> <td>.186</td> <td>543</td> <td>.454</td> <td>184</td>		rose	090	.714*	.226	062	618	.028	028	127	- 234	.379	093	.186	543	.454	184
poppy 1.38 -6.88 -2.25 0.007 7.12* -0.62 1.156 1.26 2.08 -2.27 -2.29 5.13 -7.20 bluebel -6.64 0.08 -2.10 -5.97 -0.62 7.10* -1.79 1.32 0.61 -3.13 -1.67 2.29 -2.24 -1.05 butterup 0.01 -0.27 -1.46 1.15 1.17 1.32 0.11 -1.01 -2.26 2.04 .010 1.15 1.15 1.15 1.16 2.08 2.01 -2.26 .024 .010 .011 -1.01 .208 .021 .010 .011 .012 .010 .012 .103 1.15 .015 .015 .011 .015 .015 .015 .015 .015 .015 .015 .015 .115 .115 .115 .115 .115 .115 .115 .115 .115 .115 .115 .115 .115 .115 .115 .115 .115		comflower	- 200	.226	.752*	- 373	- 225	.210	187	146	.184	.109	517	.434	- 057	145	226
bluebell 684 0.29 2.10 697 692 7.19 1.72 3.81 431 457 2.29 421 419 butterup 0.13 698 677 622 619 619 619 619 619 619 619 619 623 619 619 619 619 619 629 629 619 619 619 619 619 619 619 629 629 619 <t< td=""><td></td><td>primrose</td><td>.473</td><td>062</td><td>373</td><td>.660*</td><td>.007</td><td>597</td><td>222</td><td>.122</td><td>593</td><td>.170</td><td>.279</td><td>162</td><td>.142</td><td>.044</td><td>.156</td></t<>		primrose	.473	062	373	.660*	.007	597	222	.122	593	.170	.279	162	.142	.044	.156
buttercup 0.13 -0.26 -1.87 -2.22 1.56 -1.79 8.12* -1.30 1.15 -1.98 1.14 -2.25 -0.05 wisteria -0.05 -1.27 -1.46 1.122 -1.56 1.12 -1.10 9.22* -1.75 -0.53 0.01 -1.14 -4.32 0.07 daisy -649 -2.24 1.14 -5.93 2.08 2.01 1.15 -1.75 3.43* -6.52 -2.07 -0.05 1.16 -1.16 9.24* -1.05 3.43* -6.52 -2.07 -0.05 1.16 -1.16* 9.24* -1.05 3.43* -5.42 -2.07 -0.05 1.16*		poppy	.138	618	- 225	.007	.712*	062	.156	156	.208	358	.027	269	.513	720	.183
wisteria 085 127 146 122 158 123 130 522* 175 633 601 114 432 007 dainy 497 244 144 593 204 115 175 745 593 601 114 432 .0.07 dainy 497 244 144 593 208 115 175 743* 524 207 615 159 598 598 598 533 598 533 598 513 598 513 598 513 598 514		bluebell	654	.028	.210	597	062	.710*	179	.132	.361	313	157	.229	- 224	105	078
daity -487 -224 .184 -553 208 .361 .115 .175 .743 422 267 061 .151 .196 speedwiii .029 .379 .109 .170 .358 159 050 422 276 297 213 143 174 395 sweetpes .276 090 517 .279 021 157 157 267 279 279 133 143 174 395 sweetpes 276 093 517 279 157 157 267 279 157 157 157 157 157 157 157 279 157 157 279 157 157 218 167 157 157 218 167 157 157 218 157 157 157 157 157 157 157 157 157 157 <		buttercup	.013	028	187	- 222	.156	179	.812*	130	.115	198	.184	285	026	.019	.270
speedwell 0.29 3.79 1.09 1.70 -3.58 -3.13 -1.98 -6.53 -5.42 7.70 ⁴ -2.79 1.43 174 3.95 swestpea 2.78 -0.93 -5.57 2.79 0.27 -157 1.84 0.61 -267 -2.79 0.85 ⁶ -187 -0.60 -111 delphnum -133 1.06 .434 -162 -2.29 2.28 -114 -001 1.43 -187 -0.90 -111		wisteria	085	127	146	.122	156	.132	130	.922*	175	053	.061	114	432	.007	045
sweetpra 278 -093 -517 279 027 -157 184 061 -267 -279 .855* -187 -060 -111 delphinum -133 1.06 .434 -162 -269 229 -285 -114 -061 1.43 -187 702* -199 .090		daisy	497	234	.184	593	208	.361	.115	175	.743*	542	267	061	.151	196	017
delphinium -133 186 .434 -162 -269 229 -285 -114 -061 143 -187 702 ⁴ -199 090		speedwell	.029	.379	.109	.170	358	- 313	198	053	542	.770*	- 279	.143	174	.395	156
delphinium -133 .186 .434 -162 -269 .229 -285 -114 -061 .143 -187 .702* -199 .090		sweetpea	.278	093	- 517	279	.027	157	.184	.061	- 267	- 279	.855*	187	050	-111	.019
			133	.186	.434	-162	- 269	229	- 285	114	061	.143	187	702*	199	.090	598
																	.082
lavender - 050 454 - 145 044 - 720 - 105 019 007 - 196 395 - 111 090 - 465 754*																	094
Australiant - 0.00 A 04 - 1.00 June - 1.20 - 1.00 Juli - 0.00 - 1.10 Juli - 1.11 Juli - 1.00 - 1.00 - 1.01																	.780*

a. Measures of Sampling Adequacy(MSA)



The on-diagonal values are the KMO values for each variable, indicating the factorability. Thus, for the variable 'primrose', KMO = .660. If any variable has a KMO value of less than .5, consider dropping it from the analysis. If you carry out the analysis, you will see that the KMO for 'primrose' is the smallest, so none of the variables need be dropped. The single KMO value, in the KMO and Bartlett's Test table above, is the mean of all these KMO values.

The matrices above were all described in Section 1. The remaining tables in this factor analysis output are mostly new to you.

Com	munaliti	es	The communalities indicate how much variance in each variance is explained by the analysis.		
	Initial	Extraction			
celandine	1.000	.721	In a principal component analysis, the initial communalities		
rose	1.000	.668	calculated using all possible components, and always = 1 .		
cornflower	1.000	.848			
primrose	1.000	.752			
рорру	1.000	.837	The extraction communalities are calculated using the extra		
bluebell	1.000	.850	factors only, so these are the useful values. For 'bluebell', 8 the variance is explained by the extracted factors. If a partic		
buttercup	1.000	.606	variable has a low communality, consider dropping it from the		
wisteria	1.000	.814	analysis.		
daisy	1.000	.832			
speedwell	1.000	.786			
sweetpea	1.000	.779			
delphinium	1.000	.788	We give a note on the calculation of extraction communalitie		
aster	1.000	.710	the Component Matrix table, further down.		
lavender	1.000	.862			
helebore	1.000	.723			

This table summarises the total variance explained by the solution to the factor analysis. Here, it is shrunk to fit; each third of it is reproduced below, and annotated.

This is the first part of the output that gives a clear indication of the solution, in terms of how many factors explain how much variance. The previous tables and matrices are important, though, for indicating whether the solution is likely to be a good one.

Ļ		Initial Eigenvalu	ies	Extractio	n Sums of Square	ed Loadings	Rotation Sums of Squared Loadings ^a
Component	Total	% of Variance	Cumulative %	Total	% of Variance	Cumulative %	Total
1	6.810	45.403	45.403	6.810	45.403	45.403	4.663
2	3.540	23.602	69.004	3.540	23.602	69.004	4.785
3	1.226	8.176	77.180	1.226	8.176	77.180	5.152
4	.716	4.773	81.953				
5	.591	3.937	85.890				
6	.419	2.793	88.683				
7	.381	2.538	91.221				
8	.299	1.995	93.216				
9	.252	1.683	94.899				
10	.233	1.555	96.454				
11	.189	1.261	97.716				
12	.139	.924	98.639				
13	.108	.717	99.356				
14	.052	.347	99.703				
15	.045	.297	100.000				

Total Variance Explained

Extraction Method: Principal Component Analysis.

a. When components are correlated, sums of squared loadings cannot be added to obtain a total variance.



In this type of table, the rows relate not to variables but to factors/ components.

Below are the three sections of the Total Variance Explained table.

		Initial Eigenvalu	les
Component	Total	% of Variance	Cumulative %
1	6.810	45.403	45.403
2	3.540	23.602	69.004
3	1.226	8.176	77.180
4	.716	4.773	81.953
5	.591	3.937	85.890
6	.419	2.793	88.683
7	.381	2.538	91.221
8	.299	1.995	93.216
9	.252	1.683	94.899
10	.233	1.555	96.454
11	.189	1.261	97.716
12	.139	.924	98.639
13	.108	.717	99.356
14	.052	.347	99.703
15	.045	.297	100.000

The left section contains initial eigenvalues: the eigenvalues for all possible components, ranked in order of how much variance each accounts for.

There are 15 possible components: the same as the number of variables entered into the analysis, but that does not mean that each variable is a component.

For each component, the total variance that it explains on its own (its eigenvalue) is followed by the variance that it explains expressed as a percentage of all the variance, then by the cumulative percentage.

		Initial Eigenvalu	les
Component	Total	% of Variance	Cumulative %
1	6.810	45.403	45.403
2	3.540	23.602	69.004
3	1.226	8.176	77.180
			1

The three extracted components together explain 77.2% of variance.

Initial Eigenvalues							
onent	Total	% of Variance	Cumulative %				
	6.810	45.403	45.403				
	3.540	23.602	69.004				
	1.226	8.176	77.180				

The middle section contains information for those components with eigenvalue > 1.0: in this example there are three such components.

These values are called extraction values, because they are calculated after extraction of components. Note that, in principal component analysis, these values are the initial values (the first three rows above).

Rotation Sums of Squared Loadings ^a
Total
4.663
4.785
5.152

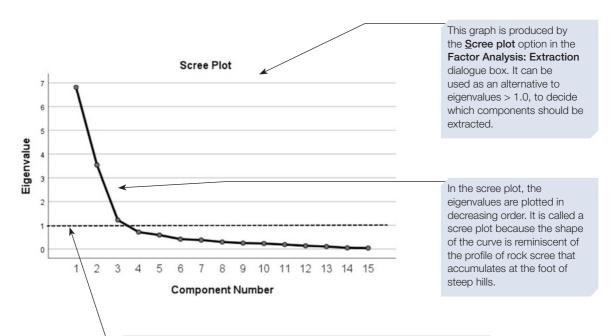
The right section of the output table shows the values for the extracted components after rotation has been carried out. (This section only appears if you request rotation, and may differ depending on the rotation method.)

Note that the eigenvalues have changed. Relative percentages and cumulative percentage of variance explained are not given for this rotation method.

a. When components are correlated, sums of squared loadings cannot be added to obtain a total variance.



The fact that three components have been extracted is neat; to that extent, the factor analysis might support our hypothetical aesthete's view. We do not yet know, however, what the factors represent. For example, they could represent not the type of flower but the colour of the flowers (pink, blue and yellow). Nor do we yet know which variables will be associated with which of the factors. So, rejoicing should be postponed.



This dotted line, which we have superimposed, indicates the approximate point at which the scree plot 'breaks' between the steep and shallow parts of the slope: components above that point would be chosen. In this example, three factors are indicated, which concurs with the choice made on the basis of selecting factors with eigenvalues greater than 1. Some authors suggest that the next factor should also be included (giving four in this example): see Cooper (2010, 287–8) for discussion.

Component Matrix^a 🖌

	C	omponent	
	1	2	3
celandine	.455	.701	.149
rose	.706	375	.170
cornflower	.709	044	587
primrose	.498	.706	074
рорру	.798	315	319
bluebell	.561	.731	034
buttercup	.353	.679	.142
wisteria	.842	300	.126
daisy	.772	.443	.197
speedwell	.724	.487	.158
sweetpea	.844	183	182
delphinium	.565	415	.544
aster	.741	377	.134
lavender	.774	316	403
helebore	.517	585	.336

Extraction Method: Principal Component Analysis.

a. 3 components extracted.

This is a table of the factor loadings *before* the rotation is carried out.

Each column shows the loading of each variable on that component. The loading can be thought of as the correlation between the component and the variable: thus, the larger the number, the more likely it is that the component underlies that variable.

Note that some loadings are positive and some negative: we will mention this in Section 3.

The variable 'poppy' has:

quite a strong loading (.798) on component 1 a medium loading (-.315) on component 2 a medium loading (-.319) on component 3.

These loadings may be useful for seeing the pattern of which variables load most strongly onto which factors: almost always a rotation will be used, however, and then the pattern becomes clearer. In particular, the negative loadings here may be an artefact of the method of calculation (Kline, 1994, 39).

The extraction communalities, in the Communalities table above, are calculated using the formula Σx^2 , where x is the factor loadings in this table. Thus, for 'lavender':

 $(.774)^2 + (-.316)^2 + (-.403)^2 = .862$

As stated above, the size of the communality indicates how much of that variable's variance is explained by the solution to the factor analysis. This table is produced by the **<u>Reproduced</u>** option in the **Factor Analysis: Descriptives** dialogue box. The upper matrix contains the reproduced correlations and the lower matrix contains the residuals. This is the whole table shrunk to fit. Underneath is an annotated section of each matrix.

¥				Rep	roduced C	orrelatio	ns									
•		celandine	rose	comflower	primrose	рорру	bluebell	buttercup	wisteria	daisy	speedwell	sweetpea	delphinium	aster	lavender	helebor
Reproduced Correlation	celandine	.721*	.084	.205	.711	.095	.763	.658	.192	.692	.694	.229	.048	.093	.071	12
	rose	.084	.668*	.417	.074	.627	.116	.019	.728	.412	.355	.634	.647	.688	.597	.64
	comflower	.205	.417	.848*	.366	.766	.386	.137	.536							
	primrose	.711	.074	.366	.752*	.199	.798	.645	.198	This	is part	of the	e reproc	duced	k	
	рорру	.095	.627	.766	.199	.837*	.228	.023	.726	corre	lation	matr	ix. Con	nare	thie	
	bluebell	.763	.116	.386	.798	.228	.850°	.690	.245							
	buttercup	.658	.019	.137	.645	.023	.690	.606*	.112	value	e, of .7	11 for	' 'primro	ose',	with	
	wisteria	.192	.728	.536	.198	.726	.249	.112	.814	'celandine', with the observed						
	daisy	.692	.412	.412	.683	.414	.750	.602	.541							
	speedwell	.694	.355	.399	.693	.374	.756	.609	.483							
	sweetpea	.229	.634	.713	.305	.789	.346	.148	.743							
	delphinium	.048	647	.099	052	408	005	005	.665							
	aster	.093	.688	.463	.093	.667	.136	.025	.754	Another point to note is that the factor						
	lavender	.071	.597	.799	.193	846	217	.002	.696							
	helebore	125	.642	.195	180	.490	149	167	.651							
esidual ^b	celandine		.016	010	145	.016	.022	139	.014	anal	/sis aiv	es a s	stronge	r rela [:]	tionsh	ain
	rose	.016		063	.030	.045	.020	.008	.016	-	0		0			
	comfower	010	- 063		.012	072	041	.051	005	petw	reen th	e two	•			
	primrose	145	.030	.012		.007	016	.011	035	.009	081	049	.067	024	007	.0
	рорру	.016	.045	072	.007		.006	003	027	007	.005	065	.045	096	.022	0
	bluebell	.022	.020	041	016	.006		037	021	083	035	027	011	.053	.026	.0
	buttercup	139	.008	.051	.011	003	037		.042	134	097	060	.052	.081	.023	0
	wisteria	.014	.016	005	035	027	021	.042		018	025	055	065	.057	018	0
	daisy	.005	015	007	.009	007	083	134	018		.045	.049	026	079	012	.0
	speedwell	035	084	006	081	.005	035	097	025	.045		.089	040	051	042	.0
	sweetpea	024	080	.012	049	065	027	060	055	.049	.089		.004	051	064	.0
	delphinium	034	122	.045	.067	.045	011	.052	065	026	040	.004		092	.048	.0
	aster	021	.055	024	024	096	.053	.081	.057	079	051	051	092		.010	1
	lavender	.039	054	064	007	.022	.026	.023	018	012	042	064	.048	.010		0
	helebore	- 001	126	484	.038	029	026	- 037	- 091	007	.037	029	.001	117	- 006	

a. Reproduced communalities

primrose

poppy

bluebell

buttercup

-.145

.016

.022

-.139

.030

.045

.020

.008

b. Residuals are computed between observed and reproduced correlations. There are 35 (33-2%) nonredundant residuals with absolute values greater than 0.05.

	celandine	rose	cornflower	primrose	
celandine	.721 ^a	.084	.205	0	nal holds the reproduced ities. They are the same
rose	.084	.668 ^a	.417	values as t	he extraction communalities
cornflower	.205	.417	.848 ^a	in the Com	nmunalities table above.
primrose	.711	.074	.366	.752 ^a	
рорру	.095	.627	.766	.199	
bluebell	.763	.116	.386	.798	
celandine		.016	010	145	
rose	.016		063	.030	
cornflower	010	063		Rather tha	n do the comparison

.012

-.072

-.041

.051

Rather than do the comparison suggested above, one can simply inspect the residuals. The residual for 'primrose' with 'celandine' is –.145 (the negative sign indicates that the reproduced correlation is the stronger as mentioned above.) This is a fairly large residual, but inspect the whole matrix on your own screen to see that the other residuals are mostly small. That most residuals are small is another indication of factorability, and it is also an indication of a good factor analysis solution.



Note that the contents of the Reproduced Correlations table are calculated after the factor extraction has been carried out, so the values will vary depending on how many factors were extracted.

Pattern Matrix^a

	Component					
	1	2	3			
celandine	006	.873	.104			
rose	.651	.039	269			
cornflower	203	.083	968			
primrose	194	.823	147			
рорру	.214	028	804			
bluebell	141	.886	130			
buttercup	046	.806	.141			
wisteria	.633	.152	363			
daisy	.304	.802	069			
speedwell	.226	.806	074			
sweetpea	.296	.155	652			
delphinium	.950	.065	.177			
aster	.634	.041	324			
lavender	.128	068	879			
helebore	.822	176	056			

Extraction Method: Principal Component Analysis. Rotation Method: Oblimin with Kaiser Normalization.

a. Rotation converged in 17 iterations.

This is the Pattern Matrix table of factor loadings *after* the rotation is carried out. (The pattern matrix is easier to interpret than the structure matrix, shown below.) If you compare these values with those in the Component Matrix table above, you will see that the factor loading values have changed.

At the foot of the table, SPSS notes the number of iterations required for the rotation.

The variable 'poppy' now has a strong loading (-.804) on component 3 and low loadings on components 1 (.214) and 2 (-.028). Whether a value is negative or positive is not relevant to the strength of the loading (as with *r* in correlation). However, if the loadings on one factor are not all in the same direction, there may be a problem; for example, if scores for reversed items were not corrected before the analysis.

Note that components 1, 2 and 3 after rotation need not necessarily be the same as components 1, 2 and 3 before rotation; they are just listed in a convenient order in each table.

For each variable, we have highlighted its strongest loading: thus, the highlights indicate which variables load most strongly on which component, as specified in the annotations below.

The variables 'cornflower', 'poppy', 'sweetpea' and 'lavender' load most strongly onto component 3.

The variables 'celandine', 'primrose', 'bluebell', 'buttercup', 'daisy' and 'speedwell' load most strongly onto component 2.

The variables 'rose', 'wisteria', 'delphinium', 'aster' and 'hellebore' load most strongly onto component 1.



We have shown you the Pattern Matrix table in the layout above to demonstrate all the factor loadings. SPSS will, however, display output in a more easily interpretable layout. See description of the **Options** dialogue box in Section 3.

Structure Matrix

	C	omponent	
	1	2	3
celandine	.019	.843	143
rose	.778	.170	579
cornflower	.249	.343	899
primrose	058	.849	293
рорру	.581	.219	894
bluebell	008	.911	318
buttercup	044	.762	067
wisteria	.813	.308	697
daisy	.403	.847	437
speedwell	.327	.846	408
sweetpea	.608	.365	832
delphinium	.874	.093	278
aster	.786	.186	627
lavender	.526	.194	919
helebore	.834	092	384

The Structure Matrix table gives the correlations between factors and variables; whereas the Pattern Matrix table (above) gives the unique relationship between each factor and each variable excluding the overlap between factors. The Pattern Matrix table is easier to interpret in terms of which variables load onto which factors.

Extraction Method: Principal Component Analysis.

Rotation Method: Oblimin with Kaiser Normalization.

Component	Correlation	Matrix	
-----------	-------------	--------	--

Component	1	2	3	
1	1.000	.083	459	
2	.083	1.000	285	
3	459	285	1.000	

Extraction Method: Principal Component Analysis.

Rotation Method: Oblimin with Kaiser Mormalization.

This table gives the factor transformation matrix that was used to carry out the rotation used to derive the rotated factors. See texts recommended in Section 1.

Under this and the two previous tables, SPSS reminds you of the rotation method that you used.

Considering the results

Once you have inspected the Pattern Matrix table (above) to see which variables load on which factors, you need to think about what the variables have in common. This will help you to identify what the underlying potential factor (or psychological construct) might be, and decide on a suitable name to describe each factor. The extracted factors may, or may not, be the same as the factors you were expecting. If we now compare the pattern of loadings in the table with the aesthete's initial suggestion, we can see that there is close agreement, with just one variable on a different factor. Use caution, however; the dimensions that were suggested (liking for *types* of flower) may not be the true factors. To be more certain, you should carefully check the characteristics of the questions/items that measure your variables, and consider all possibilities and alternative explanations. In the results section of a report, you could write about the outcome of the analysis as shown below. You could incorporate more tables (e.g. the pattern matrix) or values (e.g. the factor loadings; KMO) to illustrate your points as necessary. If you have used principal component analysis, then do ensure that you refer to that in your report, and not factor analysis.



In a report you might write:

The data were analysed by means of a principal component analysis, with direct oblimin rotation. The various indicators of factorability were good, and the residuals indicate that the solution was a good one. Three components with an eigenvalue greater than 1.0 were found; the scree plot also indicated three components. The components can be thought of as representing liking for different types of flowers: component 1 – formal garden flowers; component 2 – wild flowers; component 3 – cottage garden flowers. The component loadings are shown in Table 13.1.

 Table 13.1
 The components found by principal component analysis, and the variables that load on them

Component 1		Component 2		Component 3		
delphinium	.95	bluebell	.89	cornflower	97	
hellebore	.82	celandine	.87	lavender	88	
rose	.65	primrose	.82	рорру	80	
aster	.63	speedwell	.81	sweetpea	65	
wisteria	.63	buttercup	.81			
		daisy	.80			

Section 3: OTHER ASPECTS OF FACTOR ANALYSIS

Other options from the Factor Analysis: **Descriptives dialogue box**

Determinant

The Determinant option is in the Correlation Matrix section of the Factor Analysis: Descriptives dialogue box. If you select it, the value of the determinant will be printed underneath the Correlation Matrix table (as shown below). Its value is an indication of whether factor analysis methods other than principal component analysis can be used. Its value must not be zero. If it is zero, the correlation matrix cannot be inverted (see below).

aster	.309	.000
lavender	.224	.000
helebore	.192	.000

The determinant will be printed at the bottom left of the Correlation Matrix table. NB: The determinant may be printed to three decimal places, instead of in exponential format. If so, this determinant would appear as .000 (see tip box).

Whether the determinant is printed to three decimal places or in exponential format depends on a setting on the **Edit, Options, General** tab. On that tab, in the Output area, the No scientific notation for small numbers in tables allows you to select this; however, for numbers with three or more zeros following the decimal point, exponential format is more useful.

Inverse

The Inverse option produces the Inverse of Correlation Matrix, a complete matrix for all the variables. A section of it is shown below.

	celandine	rose	cornflower	primrose	poppy	bluebell	buttercup	wisteria	
celandine	4.755	425	918	2.469	.897	-3.904	.046	425	
rose	425	4.672	1.030	323	-3.984	.165	097	627	
cornflower	918	1.030	4.450	-1.885	-1.414	1.213	625	704	
primrose	2.469	323	- The	values in this	s table (obt	tained by ir	overting the	correlation	matrix) are used
рорру	.897	-3.984	- in th	ne calculatior	is of many	methods of	of factor ext	raction othe	er than principal
bluebell	-3.904	.165			,		ot be obtair	ned (see abo	ove), then those
buttercup	.046	097	met	hods cannot	be applied	d.			

Inverse of Correlation Matrix

CHAPTER 13

The details of matrix determinants and inverses are not covered in this book. See Tabachnick and Fidell (2014, 23–32).

Other options from the Factor Analysis: Extraction dialogue box

Method

You can choose from the various methods of extracting factors that SPSS allows. To make a sensible choice, you will need to read up about each of the methods. Here we make just a few points.

Principal component analysis, shown in Section 2, tends to be the most robust method. Remember, though, that components are distinct from factors (Kline, 1994) and that principal component analysis and factor analysis are somewhat different things. Nonetheless, with large matrices, there is little difference between the solutions from different methods (Kline, 1994). Furthermore, if a factor analysis solution is stable, then you should obtain similar results regardless of the extraction method used (Tabachnick and Fidell, 2014). Cooper (2010), however, stresses that, because principal component analysis always gives larger loadings than other methods, some rules of thumb may be misleading with this method.

The SPSS output from each of the other methods looks similar to the output from principal component analysis. The values in the particular tables that show the results of the factor extraction will, of course, differ at least slightly between methods. Some other differences and similarities are:

- 1. The word 'component' is replaced by the word 'factor' for all other methods.
- 2. The same tables are produced, except that for both generalised least squares and maximum likelihood methods, SPSS prints a Goodness-of-fit Test table that can be used to test hypotheses about the number of factors. See Kline (1994) for information about this use of the goodness-of-fit test. If you change the number of factors to be extracted, you may also need to increase the Maximum Iterations for Convergence in the Factor Analysis: Extraction dialogue box to allow the Goodness-of-fit Test table to be produced.
- 3. Communalities table:
 - a. Initial communalities have a value of 1.0 in principal component analysis. In all other methods they are less than 1.0, because those methods explain the variance that is shared between all the variables (common variance) and attempt to eliminate other variance (error variance and unique variance).
 - b. Extraction communalities are calculated from the factor matrix in the same way as they are from the component matrix, except for generalised least squares.
- 4. Total Variance Explained table:
 - a. The **Initial** part of the table will hold the same values for all methods of extraction because it shows how all the variance could be explained.
 - b. In the Extraction Sums of Squared Loadings part of the table, for principal component analysis the row values are identical to the Initial Eigenvalues row

values for each component because that method explains all the variance for the extracted factors. For the other methods, the row values here differ from the **Initial Eigenvalues** row values because these methods explain common variance and eliminate other variance.

Extract options

Two mutually exclusive options, described next, by which you can affect how many factors will be extracted are presented in the lower half of the Factor Analysis: Extraction dialogue box. Either of these options may be useful for your final analysis or for obtaining information on more of the initial factors. You will need to use one of those options if you wish to inspect 'extracted factor' information (e.g. the factor loadings) for more factors than may be extracted using the default values. That information may also be useful if you want to compare factors from your data with those obtained in previous research on the same variables. In any report, you should comment on how the number of factors to be extracted was decided upon.

The first option is to change the minimum eigenvalue, normally set at 1. The eigenvalue varies according to the number of variables entered, so it is not a robust guide for the number of factors to extract. The second option is to set the number of factors (you could do that after inspecting the scree plot).

Other options from the Factor Analysis: Rotation dialogue box Method

The default setting is <u>None</u>, for no rotation. The technique of rotation, however, was devised in order to simplify the solution to the factor analysis. Thus, you would normally select one of the rotation methods that SPSS allows, from the Factor Analysis: Rotation dialogue box. SPSS provides orthogonal rotation methods (<u>Varimax</u>, <u>Eqamax</u> and <u>Quartimax</u>) and oblique rotation methods (Direct <u>Oblimin</u> and <u>Promax</u>). As Kline (1994) pointed out, psychological constructs are likely to be correlated with one another. Of the oblique rotation methods, he recommends using Direct <u>Oblimin</u>, and we used this for the analysis described in Section 2.

If you require an orthogonal method, <u>Varimax</u> is normally thought to give rotated factors that are the easiest to interpret. The <u>Varimax</u> output differs somewhat from that for **Direct** <u>Oblimin</u>, in the following ways:

- 1. In the Total Variance Explained table, the Rotation section has three columns as percent and cumulative percent of variance are included.
- 2. There is a single Rotated Component Matrix table in place of the Pattern Matrix and the Structure Matrix tables.

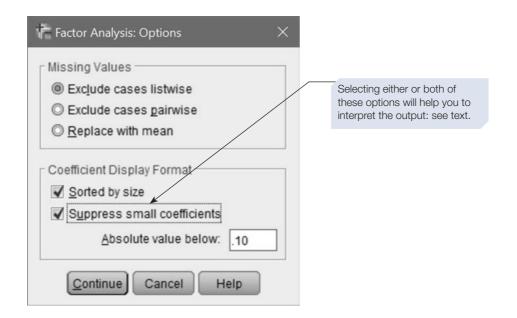
Display

If you select a rotation method, then $\underline{\mathbf{R}}$ otated solution is also selected; don't unselect it, as it provides the tables that contain the solution.

If you select $\underline{Loading plot(s)}$, SPSS draws a graph of the variables on the factor axes (up to a maximum of three factors). If you have requested a rotation, then the axes represent the rotated factors.

Factor Analysis: Options dialogue box

As mentioned in Section 2, the Factor Analysis: Options dialogue box allows you to specify missing values, and control the appearance of part of the output, as described next.



Coefficient Display Format

Two useful options here can make the output much easier to interpret. If you select **Sorted by size**, then, in the relevant tables (Component Matrix, Pattern and Structure Matrix), the variables will be sorted according to their factor loading size for each factor. If you select **Suppress absolute values less than**, then small factor loadings will be omitted from those tables. Once you have selected that option, the number can be changed, but the default of .10 seems sensible. You could try those options, and compare your output with the tables we have shown in Section 2.

Negative and positive factor loadings

Just as bivariate correlation coefficients can be positive or negative, so can factor loadings. A trivial example is when an item that taps a particular psychological construct is reversed (to avoid participant response bias). Such items should be reversed before analysis (see Chapter 4, Section 10). For factor loadings after rotation, you should note whether they are negative or positive. Negative loadings prior to rotation may be an artefact of the method of calculation used to extract factors (Kline, 1994, 39).

R factor analysis

R factor analysis is regular factor analysis, carried out on correlations between variables, as we have described in this chapter. There are other types of factor analysis in which other things are factored. For example, Q factor analysis uses correlations between participants (the rows and columns in the data file have to be reversed). More information on this and other types of factor analysis can be found in Kline (1994).

Section 4: RELIABILITY ANALYSIS FOR SCALES AND QUESTIONNAIRES

Anyone can produce a scale, and if the guidelines on writing items available in the literature are followed, then the individual items should be acceptable. There are many existing psychological scales, however, and it is unlikely that you cannot find one that assesses the construct/s in which you are interested. Moreover, it is considered better to use an existing scale than to produce another, as researchers can then compare findings from different samples and situations. For many existing scales, reliability information has been published. Nonetheless, cultural differences, or changes in language over time, or simply sample and situation differences, may affect a scale. Thus, whether you are constructing a scale or using an existing one, it is good practice to analyse the data for reliability and dimensionality. Many issues surrounding the use of scales are beyond the scope of this book. We recommend: Cooper (2010, particularly Chapters 15 and 18); Fife-Schaw (2012); Rust (2012); and John and Benet-Martinez (2014). Here, we only consider reliability, and dimensionality (Section 5).

Test-retest reliability involves testing the same participants with the same scale on two separate occasions, and calculating the correlation between the two sets of scores. It assesses the stability of a scale across time. Parallel forms, or parallel tests, describe the situation in which more than one version of a scale is available, designed to measure the same construct. To assess how similar they are, one would administer them to the same participants at the same time and correlate the scores. Internal consistency is the type of reliability we are concerned with here.

Internal consistency

If items within a scale are intended to measure aspects of the same construct, then they should all be fairly strongly correlated with each other. One way of assessing this is to correlate every item with each of the other items and inspect the matrix. Measures of internal consistency have been developed, however, which greatly simplify this process. *Split-half reliability* is an early measure, in which responses for two halves of the items are summed and then the correlation between them is calculated. Which items should go in which half, however, is a matter of debate. *Cronbach's alpha*, also called 'coefficient alpha', became easy to obtain with increasing computer power. It is related to the mean correlation between each pair of items and the number of items in the scale. It is the most commonly reported measure, with a rule of thumb that a scale should have a minimum Cronbach's alpha value of .7. It does have drawbacks, however. For example, even if a scale does have a high alpha, some individual items may be poorly correlated with the others. Thus, we should also inspect other information in the SPSS output. In the annotations below, we describe how to use three particular values, which are provided for each item:

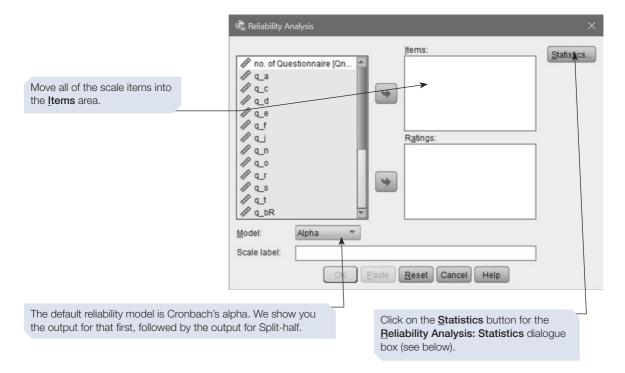
- 1. The *part-whole correlation* (or *item-total correlation*), which is the correlation between each item and the sum of the other items.
- 2. The *squared multiple correlation* for each item; that is, the R^2 obtained if the item is entered into a multiple regression as the criterion variable with all the other items as predictor variables.
- 3. The value of Cronbach's alpha for the scale if a particular item is deleted.

In addition to reliability, we should also assess whether a scale is unidimensional (see Section 5).

How to perform a reliability analysis

Before you start on this section, you should run through Chapter 4, Section 10. You will need data file 'ScaleV3.sav' based on Larsen's (1995) Attitudes Toward Recycling (ATR) scale, saved in the last exercise in that section. Remember that normally you would need many more cases. Below we demonstrate two of the ways of measuring reliability.

Start by clicking on <u>Analyze > Scale > Reliability</u> Analysis. The Reliability Analysis dialogue box will appear (as shown below). Follow the instructions.





The <u>**Reliability Analysis: Statistics**</u> dialogue box has a number of useful options for the purpose of checking scales. Otherwise, SPSS only gives output for the reliability measure that you request under <u>**Model**</u>.

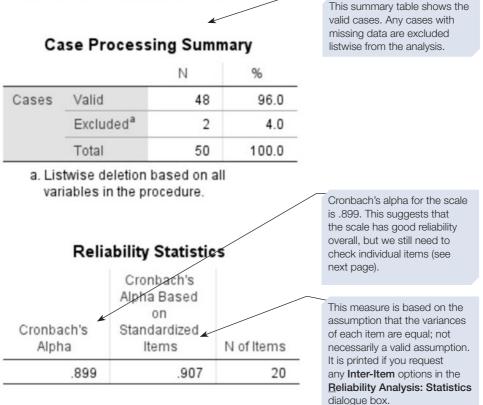
🛱 Reliability Analysis: Statistics	:	×
Descriptives for Item Scale Scale if item deleted	Inter-Item ✓ Correlations ← Covariances	Select the options shown. NB: From Inter-Item, we
Summaries <u>M</u> eans <u>V</u> ariances <u>Co</u> variances <u>Cor</u> relations	ANOVA Table <u>None</u> <u>E</u> test Friedman chi-sguare <u>Coch</u> ran chi-square	have selected Correlations, as this generates useful output in addition to the correlation matrix itself. We will not show the matrix.
Interrater Agreement: Fleiss' Kappa — Display agreement on individual ca Ignore string cases String category labels are disp Asymptotic significance level (%): 95		
Exclude <u>b</u> oth user-missing and sys User-missing values are treated as		
 Hotelling's T-square Intraclass correlation coefficient 	Tukey's test of additivity	
Model: Two-Way Mixed Confidence interval: 95 %	Type: Consistency Test value: 0	
Cance Cance	I Help	

Click on <u>Continue</u>, then on <u>ok</u>. The output is shown below.

SPSS output for reliability analysis with Cronbach's alpha

Obtained using menu item: Scale > Reliability Analysis (model = alpha)

Scale: ALL VARIABLES



The subsequent tables are obtained from options in the <u>Reliability</u> Analysis: Statistics dialogue box.

	Item Statistics							
	Mean	Std. Deviation	Ν					
q_a	3.9583	.84949	48					
q_c	3.8958	1.03635	48					
q_d	4.2708	.70679	48					
q_e	4.4583	.65097	48					
q_f	4.0208	1.06170	48					
q_j	4.3333	.88326	48					
q_n	3.6875	1.11386	48					

This table (a section only shown here) is from the <u>**Item**</u> option in **Descriptives for**.

We requested **Correlations** from **Inter-Item** in the **Reliability Analysis: Statistics** dialogue box above. A table called the Inter-Item Correlation Matrix is produced, and would normally appear here. We have omitted it, but you could inspect those correlations.

This table is from the Scale if item deleted option in Descriptives for, except for the Squared Multiple Correlation column, which is produced if you also select any Inter-Item option.

						correlation between each item and the sum of the rest of the items.
					/	The Squared Multiple Correlation
			tal Statistics	/ ¥		is the R^2 , which would be obtained
	Scale Mean if Item Deleted	Scale Variance if Item Deleted	Corrected Item-Total Correlation	Squared Multiple Correlation	Cronbach's Alpha if Item Deleted	by entering the item as the criterion variable in a multiple regression with all the other items as predictor
q_a	72.8542	107.106	.546	.726	.894	variables.
q_c	72.9167	106.418	.465	.842	.897	
q_d	72.5417	108.041	.605	.783	.893	
q_e	72.3542	108.617	.618	.700	.894	Cronbach's alpha obtained if that item
q_f	72.7917	101.147	.710	.751	.889	is omitted from the calculation.
ن_٥	72.4792	105.148	.635	.813	.892	
q_n	73.1250	114.707	.062	.380	.910	
q_o	72.5833	108.248	.530	.821	.895	
q_s	73.4583	112.466	.240	.378	.902	
q_t	72.9375	106.060	.467	.785	.897	
q_bR	72.8542	106.766	.515	.720	.895	
q_gR	72.4167	106.121	.710	.769	.891	
q_hR	73.0208	106.489	.678	.794	.892	
q_iR	73.2292	105.329	.550	.761	.894	
q_kR	73.2083	102.722	.590	.869	.893	
q_IR	72.8958	103.797	.600	.733	.892	
q_mR	73.7917	103.530	.555	.782	.894	
q_pR	72.8542	106.425	.630	.695	.892	
p_qR	73.8125	105.219	.554	.771	.894	
q_rR	73.3125	106.517	.601	.615	.893	

For item 'q_n', there are strong indications that it is not a consistent part of the scale: the part–whole correlation is very low (.062), R^2 is also low (.38), and Cronbach's alpha is increased to .91 when this item is deleted. Thus, this item would be a strong candidate for either being deleted from the scale, or being rewritten.

NB: Cronbach's alpha including 'q_n' is .899 (shown in output above). That is well above the rule of thumb of .7 for a reliable scale. Thus, this example is a good illustration of the point that Cronbach's alpha alone may be insufficient to be sure of the reliability of all items in the scale.

	Scale Statistics								
Mean	Variance	Std. Deviation	N of Items						
76.8125	117,432	10.83661	20						

This table of descriptives, for the sum of responses for the whole scale for each participant, is from the <u>Scale</u> option in **Descriptives for**.

The part-whole correlation, or the

Acting on the results

If this was the final scale, you could write: 'Cronbach's alpha for the ATR scale from the current sample was .90'. If you were assessing the scale, you would also describe other attributes from the SPSS output. For these results, however, you would consider deleting or rewriting item 'q_n', probably item 'q_s' and possibly others. If you delete items, you must repeat the reliability analyses on the remaining items. Additionally, you should consider dimensionality of the scale (Section 5).

SPSS output for reliability analysis with split-half Obtained using menu item: Scale > Reliability Analysis (model = split-half)

In the **Reliability Analysis** dialogue box, select **Split-half** instead of **Alpha** from **Model**. Only those tables that differ from the output previously illustrated are shown here.

The items in each half are shown beneath this table. The first 10 variables you moved into the **Items** box of the **Reliability Analysis** dialogue box are put into the first half, and the remaining 10 into the second half.

Reli	ability Stat			
Cronbach's Alpha	Part 1	Value	.794	
		N of Items	10 ^a	The first and third rows show the Cronbach's
	Part 2	Value	.878 🖌	alpha for each half.
		N of Items	10 ^b	
	Total N of Items		20	
Correlation Between Fo	rms		.664	
Spearman-Brown	Equal Length		.798	
Coefficient	Unequa	Unequal Length		The correlation between the sums of
Guttman Split-Half Coefficient			.790	items in each half.

a. The items are: q_a, q_c, q_d, q_e, q_f, q_j, q_n, q_o, q_s, q_t.

b. The items are: q_bR, q_gR, q_hR, q_iR, q_kR, q_IR, q_mR, q_pR, p_qR, q_rR.



You can control which items go into which half by, in the **Reliability Analysis** dialogue box, moving the variable names across one at a time in the order you wish. This procedure can be useful if you wish to compare different combinations of items. When assessing scale reliability, however, Cronbach's alpha is more often used than split-half reliability. Scale Statistics

	Mean	Variance	Std. Deviation	N of Items
Part 1	40.0833	29.142	5.39832	10 ^a
Part 2	36.7292	41.904	6.47332	10 ^b
Both Parts	76.8125	117.432	10.83661	20

The Scale Statistics table holds descriptives for the sum of responses from each half, in addition to those for the whole scale.

Section 5: DIMENSIONALITY OF SCALES AND QUESTIONNAIRES

We will describe two ways in which you can use an analysis of the dimensionality of a scale. Relevant references are given at the beginning of Section 4. First, if you want a scale in which all items assess a single construct, you can assess how strongly each item loads onto a single component. Weakly loading items would be discarded or rewritten. Second, to assess whether there is more than one construct underlying the scale, you can carry out a component or factor analysis to determine the structure. You could then use the items that load strongly on separate components as subscales. (You would need to assess the reliability of each subscale separately.) Ensuring that a scale is unidimensional, or that subscales are identified, is an aspect of construct validity.

Before you work through this section, you should be familiar with the content of Sections 1–3. Also, if you have not yet done so, you should run through Chapter 4, Section 10, as you will need data file 'ScaleV3.sav', from the last exercise in that section. We will show you the procedure with principal component analysis, but it may be more appropriate to use a factor analysis; for example, alpha factoring. Remember that in reality you would normally need many more cases than in this example, and with the data file that we are using, alpha factoring does not converge (does not produce a solution) when one factor is requested.

To identify those items that load on a single component

Enter the items into a principal component analysis (Section 2). For this purpose, the only settings you need to make are as follows:

- 1. In the Factor Analysis: Extraction dialogue box, select the option Fixed <u>number</u> of factors, then type 1 in the Factors to extract field.
- 2. In the Factor Analysis: Options dialogue box, select for Sorted by size.

For this purpose, two tables are important. An extract both of which are shown on the next page.

	Initial Eigenvalues			Extraction Sums of Squared Loadings			
Component	Total	% of Variance	Cumulative %	Total	% of Variance	Cumulative %	
1	7.654	38.272	38.272	7.654	38.272	38.272	
2	2.416	12.081	50.353		1		
3	1.463	7.313	57.666				

Total Variance Explained

Note: the largest component explains only 38% of the variance.

	Component		
	1		
q_f	.777		
q_gR	.765		
q_hR	.721		
ل_ه	.702		
q_pR	.687		
q_e	.684		
q_kR	.663		
q_d	.654		
q_rR	.644		
q_IR	.632		
q_iR	.621		
p_qR	.615		
q_mR	.614		_
q_a	.605	/	/
q_bR	.597		
q_o	.589		
q_t	.518		
q_c	.517		
q_s	.280	-	
q_n	.074		
4_11			

components extracted.

Component

The rule of thumb is that items with loadings of less than .4 should be discarded.

In this case, we can discard these questions manually – but you can also ask SPSS to this for you using the **Factor Analysis: Options** box, selecting **the Suppress small coefficients** option, and typing 0.4 into the **Absolute value below:** box.

Acting on the results

From this analysis, we would either discard items 'q_s' and 'q_n', or rewrite them. Note that these items were also identified in the reliability analysis (Item–Total Statistics table above).

If items are *discarded* (i.e. removed from the data file or questionnaire), another analysis should be carried out on the data for the remaining items, as factor loadings will change somewhat. For these data, you will find that the variance explained by the largest component increases to 42%, and the loadings of the remaining 18 items are still all above .5. You could then use the scale data (for the 18 items) in other analyses, as long as you cite the reliability and dimensionality results in your report. You must be aware, however, that another sample may give different reliability and dimensionality results.

If you were constructing a scale, you might decide to *rewrite* items instead of just deleting them; then data from a new sample must be collected with the new version of the scale. Data from the new version must then be subjected to reliability and dimensionality analyses. Thus, scale construction/modification is an iterative process.

To assess the structure of items within a scale

Enter all 20 items into a principal component analysis or factor analysis. For this purpose, we carried out the analysis in the way described in Section 2, except that:
In the Factor Analysis: Options dialogue box, we selected for Sorted by size.

We selected <u>Promax</u> in the Factor Analysis: Rotation dialogue box, as the Direct Oblimin rotation failed to converge in 25 iterations. If you wish, you could try that.

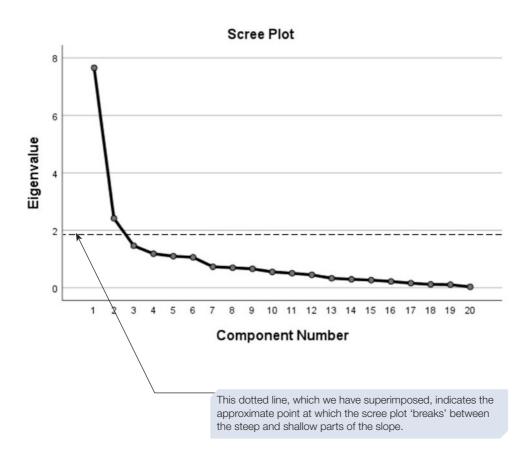
Only some of the tables from this analysis are shown here; see Section 2 for a description of all the output. Check the indicators of factorability (described in Section 2). Some of them are reasonable. The KMO value, however, is only .59. Also, some of the individual KMO values, in the diagonal of the Anti-Image Correlation matrix, are well below the value of .5 that indicates poor performance of individual items.

The section of the Total Variance Explained table below shows that there were six components with eigenvalue greater than one.

		Initial Eigenvalu	les	Extractio	n Sums of Squar	ed Loadings	Rotation Sums of Squared Loadings ^a
Component	Total	% of Variance	Cumulative %	Total	% of Variance	Cumulative %	Total
1	7.654	38.272	38.272	7.654	38.272	38.272	5.631
2	2.416	12.081	50.353	2.416	12.081	50.353	5.544
3	1.463	7.313	57.666	1.463	7.313	57.666	4.515
4	1.185	5.924	63.590	1.185	5.924	63.590	2.199
5	1.099	5.495	69.085	1.099	5.495	69.085	2.100
6	1.059	5.297	74.381	1.059	5.297	74.381	1.290
7	.731	3.655	78.037				
8	.697	3.485	81.522				
0	000	2 205	04.047				

Total Variance Explained

The Scree Plot, however, shows just two components above the break between the steep and shallow parts of the curve.



In the Pattern Matrix table below (produced by rerunning the analysis, and selecting the **Based on Eigenvalue** option in the **Factor Analysis: Extraction** dialogue) the heaviest loadings for each item are highlighted. Thirteen of the 20 items load most strongly on one of the first two components, with four items on the third. The last three components each have only one item with heaviest loading; note that items 'q_n' and 'q_s' are here.

	Component								
	1	2	3	4	5	6			
q_hR	.907	074	.068	.264	223	.054			
q_iR	.885	141	119	.267	.072	074			
q_bR	.715	.210	044	174	070	239			
q_mR	.669	.387	380	028	.082	.244			
q_f	.668	.064	.282	.065	048	029			
p_qR	.658	222	.334	141	.267	016			
q_rR	025	.845	042	052	.044	.187			
q_a	121	.640	.093	.422	078	109			
q_pR	.029	.631	038	.093	.414	072			
q_kR	.521	.617	109	191	265	076			
q_d	097	.582	.135	.496	054	.086			
q_j	.183	.560	.181	.082	096	208			
q_IR	.047	.438	.263	027	.131	.333			
q_o	.192	219	.876	.116	098	.067			
q_c	177	.215	.839	150	140	.147			
q_e	099	.400	.590	021	.148	210			
q_gR	.397	.161	.400	083	.194	.094			
q_t	.216	.025	090	.889	.162	.082			
q_s	084	.054	116	.127	.944	110			
q_n	108	.048	.113	.089	133	.881			

Pattern Matrix^a

Extraction Method: Principal Component Analysis.

Rotation Method: Promax with Kaiser Normalization.

a. Rotation converged in 17 iterations.

Acting on the results

If you obtained results like those above with a sufficient sample size, you would need to decide which option makes the most sense in terms of factor extraction... as the different methods give you the option of two, three or six factors. Unlike other statistics in this book which tend to be more 'black and white' when it comes to interpreting what they mean, interpreting the outcome of a factor analysis or PCA is greyer. In fact, at this point things become a bit more qualitative than quantitative in their

nature. While SPSS can tell you how the different variables group together, to decide how many factors to extract, you need to look at the nature and content of the items themselves to see which groupings make the most sense in real terms. Your understanding of your research question and your data should help you decide which factors to keep, and which to disregard, when your eigenvalues and scree plots don't give you a clear, single answer. The important thing is to try to account for as much variance in your data as possible, using as few factors as possible, but which group together in a meaningful way.

In the above example, you might start by checking whether a two-factor solution makes the most sense. You could do this by checking the content of the items loading onto the first two components and assessing whether they represent two distinct factors (or subscales). If you find support for the two-factor option, you might then want to collect data from a new sample to assess whether the results are replicable. You could then use exploratory factor analysis as described above, or you could use confirmatory factor analysis to test hypotheses about your measurement model. However, that is outside the scope of this book.

Important

As we have stated, a sample of 50 is small for reliability and dimensionality analyses. So if you collect data with the ATR scale (Larsen, 1995), you will probably get different results. Indeed, a different sample we analysed gave different reliability and dimensionality results from those shown above. Thus, in this and the previous section, we have shown you how to use these analyses; we have not given definitive results for the ATR scale.

Summary

- This chapter introduced you to factor analysis, and how to check the reliability and dimensionality of scales.
- Factor analysis allows you to investigate whether a factor structure underlies a set of variables.
- We have explained the output that SPSS generates for principal component analysis, which is the simplest type of factor analysis.
- Remember that, to be valid, factor analysis requires a large number of observations.
- When constructing a new scale, or using an existing scale in a new population, it is sensible to check its reliability and dimensionality.
- For guidance on incorporating SPSS output into a report, or on printing the output, see Chapter 14.

14 Using syntax and other useful features of SPSS

In this chapter

- The Syntax window
- Syntax examples
- Getting help in SPSS
- Option settings in SPSS
- Printing from SPSS
- Incorporating SPSS output into other documents
- SPSS and Excel: importing and exporting data files



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Section 1: THE SYNTAX WINDOW

- The dialogue boxes we have been using to control SPSS form an interface that allows the user to specify the analysis they require. When you click on the button, this interface translates all your selections into a series of commands telling SPSS what to do. It is possible to control SPSS using these commands directly. The language in which these commands are expressed is called the SPSS syntax.
- For the more advanced user, being able to use syntax commands will significantly increase your efficiency, allowing you to repeat complex series of analyses and to undertake analyses that would otherwise be impractical.
- A useful analogy is to think about how you record a TV programme at home. Most of the time you probably just select the programme you want from the electronic programme guide and accept all the default settings. Sometimes, however, you will want to do something a bit different, perhaps recording only the second half of a programme or deliberately overrunning the scheduled end of the programme to allow for delays in broadcast. In this situation, you will want to talk directly to the recorder and independently set the start and stop times. In the same way, using syntax commands allows you to be more flexible in your use of SPSS.
- In this section, we describe how to control SPSS directly using syntax.

An example of a syntax command

You will already have seen some SPSS syntax. Unless you have opted to suppress it, the syntax required for each command is included at the top of the output for that command (to make the output easier to follow in previous chapters we have sometimes omitted the syntax). For example, look back at the output we produced when using the **Descriptives** command in Chapter 3, Section 2 (reproduced below). You will see that the syntax for the command, including all the options you selected, appears at the start of the output. These commands have a precise structure, but it is not difficult to understand what each line does.

"Output2 [Document2] - IBM SPSS Statistics Viewer Ede Edit View Data Iransform insert Figmat Analyze Graphs Libites Edensions Window Hi Output Out						Help S SREWNESS.	instr The sele If thi oper of sy	uct S secor cted. s is th ning ti /ntax	PSS to nd and ne first ne dat	o unde d third analy a file, those	ertake tł lines re sis you l you will e shown	he Descri late to the have unde have som	commands that ptives analysis. a options we ertaken since the additional line ese lines tell you	
	2000.101100						Caral and a							
					De	escriptive	Statistics							
1		N Statistic	Range	Minimum Statistic	Maximum Statistic	Mean Statistic	Std. Deviation Statistic	Variance Statistic	Statistic	Std. Error	Kur Statistic	tosis Std. Error		
	Mem / 20	21	11.00	9.00	20.00	16.0000	3.64692	13.300	- 499	.501	-1.075	.972		
	Valid N (listwise)	21	11.00	8.00	20.00	10.0000	3.04032	13.300	433	.501	-1.070	.012		
4 5		21	1										E	

The **Descriptives** syntax command is made up of three lines of text. The first specifies the **Descriptives** command and lists the variables to be included in the analysis. Note that the second and third lines are indented and start with a slash (/) to indicate that they are continuations of the **Descriptives** command. The second line ("/SAVE") is related to the option **Save standardized values as variables** which we selected (See Chapter 3, Section 2). The third line lists the options we selected for this command and ends with a full stop indicating the end of the command.

The Paste button and the Syntax window

You may have noticed that the dialogue boxes used to execute an analysis (those that include the ok button) also contain a button marked **Paste**, which looks like this: **Paste**. If you click on the **Paste** button, the analysis is not executed, and no output is produced. Instead, control is switched to a new window called the Syntax Editor window (or 'Syntax window'), and the command lines needed to execute your analysis are pasted into this window. You can now specify a second analysis and click on the **Paste** button again. The syntax for the new command will be added to the Syntax window. In this way, you can build up a sequence of commands in the Syntax window, without executing any of them. Finally, when you have selected all the analyses you require, you can execute or run the commands. This might seem like an odd thing to want to do, but there are at least four good reasons for working in this way.

1. Duplicating actions

You may choose to work in the Syntax Editor window because you need to repeat a complex command several times. For example, when analysing the data from the adoption survey described in Chapter 4, we might need to compute 10 new variables, each of which is the mean of 10 existing variables. This would be a tedious procedure if we used the dialogue boxes, but would be easy to perform using syntax commands. This example is demonstrated later in this section.

2. Keeping a record of your analysis and repeating an analysis

Sometimes, you may want to repeat a complex series of analyses after updating your data file. This is easy to do if you saved the lines of syntax needed to execute your analysis. Alternatively, if you are working with a large, complex data set, then it is common to make errors. For example, you might undertake the wrong analysis, or use the correct command but forget to select the options you need. As a result, the Output window will fill up with erroneous output. One solution is to use the Syntax window as a notepad in which to record the details of the successful analyses. When you have the analysis working the way you want, you can save the details to the Syntax window by clicking the Pete button. In this way, you can build up a permanent record of the analysis, which you can then run to produce a 'clean' set of output. Some researchers always save the syntax of the final analysis they undertake before submitting a report or paper for publication. This means that even months later they will be able to quickly repeat an analysis and make changes if these are requested by a reviewer.

3. Describing the analysis you have undertaken

Sometimes, you will want to be able to accurately and succinctly describe exactly how you analysed a data set; for example, when emailing a colleague or in an appendix to a publication or report. Syntax is perfect in this situation as it is easy to write and is concise and precise.

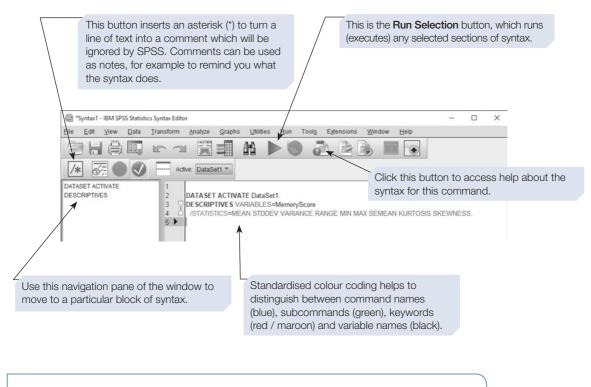
4. Tweaking the parameters of a command

Another reason for choosing to work in the Syntax window is that some of the options or parameters associated with certain commands can only be accessed using the syntax commands. In order to keep the number of buttons and options on the dialogue boxes manageable, the SPSS programmers have preset certain features of the commands. Occasionally, experienced users may want to alter one of these default settings. This can only be done using syntax. Details of the additional features of a command which can be accessed only via syntax are described in special help files available from the Help button on the dialogue box. See Section 3 for an example.

The Syntax window

The Syntax window is used to build up the syntax commands and execute them. In fact, syntax files are nothing more than simple text files, so you could use a text editor or word processing application to write these files, but the Syntax window incorporates several useful features to help you edit the syntax.

The Edit menu in the Syntax window provides access to all the normal text editing functions such as Copy, Cut, Paste and Find & Replace. Using these functions, it is possible to copy a section of syntax produced by SPSS and edit it, for example changing variable names. In this way, we can quickly produce the syntax to undertake a long sequence of commands; something that would be tedious to do using the normal dialogue boxes. The toolbar across the top of the Syntax window includes a number of useful buttons to help us construct syntax files.



Once you have constructed a block of syntax commands, select the lines you want to execute, then click on the **Run Selection** button (the large green arrow on the menu bar).

Another useful feature available in the Syntax window is that it auto-completes the names of commands and offers you lists of options to choose from. To see this in operation, open a new Syntax window (File > New > Syntax). Now start typing the command Oneway. After you have typed the first few letters, SPSS offers you a list of commands to choose from. This is illustrated in the screenshot below.

Syntax2 - IBM SP	SS Statistics S	iyntax Edito	or									-		×
		ransform	Analyze	Graphs et1 v	Utilities	<u>R</u> un	Tools	Extens	sions W	indow	Help			
Inalysis undertake Irst do a oneway A NEWAY		3 4 5	"Use full a "First do a	and final d	ata set fro ANOVA to	m Uni Da	ta Repos	itory	ing Review		^o s assgined	I to the thre	e conditio	uns.
			ONEWAY /STATISTI	ALL BRO						1	use of	0	nents t	mple of o anno
		***		HO	MOGENEIT	יר								
						IBM SI	PSS Stati	*** stics Pro	cessor is r	eady	Unicod	e:ON In 7 C	col 13	
						a sla list o STA choo	sh ch f opti FISTI	aract ons fo CS , y om. Ir	ter (/), S or this vou will n this v	SPSS comi be c	S will of mand. offered	fer you Here, if a list of	the a you c statis	

Basic rules of syntax

There are some important characteristics of syntax to note:

- 1. Each new command must start on a new line. In practice, leaving several blank lines between commands makes the syntax easier to understand.
- 2. Each command must end with a full stop or period mark (.).
- 3. Subcommands or options are usually separated by the forward slash mark (/). It is a good idea (but not essential) to start each subcommand on a new line.
- 4. You can split a command over several lines; it is safest to break the line at the start of a new subcommand.
- 5. A command or block of commands must be followed by the 'Execute' command.
- 6. It is useful to include notes to help you remember what the syntax means. These notes are called 'comments'. A comment must start with an asterisk (*) and must end with a full stop. A comment can be split over more than one line of text. SPSS will not try to interpret comments.
- 7. Make sure that you spell your variable names correctly (i.e. exactly as they appear in the Data Editor window). Misspelling a variable name is one of the most common sources of errors when running syntax commands.

In practice, it is quite rare to write a piece of syntax 'from scratch'. It is more usual to use the dialogue boxes to select an analysis and set the options, and then to paste the syntax for this command into the Syntax window using the **Pate** button. This syntax can then be copied and edited as required before being executed. Using this approach, you can be sure that the syntax and spelling will be correct. By careful use of the **<u>Find</u>** and **<u>Replace** commands (available from the **<u>E</u>dit menu)**, you can copy the syntax of a command, and change the variable(s) quickly and accurately to build up a series of analyses. An example is given below.</u>

In this example, we are seeking to compute 10 new variables. Each of these variables is the mean of a block of 10 questionnaire responses. The original variables were given names that reflect the block and question number. For example, 'b1q3' is the third question in the first block, while 'b9q8' is the eighth question in the ninth block. We could use the dialogue boxes to perform all these **Compute** commands, but this would be a laborious task. It is much easier to produce a series of syntax commands to do the work for us, and this approach is less likely to result in errors, as we can carefully check the syntax before running it. Here are the steps we could use to build up the syntax we need to perform these **Compute** commands:

- 1. Using the **Compute Variable** dialogue box, enter the details needed to compute the new variable 'b1mean' for the block 1 mean (see Chapter 4, Section 6, for details of the **Compute** command).
- 2. Click on the <u>Pate</u> button to paste the syntax commands into the Syntax Editor window (see next page).
- 3. Select this first block of text, copy it and then paste a copy of this block below the first. Leave a few blank lines between the two blocks.
- 4. Move the cursor to the start of the second block and use the <u>Replace</u> function to change all instances of the string 'b1' to 'b2'; click on the <u>Find Next</u> button, and then click the <u>Replace</u> button repeatedly until all the changes are made. Do not use the <u>Replace All</u> button, as this will replace all instances, including those in the first block.
- 5. Now repeat steps 3 and 4 until you have 10 blocks of syntax, each instructing SPSS to compute the mean of the 10 variables that make up that block. Block 1 will compute the variable 'b1mean', block 2 will compute the variable 'b2mean' and so on.
- 6. Carefully check the syntax for the **Compute** commands. Make sure you have changed all variable names systematically and that you have a full stop at the end of each line.
- 7. Make sure that you have an Execute command (with a full stop) after the last Compute.
- 8. Select all 10 **Compute** commands and click on the **Run Selection** button. The 10 new variables will be computed and appended to your data file.

These steps are illustrated below.

		This first block of syntax was produced using the Paste button on the Compute dialogue box.	
Syntax2 - IBM SPSS Statistics Syntax			- 🗆 ×
Elle Edit View Data Iran	178 J 8	tilities Bun Tools Estensions Window Help	Active: DataSet1 *
* Analysis undertake * Trist do a convery A ONEWAY * Yow concide mean re COMPUTE COMPUTE EXECUTE	Analysis undertaken i Vise full and final data First do a oneway AN ONEWIY age BY Con //STATISTICS DESCRI B	n 10 Jan 2020 to satisfy an pung Reviewer BI tet from Uni Data Repositor WA to test for ossible age differences between Ps assgined to the three conditions. Ition	This second block is a copy of the first which was edited to replace all instances of 'b1' with 'b2'.
	10 11 COMPUTE b1Mean=M 12	EAN(b1q1, b1q2, b1q3, b1q4, b1q5, b1q6, b1q7, b1q8, b1q9, b1q10).	
	14	AN(162q1, 162q2, 162q3, 162q4, 162q5, 162q6, 162q7, 162q8, 162q9, 162q10) EAN(163q1, 163q2, a, 164, 164, 164, 164, 164, 164, 164, 164	Note the blank line we have inserted between blocks.
	18 EXECUTE. 19 20	Find and Replace - Syntax2	
		Find Replace Find Signed Watch case Treat in as newline Show Options >> End Next Beplace Replace All	The third block is being edited in the same way.
I			
1		We are about to change 'b1q3' into 'b3 by clicking the Replace button. Don't u the Replace All button, as this will char every instance of 'b1' into 'b2', including those in the earlier blocks.	ise nge

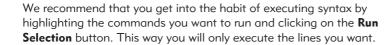
Saving syntax files

Once completed, a syntax file can be saved. If the Syntax window is the active window (i.e. if you are currently working in this window), you can simply save the contents of the window as a syntax file by selecting **Save As** from the **File** menu. SPSS will automatically add the suffix '.sps' to the end of the file name. We strongly recommend that you accept this default suffix.

It is a good idea to use the same root name for all the files relating to one project. For example, in the case of the adoption survey described in Chapter 4, the data file might be called 'Adopt.sav'. The output files produced from the analysis of this file might be saved as 'Adopt1.spv', 'Adopt2.spv' etc., and a syntax file for this research might be called 'Adopt.sps'. In this way, it is easy to see which files relate to each other. These files could then all be saved in a common folder in your file system, perhaps called 'Adoption Project Files' or something similar.

Executing syntax commands

Once you have written and saved your syntax commands, you can execute, or run, them, either by highlighting the lines of syntax you want to execute and clicking on the **Run Selection** button, or by choosing one of the options from the **<u>Run</u>** menu. It is safer to highlight the commands and use the **<u>Run</u>** Selection button.



Syntax errors

Once the syntax has been executed, any errors will be reported at the bottom of the Syntax window. These warnings are reasonably clear, and include a note of the line number at which the error occurs, so you shouldn't have too much problem tracking down the error in your file. If you have more than one error, it's a good idea to correct these one at a time, rerunning the syntax file after each error is corrected, as it is possible one error will generate several problems in the file.

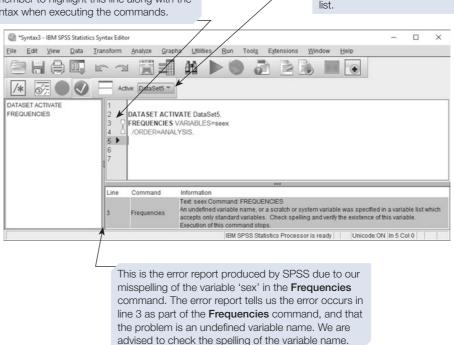
The screenshot of the Syntax window shown below illustrates an example of an error. In the second **Frequencies** command, we have deliberately misspelt the name of the variable ('seex' rather than 'sex') so that you can see what the error report looks like. SPSS reports that the error is on line 5 (which is also marked with an arrow) and explains that the variable name hasn't been defined. It even recommends you check the spelling – good advice in this case! This warning also appears in the Output window.

Selecting the correct data file

Another common error when performing data analyses using SPSS syntax is to run your syntax against the wrong data file. It is possible to have several different data files open in SPSS at the same time. This isn't something novice users are likely to do, but as you become more experienced, you may find yourself working on projects where you have several related data files, a number of which may be open simultaneously. In this situation, one of the first things you should do before running a syntax file is to check which data file is active. The active data file can be selected from a drop-down list in the Syntax window (see below), or it can be specified in a syntaxcommand, which has the form DATASET ACTIVATE Dataset1.

> If you run your syntax file on the wrong data file, SPSS will normally report errors when it encounters undefined variables. However, if you have two data files that have common variable names, SPSS may not detect any errors and your mistake could go unnoticed. Be extra careful in this situation and consider including the **DATASET ACTIVATE** syntax line at the start of each block of syntax to reduce the risk of analysing the wrong data set.

Note we have taken the sensible precaution of inserting a copy of the **DATASET ACTIVATE** command immediately prior to the syntax for our **Frequencies** analysis. This reduces the risk of analysing the wrong data set. Remember to highlight this line along with the rest of the syntax when executing the commands.



You can also select the active

data set from this drop-down

Section 2: SYNTAX EXAMPLES

One of the big advantages of being able to use SPSS syntax is that it enables you to benefit from the enormous number of helpful websites that provide short pieces of syntax to perform statistical analyses which cannot be undertaken using the dialogue boxes alone. One useful resource is the website of IBM, the company that now owns and markets SPSS. The IBM Knowledge Center for SPSS provides lots of examples of syntax routines (to find this search for 'IBM SPSS Knowledge center'). To illustrate how useful this resource is, we will now use syntax obtained from this site to compare correlation coefficients.

Comparing correlation coefficients

In Chapter 6, Section 6, we demonstrated how to compare two independent correlation coefficients in order to determine whether they were significantly different from one another. This required us to use formulae to compute Fisher's z-scores, which could then be compared against published tables of critical values. A much easier alternative is to use syntax to do these calculations for us. The section of syntax below

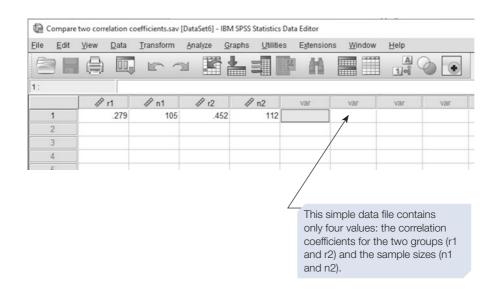
CHAPTER 14

is adapted from the IBM Knowledge Center (https://www.ibm.com/support/pages/differences-between-correlations).

*testing equality of independent correlations. *H0: R1 = R2; r1 & r2 are sample corr of x,y for groups 1 & 2 . *n1 and n2 are sample sizes for groups 1 and 2. Compute z1 = .5*ln((1+r1)/(1-r1)). Compute z2 = .5*ln((1+r2)/(1-r2)). Compute sezdiff = sqrt(1/(n1 - 3) + 1/(n2-3)). Compute ztest = (z1 - z2)/sezdiff. Compute alpha = 2*(1 - cdf.normal(abs(ztest),0,1)). Formats z1 to alpha (f8.3). List z1 to alpha. Execute.

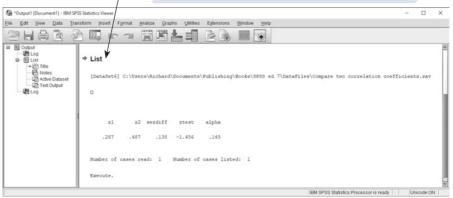
The first three lines of this syntax program are comments to help us understand the calculation. The next four lines are **Compute** statements, and these are exactly equivalent to the formula given in Chapter 6. The fifth **Compute** command calculates the alpha value (the *p* value) for the calculated *z*-score, thus avoiding the need to use tables to look up the critical value. The **Formats** command ensures that the new variables are displayed to three decimal places. The **List** command displays the values for the new variables.

Before we can run this syntax, we need a simple data file. The file will consist of just one row of data containing four variables: the correlation coefficients for the two groups to be compared, and the sample size (n) for each of these two groups. The data file is shown below.



Set up this simple data file and save it, then type the syntax commands into a Syntax Editor window. Make sure you include the Execute command at the end of the block of commands, highlight the lines of syntax and select **Run**. You should see the output given below.

The output of the **List** command shows the value for the five new variables we have computed, including the value of Fisher's *z*-test ('ztest') and the significance value ('alpha'). Compare these two values with those calculated by hand in Chapter 6 (any minor differences are due to rounding errors).



The syntax file we created to undertake this analysis demonstrates the use of the keyword **TO** in SPSS syntax. This keyword allows us to specify a list of consecutive variables. In the example above, we computed a total of five new variables, 'z1', 'z2', 'sezdiff', 'ztest' and 'alpha'. Because these are consecutive variables, occurring next to each other in the data file, we can specify all five with the syntax 'z1 to alpha'. Thus, the syntax 'List z1 to alpha' lists all five variables, as seen in the output above. This can be a useful way of specifying a large number of variables very simply.

System variables

SPSS reserves several special variables, called 'system variables', for its own use. You rarely see these, but they can be useful when writing syntax. System variable names always commence with the special character '\$'. One useful system variable is \$casenum, which is the case number – the row number in your data file. You can use this system variable in **Compute** commands. For example, the following lines of syntax will create a new variable 'ParticipantNum' and set it equal to the case number.

> Compute ParticipantNum = \$casenum. EXECUTE.

Type these lines of syntax into a Syntax window (don't forget the full stops) and then highlight and run them. You will find that the new variable has been added to the active data file. Note, however, that \$casenum is always equal to the current case number, so it is important to use this syntax to produce the \$casenum variable before doing anything that might change the order of the rows in the data file (for example, using the **Sort** command reorders the cases; see Chapter 4).

Another useful system variable is \$sysmis, which SPSS uses to indicate a system missing value. System missing values appear as a dot in the appropriate cell of the data file. An example of the use of \$sysmis is when creating a new variable. It is sometimes useful to initialise a new variable as missing, and then subsequently change this default value using **Recode** commands if certain conditions are met. An easy way to do this is with the system variable, \$sysmis. The syntax is shown below. If you run this syntax, you will see the new variable is created and all values are set to system missing.

Compute NewVariable = \$sysmis. EXECUTE.

Section 3: GETTING HELP IN SPSS

It might seem odd to wait until the last chapter of this book to describe how to use the SPSS help system, but we hope that our instructions have provided all the assistance you required so far. However, you will need to make use of the help files provided with SPSS when trying to use functions or commands not covered in this book.

The Help button in dialogue boxes

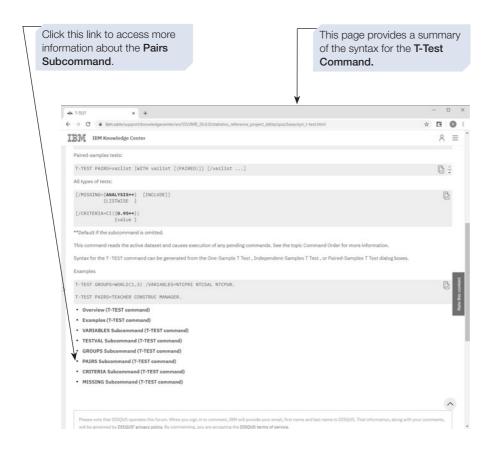
Each dialogue box includes a Help button. Click on this button to open a page of help in your internet browser. This contains a mass of useful information, including a more detailed description of the statistical procedure and the options available for the command. The page includes a number of links to related content, including a link to the syntax for the current command.

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/		our mouse over this a s for this help.	arrow to access	a table of		
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	tistics Base option. rocedure compares the means of two variables ests whether the average differs from 0.	i for a single group. The procedure o	computes the differences be	tween values of the t	two	Rate this content
measures, often called <i>before</i> data file contains the respon	blood pressure, all patients are measured at th e and after measures. An alternative design for se for the patient and also for his or her matche 5-year-old control group member).	which this test is used is a matched	d-pairs or case-control study	, in which each reco	rd in the	
	mean, sample size, standard deviation, and sta for mean difference (you can specify the confic			1000 T	in mean	s, t
Paired-Samples T Test Data						
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	Considerations pecify two quantitative variables (interval level ıbject and its matched control subject must be		easurement). For a matched	-pairs or case-contro	ol study,	

Assumptions. Observations for each pair should be made	e under the same conditions. T	The mean differences should be normally distributed. Va	riances of each variable 🔹 👻	
Hannigan and Reipdf A Mind the gap Garrpdf A	B Gerrie_et_al-2006pdf	Body worn camerapdf	Show all	

	The bottom of this page of help includes som	e useful links.
	Click here to access details of the syntax for t	his command.
A Paired-Samples T Test x		- 🗆 X
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IBM IBM Knowledge Center		8 ≡ 1
To Obtain a Paired-Samples T Test		
This feature requires the Statistics Base of	ption.	
1. From the menus choose:		
Analyze > Compare Means > Paired-	Samples T Test	
2. Select one or more pairs of variables		
3. Optionally, click Options to control the	e treatment of missing data and the level of the confidence interval.	centeers
This procedure pastes T-TEST command a	syntax.	date this contex
Paired-Samples T Test Options		
T-TEST Command Additional Feature	es	
Parent topic:	Related information	
→ T Tests	→ Detailed examples	~
	→ T-TEST	
	→ Paired-Samples T Test: Related Procedures	



The **Pairs Subcommand** is an example of an additional feature only available through the use of syntax. If we undertake a **Paired-Samples T-Test** using the SPSS dialogue boxes (as described in Chapter 6), it is necessary to specify each pair of variables to be analysed. This can be time-consuming. Using syntax, it is possible to quickly specify a large number of pairings, for example pairing the first variable in a list with every other variable in the list, or pairing each variable in one list with each variable in another list. See below for details of these options.

This page des	scribes how to use the Pairs Subcom	mand.
		- 0 3
PAIRS Subcommand (T-TEST cor x + → C i ibm.com/support/knowledgecenter,		÷ E O
-> C B ibm.com/support/knowledgecenter,	en/SSLVM8_26.0.0/statistics_reference_project_ddita/spss/base/syn_t-test_pairs.html	☆ 🖸 🔘
Table of Contents Change version or product -		🔁 Print 🖹 PDF 🗸 🕐 Help 🛛 Take a tour
PAIRS requests paired-samples t tests.		
The minimum specification for a paired-selisities two variables.	amples test is PAIRS with an analysis list. Only numeric variables can be specif	ied on the analysis list. The minimum analysis
If keyword WITH is not specified, each var	iable in the list is compared with every other variable on the list.	
 If keyword WITH is specified, every variab pairing. 	le to the left of WITH is compared with every variable to the right of WITH. WITH	can be used with PAIRED to obtain special
 To specify multiple analysis lists, use multiple used to separate each additional analysis 	tiple PAIRS subcommands, each separated by a slash. Keyword PAIRS is requi sis list.	red only for the first analysis list; a slash can
variable before WITH is compared with the first	Test, PAIEG must be enclosed in parentheses and must be used with keyword It variable after WITH, the second variable before WITH is compared with the se before and after WITH; unmatched variables are ignored and a warning messa AIES.	cond variable after WITH, and so forth. The
Example		s conte
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The first T-TEST compares TEACHER with	CONSTRUC, TEACHER with MANAGER, and CONSTRUC with MANAGER.	
The second T-TEST compares TEACHER v	ith CONSTRUC, TEACHER with ENGINEER, MANAGER with CONSTRUC, and MA	NAGER with ENGINEER. TEACHER is not
compared with MANAGER, and CONSTRUI		
 The third T-TEST compares TEACHER with 	h CONSTRUC and MANAGER with ENGINEER.	
Parent topic:	Related information	
→ T-TEST	→ T-TEST	
	→ Overview (T-TEST command)	
	→ Examples (T-TEST command)	

The Help menu

You can also access help from the **Help** menu on each window. This gives access to several sources of information. To search for help on a particular topic, select **Topics**, and then enter a few keywords into the search box on the left of the window. For example, if you search for 'chi-square', SPSS will find a long list of related help files, which it will list in the left-hand pane of the window. Select from this list to read the help information (see below).

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Chapter 4	adoption_survey	cleaned ed 7 (2).sav [DataSet1] - IBM SPSS Sta	tistics Data Editor					_					
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6	6	2.00	9.00	5.00	.00	2.00	IBM SPSS PI	ediçtive Analyti	s Community	2.00	1.00	1.00	1.00	1.00	
7	7	1.00	4.00	3.00	4.00	1.00	Give Feedba	ck		1.00	1.00	1.00	3.00	2.00	
8	8	1.00			4.00	1.00	2.00	2.00	1.00	1.00	1.00	2.00	2.00	3.00	
9	9	2.00	2.00	3.00	2.00	1.00	1.00	2.00	1.00	1.00	2.00	1.00	1.00	1.00	
10	10	2.00			3.00	2.00	2.00	1.00	2.00	3.00	2.00	2.00	3.00	2.00	
11	11	2.00	4.00	5.00	.00	5.00	4.00	5.00	5.00	5.00	5.00	5.00	5.00	5.00	
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13	13	2.00	4.00	3.00	1.00	3.00	3.00	3.00	3.00	2.00	3.00	3.00	2.00	3.00	
14	14	1.00	1.00	6.00	2.00	2.00	3.00	5.00	4.00	2.00	3.00	4.00	1.00	3.00	
15	15	1.00	3.00	3.00	1.00	9.00	1.00	2.00	1.00	1.00	1.00	1.00	2.00	3.00	
16	16	2.00			4.00	1.00	2.00	1.00	2.00	2.00	2.00	1.00	2.00	2.00	
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Help > Base Edition > Core features > Norparametric Tests > Legacy Dialogs (Norparametric Tests) > Chi-Square Test > Statistics. You can choose one or both summary statistics.		
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Quartiles. Displays values corresponding to the 25th, 50th, and 75th percentiles.		
Missing Values. Controls the treatment of missing values.		
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Exclude cases listwise. Cases with missing values for any variable are excluded from all analyses.		
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Chi-Square Test Options (One-Sample Nonparametric Tests) All categories have equal probability. This produces equal frequencies among all categories in the sa This is the default. Customize expected probability. This allows you to specify unequal frequencies for a specified list of categories. Specify	mple.	
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+ Chi-Square Test Expected Range and Expected Values		
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Other options in the <u>Help</u> menu include Command <u>Syntax Reference</u>, which provides access to a pdf file containing further details of the syntax for each command. You might like to explore all these forms of help.

What's this?

When viewing output in the Output window, double-click on a table to select it. If the pivot table opens in a **Pivot Table Editor window**, you can right-click on a heading and select <u>What's This?</u> from the drop-down menu. This will provide a short but useful description of the item. The example below shows the description of the Levene's Test for Equality of Variance in the Pivot Table for the Independent Samples Test.

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1	-1.56512	.81792	.29487	.727	.361 8.702			Equal variances not	

Section 4: OPTION SETTINGS IN SPSS

There are a number of options that can be set in SPSS. These control such things as the appearance of the various windows, the way variables are listed in dialogue boxes, the appearance of output, and the location of files. Here, we describe how to access these options and highlight a few you might like to alter.

If your screen looks different from the screenshots included in this book, this may be because some of these options settings are different. In particular, if your variables are always listed differently from ours, it may be that the **Variable Lists** options in your copy of SPSS are set differently to ours (see below).

Changing option settings

The option settings can be accessed from any of the SPSS windows. Select <u>Edit</u> > **Options**. This will bring up the **Options** dialogue box (shown below). This dialogue box has a series of tabs across the top. Click on a tab to see that set of options.

Options		×	
Charts Pivot Tables File Locations	Scripts Multiple Imputations	Syntax Editor	
Variable Lists Variable Lists Curptay lagets Alphabetical File Roles Role R	Data Currency Output Ng scientific notation for small numbe Apply localit's digit grouping format to Display a leading zero for decimal valu Measurement system: Notification:	numeric values	Click on one of these tabs to view a group of options. Here, we are looking at the General tab.
rotes to automatically assign variables (fields) to lists in dialogs. © Use custom assignments Maximum Number of Threads @ Automatic O Number of threads: 4	Parise viewer window Sport to new output Vindows Loog and feet Open syntax window at startup Open nyntax window at startup	dard 💌	If you make any changes to the options, click on the Apply button. Then click on the OK button.
OK Gance	Appy Help		

Some useful option settings

General tab

One of the most useful options in the **General** tab allows you to control how variables are listed in the dialogue boxes. By default, SPSS is set to the **Display labels** option. In this setting, SPSS lists variables by their labels (with the variable name given in brackets). When working with complex data sets, the alternative **Display names** option is

often better; this forces SPSS to list variables by name. Other options on the **General** tab allow you to control whether the variables are listed in **Alphabetical** order or in the order they are listed in the data file (**File**) or grouped by **Measurement level**. (Note that you can also choose between these options when working in an SPSS dialogue box – just right-click on the box that lists all the variables and select the options you require.)

Syntax Editor tab

In the **Syntax Editor** tab, you can change the colours applied to command names, subcommands and so on in the Syntax Editor window, and switch auto-complete on or off. You can also change where syntax is pasted into the Syntax window when you click the **Paste** button on a dialogue box (at the current cursor location or after the last command in the window).

Viewer tab

The tick box on the bottom left-hand corner of the Viewer tab controls whether the command syntax is written to the output file.

Data tab

The **Display Format for New Numeric Variables** section of the **Data** tab allows you to alter the default settings of the width and number of decimal places for a new variable. It might be useful to change this setting if you needed to create a large number of variables using the same settings. Remember, this setting alters only the way the number is displayed on screen, not the number of decimal places used when performing calculations.

Output tab

From the **Outline Labeling** section of this tab, you can select whether you want variable labels, variable names or both variable labels and variable names to appear in output. Similarly, you can choose to display either value labels, values or both value labels and values.

File Locations tab

On the **File Locations** tab, you can control the default locations for data files and other files in SPSS. You can also control the settings for the Journal file. This is a file that SPSS uses to keep a record of the syntax of all the operations you have undertaken in the current session. The Journal file can actually be really helpful if you carry out lots of analysis, and then accidentally forget to save your output or syntax file. (Something VH has done a number of times!) Because SPSS automatically keeps a running log of all of the analysis you have done (whether through the dialogue boxes or syntax), if you accidentally lose your work, you can simply look up the Journal location under the File Locations tab, navigate to the file location, and open it directly as a syntax file. Then all you need to do is highlight the section that relates to the analysis you have lost, and rerun it. Phew!

Section 5: PRINTING FROM SPSS

Here we provide some information on how to print output, data and syntax files.

Printing output from the Output viewer window

To print from the Output viewer window, either click on the printer icon at the top of the window, or select **Print** from the **File** menu. The option **All visible output** prints any output you could see by scrolling up or down in the Output window (i.e. not hidden output). Alternatively, you can print just the parts of the output you are interested in, using the **Selected output** option.

The Page Setup and Page Attributes options under the <u>File</u> menu allow you to control the paper size and orientation, the margins and to add footers and headers to your pages.

Printing data and syntax files

To obtain a printed copy of your data or syntax file, select $\underline{P}rint$ from the <u>File</u> menu of the appropriate window.



The **Eonts** option under the \underline{V} iew menu allows you to change the size and appearance of the font used to display and print the data.

Special output options for pivot tables

Most of the tables of results that appear in the Output viewer window are pivot tables. Double-clicking on a pivot table will select it, and either open a Formatting toolbar or, in the case of more complex tables, will open the table in a new a window. The tools in the Formatting toolbar can be used to adjust the appearance of the table prior to printing it. A huge number of options are available, including rotating the table (swapping rows and columns), adding or removing grid lines and scaling the table to fit the size of paper being used. Below, we describe a few of the most useful actions available from the Pivot Table window menu bar:

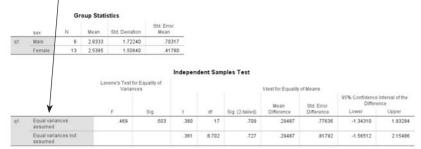
- 1. From the **Pivot** menu, select **Transpose Rows and Columns** to swap the rows and columns of a table.
- 2. From the Format menu, select <u>Table Properties</u>. The tab-style dialogue box displayed will allow you to alter the appearance of the table. The Printing tab contains two useful options (Rescale wide table to fit page and <u>Rescale long table to fit page</u>), which force SPSS to automatically adjust the size of print so that the table will fit the page without being split.
- 3. Alternatively, select **TableLooks** from the **Format** menu and change the overall style of the table using one of a number of predesigned styles. See below for an example.

Here we are going to edit the pivot table from the output of an independent samples t-test. First, double-click on the pivot table. This will open the table in the Pivot Table editor window (see below).

⇒ T-Test

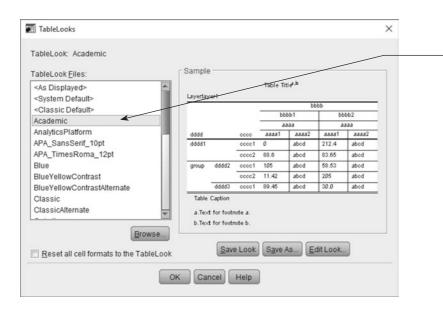
q1

assumed



Pivot Table Independent Samples Test Eile Edit View Insert Pivot Format Help example, click on Format and select TableLooks. Cell Properties. B / U A. - 51 E 10 1 S Table Properties endent Samples Test 4 TableLooks ... t-test for Equality of Means 95% Confidence Interval of the Difference Set Data Cell Widths Mean Difference Std. Error Autofit Upper Sig. (2-tailed) Lower Renumber Footnotes Equal variances assumed 17 709 29487 77636 -1.34310 1.93284 Rotate Inner Column Labels 8.702 .727 2.15486 .29487 .81792 -1.56512 Equal variances not Rotate Outer Row Labels Breakpoints

This is the Pivot Table editor window, which provides a number of tools to help change the appearance of the table. For



In the TableLooks window, you can select from a number of different table styles. The Academic style and the two APA styles are particularly useful.

Select your preferred style and then click OK. The pivot table in the output file will be reformatted to match the new style. You can then copy and paste the reformatted table into your report.

4. From the **Format** menu select **<u>A</u>utofit**. This will resize the columns and rows of the table to a size that is appropriate for their contents. This usually makes the table slightly smaller and much neater.



Before using either of the rescale options (described in point 2 above), you could apply the **<u>Autofit</u>** option. This will remove any redundant spaces from the table before it is rescaled.

- 5. From the **Insert** menu, select **Caption**. This will allow you to insert a text caption inside the table.
- 6. Select a set of table cells by clicking and dragging over them. From the Format menu, select <u>Cell Properties</u>. Using the options under the three tabs of this dialogue box, you can set the size and colour of the font and background of the cells, how the content of the cells is aligned, and the format used to display the contents. Also under the Format menu, the Set Data Cell <u>Widths</u> menu item can be used to set the width of the cells.
- 7. Double-click on any text in the pivot table, including the table title or the row or column labels to edit the text. It is also possible to edit the contents of the cells in this way.
- 8. Close the Pivot Table Editor window to apply the changes you have made. The pivot table in the Output window will be updated to reflect these changes.



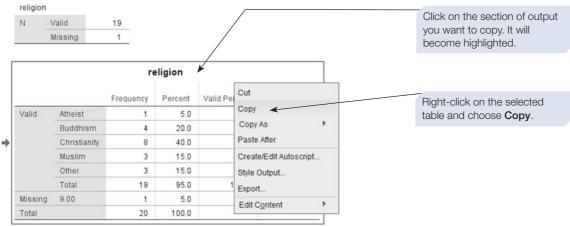
Once a pivot table is selected, it is possible to adjust the width of a column by clicking on and dragging the grid line dividing the columns. Doubleclicking on a cell allows you to change the cell contents.

Section 6: INCORPORATING SPSS OUTPUT INTO OTHER DOCUMENTS

Once you have reformatted the pivot tables in your output, you will want to incorporate it into your word-processed research report. If you are using Microsoft Word, this is simply a matter of copying and pasting the sections of the output. Follow the steps below.

Frequencies

Statistics



Cumulative

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Total

Atheist

Muslim

Other

Total

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Buddhism

Christianity

Right-click and select **Paste**. Then select one of the three **Paste Options**. Hover the cursor over each option for a preview. The first option (keep source formatting) is likely to give the best result.

document and move the

Exporting SPSS output

It is possible to export the SPSS output in one of several widely adopted file formats. For example, you can save the output as a Rich Text file, which can then be read into a word processor. This might be particularly useful if you need to send the output to someone who does not have access to a copy of SPSS. To do this, first select <u>File</u> > Export.

Export Output			×
Objects to Export			
All O All visible O Selected			Click here to select the file
Document			
Type:	Options:		format for the export.
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	Wide Pivot Tables	Wrap table to fit within page mar	
A rich text document containing both text and graphics will be created. The graphics will be	Preserve break points and groups	Yes	
embedded in the document. No graphics	Include Footnotes and Caption	Yes	
options are available.	Views of Models	Honor print setting (set in Model	
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	Page orientation	Portrait	11
	Page width	209.9699999999997	
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	Change Options	Brows	Specify the name and location
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Graphics Typg:	Optiong No options available	Broges	
Graphics Tipp: JPEG file (* Jpg)	Optiong No options available		

Section 7: SPSS AND EXCEL: IMPORTING AND EXPORTING DATA FILES

Many researchers use a spreadsheet program such as Excel to initially collate and tabulate their data. One reason for doing this is that most people will have access to an Excel-compatible spreadsheet, whereas not everyone will be able to access SPSS. A further advantage is that it is sometimes easier to pre-process a large data file in a spreadsheet program than it is to use SPSS. Fortunately, it is easy to open Excel files in SPSS. Similarly, it is quite straightforward to save an SPSS file in an Excel file format. These import and export operations are illustrated below.

Import: opening an Excel file in SPSS

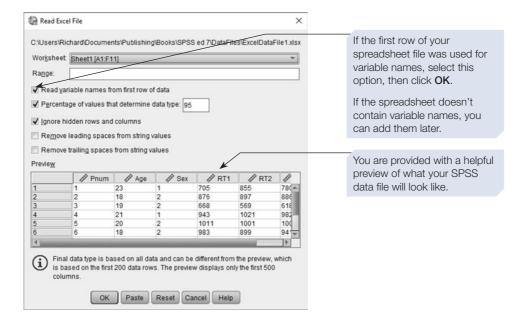
To illustrate how to open an Excel file in SPSS, we have created a simple Excel spreadsheet, shown below. The file contains some simple demographics and two variables, 'RT1' and 'RT2'. Note that in this case we have used the spreadsheet to compute 'MeanRT', the mean of 'RT1' and 'RT2'.

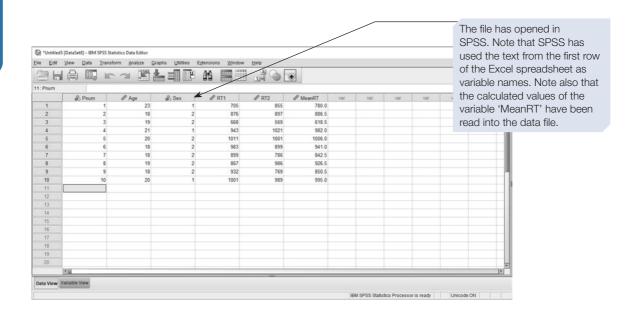
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4	A	в	С	D	E	F	G	н	-,	K L	-	
1	Pnum	Age	Sex	RT1	RT2	MeanRT						Note that the cells in column F
2	1	2	3	1 705	5 855	780						(MeanRT) contain the formula
3	2	1	8	2 876	5 897	886.5						to calculate the mean of RT1
4	3	1	9	2 668	569	618.5						
5	4	2	1	1 943	3 1021	982						and RT2.
6	5	2		2 1011	1 1001	1006						
7	6	1	8	2 983	8 899	941						
8	7	1	8	2 899	786	842.5						
9	8	1		2 867	7 986	926.5						
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We can now open this Excel file in SPSS. First, switch to SPSS and click on the menu item <u>File</u> then follow the steps below.

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	ect the Spreadsheet file n click Open.	Ensure the Excel file type is selected.
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		Help



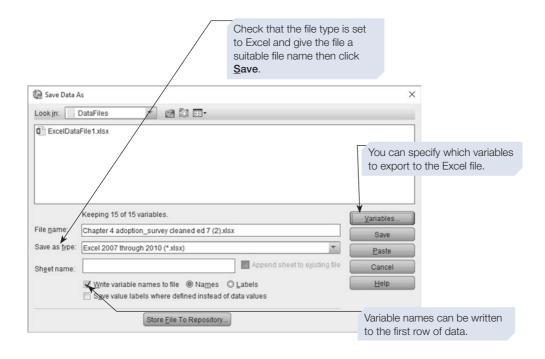


You should check the Variable View of your new file and add labels and missing values before saving the file.

Export: saving an SPSS file to Excel

It is also easy to save an SPSS data file as an Excel file.

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Collect Variable Information			Stata	a	1.00	
Stop Processor	Ctrl+Perio	bo	dBas	se	1.00	
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Exit						



The file will be saved in Excel format, and by default the variable names will be written to the first row of the spreadsheet. In this way, it is easy to move files between SPSS and Excel.

Summary

- This chapter introduced SPSS syntax. We demonstrated how to use syntax files to control SPSS, and described some of the advantages of using syntax.
- Using syntax to control SPSS can enhance your productivity and gives you an invaluable record of your analysis. Syntax can also be used to repeat a set of analyses, for example after updating a data file. Using syntax to control your analysis can be particularly useful when working with large, complex data sets.
- This chapter also showed you how to access help files, including syntax help files.
- We also demonstrated some useful option settings in SPSS.
- The final sections of the chapter showed you how to print and export SPSS files, and how to incorporate the output into Microsoft Word documents.
- Finally, we demonstrated how to move data files between a spreadsheet program, such as Excel, and SPSS. This enables you to benefit from the best features of both programs.

Appendix

All data files are available to download from macmillanihe.com/harrison-spss-7e. We recommend that you enter the first few data files to become skilled at entering data and therefore include those for Chapters 4–7 here.

Data files for	Included here	Available from website
Data handling exercises	1	✓
One-sample <i>t</i> -test	1	✓
Independent t-test	✓	✓
Paired <i>t</i> -test	1	✓
Mann–Whitney U test	✓	✓
Wilcoxon matched-pairs signed-ranks test	1	✓
Pearson's <i>r</i> correlation	✓	✓
Spearman's rho correlation	1	✓
Chi-square test	✓	✓
McNemar test	1	✓
One-way between-subjects ANOVA		✓
Two-way between-subjects ANOVA		✓
One-way within-subjects ANOVA		✓
Two-way within-subjects ANOVA		✓
Three-way mixed ANOVA		✓
Kruskal–Wallis test and Friedman test		✓
Multiple regression		✓
ANCOVA and MANOVA		✓
Discriminant analysis and logistic regression		✓
Factor analysis		✓
Log transformation exercise		✓
Scale exercise		1

DATA FOR DATA HANDLING EXERCISES: CHAPTER 4, SECTIONS 1–9

id	sex	ethnicity	religion	adopted	q1	q2	q3	q4	q5	q6	q7	q8	q9	q10
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2	1	2	3	0	4	1	5	5	1	3	2	3	1	5
3	2	4	2	1	2	3	2	2	1	2	1	1	1	1
4	2	3	2	0	2	1	1	3	3	3	3	1	1	1
5	2	2	6	0	4	5	3	4	4	2	4	5	4	3
6	2	5	5	0	2	1	9	1	2	1	1	1	1	1
7	1	4	3	4	1	1	1	2	1	1	1	3	2	5
8	1	1	6	4	1	2	2	1	1	1	2	2	3	1
9	2	2	3	2	1	1	2	1	1	2	1	1	1	1
10	2	3	3	3	2	2	1	2	3	2	2	3	2	2
11	2	4	5	0	5	4	5	5	5	5	5	5	5	5
12	1	4	5	1	4	4	3	2	4	3	4	3	4	3
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14	1	1	6	2	2	3	5	4	2	З	4	1	3	5
15	1	3	3	1	9	1	2	1	1	1	1	2	3	2
16	2	4	1	4	1	2	1	2	2	2	1	2	2	2
17	2	5	2	0	5	4	4	4	3	4	3	4	5	5
18	1	1	3	2	5	4	9	4	4	3	3	2	4	4
19	2	1	2	4	1	1	1	2	1	2	2	1	1	1
20	2	2	3	9	1	2	3	2	3	2	2	2	3	3

DATA FOR DATA HANDLING EXERCISES: CHAPTER 4, SECTION 10

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3	4	2	4	4	4	5	2	2	2	5	3	2	3	2	5	2	3	3	3	4
4	4	2	4	5	5	5	1	1	2	5	2	1	1	3	5	1	2	2	3	5
5	4	2	5	4	5	4	1	2	2	2	3	2	2	5	4	2	2	2	4	4
6	4	2	3	3	4	3	2	3	2	4	3	2	3	2	4	3	3	3	4	4
7	5	1	5	4	5	5	1	2	2	5	2	1	2	4	5	2	2	2	4	4
8	5	2	5	5	4	4	2	2	3	5	3	2	2	5	5	1	3	3	3	4
9	4	1	5	5	4	4	2	1	2	5	1	1	3	5	5	2	4	1	1	5
10	5	2	5	5	5	5	1	2	2	4	2	2	3	1	5	3	3	3	3	4
11	4	2	3	4	5	5	1	2	2	4	2	2	2	4	4	2	2	2	4	4
12	4	1	4	4	4	4	1	2	2	5	2	1	2	4	4	2	2	2	4	4
13	4	1	3	4	5	5	2	1	1	5	1	3	2	2	4	2	2	3	5	5
14	4	1	3	4	4	4	1	2	2	4	1	1	2	2	3	1	4	3	3	4
15	4	3	3	4	4	4	1	3	2	5	2	2	3	4	4	2	3	3	3	4
16	5	2	5	5	5	5	1	2	2	5	2	1	2	5	4	1	3	1	4	4
17	4	3	4	4	5	2	2	3	3	4	5	2	5	4	4	2	4	3	3	3
18	2	2	3	3	5	4	1	З	2	4	4	З	4	2	5	2	2	4	З	2
19	4	2	5	5	5	3	1	3	3	5	1	2	3	2	4	1	3	1	4	4
20	4	2	5	5	5	5	1	2	2	5	2	1	3	5	5	2	2	2	5	5
21	4	1	4	4	5	4	2	2	2	5	1	З	2	З	4	2	З	2	З	3
22	4	2	3	4	4	3	2	2	2	4	3	4	3	4	4	3	3	3	2	4
23	4	2	5	5	5	5	1	2	1	5	1	1	2	5	5	2	2	2	3	5
24	5	1	4	5	5	5	1	1	1	5	1	1	1	3	4	2	3	3	3	3
25	5	1	5	5	5	5	1	З	З	4	З	1	5	2	5	2	4	2	4	4
26	3	2	3	4	4	4	2	3	2	4	2	2	3	4	2	2	3	3	3	3

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29	4	2	4	5	5	4	2	2	2	4	2	3	2	5	4	3	2	3	4	9
30	3	2	4	4	4	3	2	2	2	4	2	2	4	4	4	2	3	2	4	4
31	4	2	5	5	5	5	1	2	2	5	3	3	4	5	5	1	2	3	4	5
32	4	3	5	5	5	9	3	3	4	4	4	4	4	3	3	4	4	4	2	4
33	3	2	3	5	4	4	1	2	2	2	5	2	2	4	5	2	3	3	5	5
34	4	2	2	5	5	4	2	2	3	5	3	4	5	4	4	2	5	3	4	4
35	4	2	5	4	4	4	2	2	4	4	2	2	4	4	4	2	2	4	4	4
36	4	3	4	5	5	5	1	2	2	5	1	1	2	5	5	3	4	2	3	4
37	4	2	5	4	5	З	2	2	З	5	2	2	4	4	5	3	З	З	З	4
38	5	2	3	4	4	4	2	2	2	4	2	3	2	3	3	2	3	2	3	4
39	5	1	4	5	5	5	1	1	1	5	1	4	3	2	4	1	1	2	З	5
40	4	1	4	4	4	3	2	2	2	4	2	2	3	4	4	2	3	2	4	4
41	4	4	4	4	4	2	4	4	4	4	4	2	4	З	4	2	5	3	З	5
42	5	1	4	5	5	5	1	3	5	5	1	1	1	5	4	1	3	1	3	1
43	5	1	5	5	5	5	1	1	1	5	1	1	2	5	5	1	2	1	З	5
44	4	2	3	4	4	5	2	2	2	5	3	1	3	4	4	2	2	3	5	5
45	2	4	2	4	4	З	2	З	З	4	З	З	З	4	З	2	4	2	З	3
46	5	4	5	5	5	5	1	2	3	4	4	2	4	5	5	2	2	3	2	5
47	З	2	5	3	5	4	2	2	2	5	2	2	3	3	5	1	2	2	4	2
48	4	2	2	4	4	5	1	2	3	5	3	1	3	3	5	2	3	2	3	5
49	3	2	2	4	3	3	2	2	2	4	3	4	2	3	4	2	3	3	3	3
50	5	5	5	5	5	1	2	4	5	4	4	2	5	5	4	2	5	2	4	4

DATA FOR ONE-SAMPLE 7-TEST: CHAPTER 5, SECTION 2

 68.00 62.00 58.00 67.00 65.00 69.00 72.00 62.00 64.00 69.00 70.00 71.00 66.00 68.00
58.00 67.00 65.00 69.00 72.00 76.00 62.00 64.00 69.00 70.00 71.00 66.00
 67.00 65.00 69.00 72.00 76.00 62.00 64.00 69.00 70.00 71.00 66.00
 65.00 69.00 72.00 76.00 62.00 64.00 69.00 70.00 71.00 66.00
 69.00 72.00 76.00 62.00 64.00 69.00 70.00 71.00 66.00
72.00 76.00 62.00 64.00 69.00 70.00 71.00 66.00
76.00 62.00 64.00 69.00 70.00 71.00 66.00
62.00 64.00 69.00 70.00 71.00 66.00
64.00 69.00 70.00 71.00 66.00
69.00 70.00 71.00 66.00
70.00 71.00 66.00
71.00 66.00
66.00
68.00
67.00
61.00
72.00
73.00
78.00

DATA FOR INDEPENDENT 7-TEST: CHAPTER 5, SECTION 3

GROUP 1 = mnemonic condition 2 = no mnemonic condition	SCORE
1	20
1	18
1	14
1	18
1	17
1	11
1	20
1	18
1	20
1	19
1	20
2	10
2	20
2	12
2	9
2	14
2	15
2	16
2	14
2	19
2	12

DATA FOR PAIRED T-TEST: CHAPTER 5, SECTION

LARGE SIZE DIFFERENCE	SMALL SIZE DIFFERENCE
936	878
923	1005
896	1010
1241	1365
1278	1422
871	1198
1360	1576
733	896
941	1573
1077	1261
1438	2237
1099	1325
1253	1591
1930	2742
1260	1357
1271	1963

DATA FOR MANN–WHITNEY UTEST: CHAPTER 5, SECTION 6

SEX 1 = male 2 = female	RATING
1	4
1	6
1	5
1	8
1	5
1	2
1	4
1	4
1	5
1	7
1	5
1	4
1	3
1	3
1	5
1	3
1	3
1	8
1	6
1	4

SEX 1 = male 2 = female	RATING
2	4
2	2
2	7
2	4
2	6
2	7
2	5
2	2
2	6
2	6
2	6
2	6
2	3
2	5
2	7
2	4
2	6
2	6
2	7
2	8

Appendix

DATA FOR WILCOXON MATCHED-PAIRS SIGNED-RANKS TEST: CHAPTER 5, SECTION

E-FIT RATING 1 (from memory)	E-FIT RATING 2 (from photograph)	E-FIT RATING	
3	6	4	4
3	4	4	2
3	5	4	5
5	6	3	3
2	3	5	3
4	3	4	3
5	3	3	2
5	3	3	3
4	3	6	4
3	3	3	3
2	3	3	2
6	6	3	3
5	3	2	4
4	3	2	5
3	3	5	6
3	5	3	5
4	5	6	4
3	2	2	3
4	5	5	5
3	5	4	2
3	2	3	5
5	6	3	2
3	4	5	6
4	3	4	2

DATA FOR PEARSON'S R CORRELATION: CHAPTER 6, SECTION 3

AGE (in years)	CFF
41	34.9
43	30.5
25	35.75
42	32.3
51	28.0
27	42.2
27	35.1
48	33.5
58	25.0
52	31.0
58	23.2
50	26.8
44	32.0
53	29.3
26	35.9
65	30.5
35	31.9
29	32.0
25	39.9
49	33.0

DATA FOR SPEARMAN'S RHO CORRELATION: CHAPTER 6, SECTION 4

CONFI-	BELIEV-	ATTRACTI-
DENCE	ABILITY	VENESS
4	4	2
4	3	3
4	6	4
4	6	4
4	4	3
4	4	4
4	3	2
5	5	4
4	4	3
6	5	4
4	6	4
4	5	5
4	4	3
6	5	4
5	5	4
4	5	3
2	4	3
6	4	4
3	5	3
3	3	3
2	5	5
5	5	4
5	6	4
5	4	4
4	5	4
5	5	4
5	5	4
4	4	5
4	4	4
3	5	4
5	6	4
5	5	4
1	5	3
5	5	4
5	5	4
5	6	4
5	5	5
4	5	4
	5	
4	5 5	4
4	5	4
5	5	3
4	4	2
5	6	5
3	0	0

CONFI- DENCE	BELIEV- ABILITY	ATTRACTI- VENESS
4	5	3
6	5	2
3	5	4
3	5	4
4	4	3
4	3	3
6	6	4
3	5	2
4	4	3
5	5	4
3	1	3
5	6	4
5	5	4
4	5	4
4	4	4
6	1	1
5	5	4
5	5	4
6	6	5
5	5	3
6	6	5
5	5	2
2	4	4
3	4	4
3	4	4
4	4	4
4	5	4
5	5	4
5	5	3
3	4	4
2	3	5
6	5	5
4	5	3
5	4	4
4	5	4
4	5	4
4	4	4
4	4	4
4	5	4
4	5	5
5	4	4
4	6	4
5	5	3
6	5	4

DATA FOR CHI-SQUARE TEST: CHAPTER 7, SECTION 4

BACKGROUND 1 = Asian 2 = Caucasian 3 = other	MOTHER'S EMPLOYMENT 1 = full time 2 = none 3 = part time	SCHOOL 1 = comprehensive 2 = private	TENDENCY TO ANOREXIA 1 = high 2 = low
2	1	1	1
2	1	1	1
2	1	1	1
2	3	1	1
2	3	2	1
2	3	2	1
2	2	2	1
2	2	2	1
2	2	2	1
2	1	2	1
2	1	2	1
2	1	2	1
2	3	2	1
2	3	2	1
2	3	2	1
2	3	2	1
2	2	2	1
2	2	2	1
2	2	2	1
2	2	2	1
2	3	2	1
2	3	2	1
1	3	2	1
1	1	2	1
1	1	2	1

BACKGROUND 1 = Asian 2 = Caucasian 3 = other	MOTHER'S EMPLOYMENT 1 = full time 2 = none 3 = part time	SCHOOL 1 = comprehensive 2 = private	TENDENCY TO ANOREXIA 1 = high 2 = low
1	1	2	1
3	2	2	1
3	3	2	1
3	2	2	1
3	2	2	1
3	1	2	1
3	1	2	1
3	1	2	1
3	1	2	1
3	1	2	1
3	2	2	1
3	2	2	1
3	2	2	1
2	2	1	2
2	1	1	2
2	1	1	2
2	3	1	2
2	3	1	2
2	2	1	2
2	2	1	2
2	2	1	2
2	2	1	2
2	3	1	2
2	3	1	2
2	3	1	2
2	3	1	2
2	3	1	2
2	3	1	2

BACKGROUND 1 = Asian 2 = Caucasian 3 = other	MOTHER'S EMPLOYMENT 1 = full time 2 = none 3 = part time	SCHOOL 1 = comprehensive 2 = private	TENDENCY TO ANOREXIA 1 = high 2 = low
2	3	1	2
2	3	1	2
2	2	1	2
2	2	1	2
2	2	1	2
2	3	1	2
2	3	1	2
2	2	1	2
2	2	1	2
2	2	1	2
2	2	1	2
2	3	1	2
2	1	1	2
2	1	1	2
2	1	2	2
2	1	2	2
2	1	2	2
2	1	2	2
2	1	2	2
2	1	2	2
2	1	2	2
1	1	2	2
1	1	2	2
3	1	2	2
3	1	2	2
3	1	2	2
3	1	2	2

DATA FOR MCNEMAR TEST: CHAPTER 7, SECTION 5

NORMAL HANDWRITING 1 = correct 2 = incorrect	HANDWRITING AS IF OPPOSITE SEX 1 = correct 2 = incorrect	NORMAL HANDWRITING 1 = correct 2 = incorrect	HANDWRITING AS IF OPPOSITE SEX 1 = correct 2 = incorrect
2	2	1	2
1	1	1	2
1	1	2	1
1	2	1	2
1	1	1	1
2	2	1	1
1	1	1	2
1	2	1	2
2	2	2	2
1	1	1	1
2	2	1	1
1	2	1	1
2	2	2	2
2	2	2	1
1	1	1	2
2	2	2	2
1	1	1	1
2	2	1	2
1	2	1	1
2	2	1	1
1	2	2	2
1	2	1	1
1	2	1	2
1	2	2	2
1	1		

Glossary

This glossary provides an explanation of many of the terms used in the book. We have used italics to indicate terms that have their own entry in the glossary. For further information about statistical or experimental design concepts, we encourage you to consult a statistics text.

ANCOVA (analysis of covariance)

An extension of analysis of variance (*ANOVA*) in which at least one covariate included in the design. In ANCOVA, a covariate is a *variable* that has a statistical association with the *dependent variable*. ANCOVA allows us to estimate the impact of the *independent variable/s* after allowing for the influence of the covariates.

ANOVA (analysis of variance)

An *inferential statistical test* that allows analysis of data from designs with more than two experimental *conditions* and/or with more than one *factor* (or *independent variable*). ANOVA is appropriate for use with *parametric* data. The absence, however, of *nonparametric* equivalents for two or more factor designs means that ANOVA is often used in such circumstances. Fortunately, it is said to be fairly robust to violations of the assumptions for *parametric* tests, provided that the *cell* sizes are equal.

Introduction to ANOVA One-way between-subjects ANOVA (see also Kruskal–Wallis) One-way within-subjects ANOVA (see also Friedman) Factorial between-subjects ANOVA Factorial within-subjects ANOVA Factorial mixed ANOVA

Association

See correlation and chi-square.

Asymptotic significance

The *p* value calculated under the assumption that the *sample* is large and has an asymptotic distribution. For the vast majority of *inferential statistical tests*, the *p* value given by SPSS is the asymptotic significance. For many tests, SPSS now gives an option to also calculate the *exact significance*.

Bar chart

A graph used to display summary statistics such as the *mean* (in the case of a *scale* variable) or the *frequency* (in the case of a nominal variable). See also *chart*.

Between-participants design

See between-subjects design.

Between-subjects design

An *experimental design* in which all *factors* are between-subjects factors; that is, each participant contributes data to only one *level* of a factor (each participant only experiences one condition). See also *independent groups design* and *natural independent groups design*.

Binary logistic regression

See logistic regression.

Bivariate

An analysis involving two *variables*, for example, *correlation*. See also *univariate* and *multivariate*.

Bivariate regression

An inferential statistical procedure used to investigate a *linear relationship* between two *variables*. It can indicate the extent to which one variable can be explained or predicted by the other variable. See also *regression*.

Case

A case is the unit of analysis. In psychology, this is normally the data deriving from a single *participant*. An exception is in a *matched-subjects design*, when the pair of matched participants form the case. Each case should be entered into a separate row in the SPSS *Data window*. In some research, the cases will not be people. For example, we may be interested in the average academic attainment for pupils from different schools. Here, the cases would be the schools.

Cell

An element in the *Data Editor window* table, into which a value is entered. In *ANOVA* and *chi-square*, the combination of one *level* of one *factor* and one level of another factor. The cell size is the number of *cases* (normally *participants*) that fall into that cell.

Chart

The name that SPSS gives to a graph. A wide range of graph types are available from the <u>**G**</u>raphs *menu item*. In addition, the output of some statistical procedures include optional charts.

Chi-square

An *inferential statistical test* that is used to analyse *frequencies* of *nominal* data. Chi-square allows comparison between the observed frequencies and the pattern that would be expected by chance. In psychology, the multidimensional chi-square is most often used. It can be thought of as a test of association between two *variables*, or as a test of difference between two independent groups.

The chi-square distribution is used to assess *significance* for some other statistical tests, for example *Friedman* and *Kruskal–Wallis*.

Cleaning

See data cleaning.

Compute

An SPSS procedure that allows us to compute (calculate) a new *variable* based on one or more existing *variables*. The new *variable* is added to the data file.

Condition

See level.

Confidence interval

A pair of values that define a range within which we expect the *population parameter*, such as the *mean*, to fall. In the case of the 95% confidence interval, these values define the range within which there is a 95% probability that the parameter will fall.

Confounding variable

Any uncontrolled *variable* that changes systematically across the *levels* of the *independent variable* or *factor*. If a study includes a confounding variable, it is not possible to determine whether the results of an experiment are due to the *independent variable* alone, to the confounding variable alone or to some *interaction* between those two variables.

Contingency table

Displays the *frequencies* or counts for the levels of one or more variables.

Correlation

Describes a *linear relationship*, or association, between two *variables* (measured on ordinal, interval or ratio *level of measurement*). *Pearson's r, Spearman's rho* and *Kendall's tau* are *inferential statistical tests* of correlation. See also *scatterplot*.

Count

An SPSS procedure that allows us to count the number of times a particular value occurs, in one or more *variables*.

Covariate

See ANCOVA.

Criterion variable

The variable that is explained by the *predictor variable*/s in a *regression* analysis. Some sources use the term 'dependent variable' or 'outcome variable' instead of criterion variable.

Data

A set of values. A data set is typically made up of a number of *variables*. In *quantitative research*, data are numeric.

Data cleaning

The process of checking the accuracy of a *data file* and correcting any errors that are found.

Data Editor window

The SPSS window in which data are entered and edited. It has the appearance, but not the functionality, of a spreadsheet window.

Data file

SPSS records a set of *data* in a data file. Data files in SPSS usually have the file extension '.sav'.

Data handling

A range of operations performed on the data after they have been entered into SPSS. The different types of data handling are accessed through the *menu items* **Data** and **Transform**. See also *compute*, *count*, *rank cases*, *recode*, *select cases*, *sort cases*, *split*.

Data transformation

In data transformation, each data value is replaced by a new value that is computed from the original value. A common data transformation is the *logarithmic transformation*.

Data View

In SPSS, the *Data Editor window* has two display settings. The Data View shows the data table, with the variables in columns and the cases in rows. See also *Variable View*.

Degrees of freedom

A value related to the number of *participants* who took part in an experiment (*t-test, ANOVA*) or to the number of *factors* (*independent variables*) in an experiment (ANOVA, *chi-square*). The degrees of freedom are required when using statistical tables of *significance*. Although SPSS gives the exact *p* value, degrees of freedom should still be reported as shown on the annotated output pages of those *inferential statistical tests*.

Dependent variable

The *variable* that is measured in an experiment, and whose values are said to depend on those of the *independent variable* (or *factor*).

Descriptive statistics

Procedures that allow you to describe data by summarising, displaying or illustrating them. Often used as a general term for summary descriptive statistics: *measures of central tendency* and *measures of dispersion*. Graphs (see *chart*) are descriptive statistics used to illustrate the data.

Dialogue box

A box that appears on the computer screen, normally after you have clicked on a sequence of *menu items*. SPSS uses dialogue boxes to allow you to control the details of a statistical procedure. Each chapter includes illustrations showing the dialogue boxes and explaining how to complete them.

Dichotomous variable

A *variable* that can only take one of two values, for example indicating the presence or absence of something.

Discriminant analysis

An inferential statistical procedure used to determine which *variables* predict membership of (or discriminate between) different categories of another *variable*.

Effect size

A measure of the magnitude of an effect. Can be expressed in the units used to measure the *dependent variable*, or in standardised units such as Cohen's *d*.

Equality of variance

See homogeneity of variance.

Error bar graph

A graph in which the *mean* of each *condition* is plotted with a vertical bar that provides an indication of the magnitude of the error associated with the measurement of the *mean*, and hence an indication of how accurate the measurement is likely to be. In SPSS, we can use error bars to indicate several different measures of error, including *standard error, standard deviation* and *confidence interval*.

Exact significance

The *p* value calculated on the assumption that our data are a small *sample* of the *population* and/or do not have an asymptotic distribution. For many tests, SPSS now gives an option to calculate the exact significance in addition to the default *asymptotic significance*. Fisher's Exact test, an alternative in *chi-square*, only gives the exact *p* value.

Experimental design

Describes specific methods by which experiments are carried out and which are intended to prevent *participant irrelevant variables* from *confounding* the experiment; for example, *within-subjects* and *between-subjects designs*. One or more variables are systematically manipulated to see whether this impacts another variable. This allows researchers to establish cause and effect. Basic designs are described in Chapter 1, and other designs are described where relevant for particular statistical tests.

Experimental hypothesis

See hypothesis.

Execute

Part of SPSS *syntax*. The execute command instructs SPSS to perform the command specified in the preceding lines of syntax.

Factor

In ANOVA: another term for *independent variable*. The term 'factor' is used particularly when discussing ANOVA statistical tests and designs, whereas the term 'independent variable' is used more often for *two-sample designs*. See also *betweensubjects design* and *within-subjects design*.

In *factor analysis*: a dimension (or a psychological construct) that underlies several measured *variables*.

Factor analysis

Statistical procedure used to identify whether a *factor* structure underlies *correlations* between a number of *variables*.

Factorial ANOVA

A type of *ANOVA* (*Analysis of Variance*) that includes more than one *independent variable* (or factor).

Factorability

Indicators of whether it is likely that there are any *factors* underlying a set of *variables* that are entered into a *factor analysis*.

F-ratio

The statistic obtained in *ANOVA* calculations. It can be described as the *variance* due to manipulation of the *factor* divided by the *variance* due to error.

Frequency/ies

The number of times a particular event or value occurs. Also an SPSS command available from the *menu item* <u>Analyze</u>, which will produce tables of frequencies showing the number of times a particular value occurs in each *variable*. See also *bar chart*.

Friedman

A nonparametric equivalent of the one-way within-subjects ANOVA.

Graph

See chart.

Graph Board Template Chooser

A system that helps the user construct *graphs* by picking from a large collection of templates that can then be modified.

Grouping variable

SPSS uses this term for the *variable* that specifies two or more groups of *cases* to be compared. For *independent groups design*, the grouping variable is the *independent variable*. For example, in an experiment to compare the performance of participants in either quiet or noisy conditions, the grouping variable will be the independent variable, noise level. In other cases, the grouping variable is an independent variable that is not manipulated by the experimenter. For example, if we wish to compare the performance of men and women, the independent variable will be sex. Grouping variables are nominal variables. See *levels of measurement*.

Help

There are a number of different sources of help available within SPSS, including the **Help** menu and the Help button on most *dialogue boxes*.

Homogeneity of regression slopes

An assumption of *ANCOVA* that the relationship between the *dependent variable* and the covariate is similar across all levels of the *independent variable*.

Homogeneity of variance

Also referred to as *equality of variance*. One of the requirements for using *parametric* statistical tests: the *variance* of the data for one *group or condition* should be relatively similar to that of the other groups or conditions being analysed, even when they come from *populations* with different *means*.

Homogeneity of variance-covariance matrices

An assumption of *multivariate analysis of variance (MANOVA)*, which is an extension of the assumption of *homogeneity of variance*.

Hypothesis

A prediction about the outcome of a study. The experimental or research hypothesis predicts that a difference between *conditions* will occur, that a relationship will be found or that an *interaction* will occur. The *null hypothesis* predicts that there will be no difference between conditions, that a relationship will not be found or that an interaction will not occur.

Independent groups design

An experimental design in which each participant experiences only one level of the independent variable. See also natural independent groups design and between-subjects designs.

Independent variable

A *variable* either that is systematically manipulated by the experimenter to have different values (true experiments), or the values of which are chosen by the experimenter (*natural independent groups designs*). Each value of the independent variable is called a *level*. See also *factor*.

Inferential statistical tests

Procedures that allow you to draw inferences from the data collected. The outcome of an inferential statistical test estimates the probability of obtaining these data if the null hypothesis was true. If that probability is sufficiently small ($p \leq .05$ in psychology), the null hypothesis is rejected; otherwise it is retained. Various inferential statistics are covered in this book.

Interaction

An interaction is present between two *variables* When an interaction is present, the impact of one variable depends on the level of the other variable. For example, a high dose of drug A may impair performance more than a low dose, but drugs A and B may interact such that the effect of drug A is reversed in the presence of drug B.

Interaction graph

A *line graph* showing each *level* of two *factors*. The *dependent variable* is on the *y*-axis and the *levels* of one *factor* on the *x*-axis; the *levels* of a second *factor* are indicated by separate lines on the graph. See also *chart*.

Interval data

Data that have been measured at the interval level. Also referred to as continuous data. These are data that are measured along a scale, where the differences (or intervals) on adjacent points on the scale are the same. SPSS refers to these data type as 'scale'.

Irrelevant variable

Any *variable* other than the *independent variable* (or *factor*) and the *dependent variable/s*. Good *experimental design* should ensure that irrelevant variables are controlled so that they do not become *confounding variables*.

Kendall's tau

An inferential statistical test of correlation used to analyse nonparametric data.

Kruskal-Wallis

A nonparametric equivalent of the one-way between-subjects ANOVA.

Level

An *independent variable* or *factor* will have two or more levels. For example, the variable temperature may have two levels: hot and cold. If there is only one *factor*, then its levels are equivalent to the *conditions* of the experiment. When there are two or more *factors*, the conditions are defined by the particular combination of the levels of the *factor* (for example, male participants tested under hot conditions).

Levels of measurement

The type of scale used to measure *variables*. Usually, four levels of measurement are described: *nominal*, *ordinal*, *interval* and *ratio*. The first two are classified as *nonparametric* levels of measurement, and the last two as *parametric* levels of measurement. SPSS uses the term *scale* to describe interval and ratio levels of measurement. See also *measure*.

Linear relationship

A relationship between two *variables* is said to be linear if the data points follow a straight line (or close to a straight line) when plotted in a *scatterplot*.

Linear trend

Describes the situation in which the *mean* values from a *factor* with three or more *levels* follow a straight line; this can be assessed for significance. See also *quadratic trend*.

Line graph

A graph in which the points plotted are joined by a line. The points could each represent the *mean* of one sample, or they could represent the *frequency* of particular values in an SPSS *variable*. See also *interaction graph* and *chart*.

Logarithmic transformation

A *data transformation* in which each point in a data set is replaced with its logarithmic (either natural or common) value. Logarithmic transformations are typically applied to a data set to reduce the influence of *outliers*, so the data more closely match the assumptions of a statistical procedure.

Logistic regression

A statistical procedure that examines the impact of a number of *predictor variables* on a categorical (or nominal) dependent variable. SPSS distinguishes between binary logistic regression, in which the dependent variable has two categories, and multinomial logistic regression, in which the dependent variable can have more than two categories.

Log transformation

See logarithmic transformation.

Mann–Whitney

An inferential statistical test used to analyse *nonparametric* data from *two-sample independent groups designs*.

Matched-subjects design

An *experimental design* in which each *participant* is matched closely with another participant, to give a participant pair. Each member of the pair is then allocated, by a random process, to different *levels* of the *independent variable*. Also called matched pairs design. It is a type of *related design*.

McNemar

An inferential statistical test used to analyse *nominal* data obtained by measuring a *dichotomous variable* for a *two-sample related design*.

Mean (M)

A *measure of central tendency*: the scores are summed and the total is divided by the number of scores.

Measure

SPSS uses the term measure, in the *Variable View*, to refer to the level of measurement for a *variable*. See *levels of measurement*.

Measure of central tendency

The average or typical score for a sample. See *mean, median* and *mode*.

Measure of dispersion

A measure of variability within a sample. See *range*, *standard deviation*, *standard error* and *variance*.

Median

A *measure of central tendency*: the scores are put into rank order and the middle score is the median.

Menu items

In SPSS, the menu items are the words in the highlighted bar across the top of the window, which give the user access to drop-down lists of options.

Mixed design

A design in which at least one *factor* is *between-subjects* and at least one is *within-subjects*. This is an aspect of the terminology used to describe *ANOVA* designs.

Missing values

A data set may be incomplete, for example, if some observations or measurements failed or if participants didn't respond to some questions. It is important to distinguish these missing data points from valid data. Missing values are the values SPSS has reserved for each variable to indicate that a data point is missing. These missing values can either be specified by the user (*user missing*) or automatically set by SPSS (*system missing*).

Mode

The most common value in a *sample* of scores: a *measure of central tendency*. If a *sample* of scores has more than one mode, SPSS shows the lowest value.

Multinominal logistic regression

See logistic regression.

Multiple regression

An inferential statistical procedure used to investigate *linear relationships* between three or more *variables*. Multiple regression can indicate the extent to which one variable can be explained or predicted by one or more of the other variables. See also *regression*.

Multivariate

Analysis of data in which either two or more *dependent variables* are measured while one or more factors are manipulated, for example, *multivariate analysis of variance*; or three or more *variables* are measured, for example *multiple regression*. See also *univariate*, *bivariate* and *dependent variable*.

Multivariate analysis of variance

An inferential statistical procedure used to analyse data collected using designs in which two or more *dependent variables* are measured, while one or more *factors* are manipulated. See also *ANOVA*.

Ν

In statistics, the character 'N' (uppercase) is typically used to indicate the size of a *population*, while 'n' (lowercase) is used to indicate the *sample* size. SPSS uses 'N' to refer to the number of cases being analysed.

Natural independent groups design

An *independent groups design* or *between-subjects design* in which the groups are chosen by the experimenter from pre-existing (natural) groups. For example: male and female; smoker, ex-smoker and non-smoker. The results of natural groups studies cannot be used to draw conclusions about cause-and-effect relationships, but only about differences or associations that might be a result of variables other than those used to define the natural groups. Studies using a natural independent groups design are sometimes called 'quasi-experimental studies'.

Nominal data

Data collected at a *level of measurement* that yields nominal data (nominal just means 'named'), also referred to as 'categorical data', where the value does not imply anything other than a label; for example, 1 = male and 2 = female.

Nonparametric

A term used to denote:

- 1. nominal and ordinal levels of measurement
- 2. data that may be measured on ratio or interval scales but do not meet the other assumptions (*homogeneity of variance* and normality of distribution) underlying *parametric* statistical tests
- 3. the *inferential statistical tests* used to analyse nonparametric data.

Nonparametric statistics make use of rank order, either of scores or the differences between scores, unlike *parametric* statistical tests. Because nonparametric tests make no assumptions about normality of distribution in the data, they are sometimes called 'distribution-free tests'.

Null hypothesis

The opposite of your research hypothesis. It predicts that there will be no significant effect or relationship between your variables.

One-tailed test

You use a one-tailed test when you have a directional *hypothesis*; that is, when your hypothesis predicts the direction of the relationship between your variables or conditions (i.e. if you predict that one group will score higher than another; or that scores will increase on one variable as they decrease on another).

Options

Options in *dialogue boxes* can be set to request additional statistics or control the appearance of *charts*. Also, selecting **Options** from the **Edit** *menu item* allows you to set options that will be applied more generally.

Ordinal data

Ordinal data are data measured on a scale that only indicates rank and not absolute value. An example of an ordinal scale is academic performance measured by class rank. See also *levels of measurement*.

Outlier

A value recorded for a *variable* that is far from the majority of values.

Output window

See Viewer window.

P value

The *p* value is the probability of obtaining the observed results if the null hypothesis were true. In psychology, it is conventional to declare a result to be statistically significant if the *p* value is less than .05. See also *significance level*.

Parameters

A characteristic of an entire *population*, such as the *mean*. Normally, we estimate a *population* parameter based on the statistics of the *sample*.

Parametric

A term used to denote:

- 1. ratio and interval levels of measurement
- data that are measured on one of those scales and also meet the two other requirements (*homogeneity of variance* and normality of distribution) for parametric statistical tests
- 3. the *inferential statistical tests* used to analyse parametric data.

Parametric statistics make use of the actual values of scores in each *sample*, unlike *nonparametric* statistical tests.

Paste

In SPSS, the Paste button pastes the lines of *syntax* required to perform a command into the *Syntax Editor window*.

Participant

People who take part in an experiment or research study. Previously, the word 'subject' was used, and still is in many statistics books. The word 'subject' is often still used to describe *ANOVA* experimental designs and analyses (e.g. '2*2 within-subjects design') as in this book.

Participant irrelevant variable

Any *irrelevant variable* that is a property of the *participants* in an experiment. The term 'subject irrelevant variables' is sometimes still used.

Pearson's r

An inferential statistical test of correlation used to analyse parametric data.

Pivot table

The name that SPSS gives to a table of results displayed in the *Output viewer window*. The appearance of a pivot table can be altered for the purposes of presentation in a report.

Planned comparisons

A group of statistical tests used to compare particular *conditions* from *ANOVA* designs. Planned comparisons must be specified before data are collected. If used inappropriately, then the *Type 1 error* will be inflated. See also *unplanned comparisons*.

Population

The total set of all possible scores for a particular variable. See also N.

Power

The power of an inferential statistical procedure is the probability that it will yield a statistically significant result.

Predictor variable

The variable/s used in a *regression* analysis to explain another variable (the *criterion variable*). Some sources use the term independent variable/s instead of predictor variable/s.

Principal components analysis (PCA)

PCA is a statistical technique that aims to identify patterns in a large data set based on the patterns and correlations between individual data points. It aims to reduce a set of variables by looking for clusters of variables that appear to be related to one another (and therefore may be tapping into the same underlying factor). As such, PCA is primarily concerned with identifying variables that share variance with one another.

Print

The content, or a selection, of all SPSS windows can be printed by selecting **Print** from the **<u>File</u>** *menu item* while the appropriate window is open.

Quadratic trend

Describes the situation in which the *mean* values from a *factor* with three or more *levels* follow a curved line (for example, a U shape); this can be assessed for significance. See also *linear trend*.

Quantitative data

Is used to describe numeric data measured on any of the four *levels of measurement*. Sometimes though, the term 'qualitative data' is then used to describe data measured with nominal scales.

Quantitative research

In psychology, describes research that involves the analysis of numeric or *quantitative data* measured on any of the four *levels of measurement*. In contrast, qualitative research often involves the detailed analysis of non-numeric data, collected from small, non-random *samples*. SPSS is designed for use in quantitative research.

Range

A *measure of dispersion*: the scores are put into rank order and the lowest score is subtracted from the highest score.

Rank cases

An SPSS procedure that converts interval or ratio data into ordinal data. A new *variable* is added to the data file that gives the rank of one of the existing variables.

Ratio data

Ratio data has all the properties of an interval variable, but also has an absolute value of 0. That is, when the variable equals 0 it means that there is none of that variable, which also gives more meaning to the relationship between different measurements on a ratio scale. For example, as weight is a ratio variable, a weight of 6g is twice as heavy as 3g. Whereas a temperature of 40°C is not twice as hot as 20°C, so while Celsius represents interval data, it is not ratio data as it does not have an absolute zero (i.e. 0°C does not represent an absence of temperature).

Recode

An SPSS procedure that allows the user to systematically change a particular value or range of values in a *variable*.

Regression

If two *variables* have been measured, *bivariate regression* can be used to allow prediction of a participant's score on one variable from their score on the other variable. If three or more variables have been measured, then *multiple regression* can be used to analyse the data. A regression line is the line drawn using the regression formula, and represents the 'best fit' to the data points in a *scatterplot*.

Related designs

A term that includes *repeated measures* and *matched subjects designs*. See also *within-subjects designs*.

Repeated measures design

An *experimental design* in which each *participant* experiences every *level* of the *independent variable*. It is a type of *related design*.

Sample

A subset of a *population*. A smaller set of scores we hope is representative of the *population*. See also *N*.

Scale

In SPSS, describes interval and ratio levels of measurement.

Scatterplot

Sometimes called a 'scattergram' or 'scattergraph'. Two variables are plotted, one on the *x*-axis the other on the *y*-axis, and a point is plotted for each case. Used to display the results of a *correlation* analysis. See also *chart*.

Select cases

SPSS procedure that allows the user to select a subsample of the *cases* on the data file. Cases can be selected on the basis of the values of a particular *variable*. Subsequent analyses will only be performed on the selected cases.

Sig

SPSS uses the shorthand 'sig' to indicate a p value. See also significance level.

Significance level

The outcome of an inferential statistical test estimates the probability of obtaining the observed results assuming the null *hypothesis* is true. If that probability is equal to or less than the significance level, the null hypothesis is rejected; otherwise it is retained. By convention in psychology, the significance level is set to .05.

Situational irrelevant variable

Any *irrelevant variable* that relates to the situation in which an experiment is carried out.

Skewed data

If a data *sample* is not normally distributed but has a 'tail' of *cases* that are either particularly low or particularly high compared with most of the scores, the *sample* is said to be skewed. Such a *sample* does not meet the assumption of normality. See *parametric*.

Sort cases

SPSS procedure by which the *cases* in the *Data window* can be sorted into a desired order based on the values of one or more *variables*.

Spearman's rho

An inferential statistical test of correlation used to analyse nonparametric data.

Sphericity

An assumption when performing a within-subjects or mixed *ANOVA* that the variances of the differences between all possible pairs of levels of a *within-subjects* factor are equal.

Split

SPSS procedure that allows the user to split the *cases* into two or more groups based on the values of a *grouping variable*; subsequent analyses will be performed separately for each group or the groups will be compared.

Standard deviation (SD)

A measure of dispersion that gives an indication of the average difference from the mean. SPSS uses the formula that is designed to estimate the standard deviation of a population based on a sample (i.e. N - 1 is used as the denominator rather than N).

Standard error (SE)

A measure of dispersion that is equal to the standard deviation divided by the square root of N. The full name is 'standard error of the mean'. (The 'standard error of differences between means' is obtained as part of calculations for the t-test; the 'standard error of the estimate' is used in regression.)

Statistics

A general term for procedures for summarising or displaying data (*descriptive statistics*) and for analysing data (*inferential statistical tests*). A characteristic of a *sample*, such as the *mean*, used to estimate the *population parameters*.

Syntax

The program language that can be used to directly control SPSS. For more experienced users, controlling SPSS in this way can sometimes be a useful alternative to using the *dialogue boxes*. Syntax commands may appear in the *Output window*. Syntax commands can be pasted into and edited in the *Syntax window*.

Syntax window

The Syntax Editor window can be used to write *syntax* files to control an analysis. Beginners will not need to use the Syntax window.

System missing

A *missing value* automatically assigned by SPSS. System missing values are shown as dots in the relevant data cells. See also *user missing*.

System variables

Variables reserved by SPSS. System variables have names starting with the '\$' symbol.

t-test

An *inferential statistical test* used to analyse *parametric* data. The one-sample *t*-test compares the *mean* of a single *sample* with a known *population* or score, for example average IQ score. For *two-sample designs*, there are two versions, both comparing two *means*: the independent *t*-test for *independent groups designs*, and the paired *t*-test for *related designs*.

Transformation

See data transformation; logarithmic transformation.

Two-sample designs

Experimental designs with two *levels* of one *independent variable*. See also *independent groups design* and *related designs*.

Two-tailed test

You use a two-tailed test when your *hypothesis* predicts a difference between groups/conditions (or a relationship between variables), but it makes no reference to the direction of the effect. If your hypothesis is directional, then you should use a *one-tailed test*.

Type 1 error

A Type 1 error occurs when the *null hypothesis* is rejected in error (i.e. when we accept that there is an effect, when in reality there isn't). This is a false positive. If the *significance level* is set at .05 (the convention in psychology), we will make a Type 1 error on an average of 1 in 20 occasions. If the significance level is reduced, the chance of Type 1 errors will fall, but the Type 2 error rate will increase. Repeated, unplanned testing of a data set can also increase the Type 1 error rate. See also *planned* and *unplanned comparisons*.

Type 2 error

The situation in which the *null hypothesis* is retained in error (i.e. incorrectly accepting the null hypothesis, when in reality there is a genuine effect in the population). This is a false negative. The Type 2 error rate depends partly on the *significance level* and partly on the power of the *inferential statistical test*.

Univariate

An analysis involving just one *dependent variable*, for example, *Mann–Whitney*, *paired* t-*test* and *ANOVA*. See also *bivariate* and *multivariate*.

Unplanned comparisons

A group of *inferential statistical tests* that may be used to make all the possible comparisons between *conditions* from *ANOVA* designs, as they control for the increased chance of obtaining *Type 1 errors*. See also *planned comparisons*.

User missing

A missing value assigned by the user. See also system missing.

Value label

The label assigned to a particular value of a *variable* in an SPSS data file. Value labels are particularly useful in the case of *nominal* variables. Value labels are included in the SPSS output and will help you to interpret your analysis.

Variable

In *experimental design*, anything that varies that can have different values at different times or for different *cases*. See also *confounding variable*, *dependent variable*, *independent variable* and *irrelevant variable*. A variable in SPSS is represented by a column in the *Data Editor window*.

Variable label

Explanatory label that you can give to an SPSS *variable* when you define it. The variable label is printed in the SPSS output and often is shown in *dialogue boxes*.

Variable name

Name given to an SPSS *variable* when it is defined. The variable name will appear at the top of the column in the *Data Editor window*, and may appear in the SPSS output.

Variable View

In SPSS, the *Data Editor window* has two display settings. The Variable View shows details of the settings for each variable in the data file. See also *Data View*.

Variance

A *measure of dispersion* equal to the square of the *standard deviation*. Homogeneity of variance between the *samples* is one of the requirements for using *parametric* statistical tests. SPSS will test for equality of variance (e.g. when performing the independent t-*test*). A rule of thumb is that the larger variance should be no greater than three times the smaller variance.

Viewer window

Displays the output of statistical procedures in SPSS. Also referred to as the *Output window*.

Wilcoxon matched-pairs signed-ranks test

An *inferential statistical test* used to analyse *nonparametric* data from *two-sample related designs*.

Within-participants design

See within-subjects design.

Within-subjects design

A design in which each *participant* experiences every *level* of every *factor* (or if there are *matched subjects*, in which each pair experiences every level of every factor). See also *repeated measures design* and *related designs*.

z-score

A way of expressing a score in terms of its relationship to the *mean* and *standard deviation* of the *sample*. A *z*-score of -2.5 represents a number that is 2.5 *standard deviations* below the *mean*.

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